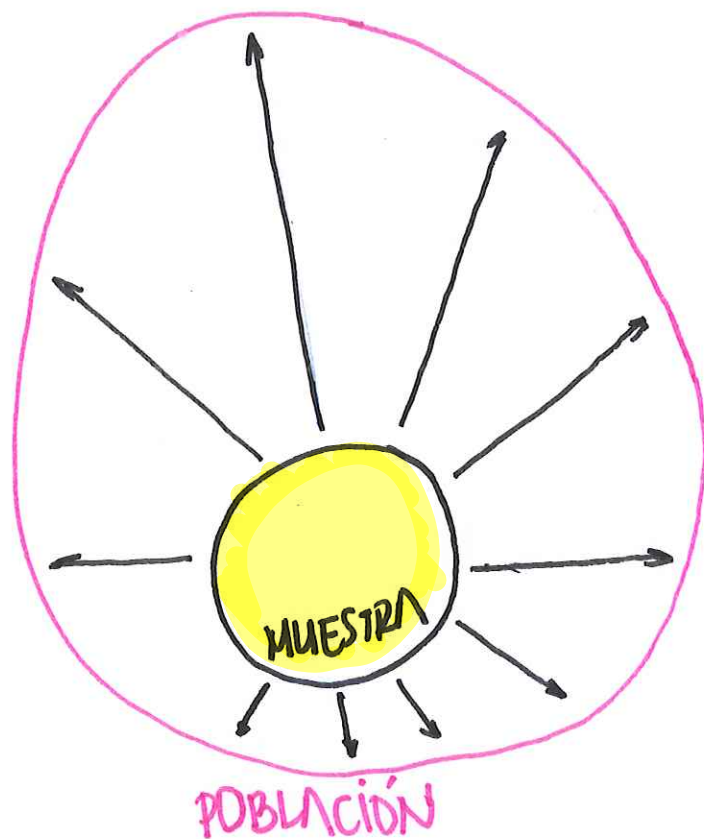


INTERVALOS DE CONFIANZA

Daniel González

INFERENCIA ESTADÍSTICA



Estimación Puntual

$$\hat{\theta}$$

Por intervalos
de confianza

Pruebas de Hipótesis

$$H_0: \theta = \theta_0$$

$$H_a: \theta \neq \theta_0$$

CUANDO NO SE CONOCE EL VALOR DE UN PARÁMETRO, SE UTILIZA LA ESTIMACIÓN PARA ENCONTRAR UN VALOR APROXIMADO A PARTIR DE LOS VALORES DE UNA MUESTRA

CUANDO SE QUIERE VALIDAR UNA AFIRMACIÓN SOBRE UN PARÁMETRO DE UNA POBLACIÓN

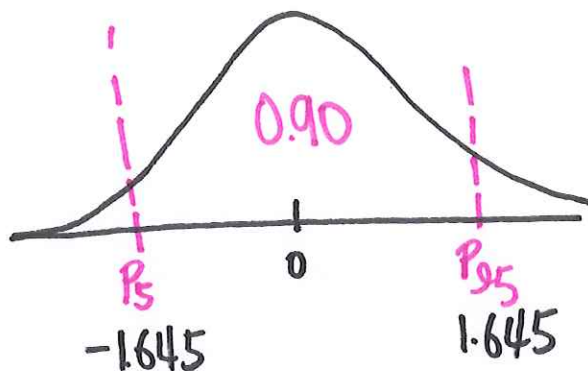
INFERENCIA ESTADÍSTICA
ESTIMACIÓN

ESTIMACIÓN POR INTERVALOS
DE CONFIANZA

(LIC ; LSC)

INTERVALOS DE CONFIANZA

μ :



$$P(-1.645 < Z < 1.645) = 0.90$$

$$P(z_{\alpha/2} < Z < z_{\alpha/2}) = 1 - \alpha$$

$$P(-1.645 < \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} < 1.645) = 0.90$$

$$P(-1.645 \frac{\sigma}{\sqrt{n}} < \bar{X} - \mu < 1.645 \frac{\sigma}{\sqrt{n}}) = 0.90$$

$$P(-1.645 \frac{\sigma}{\sqrt{n}} - \bar{X} < -\mu < 1.645 \frac{\sigma}{\sqrt{n}} - \bar{X}) = 0.90$$

$$P(\bar{X} - 1.645 \frac{\sigma}{\sqrt{n}} < \mu < \bar{X} + 1.645 \frac{\sigma}{\sqrt{n}}) = 0.90$$

$$IC_{\mu: 1-\alpha} \\ \bar{X} \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

