

Homework - 4

Q.1

$$X_i = \begin{cases} 1 & \text{if } H \\ 0 & \text{if } T \end{cases}$$

N denotes flip until we get 100 H
 \therefore Given $y = x_1 + x_2 + \dots + x_N$

Since N and x_i are dependent, we can clearly see, $\therefore Y$ is not a random sum of random variable.

For any $N = n$, we know we will have 100 H.

$\therefore Y = x_1 + x_2 + \dots + x_N = 100$ always.

\therefore pmf of $Y =$

$$P_Y(y) = \begin{cases} 1 & y = 100 \\ 0 & \text{otherwise} \end{cases}$$

Q.2

Given:

$$P[V] = \frac{3}{4}$$

$$P[D] = \frac{1}{4}$$

a) $E[K_{100}]$

K_n is number of video packets in a collection of n packets.

\therefore it follows binomial distribution

$$\begin{aligned} \therefore E[K_{100}] &= 100 E[K] \\ &= 100 \left[\frac{3}{4} \times 1 + \frac{1}{4} \times 0 \right] \end{aligned}$$

$$E[K_{100}] = 75$$

$$b) \sigma_{K_{100}} = \left[E[K_{100}^2] - (E[K_{100}])^2 \right]^{1/2}$$

for a binomial distribution.

$$\sigma = \sqrt{npq} = \sqrt{100 \times \frac{3}{4} \times \frac{1}{4}}$$

$$\sigma_{K_{100}} = 4.33$$

c) Central limit theorem to estimate

$$P[K_{100} \geq 100] = 1 - P[K_{100} \leq 100]$$

$$= 1 - \Phi\left(\frac{100 - 75}{4.33}\right) = 1 - \Phi(-13.1639)$$

$$= 1 - (1 - \Phi(13.1639))$$

$$\therefore P[K_{100} \geq 100] = \Phi(13.1639)$$

$$d) P[16 \leq K_{100} \leq 24]$$

$$\frac{24 + 0.5 - 75}{4.33}$$

$$P[16 \leq K_{100} \leq 24] = \Phi\left(\frac{24 + 0.5 - 75}{4.33}\right) - \Phi\left(\frac{16 + 0.5 - 75}{4.33}\right)$$

$$= \Phi(-11.662) - \Phi(-13.510) \approx 0$$

$$= 0$$

∴ This was using De-Moivre-Laplace approximation.

Q.3 $Y_n = X_{2n-1} + (-X_{2n})$ (a)

a) $E[Y_n] = E[X_{2n-1} - X_{2n}]$

$$= E[X_{2n-1}] - E[X_{2n}]$$

$$= E[X] - E[X] = 0$$

{ This is according to theorem that expected value of sum equals sum of expected values.

$$E[W_n] = E[X_1] + E[X_2] + \dots + E[X_n]$$

$$\text{where } W_n = X_1 + \dots + X_n$$

$$\text{Var}[Y_n]$$

according to theorem, variance of sum = sum of variances for uncorrelated X_1, \dots, X_n .

$$\text{Var}[Y_n] = \text{Var}[X_{2n-1}] + \text{Var}[-X_{2n}]$$

$$= 2 \text{Var}[X]$$

b) $M_n(Y)$

$$E[M_n(Y)] = E[Y]$$

according to sample mean theorem.

$$E[M_n(Y)] = E[Y_n] = 0$$

$$\text{Var}[M_n(Y)] = \frac{\text{Var}[Y_n]}{n} = \frac{2 \text{Var}[X]}{n}$$

Q 4 Chebyshev's inequality states \rightarrow work of 2.9

$$P[|x - E[x]| > c] \leq \frac{\sigma^2}{c^2}$$

we substitute $c = k\sigma$

$$\Rightarrow P[|x - E[x]| \geq k\sigma] \leq \frac{1}{k^2}$$

Actual probability that Gaussian variable Y is more than k standard deviation from expected value is

$$\begin{aligned} P[|Y - E[Y]| \geq k\sigma_y] &= P[Y - E[Y] \leq -k\sigma_y] + P[Y - E[Y] \geq k\sigma_y] \\ &= 2P\left[\frac{Y - E[Y]}{\sigma_y} \geq k\right] \\ &= 2Q(k) \end{aligned}$$

Comparison of result of upper bound

	$k=1$	$k=2$	$k=3$	$k=4$	$k=5$
Chebyshev's bound	1	0.25	0.11	0.0625	0.04
$2Q(k)$	0.317	0.054	0.0044	6.3×10^{-5}	5.7×10^{-7}

Chebyshev bound gets weaker for higher 'k' values.

