

## **Distributed Source Location in Wireless Sensor Networks**

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## **Introduction**

What?  
Why?  
How?  
Why only this?

## **Sub-gradient methods**

Nowak's decentralized algo.(ICASSP'04)  
Improvements  
Sequential Bayesian analysis

## **Conclusions/ Future work**

Who leads the race?  
Oversimplified assumptions?  
What else can be done?

## Introduction

What?      Given the measurements of a source/ target  
by a set of sensors,  
obtain/ estimate its location/ position

Why?      Tracking  
Surveillance

How?

- Triangulation
  - L**ateration (from distances)
  - Angulation (from angles/ bearings)
- Scene analysis
  - Image processing
- Proximity analysis
  - Physical contact....

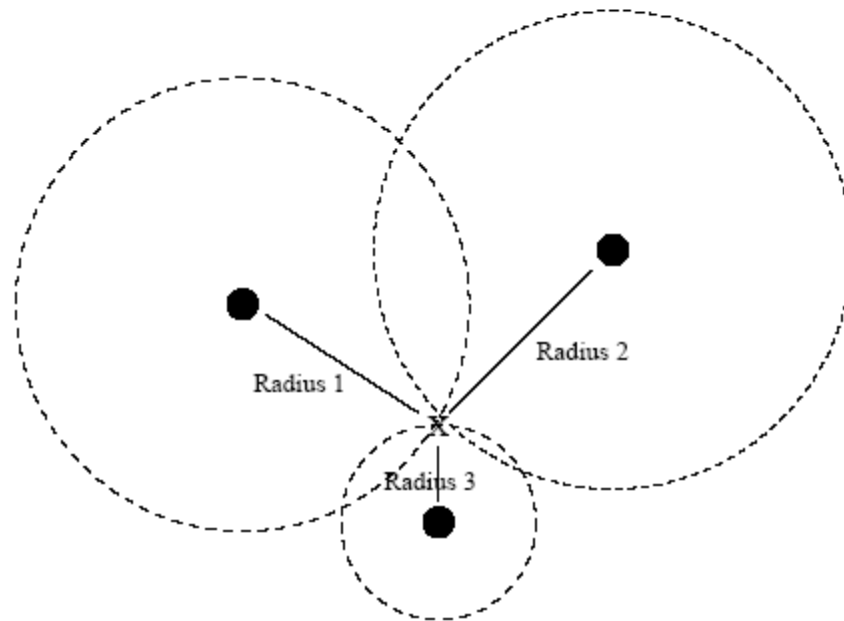
- Centralized
- **Distributed**

## Introduction

Why only this?

**Lateralation**

**Distributed**



## Sub-gradient methods

Model :

$$s_j(i) = \frac{A}{((\theta_1 - x_j)^2 + (\theta_2 - y_j)^2)^{\beta/2}} + e_j(i)$$

$s_j(i)$  is the  $i^{th}$  measurement at the  $j^{th}$  sensor

$A$  is the acoustic energy

$\begin{bmatrix} \theta_1 \\ \theta_2 \end{bmatrix}$  and  $\begin{bmatrix} x_j \\ y_j \end{bmatrix}$  are the locations of the source and the  $j^{th}$  sensor,

$e_j(i)$  are i.i.d. samples of zero-mean Gaussian noise process with variance  $\sigma^2$

$\beta$  is a parameter dependent on the transmission medium

Assumptions:

$$\beta = 2$$

$$\sigma^2$$

$$A$$

are known

$$\begin{bmatrix} x_j \\ y_j \end{bmatrix}$$

## Sub-gradient methods

Objective function to be minimized under least-squared cost is:

centralized

$$f = \frac{1}{MN} \sum_{j=1}^M \sum_{i=1}^N \left[ s_j(i) - \frac{A}{(\theta_1 - x_j)^2 + (\theta_2 - y_j)^2} \right]^2$$