INTERVALOS DE CONFIANZA

$$\mu. \quad \bar{x} \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

SUPUESTO

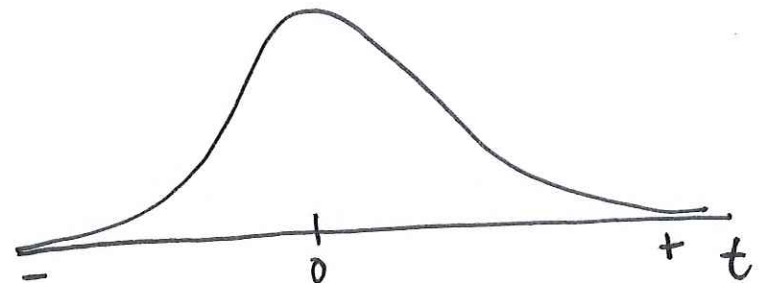
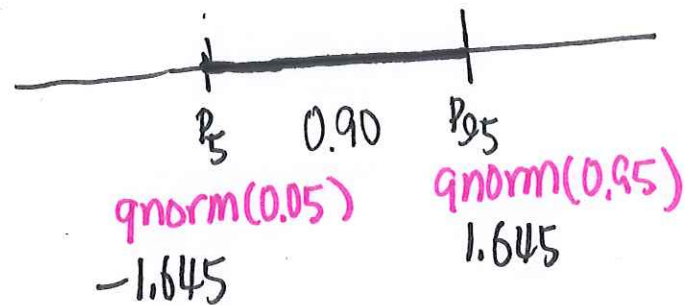
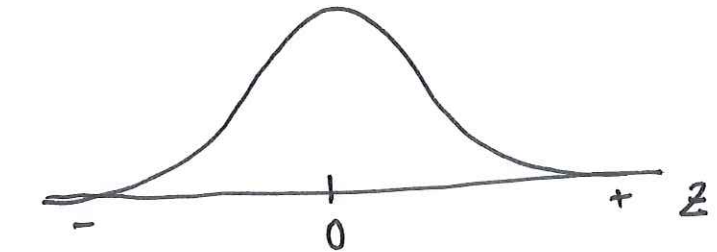
- $X \sim N(\mu, \sigma^2)$
- σ^2 CONOCIDA

$$\bar{x} \pm t_{\alpha/2, \nu=n-1} \frac{s}{\sqrt{n}}$$

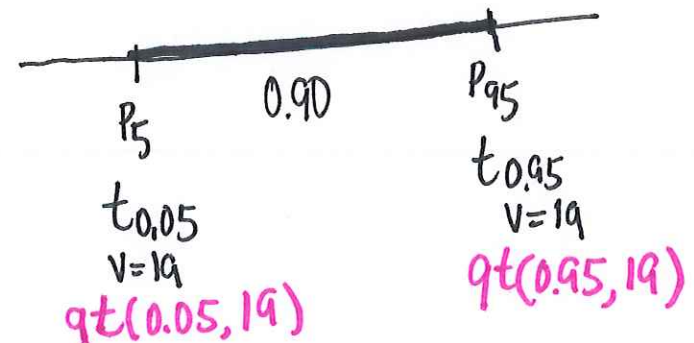
- $X \sim \text{NORMAL}$
- σ^2 DESCONOCIDA

$$\bar{x} \pm z_{\alpha/2} \frac{s}{\sqrt{n}}$$

- $X \sim \text{DESCONOCIDA}$
- $n \gg \text{TCL} \rightarrow \bar{X} \sim \text{NORMAL}$

