

Homework 3

1) We know that

$$S_{xy} = \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})$$

To show S_{xy} can be also written as

$$S_{xy} = \sum_{i=1}^n (x_i - \bar{x}) y_i$$

Proof:

$$S_{xy} = \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})$$

Expand product inside the summation

$$S_{xy} = \sum_{i=1}^n \underbrace{(x_i - \bar{x})}_{(1)} \underbrace{y_i - \bar{y}}_{(2)} = \sum_{i=1}^n \underbrace{(x_i - \bar{x})}_{(1)} \underbrace{y_i}_{(3)} - \sum_{i=1}^n \underbrace{(x_i - \bar{x})}_{(1)} \underbrace{\bar{y}}_{(2)}$$

We know \bar{y} is constant, so factoring out of summation

~~$S_{xy} = \sum_{i=1}^n (x_i - \bar{x}) y_i - \sum_{i=1}^n (x_i - \bar{x}) \bar{y}$~~

Part (2) will be $\bar{y} \sum_{i=1}^n (x_i - \bar{x})$

(3)

lets expand part (3)

We know $\sum_{i=1}^n x_i = n\bar{x}$

$$\sum_{i=1}^n x_i - \sum_{i=1}^n \bar{x} = n\bar{x} - n\bar{x} = 0$$

$$\Rightarrow \sum_{i=1}^n (x_i - \bar{x}) = 0 \Rightarrow \bar{y} \times 0 = 0$$

\Rightarrow Part ② is 0

$$\therefore S_{xy} = \sum_{i=1}^n (x_i - \bar{x}) y_i \quad \text{--- eq ①}$$

We know,

$$\hat{B}_1 = \frac{S_{xy}}{S_{xx}}, \quad \text{and } c_i = \frac{x_i - \bar{x}}{S_{xx}}$$

$$S_{xy} = \left(\sum_{i=1}^n (x_i - \bar{x}) y_i \right) \quad \text{from eq ①}$$

$$\hat{B}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x}) y_i}{S_{xx}}$$

$$\hat{B}_1 = \sum_{i=1}^n c_i y_i$$

5A

$$y = \beta_0 + \beta_1 x + \epsilon \quad \text{[True Relationship]}$$

for a collected dataset, (x_i, y_i)
the relationship can be expressed as

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

$$\sum_{i=1}^n y_i = \sum_{i=1}^n \beta_0 + \sum_{i=1}^n \beta_1 x_i + \sum_{i=1}^n \epsilon_i$$

(a) let's try to prove $E(\hat{\beta}_0) = \beta_0$,

To problem ① we pm

We know that $\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$.

$$E[\hat{B}] = \underbrace{E[Y]}_{\text{Part (1)}} - \underbrace{\lambda E[B_1]}_{\text{Part (2)}} \rightarrow \text{eqn (1)}$$

Expand Part ① :-

finding expectation of θ

$$\text{Ex. 1. } y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

$$\bar{y} = \frac{1}{n} \sum_{i=1}^n y_i = \frac{\cancel{\alpha} \beta_0}{\cancel{\alpha}} + \frac{\cancel{\alpha} \beta_1}{\cancel{\alpha}} \times \bar{x} + \frac{1}{n} \sum_{i=1}^n \epsilon_i$$

$$\bar{y} = \beta_0 + \beta_1 \bar{X} + \frac{1}{n} \sum_{i=1}^n \epsilon_i$$

lets compute ~~consider~~

Expectation along Y.

$$E[\bar{y}] = E[\beta_0] + E[\beta_1 \bar{x}] + E\left[\frac{1}{n} \sum_{i=1}^n \epsilon_i\right]$$

lets consider $E[\epsilon] = 0$.

Then $E(\hat{y}) = \beta_0 + \beta_1 \bar{x} \rightarrow \text{Equation (1)}$

Expand part (2)

$$\hat{\beta}_1 = \frac{S_{xy}}{S_{xx}} = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

\Rightarrow We proved in problem (1) that $S_{xy} = \sum_{i=1}^n (x_i - \bar{x}) y_i$

$$\Rightarrow \hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x}) (y_i)}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x}) (\beta_0 + \beta_1 x_i + \epsilon_i)}{S_{xx}}$$

$$= \frac{\sum_{i=1}^n (x_i - \bar{x}) \cdot \beta_0}{S_{xx}} + \frac{\sum_{i=1}^n (x_i - \bar{x}) \beta_1 x_i}{S_{xx}} + \frac{\sum_{i=1}^n (x_i - \bar{x}) \epsilon_i}{S_{xx}}$$

$$= \underbrace{0}_{\text{Equation (2)}} + \beta_1 + \epsilon \left[\frac{\sum_{i=1}^n (x_i - \bar{x}) \epsilon_i}{S_{xx}} \right]$$

Since we considered $E(\epsilon_i) = 0$

$$\Rightarrow E[\hat{\beta}_1] = \beta_1$$

Now $E(\epsilon_i)$ will be

$$\begin{aligned} E[\hat{\beta}_0] &= E[\bar{y}] - E[\hat{\beta}_1] \bar{x} \\ &= \beta_0 + \beta_1 / \bar{x} - \beta_1 / \bar{x} \\ \underline{E[\hat{\beta}_0] = \beta_0} \quad \text{when} \quad \underline{E(\epsilon_i) = 0} \end{aligned}$$

