

Problem Set 4 *Solutions*

Question 1. This problem will use the panel dataset of Texas elementary schools used in class (*texas_elementary_panel.dta*) to estimate the effects of student mobility on school average performance on standardized tests. (29 points)

use https://github.com/spcorcor18/LP0-8852/raw/main/data/Texas_elementary_panel_2004_2007.dta

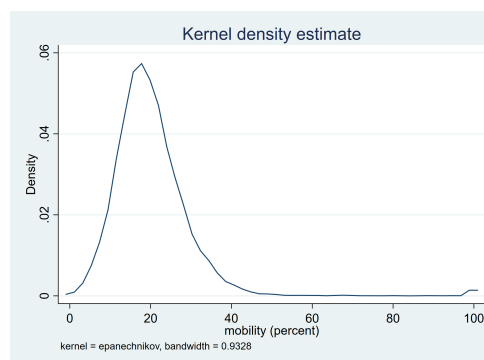
- (a) The variable *cpemallp* is defined as the percentage of students in a school who were enrolled less than 83% of the school year (i.e., were not present 6 or more weeks at that school). Rename this variable *mobility*, report the overall mean and standard deviation for this variable, and produce a kernel density plot for this variable (use the *kdensity* command). The kernel density is like a smooth approximation to a histogram. Describe what this distribution looks like. (3 points)

```
. rename cpemallp mobility
```

```
. sum mobility avgpassing
```

Variable	Obs	Mean	Std. Dev.	Min	Max
mobility	16,072	20.15625	9.892888	0	100
avgpassing	16,225	75.4024	13.83064	5	99

```
. kdensity mobility
```



The mean *mobility* share is 20.2%, meaning in the average school about 1 in 5 students were enrolled less than 83% of the school year. The standard deviation is 9.9 percentage points (based on the state-by-year observations). The kernel density shows this is very right-skewed distribution, with some schools having unusually high mobility rates.

- (b) Declare this dataset to be a panel using `xtset`. Use the same cross-sectional unit and time dimension variables used in class. Use `xtsum` to get a set of descriptive statistics for *mobility*. Does it appear that school mobility is primarily a between-school phenomenon, or something that varies more within schools over time? Explain how you know, and explain in words how the standard deviations (overall, within, and between) are calculated. (4 points)

```
. xtset campus year
      panel variable:  campus (unbalanced)
      time variable:  year, 2004 to 2007, but with gaps
              delta:  1 unit
```

```
. xtsum mobility
```

Variable		Mean	Std. Dev.	Min	Max	Observations
mobility overall		20.15625	9.892888	0	100	N = 16072
between			10.56573	0	100	n = 4302
within			2.926674	-7.643754	70.18124	T-bar = 3.73594

At 10.6, the between-school standard deviation of *mobility* is considerably larger than the within-school standard deviation of 2.9. The latter is calculated using deviations from school-specific means, while the former is calculated using deviations of school-specific means from the grand mean. The “overall” standard deviation uses deviations of each data point from the grand mean. The finding that school mobility is primarily a between-school phenomenon is not surprising. Some schools likely suffer from persistently high mobility year after year. Annual deviations from this long-run average are likely to be small.

- (c) Estimate a simple regression of the average TAKS exam passing rate (*avgpassing*) on mobility (refer to the lecture notes for the *avgpassing* variable). How are these variables related? Report your results and interpret your coefficient estimate in words. Is the coefficient statistically significant? Practically significant? (4 points)

Results are below. There is a strong negative relationship between school mobility and the average passing rate on state tests. The estimated coefficient is statistically ($p < 0.001$) and practically significant. A one-standard deviation increase in school mobility rates (9.9) is associated with a 7.5 percentage point lower passing rate. When benchmarked against the standard deviation in passing rates in the data (13.8), this is a large effect.

```
. rename ca311tar avgpassing
. reg avgpassing mobility
```

Source	SS	df	MS	Number of obs	=	15,831
Model	525812.917	1	525812.917	F(1, 15829)	=	3341.83
Residual	2490582.58	15,829	157.343015	Prob > F	=	0.0000
				R-squared	=	0.1743
				Adj R-squared	=	0.1743
Total	3016395.5	15,830	190.549305	Root MSE	=	12.544

avgpasing	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
mobility	-.7570524	.0130959	-57.81	0.000	-.7827218 -.731383
_cons	90.29461	.276085	327.05	0.000	89.75345 90.83576

- (d) Should the coefficient estimated in part (c) be interpreted as the causal effect of student mobility on school performance? Briefly explain why or why not. (2 points)

For the regression in (c) to have a causal interpretation, we have to believe that the covariance between the population error term u and *mobility* is zero. This seems unlikely if there are omitted variables correlated with both mobility and passing rates. Chances are, schools with high mobility rates are disadvantaged in other ways that would lead us to predict lower achievement in those schools.

- (e) Add the following explanatory variables to your regression in (c): percent black, white, Hispanic, Asian or Pacific Islander (API), Limited English Proficient (LEP), and economically disadvantaged. Also include year effects and a dummy variable for charter schools (*charter*, which may need to be encoded as numeric). How does the inclusion of these covariates affect your estimated coefficient on *mobility*? Is it still statistically significant? Does the change make sense to you (explain)? Finally, provide a written interpretation of the estimated coefficients for the three year dummies (2005, 2006 and 2007). (4 points)

Results shown below. Perhaps not surprisingly, the coefficient on *mobility* is much smaller in absolute value (-0.152). This was anticipated given our answer in part (d). Omitted variables were likely positively correlated with *mobility* and negatively correlated with *avgpasing*, suggesting our “short” regression coefficient was upwardly biased. That is, it likely over-stated the negative relationship between *mobility* and *avgpasing*.

```
. encode charter, gen(charter2)
```

```
. reg avgpasing mobility cpetblap cpetwhip cpethisp cpetpacp cpetecop i.year i.charter2
```

Source	SS	df	MS	Number of obs	=	15,831
Model	1458455.37	10	145845.537	F(10, 15820)	=	1480.98
				Prob > F	=	0.0000

