

# HW-3

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## Homework 3 - Predictive Modeling in Finance and Insurance

### 1. Likelihood Function for mean of normal distribution

#### a. Joint Density Function

Note that  $Y_1, Y_2$ , and  $Y_3$  are independent. Therefore, their joint probability density function (p.d.f) is a product of their marginal probability density functions:

$$\begin{aligned} f_{(Y_1, Y_2, Y_3)}(y_1, y_2, y_3) &= f_{Y_1}(y_1)f_{Y_2}(y_2)f_{Y_3}(y_3) \\ &= \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(y_1-\mu_1)^2} * \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(y_2-\mu_2)^2} * \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(y_3-\mu_3)^2} \\ &= \frac{1}{(2\pi\sigma^2)^{\frac{3}{2}}} * e^{-\frac{1}{2\sigma^2}(\sum_{i=1}^3 (y_i-\mu_i)^2)} \end{aligned}$$

#### b. Likelihood function and Log-Likelihood

The likelihood function is just the joint p.d.f, given parameter of interest  $\vec{\mu} = (\mu_1, \mu_2, \mu_3)$ :

$$L(\vec{\mu}) = f_{(Y_1, Y_2, Y_3)}(y_1, y_2, y_3; \mu) = \frac{1}{(2\pi\sigma^2)^{\frac{3}{2}}} * e^{-\frac{1}{2\sigma^2}(\sum_{i=1}^3 (y_i-\mu_i)^2)}$$

The log-likelihood is just the natural log of this function:

$$\begin{aligned} \ell(\vec{\mu}) &= \ln(L(\sigma^2)) = \ln\left(\frac{1}{(2\pi\sigma^2)^{\frac{3}{2}}}\right) + \ln\left(e^{-\frac{1}{2\sigma^2}(\sum_{i=1}^3 (y_i-\mu_i)^2)}\right) \\ &= -\frac{3}{2}\ln(2\pi\sigma^2) - \frac{1}{2\sigma^2}\left(\sum_{i=1}^3 (y_i-\mu_i)^2\right) \end{aligned}$$

#### c. Score function, Observed Information, Expected Information

The score function is simply the derivative of the log likelihood with respect to the parameter of interest,  $\sigma^2$ : (I call  $\vec{y} = (y_1, y_2, y_3)$  for simplicity)

$$S(\sigma^2; \vec{y}) = \frac{d}{d\sigma^2}\ell(\sigma^2) = -\frac{3}{2\sigma^2} + \frac{1}{2(\sigma^2)^2}((y_1-700)^2 + (y_2-750)^2 + (y_3-850)^2)$$

The observed information is simply the second derivative of the log likelihood multiplied by  $-1$ :

$$\begin{aligned} j(\sigma^2; \vec{y}) &= -\frac{d^2}{d(\sigma^2)^2}\ell(\sigma^2) = -\left(\frac{3}{2(\sigma^2)^2} - \frac{1}{(\sigma^2)^3}((y_1-700)^2 + (y_2-750)^2 + (y_3-850)^2)\right) \\ &= -\frac{3}{2(\sigma^2)^2} + \frac{1}{(\sigma^2)^3}((y_1-700)^2 + (y_2-750)^2 + (y_3-850)^2) \end{aligned}$$

## 2. Fun with Distributions

### a. Distribution of $Y_1^2$

Since  $Y_1 \sim N(0, 1)$ ,  $Y_1^2 \sim \chi^2(1)$ , or the chi-squared distribution with 1 degree of freedom.

### b. Combination of $Y_1$ and $Y_2$

Note  $\frac{Y_2 - \mu_2}{\sigma_2} = \frac{Y_2 - 3}{2} \sim N(0, 1)$ ; therefore:

$$\left(\frac{Y_2 - 3}{2}\right)^2 \sim \chi^2(1)$$

Using the independence of  $Y_1$  and  $Y_2$  and Cochran's Theorem:

$$y^T y = \begin{bmatrix} Y_1 & \frac{Y_2 - 3}{2} \end{bmatrix} * \begin{bmatrix} Y_1 \\ \frac{Y_2 - 3}{2} \end{bmatrix} = Y_1^2 + \left(\frac{Y_2 - 3}{2}\right)^2 = \chi^2(1 + 1) = \chi^2(2)$$

So,  $y^T y$  has the chi-squared distribution with 2 degrees of freedom.

### c. Multivariate Normal

Note that  $V$  in this case is the Variance-Covariance matrix. Since  $Y_1$  and  $Y_2$  are independent, the off-diagonal elements, which represent covariance, are 0. There diagonal elements are just  $\sigma_1^2 = 1$  and  $\sigma_2^2 = 4$ , respectively, so:

$$V = \begin{bmatrix} 1 & 0 \\ 0 & 4 \end{bmatrix}$$

I find the inverse of this 2 by 2 matrix:

$$V^{-1} = \frac{1}{1(4) - 0(0)} \begin{bmatrix} 4 & 0 \\ 0 & 1 \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & \frac{1}{4} \end{bmatrix}$$

Therefore:

$$\begin{aligned} y^T V^{-1} y &= \begin{bmatrix} Y_1 & Y_2 \end{bmatrix} * \begin{bmatrix} 1 & 0 \\ 0 & \frac{1}{4} \end{bmatrix} * \begin{bmatrix} Y_1 \\ Y_2 \end{bmatrix} \\ &= \begin{bmatrix} Y_1 & \frac{Y_2}{4} \end{bmatrix} * \begin{bmatrix} Y_1 \\ Y_2 \end{bmatrix} \\ &= Y_1^2 + \left(\frac{Y_2}{2}\right)^2 \end{aligned}$$

## 4. Linear Regression

### a. Fitting model B

```
library(ggplot2)
library(readxl)
```

I first import the data:

```
carbData <- read_excel("Table 6.3 Carbohydrate diet-1.xls", skip = 2, sheet = "Sheet1")
```