$M$  is the number of sensors

$N$  is the number measurements in one cycle  
(a sensor contributes only once in a cycle)

Sub-gradient method (Nowak et al.'s algo.)

$$\hat{\theta}_{j+1}^k = \hat{\theta}_j^k - \alpha \nabla f_{j+1}(\hat{\theta}_j^k)$$

$$\hat{\theta}_j^k = \begin{bmatrix} \hat{\theta}_1 \\ \hat{\theta}_2 \end{bmatrix} \text{ at the } j^{th} \text{ sensor during the } k^{th} \text{ cycle,}$$

$$f_j = \frac{1}{N} \sum_{i=1}^N \left[ s_j(i) - \frac{A}{(\theta_1 - x_j)^2 + (\theta_2 - y_j)^2} \right]^2,$$

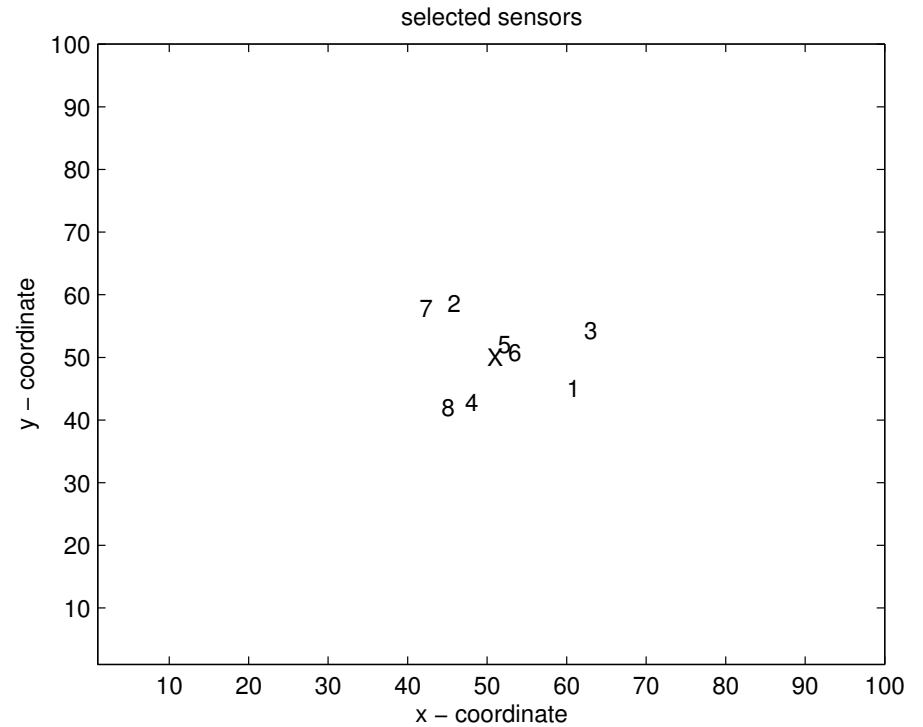
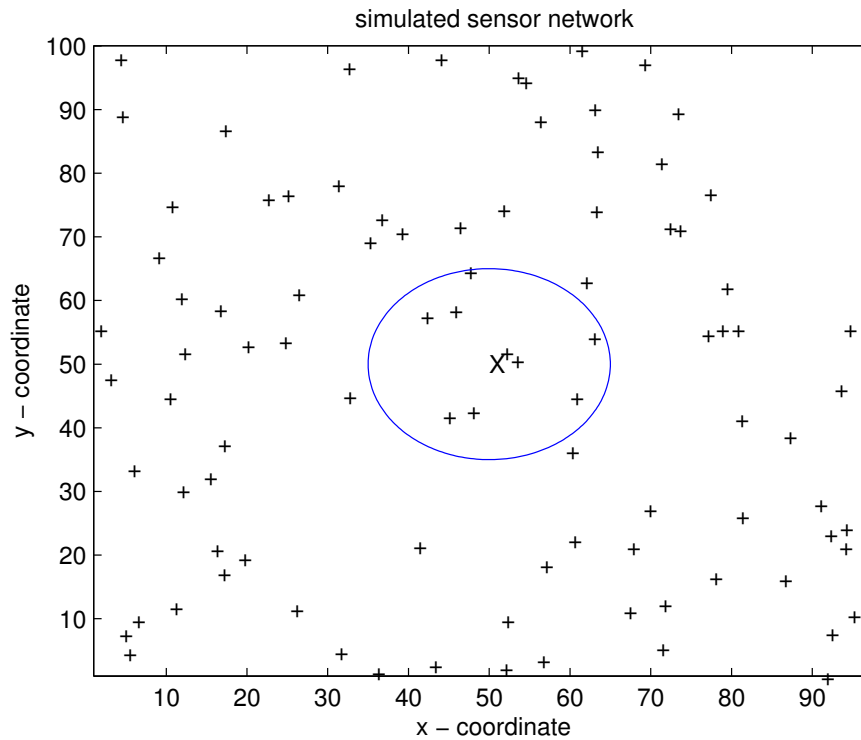
$$\nabla f_j = \frac{4A \sum_{i=1}^N \left[ s_j(i) - \frac{A}{(\theta_1 - x_j)^2 + (\theta_2 - y_j)^2} \right]}{N((\theta_1 - x_j)^2 + (\theta_2 - y_j)^2)} \begin{bmatrix} \theta_1 - x_j \\ \theta_2 - y_j \end{bmatrix};$$