Residual		1557940.13	15,820	98.4791487	R-squared	=	0.4835
-----+-----					Adj R-squared	=	0.4832
Total		3016395.5	15,830	190.549305	Root MSE	=	9.9237

avgpassing		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----							
mobility		-.1524116	.0129286	-11.79	0.000	-.1777531	-.12707
cpetblap		-.0949501	.1328527	-0.71	0.475	-.3553565	.1654564
cpetwhip		.021062	.1338113	0.16	0.875	-.2412234	.2833475
cpethisp		-.0208526	.1328268	-0.16	0.875	-.2812084	.2395031
cpetpacp		.1701193	.1345458	1.26	0.206	-.0936058	.4338443
cpetecop		-.2318625	.0062852	-36.89	0.000	-.2441822	-.2195428
year							
2005		5.127777	.2243313	22.86	0.000	4.688062	5.567492
2006		6.194567	.2246185	27.58	0.000	5.75429	6.634845
2007		8.183221	.2229574	36.70	0.000	7.746199	8.620243
charter2							
Y		-9.452535	.5655251	-16.71	0.000	-10.56103	-8.344042
_cons		89.32539	13.3053	6.71	0.000	63.24549	115.4053

- (f) Should the coefficient estimated in (e) be interpreted as the causal effect of student mobility on school performance? Briefly explain why or why not. How might a regression model with school fixed effects improve upon the model in (e)? (3 points)

Again, for the regression in (e) to have a causal interpretation, we have to believe that the covariance between the population error term u and *mobility* is zero, conditional on the other explanatory variables. While we have now controlled for several school characteristics that made this assumption more plausible, there may be other unobserved school characteristics that are omitted from the regression that are systematically related to *mobility* and *avgpasing*.

- (g) Estimate the regression in (e) with school fixed effects. How does this approach affect the estimated coefficient on *mobility*? Is it still statistically significant? Does the change make sense to you? Provide an intuitive explanation of the finding. Were any explanatory variables dropped from the model (or are there any that you expected would fall out that didn't)? (5 points)

Results are shown below. Interestingly, the coefficient on *mobility* is now very small and statistically insignificant. This change makes sense if we believe the school fixed effect is capturing unobserved school characteristics that are systematically associated with high mobility rates and low achievement. The fixed effects model relies entirely on *within-school* variation in

mobility rates over time to estimate the slope coefficients. Note that charter status falls out of the model, since it is time-invariant.

```
. xtreg avgpassing mobility cpetblap cpetwhip cpethisp cpetpacp cpetecop i.year i.charter2, fe
note: 2.charter2 omitted because of collinearity
```

```
Fixed-effects (within) regression      Number of obs   =    15,831
Group variable: campus                 Number of groups =     4,230
```

```
R-sq:                                Obs per group:
    within = 0.2684                    min =          1
    between = 0.3627                   avg =         3.7
    overall = 0.3351                   max =          4
```

```
corr(u_i, Xb) = -0.1838                F(9,11592)       =    472.42
                                          Prob > F         =     0.0000
```

avgpassing	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
mobility	.0002792	.016423	0.02	0.986	-.0319127	.0324712
cpetblap	.4700703	.1832306	2.57	0.010	.1109074	.8292332
cpetwhip	.8211552	.1828463	4.49	0.000	.4627456	1.179565
cpethisp	.5235963	.1826816	2.87	0.004	.1655095	.881683
cpetpacp	.4689925	.1946184	2.41	0.016	.0875076	.8504774
cpetecop	-.007252	.0148914	-0.49	0.626	-.0364416	.0219377
year						
2005	5.144047	.1291655	39.83	0.000	4.89086	5.397233
2006	6.222979	.1366713	45.53	0.000	5.955081	6.490878
2007	8.739497	.1437972	60.78	0.000	8.457631	9.021364
charter2						
Y	0	(omitted)				
_cons	8.935894	18.19651	0.49	0.623	-26.73234	44.60413
sigma_u	10.502657					
sigma_e	5.575046					
rho	.78016966	(fraction of variance due to u_i)				

```
F test that all u_i=0: F(4229, 11592) = 9.32                Prob > F = 0.0000
```

- (h) What statistical assumptions must hold in order to interpret the coefficient estimate in (g) as causal? (4 points)

The fixed effects assumptions should hold, as described in the Wooldridge text. These include FE1 (linear model), FE2 (cross-sectional units are a random sample), FE3 (variation in x over time, with no perfect collinearity), and FE4 (strict exogeneity). The last assumption is rather important: there can effectively be no relationship between the population error term u and

the x in any time period. In this context, this assumption would be violated if, for example, unusually low achievement in one year affected the mobility rate in another year, perhaps through a changing composition of students in the school. Assumptions 5-6 in Wooldridge relate to the error variance, and thus the appropriate calculation of standard errors. It would make sense to adjust standard errors for clustering at the school level in this context.

Question 2. This problem will examine teacher effects on students' math and reading achievement using student-level data from a large urban school district. You will use methods that are closely related to those used in practice for estimating teacher "value-added." You can find the necessary data on Github under the name *LUSD4_5.dta*. All students in this database are in grades 4 and 5, and the test results are from 2005 and 2006. (26 points)

use https://github.com/spcorcor18/LP0-8852/raw/main/data/LUSD4_5.dta

- (a) First provide some descriptive information about the contents of this panel database. How many student observations are there in each grade and year? How many students appear in *both* grades 4 and 5 in this data? How many unique schools are in the data? How many unique teachers? The variable *school* is a unique school identifier, and *teacher* is the unique teacher identifier. Be clear in your Stata code how you answered these questions. (3 points)

See below. By using `xtset` with *id* and *grade* we can easily see how many students appear in both grades (N=9,728). There are other ways one can do this. There are 190 unique schools and 1,856 unique teachers.

```
. table grade year, row col
```

```
-----
grade      |      year (spring)
level      |      2005      2006      Total
-----+-----
          4 | 12,116   11,556   23,672
          5 | 11,919   11,570   23,489
          |
        Total | 24,035   23,126   47,161
-----
```

```
. xtset id grade
      panel variable:  id (unbalanced)
      time variable:  grade, 4 to 5
                  delta:  1 unit
```

```
. xtdescribe
```

```
id:  9.000e+09, 9.000e+09, ..., 9.001e+09      n =      37433
```

```

grade: 4, 5, ..., 5                                T =                2
      Delta(grade) = 1 unit
      Span(grade)  = 2 periods
      (id*grade uniquely identifies each observation)