$E[\hat{\beta}_0]$ $\hat{\beta}_0$ is an unbiased estimator of β_0 when $E(\epsilon_i) = 0$

(b) ~~Var~~ to prove $\text{Var}(\hat{\beta}_1) = \frac{\sigma^2}{S_{xx}}$

We know $\hat{\beta}_1 = \beta_1 + \underbrace{\sum_{i=1}^n (x_i - \bar{x}) \epsilon_i}_{S_{xx}}$

$$\text{Var}(\hat{\beta}_1) = \underbrace{\text{Var}(\beta_1)}_0 + \text{Var}\left(\sum_{i=1}^n (x_i - \bar{x}) \epsilon_i\right)$$

↓
0
Variance of constant is zero.

W.K.T $\text{Var}(ax) = a^2 \text{Var}(x)$

$$\Rightarrow \text{Var}(\hat{\beta}_1) = \left(\frac{1}{S_{xx}} \right) \text{Var} \left(\sum_{i=1}^n \underbrace{(x_i - \bar{x})}_{a} \underbrace{G_i}_{x} \right)$$

\downarrow
 a^2

$$\Rightarrow \text{Var}(\hat{\beta}_1) = \frac{1}{S_{xx}} \sum_{i=1}^n (x_i - \bar{x})^2 \text{Var}(G_i)$$

Let's assume $\text{Var}(G_i) = \sigma^2$.

$$\text{Var}(\hat{\beta}_1) = \frac{\sigma^2}{S_{xx}} \sum_{i=1}^n (x_i - \bar{x})^2$$

$$\left[\text{Var}(\hat{\beta}_1) = \frac{\sigma^2}{S_{xx}} \right] \text{ when } \text{Var}(G_i) = \sigma^2$$

(c) We know $\hat{\beta}_1 = \beta_1 + \sum_{i=1}^n \frac{(x_i - \bar{x})}{S_{xx}} G_i$

Let's assume $G_i \sim N(0, \sigma^2)$.

Then according to linear property of

Normal distribution, the linear combination $\sum_{i=1}^n (x_i - \bar{x}) G_i$ is also normal distribution.

for $\hat{\beta}_1$, we found $E[\hat{\beta}_1] = \beta_1$,

$$\text{Var}[\hat{\beta}_1] = \frac{\sigma^2}{S_{xx}}$$

$\Rightarrow \therefore$ We have

$$\hat{\beta}_1 \sim N\left(\beta_1, \frac{\sigma^2}{S_{xx}}\right) \text{ when } \epsilon_i \sim N(0, \sigma^2)$$

2A Let's consider $E(\hat{\epsilon}_i) = E(\epsilon_i) = \alpha$.

non zero
constant.

from eq (2) in Question (5) pt (a)

$$\hat{\beta}_1 = \beta_1 + \frac{\sum_{i=1}^n (x_i - \bar{x}) \epsilon_i}{S_{xx}}$$

$$E[\hat{\beta}_1] = E[\beta_1] + E\left[\frac{\sum_{i=1}^n (x_i - \bar{x}) \epsilon_i}{S_{xx}}\right]$$

Constant

$$= \beta_1 + \frac{\sum_{i=1}^n (x_i - \bar{x}) E[\epsilon_i]}{S_{xx}}$$

$$= \beta_1 + \frac{\sum_{i=1}^n (x_i - \bar{x}) \alpha}{S_{xx}} \rightarrow \text{non zero constant,}$$

lets factor it out.

$$= \beta_1 + \left(\alpha \frac{\sum_{i=1}^n (x_i - \bar{x})}{S_{xx}} \right) \rightarrow 0.$$

$E[\hat{\beta}_1] = \beta_1$ $\hat{\beta}_1$ is unbiased even when $E(\epsilon_i) = \alpha \neq 0$.

lets calculate $E[\hat{\beta}_0]$.

We know $\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$

$$\bar{y} = \beta_0 + \beta_1 \bar{x} + \bar{e}$$

$$\hat{\beta}_0 = \underbrace{\beta_0 + \beta_1 \bar{x}}_{\text{constant}} + \bar{e} - \hat{\beta}_1 \bar{x}$$

$$E[\hat{\beta}_0] = \beta_0 + \beta_1 \bar{x} + E[\bar{e}] - E[\hat{\beta}_1 \bar{x}]$$

We know

$$E[\hat{\beta}_1] = \beta_1$$

$$\Rightarrow E[\hat{\beta}_0] = \beta_0 + \beta_1 \bar{x} + \alpha - \beta_1 \bar{x}$$

$$\Rightarrow E[\hat{\beta}_0] = \beta_0 + \underbrace{E[\bar{e}]}_{\text{Part (a)}}$$

• Lets expand Part (a)

$$E[\bar{e}] = E\left[\frac{1}{n} \sum_{i=1}^n e_i\right] = \frac{1}{n} \sum_{i=1}^n E[e_i]$$

$$= \frac{1}{n} \times n \times \alpha$$

$$= \alpha$$

$$\Rightarrow E[\hat{\beta}_0] = \beta_0 + \alpha$$

The expectation of $\hat{\beta}_0$ is biased estimator, which is biased by α .