$\alpha$  is the step size in the steepest decent direction.

distributed

## Sub-gradient methods

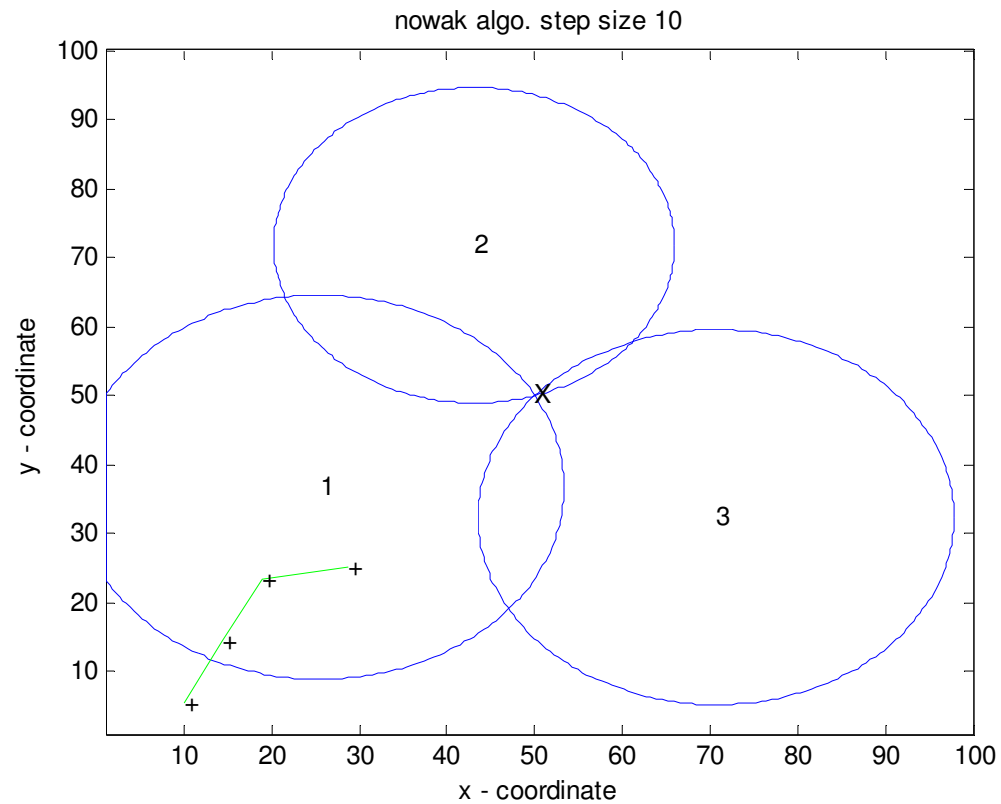
### Set-up:



All the sensor within the radius of 15 units and not less than 1 unit from the source collaborate in estimating the location

## Sub-gradient methods

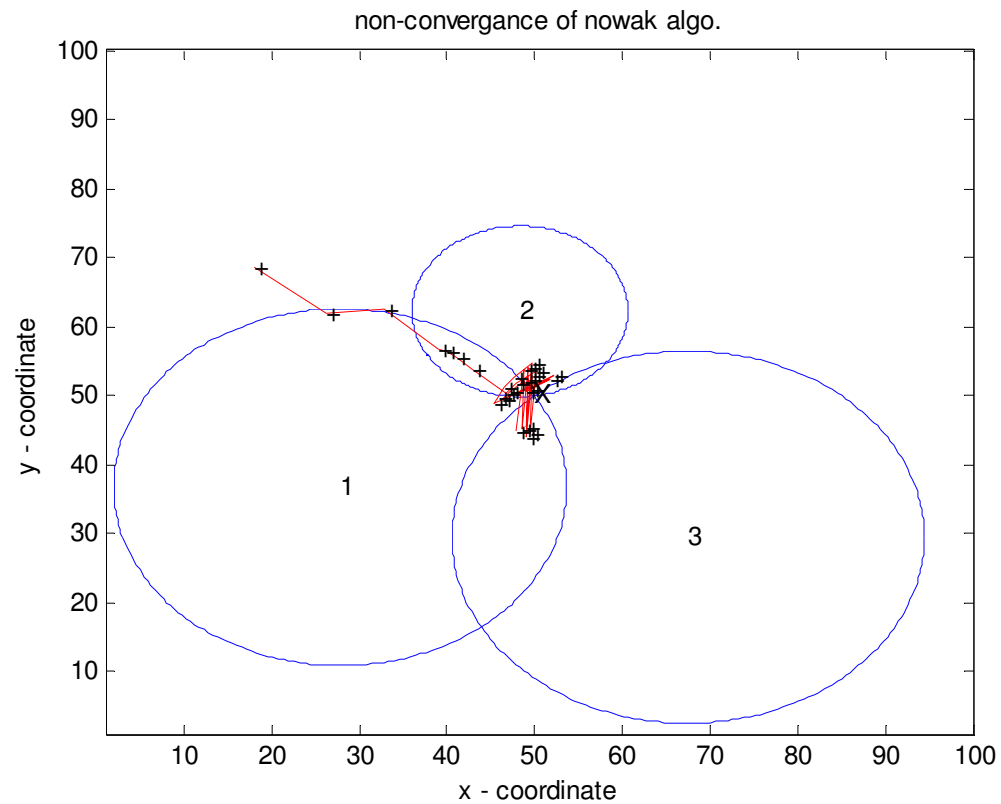
(Nowak et al.'s algo.)





## Sub-gradient methods

(Nowak et al.'s algo.)



## Sub-gradient methods

(Modified Nowak et al.'s algo.)