```

```

Distribution of T_i:  min      5%      25%      50%      75%      95%      max
                   1         1         1         1         2         2         2

```

```

      Freq.  Percent   Cum. | Pattern
-----+-----
    13944    37.25   37.25 |  1.
    13761    36.76   74.01 |  .1
     9728    25.99  100.00 | 11
-----+-----
    37433   100.00           |  XX

```

```

. unique school
Number of unique values of school is 190
Number of records is 47161

. unique teacher
Number of unique values of teacher is 1856
Number of records is 47161

```

- (b) Estimate four separate regressions: by grade (4 and 5) and by subject (math and reading). The dependent variable will be either the standardized math score (*mathz*) or standardized reading score (*readz*). Both are z-scores with a mean of zero and standard deviation of 1 (standardized for the grade, subject, and year). Use the following explanatory variables: age, female, LEP, special ed, immigrant, economically disadvantaged, black, Hispanic, Asian, and a year effect (i.e., a dummy variable for 2006). At this point, do not include any fixed effects. Provide a brief interpretation of your regression results. (5 points)

Results below. Across models, older students, special education, economically disadvantaged, LEP, black, and Hispanic students tend to score lower than their younger, non-special education, non-economically disadvantaged, non-LEP, white, and Asian counterparts. Girls tend to score lower in math than boys, but higher in reading. Scores tend to be higher in 2006 than in 2005. (This may seem unusual since these are standardized by year, but it may have to do with sample composition).

```

. foreach g in 4 5 {
2.  foreach s in math read {
3.    reg 's'z age female lep speced immig econdis black hispanic asian i.year if grade=='g'
4.  }
5.  }

```

```

Source |          SS          df          MS      Number of obs      =      23,611

```


Source	SS	df	MS	Number of obs	=	23,225
				F(10, 23214)	=	599.20
Model	3708.88513	10	370.888513	Prob > F	=	0.0000
Residual	14368.7779	23,214	.61897036	R-squared	=	0.2052
				Adj R-squared	=	0.2048
Total	18077.6631	23,224	.778404369	Root MSE	=	.78675

mathz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	-.2383004	.0089102	-26.74	0.000	-.2557651	-.2208358
female	-.095322	.0104105	-9.16	0.000	-.1157274	-.0749167
lep	-.3295649	.013557	-24.31	0.000	-.3561376	-.3029922
speced	-.6832481	.0236204	-28.93	0.000	-.7295457	-.6369505
immig	.0375388	.0324508	1.16	0.247	-.0260669	.1011444
econdis	-.2781333	.0167387	-16.62	0.000	-.3109423	-.2453243
black	-.6007952	.0224095	-26.81	0.000	-.6447193	-.5568712
hispanic	-.2916257	.0222087	-13.13	0.000	-.3351562	-.2480952
asian	.2373296	.0338157	7.02	0.000	.1710486	.3036106
year						
2006	.2088527	.0103401	20.20	0.000	.1885853	.2291201
_cons	3.474444	.1005276	34.56	0.000	3.277403	3.671485

Source	SS	df	MS	Number of obs	=	22,699
				F(10, 22688)	=	822.96
Model	5973.09754	10	597.309754	Prob > F	=	0.0000
Residual	16467.1657	22,688	.725809489	R-squared	=	0.2662
				Adj R-squared	=	0.2659
Total	22440.2632	22,698	.988644957	Root MSE	=	.85194

readz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	-.221383	.0098717	-22.43	0.000	-.2407321	-.2020338
female	.0124986	.0113864	1.10	0.272	-.0098195	.0348166
lep	-.7163557	.0148404	-48.27	0.000	-.745444	-.6872674
speced	-.4139044	.0316657	-13.07	0.000	-.4759713	-.3518375
immig	.1305244	.0404587	3.23	0.001	.0512226	.2098262
econdis	-.4631401	.0182894	-25.32	0.000	-.4989887	-.4272916
black	-.6213506	.0244451	-25.42	0.000	-.6692646	-.5734365
hispanic	-.4472469	.0242336	-18.46	0.000	-.4947465	-.3997474
asian	-.0723524	.0369619	-1.96	0.050	-.1448002	.0000954
year						
2006	.0445679	.0113193	3.94	0.000	.0223813	.0667546
_cons	3.548013	.1114301	31.84	0.000	3.329602	3.766423

- (c) Now estimate the same regressions as in part (b), but add as an additional control the lagged math score (in the math regressions) and the lagged reading score (in the reading regressions). These variables are already in the dataset as *mathz_1* and *readz_1*. How do the results change, and how should our interpretation of these results change, given the inclusion of lagged (prior grade) achievement? (5 points)

Results shown below. Not surprisingly, the coefficient on the lagged score is positive and highly significant. (A student's score in the prior grade is a strong predictor of their score in the current grade). The interpretation of the other slope coefficients now differs since achievement in the prior grade is being controlled for. For example, the coefficient on *econdis* is now the predicted difference between the average scores of economically disadvantaged students and non-economically disadvantaged students, holding constant the other predictor variables in the model and prior achievement. For example, 4th grade students who are economically disadvantaged do worse in math than their prior year's math score would predict. Some analysts think of this in terms of "gains," although we are not strictly modeling year-to-year gains.

```
. foreach g in 4 5 {
  2. foreach s in math read {
  3.   reg 's'z 's'z_1 age female lep speced immig econdis black hispanic asian i.year if grade==
  4.   }
  5.   }
```

Source	SS	df	MS	Number of obs	=	23,453
				F(11, 23441)	=	1752.16
Model	8622.61122	11	783.873747	Prob > F	=	0.0000
Residual	10486.9181	23,441	.447375029	R-squared	=	0.4512
				Adj R-squared	=	0.4510
Total	19109.5293	23,452	.814835804	Root MSE	=	.66886

mathz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
mathz_1	.542333	.0047903	113.21	0.000	.5329437 .5517223
age	-.1578989	.00832	-18.98	0.000	-.1742066 -.1415912
female	-.0366597	.0088193	-4.16	0.000	-.053946 -.0193734
lep	-.0777164	.0113459	-6.85	0.000	-.0999552 -.0554776
speced	-.2977279	.0211226	-14.10	0.000	-.3391295 -.2563262
immig	.0804139	.0260006	3.09	0.002	.0294509 .1313769
econdis	-.1438634	.0144808	-9.93	0.000	-.1722466 -.1154801
black	-.3174066	.0191312	-16.59	0.000	-.354905 -.2799083
hispanic	-.1517818	.0189552	-8.01	0.000	-.1889352 -.1146283
asian	.1002384	.0287294	3.49	0.000	.0439268 .1565499
year					
2006	.092809	.0087483	10.61	0.000	.0756618 .1099561