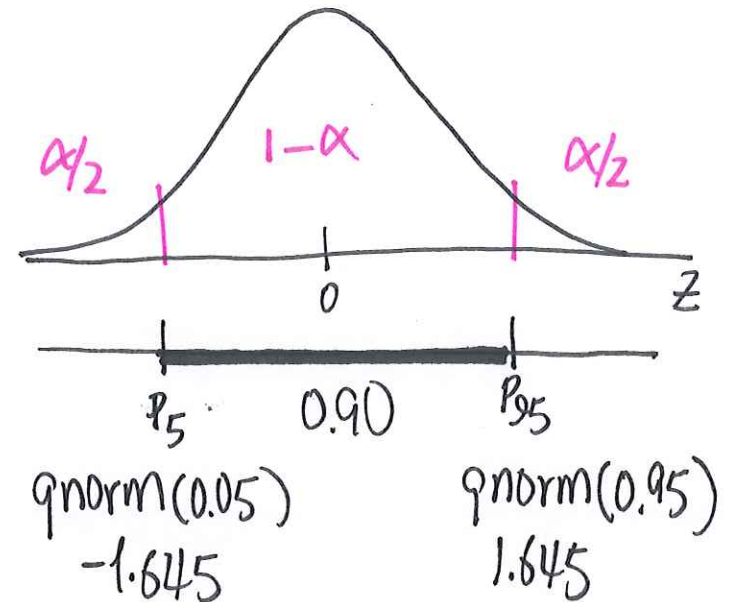


$n=20$

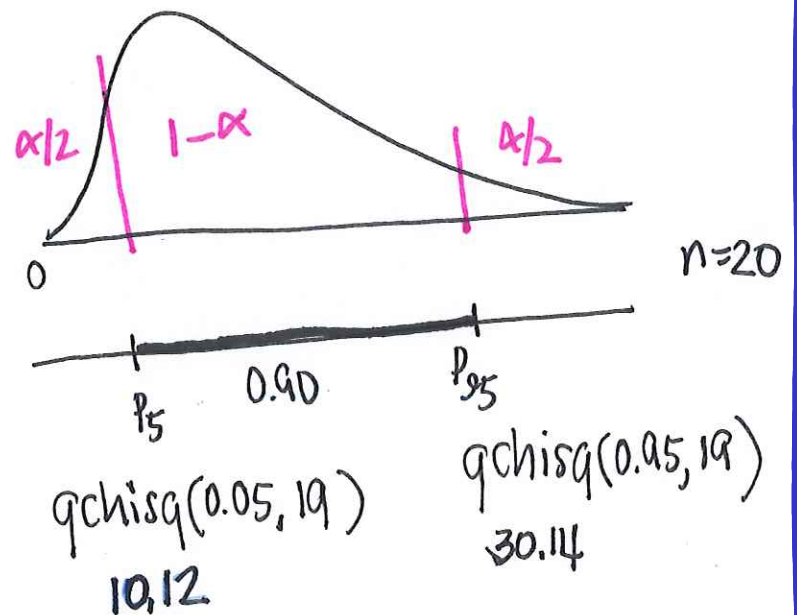


$$P \quad \hat{p} \pm z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

SUPUESTO
 $n \gg$



$$\sigma^2 \left(\frac{(n-1)S^2}{\chi^2_{\alpha/2, \nu=n-1}} ; \frac{(n-1)S^2}{\chi^2_{1-\alpha/2, \nu=n-1}} \right)$$



IC μ .

SUPUESTO₁

SUPUESTO₂

IC

$X \sim N(\mu, \sigma^2)$

σ^2 CONOCIDA : $\bar{X} \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$

σ^2 DESCONOCIDA : $\bar{X} \pm t_{v=n-1} \frac{s}{\sqrt{n}}$

$X \sim ?$

$n \gg$

: TCL $\bar{X} \pm z_{\alpha/2} \frac{s}{\sqrt{n}}$

$n <$

: MÉTODO NO PARAMETRICO
(REMUESTREO)

DIFERENCIA DE MEDIAS

$\mu_1 - \mu_2$ • GRUPOS PAREADOS O EMPAREJADOS

$$X_1 \sim N(\mu_1, \sigma_1^2)$$

$$X_2 \sim N(\mu_2, \sigma_2^2)$$

$$\bar{d} \pm t_{\alpha/2} \frac{S_d}{\sqrt{n}}$$

$$v = n - 1$$

| X_1 | X_2 | $d = X_1 - X_2$ | |
|----------|----------|-----------------|----------------------------------|
| x_{11} | x_{21} | d_1 | } $\bar{d} = \frac{\sum d_i}{n}$ |
| x_{12} | x_{22} | d_2 | |
| \vdots | \vdots | \vdots | |
| \vdots | \vdots | \vdots | |
| x_{1n} | x_{2n} | d_n | |
| | | | S_d |