Q.5 To show, $M_n(X) = X_1 + \dots + X_n$

Satisfies \Rightarrow

$$P[M_n(X) \geq c] \leq \left(\min_{s \geq 0} e^{-sc} \phi_X(s) \right)^n$$

Let $W_n = X_2 + \dots + X_n$

$$\therefore M_n(X) = W_n$$

$$\therefore P[M_n(X) \geq c] = P[W_n \geq nc]$$

$$\phi_{W_n}(s) = (\phi_X(s))^n$$

applying Chebyshev bound to $W_n \Rightarrow$

$$P[W_n \geq nc] \leq \min_{s \geq 0} e^{-snc} \phi_{W_n}(s)$$

$$P[W_n \geq nc] = \min_{s \geq 0} (e^{-sc} \phi_X(s))^n$$

For $y \geq 0$, y^n is increasing

\therefore Value of s that minimizes $e^{-sc} \phi_X(s)$
also minimizes $(e^{-sc} \phi_X(s))^n$

$$P[M_n(X) \geq c] = P[W_n \geq nc] \leq \left(\min_{s \geq 0} e^{-sc} \phi_X(s) \right)^n$$

Q. 6

$$P[A] = 0.8$$

Since X_A is bernoulli $(1, 0)$

$$E[X_A] = 1 \cdot P[A] + 0 \cdot (1 - P[A])$$

$$E[X_A] = 0.8$$

$$\text{Var}[X_A] = P[A](1 - P[A]) = 0.16$$

$$b) \text{Var}[\hat{P}_n(A)]$$

Let $X_{A,i}$ denote X_A for i th trial

We know, $E[\sum_{i=1}^n X_{A,i}] = n P[A]$

$$\hat{P}_n(A) = M_n(X_A) = \frac{1}{n} \sum_{i=1}^n X_{A,i}$$

$$\therefore \text{Var}[\hat{P}_n(A)] = \frac{1}{n^2} \sum_{i=1}^n \text{Var}[X_{A,i}]$$

$$= \frac{1}{n^2} \cdot n P[A](1 - P[A])$$

$$= \frac{0.16}{n}$$

$$= \frac{0.16}{n} \quad \text{after substitution}$$

$$= \frac{0.16}{100}$$

$$a) P[|\hat{P}_n(A) - P[A]| \leq 0.12] \geq 1 - \alpha$$

$$\hat{P}_{100}(A) = M_{100}(X_A)$$

we can write

$$P[|\hat{P}_{100}(A) - P[A]| < c] \geq 1 - \frac{\text{Var}(X_A)}{100 c^2}$$

$$= 1 - \frac{0.16}{100 c^2} = 1 - \alpha$$

for $C=0.1$

$$d = \frac{0.16}{100(0.1)^2} = 0.16$$

Thus for 100 samples, confidence coefficient

$$\Rightarrow 1 - \alpha = 0.84$$

d) Find n , such that

$$P[|\hat{p}_n(A) - P(A)| \leq 0.1] \geq 0.95$$

$$P[|\hat{p}_n(A) - P(A)| < c] \geq 1 - \frac{\text{var}(X_n)}{n c^2}$$

$$1 - \frac{0.16}{n c^2} \geq 1 - \alpha$$

$$\frac{0.16}{n c^2} \leq \alpha$$

\therefore for $(\alpha=0.05)$ we have confidence coefficient $1 - \alpha = 0.95$

$$\text{if } \alpha = \frac{0.16}{n(0.1)^2} = 0.05$$

$$\text{or } n \geq \frac{0.16}{0.05(0.1)^2} = 320$$

$$(nX)_{0.05/2} = (n)_{0.05/2}$$

$$(nX)_{0.05/2} = [n]_{0.05/2}$$

$$n = \frac{0.16}{0.05(0.1)^2} = 320$$