_cons	1.992013	.0859064	23.19	0.000	1.823631	2.160395
-------	----------	----------	-------	-------	----------	----------

Source	SS	df	MS	Number of obs	=	22,792
				F(11, 22780)	=	1508.25
Model	9446.96228	11	858.814753	Prob > F	=	0.0000
Residual	12971.1558	22,780	.569409826	R-squared	=	0.4214
				Adj R-squared	=	0.4211
Total	22418.1181	22,791	.983639073	Root MSE	=	.75459

readz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
readz_1	.5962596	.0056649	105.26	0.000	.585156	.6073632
age	-.1141489	.0096121	-11.88	0.000	-.1329893	-.0953085
female	.0781508	.0100935	7.74	0.000	.0583668	.0979347
lep	-.1622692	.0129857	-12.50	0.000	-.1877221	-.1368162
speced	-.1671733	.0285109	-5.86	0.000	-.2230566	-.11129
immig	.2316611	.0330387	7.01	0.000	.166903	.2964193
econdis	-.2313162	.0165799	-13.95	0.000	-.2638139	-.1988185
black	-.403942	.0217743	-18.55	0.000	-.4466211	-.3612628
hispanic	-.2514092	.0217181	-11.58	0.000	-.2939782	-.2088403
asian	.0187075	.0328078	0.57	0.569	-.0455981	.0830131
year						
2006	-.0005656	.0100142	-0.06	0.955	-.020194	.0190629
_cons	1.609512	.0991515	16.23	0.000	1.415169	1.803856

Source	SS	df	MS	Number of obs	=	23,152
				F(11, 23140)	=	1853.74
Model	8238.08421	11	748.916746	Prob > F	=	0.0000
Residual	9348.65566	23,140	.404004134	R-squared	=	0.4684
				Adj R-squared	=	0.4682
Total	17586.7399	23,151	.759653573	Root MSE	=	.63561

mathz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
mathz_1	.511132	.0047122	108.47	0.000	.5018958	.5203682
age	-.1221437	.0072934	-16.75	0.000	-.1364392	-.1078482
female	-.0350648	.0084422	-4.15	0.000	-.0516119	-.0185176
lep	-.1784876	.0110541	-16.15	0.000	-.2001544	-.1568209
speced	-.321753	.0197192	-16.32	0.000	-.3604039	-.2831022
immig	.0903732	.026243	3.44	0.001	.0389351	.1418112
econdis	-.1033876	.0136281	-7.59	0.000	-.1300997	-.0766755
black	-.2482319	.0184122	-13.48	0.000	-.2843211	-.2121428
hispanic	-.1219343	.0180299	-6.76	0.000	-.1572741	-.0865946
asian	.1048426	.0273801	3.83	0.000	.0511757	.1585095

year							
2006		.2150489	.0083652	25.71	0.000	.1986525	.2314452
_cons		1.680642	.0830348	20.24	0.000	1.517888	1.843396

Source		SS	df	MS	Number of obs	=	22,595
					F(11, 22583)	=	2078.42
Model		11231.9672	11	1021.08792	Prob > F	=	0.0000
Residual		11094.5981	22,583	.491280966	R-squared	=	0.5031
					Adj R-squared	=	0.5028
Total		22326.5652	22,594	.98816346	Root MSE	=	.70091

readz		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
readz_1		.5459355	.0052757	103.48	0.000	.5355948	.5562762
age		-.123624	.0082074	-15.06	0.000	-.139711	-.107537
female		-.0428927	.0094032	-4.56	0.000	-.0613237	-.0244617
lep		-.5386067	.0123582	-43.58	0.000	-.5628296	-.5143838
speced		-.1946905	.0267903	-7.27	0.000	-.2472012	-.1421797
immig		-.061676	.0345765	-1.78	0.074	-.1294483	.0060963
econdis		-.2332101	.0152689	-15.27	0.000	-.2631381	-.203282
black		-.3151816	.0204038	-15.45	0.000	-.3551746	-.2751887
hispanic		-.2482304	.0201228	-12.34	0.000	-.2876726	-.2087882
asian		-.0199688	.0305672	-0.65	0.514	-.0798826	.039945
year							
2006		.0383332	.0093364	4.11	0.000	.0200332	.0566333
_cons		1.972282	.093289	21.14	0.000	1.789429	2.155135

- (d) Next, estimate the regressions in part (c) (with the lagged score), but this time use `xtreg` and include a fixed effect for the classroom teacher. (Instead of using `xtset`, you can include the options `fe` and `i(teacher)` in the `xtreg` command. This is equivalent to `xtset` without officially setting the panel variables). How should our interpretation of these results change, given the inclusion of teacher fixed effects? Report your results. (5 points)

Results below. The interpretations of the slope coefficients do not have a fundamentally different interpretation, but it is important to keep in mind that they are estimated using *within-teacher* variation in the covariates and achievement. So, for example, the achievement of girls is effectively compared with the achievement of boys in the same class.

```
. foreach g in 4 5 {
  2. foreach s in read math {
```

```

3.      xtreg 's'z 's'z_1 age female lep speced immig econdis black hispanic asian ///
>      i.year if grade=='g', fe i(teacher)
4.      }
5.      }

```

warning: existing panel variable is not teacher

```

Fixed-effects (within) regression      Number of obs      =      22,792
Group variable: teacher                Number of groups   =       1,065

```

```

R-sq:                                Obs per group:
    within = 0.3308                      min =           1
    between = 0.5839                      avg  =          21.4
    overall = 0.4132                      max  =           46

```

```

corr(u_i, Xb) = 0.1508                  F(11,21716)        =       975.82
                                      Prob > F            =       0.0000