Rewrite the objective function w.r.t the current sensor:

$$f = \frac{1}{MN} \sum_{j=1}^M \sum_{i=1}^N \left[ s_j(i) - \frac{A}{r_j^2} \right]^2$$

Then, solution to the normal eqn. is,

ML estimate

$$r_j^2 = \frac{A}{\bar{s}_j}$$

$\bar{s}_j$  is the sample average of the measurements  
at the  $j^{th}$  sensor in the  $k^{th}$  cycle

$$\begin{aligned} r_j^k &= \sqrt{\frac{A}{\bar{s}_j^k}} \\ \phi_j^k &= \tan^{-1} \left( \frac{\hat{\theta}_{j2}^{k-1} - y_j}{\hat{\theta}_{j1}^{k-1} - x_j} \right) \\ \hat{\theta}_{j1}^k &= y_j + r_j^k \cos(\phi_j^k) \\ \hat{\theta}_{j2}^k &= x_j + r_j^k \sin(\phi_j^k) \end{aligned}$$

## Sub-gradient methods

(Modified Nowak et al.'s algo.)

$$d_j^{k-1} = \left( (\hat{\theta}_{j1}^{k-1} - x_j)^2 + (\hat{\theta}_{j2}^{k-1} - y_j)^2 \right)^{\frac{1}{2}}$$

$$\begin{bmatrix} \hat{\theta}_{j1}^k \\ \hat{\theta}_{j2}^k \end{bmatrix} = \begin{bmatrix} \hat{\theta}_{j1}^{k-1} \\ \hat{\theta}_{j2}^{k-1} \end{bmatrix} - \frac{d_j^{k-1} - r_j^{k-1}}{d_j^{k-1}} \begin{bmatrix} \hat{\theta}_{j1}^{k-1} - x_j \\ \hat{\theta}_{j2}^{k-1} - y_j \end{bmatrix}$$

$$\begin{bmatrix} \hat{\theta}_{j1}^k \\ \hat{\theta}_{j2}^k \end{bmatrix} = \begin{bmatrix} \hat{\theta}_{j1}^{k-1} \\ \hat{\theta}_{j2}^{k-1} \end{bmatrix} - \frac{d_j^{k-1} - r_j^{k-1}}{d_j^{k-1}} \begin{bmatrix} \hat{\theta}_{j1}^{k-1} - x_j \\ \hat{\theta}_{j2}^{k-1} - y_j \end{bmatrix}$$

Nowak's algo  
(adaptive step-size)