- GRUPOS INDEPENDIENTES
- SUPUESTOS: NUNCA $\sigma_1^2 = \sigma_2^2$

$$(\bar{x}_1 - \bar{x}_2) \pm t_{\alpha/2} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

$$v = n_1 + n_2 - 2$$

$$\text{donde } S_p^2 = \frac{(n_1 - 1) S_1^2 + (n_2 - 1) S_2^2}{n_1 + n_2 - 2}$$

- $\sigma_1^2 \neq \sigma_2^2$

$$(\bar{x}_1 - \bar{x}_2) \pm t_{\alpha/2} \sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}$$

$$v^*$$

$$v^* = \frac{\left(\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2} \right)^2}{\frac{(S_1^2/n_1)^2}{n_1 - 1} + \frac{(S_2^2/n_2)^2}{n_2 - 1}}$$

NOTA:

(-, -)
 $\mu_1 < \mu_2$

(-, +)
 $\mu_1 = \mu_2$

(+, +)
 $\mu_1 > \mu_2$

COMPARACIÓN DE PROPORCIONES

$P_1 - P_2$

$$(\hat{P}_1 - \hat{P}_2) \pm Z_{\alpha/2} \sqrt{\frac{P_1(1-P_1)}{n_1} + \frac{P_2(1-P_2)}{n_2}}$$

NOTA:

$(-, -): P_1 < P_2$

$(-, +): P_1 = P_2$

$(+, +): P_1 > P_2$

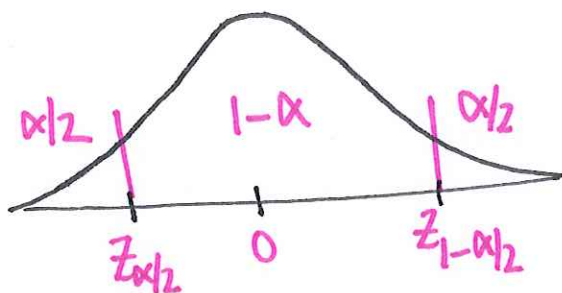
RAZÓN DE VARIANZAS

$$\frac{\sigma_1^2}{\sigma_2^2} \left(\frac{S_1^2/n_1-1}{S_2^2/n_2-1} f_{\alpha/2, v_1, v_2} ; \frac{S_1^2/n_1-1}{S_2^2/n_2-1} f_{1-\alpha/2, v_1, v_2} \right)$$

;

TAMANO DE MUESTRA

- ESTIMACIÓN DE μ .



$$P(z_{\alpha/2} \leq Z \leq z_{1-\alpha/2}) = 1-\alpha$$

$$P\left(\bar{X} - z_{\alpha/2} \frac{\sigma}{\sqrt{n}} \leq \mu \leq \bar{X} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}}\right) = 1-\alpha$$

$$\mu = \bar{X} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

DESPEJAMOS n

$$n = \frac{z_{\alpha/2}^2 \times \sigma^2}{e^2}$$

$$e = |\mu - \bar{x}| < e$$

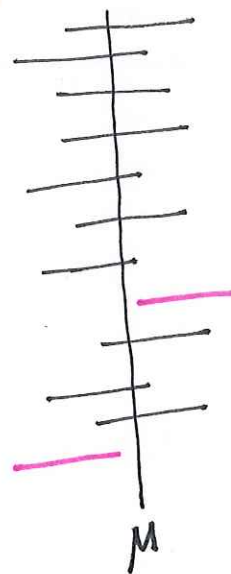
$$n = \frac{z_{\alpha/2}^2 \sigma^2}{e^2}$$

CONFIANZA (1)

VARIANZA (2)

ERROR DE MUESTREO (3)

(1) CONFIANZA



(2) VARIANZA

- PRUEBA PILOTO
 - ESTUDIO PREVIO
 - EXPERTO
- $$\sigma \approx \frac{\text{MAX} - \text{MÍN}}{4}$$

(2) ERROR DE MUESTREO

(1) y (2) A CARGO DEL INVESTIGADOR

TAMAÑO DE MUESTRA

- ESTIMACIÓN DE P

$$n = \frac{Z_{\alpha/2}^2 pq}{e^2}$$

CONFIANZA (1)
 VARIANZA (2)
 ERROR DE MUESTREO (3)

| (1) CONFIANZA | $Z_{\alpha/2}$ |
|---------------|----------------|
| 90% | 1.645 |
| 95% | 1.96 |
| 99% | 2.576 |