```

readz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
readz_1	.5593767	.005819	96.13	0.000	.547971	.5707824
age	-.0927569	.0092881	-9.99	0.000	-.1109623	-.0745514
female	.0797181	.0096228	8.28	0.000	.0608567	.0985794
lep	-.2897263	.0254108	-11.40	0.000	-.3395333	-.2399193
speced	-.1660317	.0276888	-6.00	0.000	-.2203038	-.1117597
immig	.2069517	.0318912	6.49	0.000	.1444425	.2694608
econdis	-.1093277	.0178174	-6.14	0.000	-.1442511	-.0744043
black	-.2515821	.0247615	-10.16	0.000	-.3001164	-.2030479
hispanic	-.1448184	.0235422	-6.15	0.000	-.1909628	-.0986741
asian	.010661	.0328858	0.32	0.746	-.0537977	.0751196
year						
2006	-.0085693	.0117874	-0.73	0.467	-.0316735	.0145349
_cons	1.24184	.0978645	12.69	0.000	1.050019	1.433662
sigma_u	.35924206					
sigma_e	.70436283					
rho	.20642771	(fraction of variance due to u_i)				

F test that all u_i=0: F(1064, 21716) = 4.16 Prob > F = 0.0000

```

Fixed-effects (within) regression      Number of obs      =      23,453
Group variable: teacher                Number of groups   =       1,069

```

```

R-sq:                                Obs per group:
    within = 0.3763                      min =           1
    between = 0.5623                      avg  =          21.9
    overall = 0.4478                      max  =           47

```

F(11,22373) = 1227.36

corr(u_i, Xb) = 0.1465 Prob > F = 0.0000

mathz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
mathz_1	.5197333	.004928	105.46	0.000	.510074	.5293926
age	-.1334734	.0078374	-17.03	0.000	-.1488352	-.1181116
female	-.0396879	.0082003	-4.84	0.000	-.0557612	-.0236146
lep	-.1174284	.0207741	-5.65	0.000	-.1581472	-.0767097
speced	-.2785083	.0200028	-13.92	0.000	-.3177153	-.2393014
immig	.0548364	.0245967	2.23	0.026	.0066251	.1030477
econdis	-.0509026	.0151735	-3.35	0.001	-.0806437	-.0211614
black	-.2164042	.0210678	-10.27	0.000	-.2576985	-.1751098
hispanic	-.1043176	.0199644	-5.23	0.000	-.1434492	-.0651861
asian	.063301	.0280419	2.26	0.024	.0083369	.1182651
year						
2006	.0813267	.0100288	8.11	0.000	.0616696	.1009837
_cons	1.636226	.0826756	19.79	0.000	1.474176	1.798276
sigma_u	.36165127					
sigma_e	.60862724					
rho	.26094723	(fraction of variance due to u_i)				

F test that all u_i=0: F(1068, 22373) = 5.56 Prob > F = 0.0000

Fixed-effects (within) regression Number of obs = 22,595
Group variable: teacher Number of groups = 894

R-sq: Obs per group:
within = 0.3515 min = 1
between = 0.7229 avg = 25.3
overall = 0.4917 max = 60

corr(u_i, Xb) = 0.3050 F(11,21690) = 1068.63
Prob > F = 0.0000

readz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
readz_1	.5321018	.0055098	96.57	0.000	.5213022	.5429015
age	-.0998047	.0080284	-12.43	0.000	-.115541	-.0840684
female	-.0481933	.0090702	-5.31	0.000	-.0659715	-.0304151
lep	-.327592	.0191981	-17.06	0.000	-.3652217	-.2899622
speced	-.1791101	.0261651	-6.85	0.000	-.2303956	-.1278247
immig	-.0770656	.0337926	-2.28	0.023	-.1433015	-.0108296
econdis	-.1130444	.0162896	-6.94	0.000	-.1449732	-.0811155
black	-.1930209	.0227613	-8.48	0.000	-.2376347	-.1484072
hispanic	-.1361477	.0216871	-6.28	0.000	-.1786559	-.0936394
asian	-.0317498	.0305681	-1.04	0.299	-.0916655	.0281659

year							
2006		.0172887	.0109689	1.58	0.115	-.004211	.0387885
_cons		1.464472	.0928889	15.77	0.000	1.282403	1.646541

sigma_u		.33181699					
sigma_e		.65984112					
rho		.20184045	(fraction of variance due to u_i)				

F test that all u_i=0: F(893, 21690) = 4.25					Prob > F = 0.0000		
Fixed-effects (within) regression					Number of obs	=	23,152
Group variable: teacher					Number of groups	=	898
R-sq:					Obs per group:		
within	=	0.3882			min	=	1
between	=	0.6437			avg	=	25.8
overall	=	0.4660			max	=	59
					F(11,22243)	=	1283.27
corr(u_i, Xb) = 0.1237					Prob > F	=	0.0000

mathz		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	

mathz_1		.5014138	.0048905	102.53	0.000	.4918281	.5109994
age		-.1038846	.0069652	-14.91	0.000	-.117537	-.0902323
female		-.0390795	.0079636	-4.91	0.000	-.0546887	-.0234704
lep		-.1188986	.0163934	-7.25	0.000	-.1510308	-.0867663
speced		-.2934953	.01881	-15.60	0.000	-.3303642	-.2566264
immig		.0893871	.0252091	3.55	0.000	.0399754	.1387988
econdis		-.0630879	.0142334	-4.43	0.000	-.0909864	-.0351894
black		-.2064401	.0200474	-10.30	0.000	-.2457344	-.1671457
hispanic		-.088959	.0190129	-4.68	0.000	-.1262257	-.0516923
asian		.0940341	.0267766	3.51	0.000	.0415501	.1465181
year							
2006		.1579761	.00962	16.42	0.000	.1391201	.1768321
_cons		1.424638	.0807365	17.65	0.000	1.266389	1.582887

sigma_u		.32092983					
sigma_e		.58513872					
rho		.23125241	(fraction of variance due to u_i)				

F test that all u_i=0: F(897, 22243) = 5.64					Prob > F = 0.0000		

- (e) Teacher fixed effects—systematic variation in achievement after controlling for prior student achievement and other student characteristics—are often referred to as the

teacher's "value added." How much of the variance in achievement is due to the teacher effect? (This is reported as the "rho" in the `xtreg` output). (3 points)

The values of *rho* in the above regressions are 0.206, 0.261, 0.202, and 0.231. After controlling for prior achievement and other student characteristics, roughly 20-25% of the variation in achievement is attributable to variation across teachers. This provides some indication of the "importance" of teachers to student outcomes.