Only quantity being computed  
At the current sensor in the current cycle

$$r_j^k = \sqrt{\frac{A}{\bar{s}_j^k}}$$

Improve the estimate by  
using moving average

$$\bar{s}_j^k = \frac{\bar{s}_j + (k-1)\bar{s}_j^{k-1}}{k}$$

## Sub-gradient methods

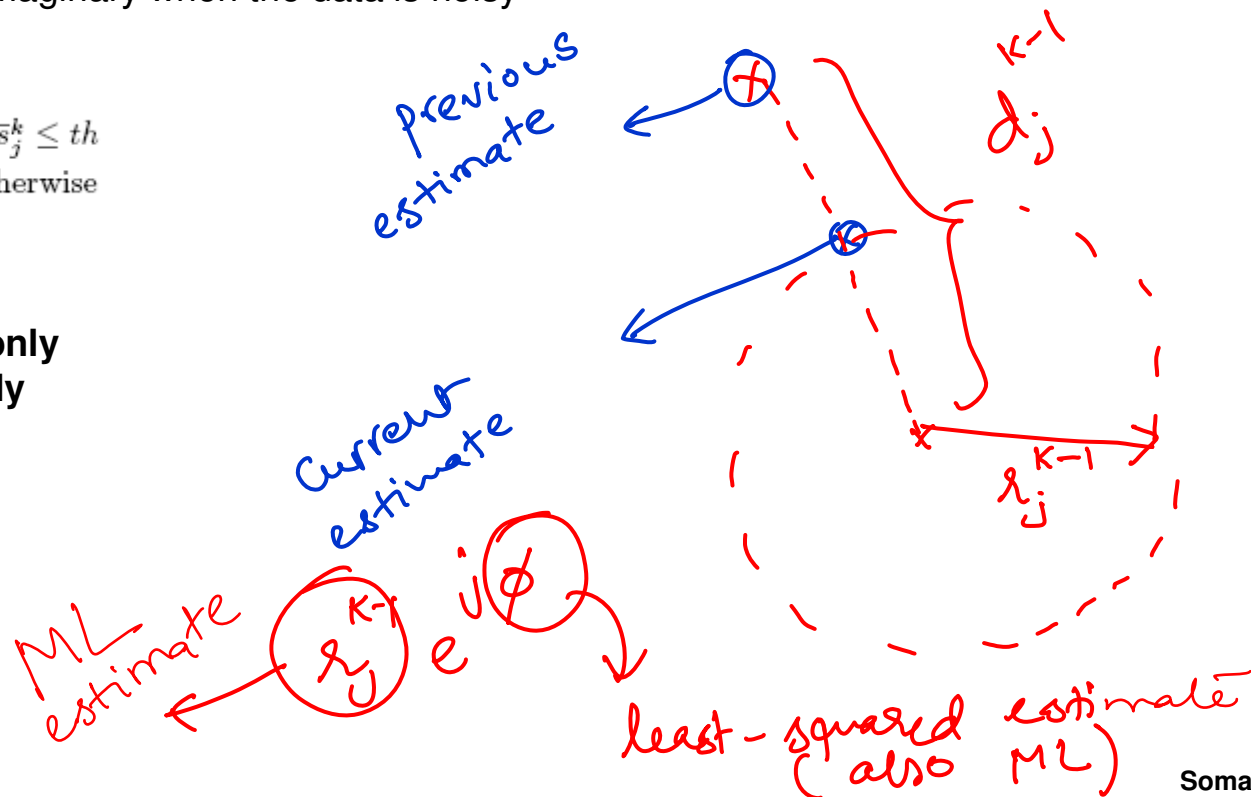
(Modified Nowak et al.'s algo.)

$$\begin{bmatrix} \hat{\theta}_{j1}^k \\ \hat{\theta}_{j2}^k \end{bmatrix} = \begin{bmatrix} \hat{\theta}_{j1}^{k-1} \\ \hat{\theta}_{j2}^{k-1} \end{bmatrix} - \underbrace{\frac{d_j^{k-1} - r_j^{k-1}}{d_j^{k-1}}}_{\text{magnitude}} \underbrace{\begin{bmatrix} \hat{\theta}_{j1}^{k-1} - x_j \\ \hat{\theta}_{j2}^{k-1} - y_j \end{bmatrix}}_{\text{direction vector}}$$

