

Project 2

Daniel Fredin, Junhan Li, & Eric Chen

Problem 1

Problem 3

a)

ii) 10-fold Cross-Validation

Model Accuracy

```
# Predict outcome using model from training data based on testing data
predictions <- predict(cvmodel, newdata=test)

# Create confusion matrix to assess model fit/performance on test data
con_matx <- confusionMatrix(data=predictions, test$COMMIT)
con_matx
```

```
## Confusion Matrix and Statistics
##
##           Reference
## Prediction no yes
##      no  13   4
##      yes   2  11
##
##              Accuracy : 0.8
##              95% CI : (0.6143, 0.9229)
##      No Information Rate : 0.5
##      P-Value [Acc > NIR] : 0.0007155
##
##              Kappa : 0.6
##
##  McNemar's Test P-Value : 0.6830914
##
##              Sensitivity : 0.8667
##              Specificity : 0.7333
##              Pos Pred Value : 0.7647
##              Neg Pred Value : 0.8462
##              Prevalence : 0.5000
##              Detection Rate : 0.4333
##      Detection Prevalence : 0.5667
##              Balanced Accuracy : 0.8000
```

```
##  
##      'Positive' Class : no  
##
```

b)

c)

```
compute_weibull_stats <- function(shape, scale) {  
  # Compute the mean  
  mean_value <- scale * gamma(1 + (1/shape))  
  
  # # Compute the median  
  # median_value <- scale * qweibull(0.5, shape, scale)  
  
  # Compute the median  
  median_value <- scale * (log(2)^(1/shape))  
  
  # Compute the mode  
  if (shape > 1)  
    mode_value <- scale * ((shape - 1) / shape)^(1/shape)  
  else  
    mode_value <- 0  
  
  # Compute the variance  
  variance_value <- scale^2 * (gamma(1 + (2/shape)) - (gamma(1 + (1/shape)))^2)  
  
  # Return the computed statistics as a named list  
  return(list(mean = mean_value,  
              median = median_value,  
              mode = mode_value,  
              variance = variance_value))  
}  
  
# lambda = scale  
# K = shape  
  
# Example usage  
scale_param <- 6  
shape_param <- 1  
  
stats <- compute_weibull_stats(shape_param, scale_param)  
  
# Accessing the computed statistics  
mean_value <- stats$mean  
median_value <- stats$median  
mode_value <- stats$mode  
variance_value <- stats$variance
```

```
# Printing the computed statistics
cat("Mean:", mean_value, "\n")
```

```
## Mean: 6
```

```
cat("Median:", median_value, "\n")
```

```
## Median: 4.158883
```

```
cat("Mode:", mode_value, "\n")
```

```
## Mode: 0
```

```
cat("Variance:", variance_value, "\n")
```

```
## Variance: 36
```

Maximum likelihood estimator (MLE) of the parameters from the Weibull distribution

Maximum Likelihood Estimator (MLE) of σ from the normal distribution assuming μ is constant.

Step 1: Probability Density Function for a normal distribution.

$$f(x; \lambda, k) = \begin{cases} \frac{k}{\lambda} \left(\frac{x}{\lambda}\right)^{k-1} e^{-\left(\frac{x}{\lambda}\right)^k}, & x \geq 0 \\ 0, & x < 0 \end{cases}$$

Step 2: The likelihood function (L).

$$\begin{aligned} L_{\hat{x}}(\lambda, k) &= \prod_{i=1}^n \frac{k}{\lambda} \left(\frac{x_i}{\lambda}\right)^{k-1} e^{-\left(\frac{x_i}{\lambda}\right)^k} \\ &= \left(\frac{k}{\lambda}\right)^n \left(\frac{1}{\lambda^{k-1}}\right)^n \left(\prod_{i=1}^n x_i^{k-1}\right) \left(e^{-\sum_{i=1}^n \left(\frac{x_i}{\lambda}\right)^k}\right) \\ &= \frac{k^n}{\lambda^{nk}} e^{-\sum_{i=1}^n \left(\frac{x_i}{\lambda}\right)^k} \prod_{i=1}^n x_i^{k-1} \end{aligned}$$

Step 3: The log-likelihood function (ln(L)).

$$\ln L_{\hat{x}}(\lambda, k) = n \ln(k) - nk \ln(\lambda) - \sum_{i=1}^n \left(\frac{x_i}{\lambda}\right)^k + (k-1) \sum_{i=1}^n \ln x_i$$

Step 4: Derivative of the log-likelihood function ($\ln(L)$).

$$\begin{aligned}
\frac{d \ln L}{d\sigma^2} &= \frac{d}{d\sigma^2} \left(-\frac{n}{2} \ln(2\pi) - \frac{n}{2} \ln(\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2 \right) \\
&= -\frac{n}{2\sigma^2} - \left[\frac{1}{2} \sum_{i=1}^n (x_i - \mu)^2 \right] \frac{d}{d\sigma^2} \left(\frac{1}{\sigma^2} \right) \\
&= -\frac{n}{2\sigma^2} - \left[\frac{1}{2} \sum_{i=1}^n (x_i - \mu)^2 \right] \left(-\frac{1}{(\sigma^2)^2} \right) \\
&= \frac{1}{2\sigma^2} \left[\frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \mu)^2 - n \right]
\end{aligned}$$

Step 5: Solving for the desired parameter.

$$\begin{aligned}
\frac{d \ln L}{d\sigma^2} &= 0 \\
\frac{d \ln L}{d\sigma^2} &= \frac{1}{2\sigma^2} \left[\frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \mu)^2 - n \right] = 0 \\
&= \frac{1}{\sigma^2} \sum_{i=1}^n (x_i - \mu)^2 - n = 0 \\
&= \frac{1}{n} \sum_{i=1}^n (x_i - \mu)^2 = \sigma^2
\end{aligned}$$

Therefore the estimator $\hat{\sigma}$ is equal to the unadjusted sample variance.

$$\hat{\sigma} = \sqrt{\frac{1}{n} \sum_{i=1}^n (x_i - \hat{\mu})^2}$$

Since $\hat{\mu} = \bar{x}$ then the estimator $\hat{\sigma}$ is equal to,

$$\hat{\sigma} = \sqrt{\frac{1}{n} \sum_{i=1}^n (x_i - \bar{x})^2}$$