- (f) Save the estimated teacher fixed effects using `predict`, as shown in class. Keep one observation per teacher (you can use `duplicates drop` to do this) and create a histogram of the estimated teacher fixed effects. What is the standard deviation of these teacher fixed effects? What is the difference between a teacher at the 75th percentile of the teacher effect distribution and a teacher at the 25th percentile? (5 points)

Stata syntax and results are shown below (only one histogram is pictured, for 5th grade math). The standard deviation in teacher effects ranges from 0.32 - 0.34, depending on the grade and subject. The difference between the 25th and 75th percentiles ranges from 0.38 to 0.45, depending on the grade and subject. What do these numbers mean? Recall that the fixed effects are estimates of unique intercepts for each teacher. In the case of 4th grade reading, a standard deviation of 0.35 means the students of a teacher one standard deviation above average perform 0.35 better than average than the students of the average teacher.

```
. foreach g in 4 5 {
  2. foreach s in read math {
  3.     qui xtreg 's'z 's'z_1 age female lep speced immig econdis black hispanic asian ///
>     i.year if grade=='g', fe i(teacher)
  4.
  .     predict tcheff's'g', u
  5.     preserve
  6.         duplicates drop teacher, force
  7.         summ tcheff's'g', detail
  8.         tabstat tcheff's'g', stat(p25 p75 iqr)
  9.         histogram tcheff's'g'
  10.    restore
  11.    }
  12.    }
(24,369 missing values generated)
```

Duplicates in terms of teacher

(45,305 observations deleted)

u[teacher]

	Percentiles	Smallest		
1%	-.863707	-1.925172		
5%	-.5911111	-1.28669		
10%	-.4549403	-1.102798	Obs	974
25%	-.2620801	-1.024002	Sum of Wgt.	974
50%	-.0418055		Mean	-.0292627
		Largest	Std. Dev.	.3520821
75%	.1851209	1.11441		
90%	.3914686	1.243015	Variance	.1239618
95%	.5294719	1.418795	Skewness	.0488429
99%	.9337133	1.566302	Kurtosis	4.580015

variable	p25	p75	iqr
-----+-----			
tcheffread4	-.2620801	.1851209	.447201
-----+-----			

(bin=29, start=-1.9251719, width=.12039566)
(23,708 missing values generated)

Duplicates in terms of teacher

(45,305 observations deleted)

	Percentiles	Smallest		
1%	-.9152396	-1.986226		
5%	-.5875053	-1.594639		
10%	-.4370334	-1.163883	Obs	1,004
25%	-.2358474	-1.062243	Sum of Wgt.	1,004
50%	-.0311564		Mean	-.0344432
		Largest	Std. Dev.	.3442303
75%	.169738	.890529		
90%	.3790837	.9583023	Variance	.1184945
95%	.5267823	1.196018	Skewness	-.1834778
99%	.8022034	1.493365	Kurtosis	4.888053

variable	p25	p75	iqr
-----+-----			
tcheffmath4	-.2358474	.169738	.4055854
-----+-----			

(bin=30, start=-1.9862257, width=.11598635)
(24,566 missing values generated)

Duplicates in terms of teacher

(45,305 observations deleted)

u[teacher]

```

-----
      Percentiles      Smallest
  1%      -.8895572      -1.868376
  5%      -.5953411      -1.559686
 10%      -.4524955      -1.512358      Obs              806
 25%      -.219957       -1.300198      Sum of Wgt.         806

 50%      -.0207986
                        Largest      Mean              -.0439087
                        .8488992      Std. Dev.         .3286723
 75%      .1577055      .8488992
 90%      .3190411      .8810328      Variance          .1080255
 95%      .4412906      1.081512      Skewness          -.5181577
 99%      .6588145      1.588887      Kurtosis          5.845002

```

```

      variable |      p25      p75      iqr
-----+-----
  tcheffread5 | -.219957 .1577055 .3776625
-----

```

(bin=28, start=-1.8683757, width=.12347366)
(24,009 missing values generated)

Duplicates in terms of teacher

(45,305 observations deleted)

```

                        u[teacher]
-----
      Percentiles      Smallest
  1%      -.8771342      -1.827223
  5%      -.5433015      -1.528425
 10%      -.4240153      -1.386065      Obs              823
 25%      -.2257731      -1.339902      Sum of Wgt.         823

 50%      -.0290482
                        Largest      Mean              -.0414026
                        .8242272      Std. Dev.         .3205586
 75%      .1641627      .8242272
 90%      .3249792      .8543559      Variance          .1027578
 95%      .4475881      .8997599      Skewness          -.5513376
 99%      .6830953      .9163185      Kurtosis          5.204493

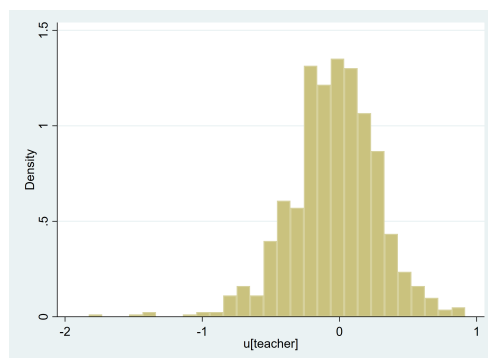
```

```

      variable |      p25      p75      iqr
-----+-----
  tcheffmath5 | -.2257731 .1641627 .3899358
-----

```

(bin=28, start=-1.8272233, width=.09798364)



Question 3. This problem will use the same student-level data from a large urban school district to estimate the impact of having a same-race teacher on achievement. (That is, how a student performs when they share the same race/ethnicity as their teacher, relative to when they don't.) For a study that tackles this very question see Dee (2004). **(20 points)**

use https://github.com/spcorcor18/LP0-8852/raw/main/data/LUSD4_5.dta

- (a) Create a variable called *same_race* that equals zero unless the student and teacher share the same race/ethnicity, in which case *same_race* should be coded as one. Use the white, black, Hispanic, and Asian categories, but not the “other” race category. In what percent of cases (i.e., student-year observations) are students assigned to a teacher of the same race/ethnicity? How does this rate of same race exposure vary by student race/ethnicity? **(4 points)**