Step-size can be imaginary when the data is noisy

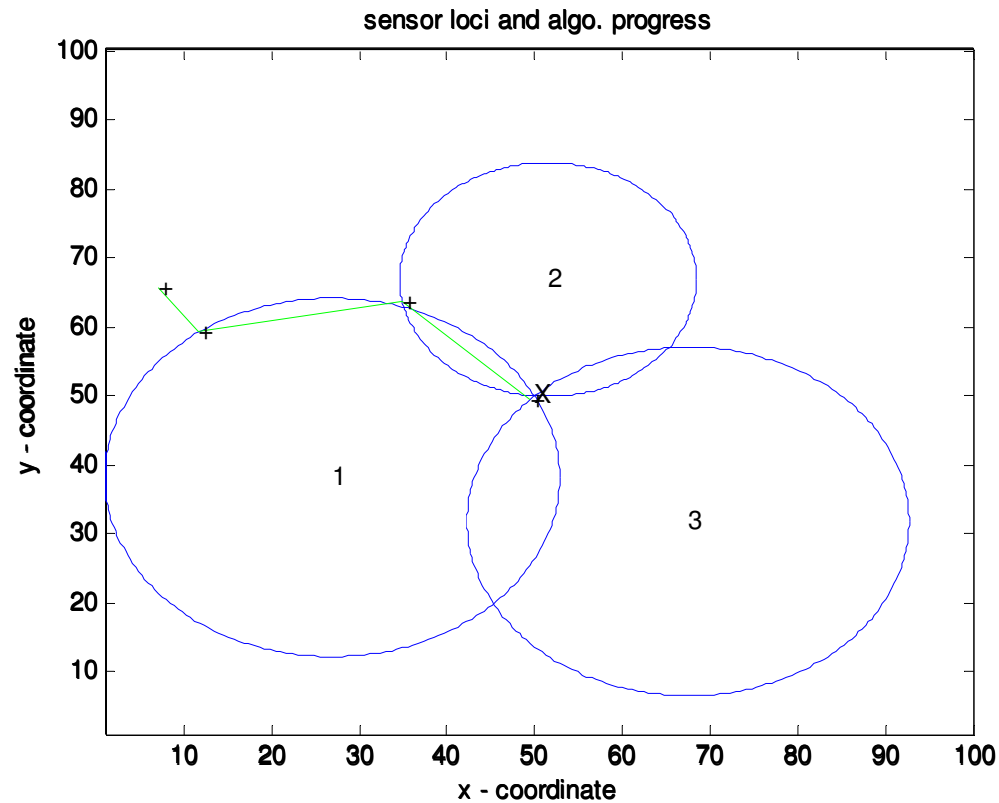
$$r_j^k = \begin{cases} 0 & \text{if } \bar{s}_j^k \leq th \\ \sqrt{\frac{A}{\bar{s}_j^k}} & \text{otherwise} \end{cases}$$

- forward-only
- not greedy



## Sub-gradient methods

(Modified Nowak et al.'s algo.)



## Sequential Bayesian analysis

Posterior density at J-th sensor during kth cycle becomes  
Prior for the (j+1)-th sensor in the same cycle.



$$\Pi_{j+1}^k(\theta_1, \theta_2) = P_j^k(\theta_1, \theta_2)$$

Normal **prior**

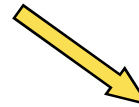
$$\Pi(\theta_1, \theta_2) = \frac{\exp(-\frac{(\theta_2 - \mu_2)^2}{2\sigma_2^2}) \exp(-\frac{(\theta_1 - \mu_1 - \rho \frac{\sigma_1}{\sigma_2}(\theta_2 - \mu_2))^2}{2\sigma_1^2(1 - \rho^2)})}{(2\pi\sigma_2^2)^{0.5} (2\pi\sigma_1^2(1 - \rho^2))^{0.5}}$$



$$\begin{aligned} \mu_1 &= E[\theta_1] \\ \mu_2 &= E[\theta_2] \\ \sigma_1^2 &= E[(\theta_1 - \mu_1)^2] \\ \sigma_2^2 &= E[(\theta_2 - \mu_2)^2] \\ \rho &= \frac{E[(\theta_1 - \mu_1)(\theta_2 - \mu_2)]}{\sigma_1 \sigma_2} \end{aligned}$$

**Likelihood function** of the data

$$L(\theta_1, \theta_2) = \frac{1}{(2\pi\sigma^2)^{0.5}} \exp(-\frac{\sum_{i=1}^N (s_j^i - \frac{A}{(\theta_1 - x_j)^2 + (\theta_2 - y_j)^2})^2}{2\sigma^2})$$



**Posterior density**

$$P(\theta_1, \theta_2) = \frac{L(\theta_1, \theta_2) \Pi(\theta_1, \theta_2)}{\iint L(\theta_1, \theta_2) \Pi(\theta_1, \theta_2) d\theta_1 d\theta_2}$$