ERROR DE MUESTREO

$$|p - \hat{p}| < e$$

VARIANZA

- PRUEBA PILOTO,
- VARIANZA MÁXIMA
- EXPERTO

| | p | q | pq |
|----------------|-----|-----|------|
| | 0.1 | 0.9 | 0.09 |
| | 0.2 | 0.8 | 0.16 |
| | 0.3 | 0.7 | 0.21 |
| | 0.4 | 0.6 | 0.24 |
| VARIANZA MÁX → | 0.5 | 0.5 | 0.25 |
| | 0.6 | 0.4 | 0.24 |
| | 0.7 | 0.3 | 0.21 |
| | 0.8 | 0.2 | 0.16 |
| | 0.9 | 0.1 | 0.09 |

TAMAÑO DE MUESTRA

μ .

$$n = \frac{Z_{\alpha/2}^2 \sigma^2}{e^2}$$

CONFIANZA (1)

VARIANZA (2)

• PRUEBA PILOTO

ERROR DE MUESTREO (3)

P.

$$n = \frac{Z_{\alpha/2}^2 pq}{e^2}$$

CONFIANZA (1)

VARIANZA (2)

• PRUEBA PILOTO

• VARIANZA MAX

ERROR DE MUESTREO (3)



(1) y (3)

A CARGO DEL
INVESTIGADOR

SI $\frac{n}{N} > 0,05$

SE DEBE CORREGIR
EL TAMAÑO DE
MUESTRA POR
POBLACION FINITA

$$n = \frac{n_0 N}{n_0 + N - 1}$$

INTERVALOS DE CONFIANZA NO PARAMÉTRICOS

CUANDO $n < 30$

$X \sim ?$ (NO NORMAL)

MÉTODO NO PARAMÉTRICO

- MUESTRA: X_1, X_2, \dots, X_n
- REMUESTREO
MUESTREO ALATORIO CON REPETICIÓN
- SE RECONSTRUYE POBLACIÓN SIMULANDO
UNA GRAN CANTIDAD DE VALORES
DEL ESTIMADOR
- SE CALCULAN LOS PERCENTILES

MÉTODO 1: $(P_{\alpha/2}; P_{1-\alpha/2})$

MÉTODO 2:

$(2\bar{X} - P_{1-\alpha/2}; 2\bar{X} - P_{\alpha/2})$

NAVIDI
(2006)

INTERVALOS DE CONFIANZA

$$IC_{\mu}: \bar{X} \pm z_{\alpha/2} \sigma / \sqrt{n} \quad (1)$$

$$\bar{X} \pm t_{\alpha/2} \quad s / \sqrt{n} \quad (2)$$

$v = n - 1$

$$\bar{X} \pm z_{\alpha/2} s / \sqrt{n} \quad (3)$$

$$IC_p: \hat{p} \pm z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \quad (4)$$

$$IC_{\sigma^2}: \left(\frac{(n-1)S^2}{\chi_{1-\alpha/2}} ; \frac{(n-1)S^2}{\chi_{\alpha/2}} \right) \quad (5)$$

$$IC_{\mu_1 - \mu_2}: \bar{d} \pm t_{\alpha/2} \quad s_d / \sqrt{n} \quad (6)$$

$v = n - 1$

$$(\bar{x}_1 - \bar{x}_2) \pm t_{\alpha/2} \quad s_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}} \quad (7)$$

$v = n_1 + n_2 - 2$

$$(\bar{x}_1 - \bar{x}_2) \pm t_{\alpha/2} \quad \sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}} \quad (8)$$

v^0

$$IC_{p_1 - p_2}:$$

$$(\hat{p}_1 - \hat{p}_2) \pm z_{\alpha/2} \sqrt{\frac{\hat{p}_1(1-\hat{p}_1)}{n_1} + \frac{\hat{p}_2(1-\hat{p}_2)}{n_2}} \quad (9)$$

$$IC_{\sigma_1^2 / \sigma_2^2}: \left(\frac{S_1^2}{S_2^2} \cdot \frac{1}{f_{1-\alpha/2}} ; \frac{S_1^2}{S_2^2} \cdot \frac{1}{f_{\alpha/2}} \right) \quad (10)$$

$\frac{v_1}{v_2}$