Results below. In about 52% of cases (student x year observations) the student had a teacher with the same race or ethnicity. This percentage was higher for black and white students (at 73-74%) and lower for Hispanic (42%) and Asian (6%) students.

```
. gen same_race = 0

. replace same_race = 1 if tch_black==1 & black==1
(8,708 real changes made)

. replace same_race = 1 if tch_white==1 & white==1
(3,295 real changes made)

. replace same_race = 1 if tch_hisp==1 & hisp==1
(12,341 real changes made)

. replace same_race = 1 if tch_asian==1 & asian==1
(87 real changes made)

. tabulate same_race
```

same_race	Freq.	Percent	Cum.
0	22,730	48.20	48.20
1	24,431	51.80	100.00
Total	47,161	100.00	

```
. foreach j in black white hisp asian {
2.   tabulate same_race if 'j'==1
3. }
```

same_race	Freq.	Percent	Cum.
0	3,221	27.00	27.00
1	8,708	73.00	100.00
Total	11,929	100.00	

same_race	Freq.	Percent	Cum.
0	1,182	26.40	26.40
1	3,295	73.60	100.00
Total	4,477	100.00	

same_race	Freq.	Percent	Cum.
0	16,915	57.82	57.82
1	12,341	42.18	100.00
Total	29,256	100.00	

same_race	Freq.	Percent	Cum.
0	1,391	94.11	94.11
1	87	5.89	100.00
Total	1,478	100.00	

- (b) Estimate two regressions where the dependent variables are the math and reading z-scores, respectively, and *same_race* is the explanatory variable. Explain why the estimated coefficient on *same_race* should not be interpreted as causal. (4 points)

Results below, with separate models by subject and grade. In all cases, students with a same race/ethnicity teacher tend to perform worse, on average, than students who do not. For these regressions to have a causal interpretation, we have to believe that the covariance between the population error term u and *same_race* is zero. This seems unlikely if there are omitted variables correlated with both test scores and a match with a same

race/ethnicity teachers. As the correlation matrix shows, black and LEP students are more likely to have a same race teacher. But these students also tend to have lower achievement, on average.

```
. foreach g in 4 5 {
  2. foreach s in math read {
  3.   reg 's'z same_race if grade=='g'
  4.   }
  5. }
```

Source	SS	df	MS	Number of obs	=	23,611

Model	34.3198109	1	34.3198109	F(1, 23609)	=	40.26
Residual	20124.0472	23,609	.8523888	Prob > F	=	0.0000

				R-squared	=	0.0017

				Adj R-squared	=	0.0017
Total	20158.367	23,610	.853806311	Root MSE	=	.92325

mathz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
same_race	-.0767708	.0120988	-6.35	0.000	-.1004852	-.0530563
_cons	.170034	.0090384	18.81	0.000	.1523181	.1877499

Source	SS	df	MS	Number of obs	=	22,963

Model	23.0795803	1	23.0795803	F(1, 22961)	=	23.42
Residual	22626.2559	22,961	.98542119	Prob > F	=	0.0000

				R-squared	=	0.0010

				Adj R-squared	=	0.0010
Total	22649.3355	22,962	.986383395	Root MSE	=	.99268

readz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
same_race	-.0638795	.0131995	-4.84	0.000	-.0897515	-.0380075
_cons	.1045238	.0098844	10.57	0.000	.0851497	.1238979

Source	SS	df	MS	Number of obs	=	23,225

Model	93.7662193	1	93.7662193	F(1, 23223)	=	121.08
Residual	17983.8969	23,223	.774400244	Prob > F	=	0.0000

				R-squared	=	0.0052

				Adj R-squared	=	0.0051
Total	18077.6631	23,224	.778404369	Root MSE	=	.88

mathz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
same_race	-.1272158	.0115611	-11.00	0.000	-.1498764	-.1045552
_cons	.2204888	.0079835	27.62	0.000	.2048406	.236137

Source	SS	df	MS	Number of obs	=	22,699
Model	156.923336	1	156.923336	F(1, 22697)	=	159.84
Residual	22283.3399	22,697	.981774679	Prob > F	=	0.0000
				R-squared	=	0.0070
				Adj R-squared	=	0.0069
Total	22440.2632	22,698	.988644957	Root MSE	=	.99085

readz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
same_race	-.1664715	.0131675	-12.64	0.000	-.1922807	-.1406624
_cons	.1438336	.0090919	15.82	0.000	.1260129	.1616542

```
. corr same_race black white hisp asian lep speced econdis
(obs=47,161)
```

	same_race	black	white	hisp	asian	lep	speced	econdis
same_race	1.0000							
black	0.2468	1.0000						
white	0.1413	-0.1884	1.0000					
hisp	-0.2461	-0.7438	-0.4140	1.0000				
asian	-0.1653	-0.1047	-0.0583	-0.2299	1.0000			
lep	0.2429	-0.3901	-0.2158	0.5182	-0.1054	1.0000		
speced	-0.0164	0.0115	0.0511	-0.0335	-0.0205	-0.0273	1.0000	
econdis	0.0224	0.0237	-0.5321	0.3632	-0.1759	0.2816	-0.0146	1.0000

- (c) Briefly explain how a regression model with *student fixed effects* might improve upon the regressions in part (b). What problem might this solve? (2 points)

There are likely to be observable and unobservable factors correlated with achievement and assignment to a same-race teacher. Some of this may have to do with geography and the local teacher labor market—that is, whether or not teachers share the same demographics as their students. Student fixed effects estimate the “same race” effect using *within-student* variation over time. Students would effectively be compared against themselves, in states in which they are and are not exposed to a same-race teacher. Importantly, students that experience no variation in this explanatory variable do not contribute to the coefficient estimates. This is relevant if we are concerned about generalizing to the full population of students.

- (d) Use `xtset` to designate student as the panel variable, and year as the time dimension. Estimate the same regressions as in Question #2 part (d) (with student covariates and lagged score), and use `xtreg, fe` to include student fixed effects. Also include *same_race* among your explanatory variables. Do **not** run the model separately by

grade; you need multiple observations per student for this model to make sense. Describe what you find for the “same race” coefficient. Is it statistically significant? Practically significant? Can one make a strong claim for causal inference in this case? Explain why or why not. (6 points)

Results below. Interestingly, in all cases the coefficient on *same_race* is positive and statistically significant. When students share the same race/ethnicity as their teacher, they score 0.09 sd higher in reading and 0.04 sd higher in math. Both are statistically and (I would argue) practically significant. It is easier to make a casual claim in this case. One would be concerned about omitted variables bias if there were a time-varying omitted variable that is correlated with changes in both *same_race* and test scores. (This would represent a violation of the strict exogeneity assumption). If, for example, parents responded to a worse- or better-than-expected test result by purposefully moving their student into a classroom with a same-race teacher, this would be a violation of strict exogeneity. It’s not clear whether this is likely to occur in practice, however.

```
. xtset id year
      panel variable:  id (unbalanced)
      time variable:  year, 2005 to 2006
              delta:  1 unit

.
.  foreach s in read math {
2.      xtreg 's'z 's'z_1 age female lep speced immig econdis black hispanic asian i.year same_race
3.      }
```

```
Fixed-effects (within) regression              Number of obs   =       45,387
Group variable: id                           Number of groups =       35,987
```

```
R-sq:                                         Obs per group:
      within = 0.1598                        min =           1
      between = 0.1175                       avg =          1.3
      overall = 0.1082                       max =           2
```

```
corr(u_i, Xb)  = -0.6685                     F(12,9388)       =       148.80
                                         Prob > F         =       0.0000
```

	readz	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
readz_1		-.3575557	.0087662	-40.79	0.000	-.3747392	-.3403721
age		-.3907786	.1608187	-2.43	0.015	-.7060181	-.0755392
female		-.0691433	.2348174	-0.29	0.768	-.5294362	.3911496
lep		.0530388	.0259626	2.04	0.041	.0021465	.1039311
speced		-.0738881	.072518	-1.02	0.308	-.2160391	.0682628
immig		.3530726	.058952	5.99	0.000	.2375138	.4686314

econdis		-.0803718	.0346397	-2.32	0.020	-.1482732	-.0124704
black		.2817545	.6501012	0.43	0.665	-.9925848	1.556094
hispanic		.1130556	.5752511	0.20	0.844	-1.014561	1.240672
asian		.8782935	.5752299	1.53	0.127	-.2492817	2.005869
year							
2006		.3515197	.1610366	2.18	0.029	.035853	.6671863
same_race		.0850675	.0148206	5.74	0.000	.056016	.114119
_cons		4.011723	1.734289	2.31	0.021	.6121403	7.411306

sigma_u		1.2400891					
sigma_e		.52496864					
rho		.84802577	(fraction of variance due to u_i)				

F test that all u_i=0: F(35986, 9388) = 2.22 Prob > F = 0.0000

Fixed-effects (within) regression
 Group variable: id

Number of obs = 46,605
 Number of groups = 37,022

R-sq:
 within = 0.2411
 between = 0.4617
 overall = 0.3945

Obs per group:
 min = 1
 avg = 1.3
 max = 2

corr(u_i, Xb) = -0.8502

F(12,9571) = 253.41
 Prob > F = 0.0000

mathz		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	

mathz_1		-.4139616	.0081244	-50.95	0.000	-.4298872	-.3980361
age		.0713948	.1413035	0.51	0.613	-.20559	.3483795
female		.1176982	.2064662	0.57	0.569	-.2870193	.5224157
lep		.1201043	.0227077	5.29	0.000	.0755924	.1646162
speced		-.1588845	.0523051	-3.04	0.002	-.2614137	-.0563553
immig		.0420339	.0507581	0.83	0.408	-.0574628	.1415305
econdis		-.0212018	.0302013	-0.70	0.483	-.0804028	.0379992
black		.5786446	.5839093	0.99	0.322	-.5659413	1.723231
hispanic		.3287884	.5057374	0.65	0.516	-.6625641	1.320141
asian		-.188493	.5058222	-0.37	0.709	-1.180012	.8030258
year							
2006		.0766177	.1414914	0.54	0.588	-.2007355	.3539709
same_race		.0414423	.0128886	3.22	0.001	.0161779	.0667067
_cons		-1.042015	1.524875	-0.68	0.494	-4.031092	1.947062

sigma_u		1.2695828					
sigma_e		.46158338					
rho		.88324881	(fraction of variance due to u i)				

 F test that all u_i=0: F(37021, 9571) = 2.29

Prob > F = 0.0000

- (e) Are there any explanatory variables that are dropped in the models in (d)? Are there any explanatory variables that should be dropped that weren't? What does the latter indicate to you? **(2 points)**

There are no explanatory variables dropped in the above models. One would expect time-invariant variables such as gender and student race/ethnicity to fall out of the regression, but they appear not to have done so in this case. This suggests there is unexpected variation in these variables, perhaps due to miscoding or other errors.

- (f) Finally, use the command `xttrans` to describe the frequency of changes in exposure to a same-race teacher over time. Interpret the results of this command. **(2 points)**

Results below. The panel used in the above regressions is unbalanced—some students are observed in two years, but many are only observed in one. Identification of the *same_race* coefficient only comes from students observed in more than one year, who experience a change in *same_race*. The `xttrans` output only pertains to the students observed in more than one year.

Note the row percentages of the `xttrans` output sum to 100, and cell frequencies sum to 9,728, the total number of students observed in both periods. Of the 4,336 students who do *not* have a same race teacher in year 1, 78% again do not have a same race teacher in year 2. 22% do. Of the 5,392 students who *do* have a same race teacher in year 1, 67% continue to do so in year 2. 33% do not. Taken together, only 973+1,795 of the students experienced a switch in the *same_race* variable, or about 28% of all students. If you were concerned that these students represent an unusual population, you could look descriptively at these students and contrast them with students that did not experience such a change. For example, are they more likely to live in urban areas? Did they change schools or districts?

```
. egen count=count(id),by(id)
. table year if count==2
```

```
-----
year      |
(spring)  |      Freq.
-----+-----
      2005 |      9,728
      2006 |      9,728
-----
```

```
. tabulate same_race if year==2005 & count==2
```

same_race	Freq.	Percent	Cum.
0	4,336	44.57	44.57
1	5,392	55.43	100.00
Total	9,728	100.00	

```
. xttrans same_race, freq
```

same_race	same_race		Total
	0	1	
0	3,363	973	4,336
	77.56	22.44	100.00
1	1,795	3,597	5,392
	33.29	66.71	100.00
Total	5,158	4,570	9,728
	53.02	46.98	100.00