## Sequential Bayesian analysis

### Bayesian analysis

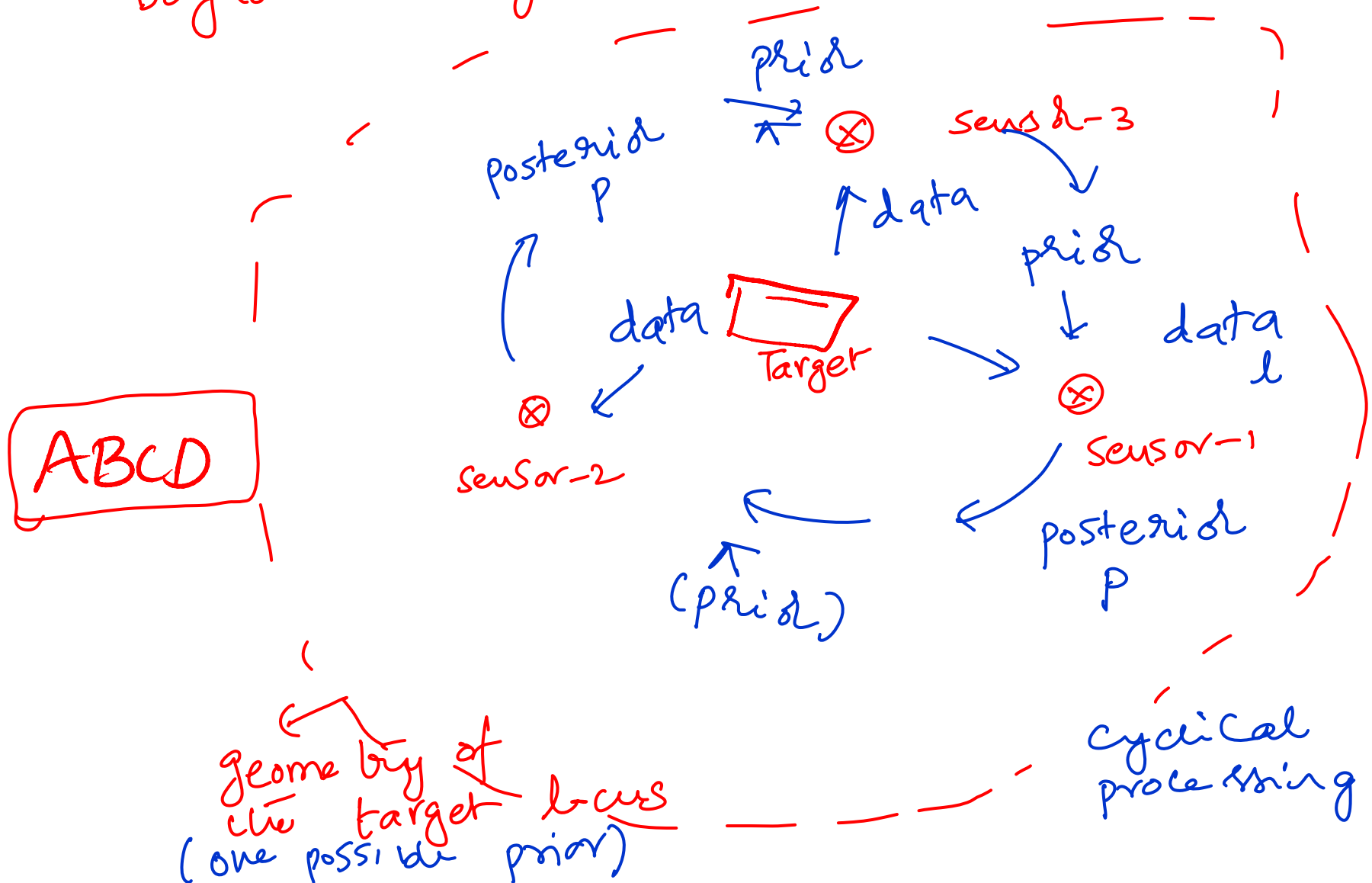
- incorporate subjective knowledge into the problem
- can make inference even when data is insufficient
- suitable for on-line processing (sequential processing)

posterior  $\propto$   
(for inference)

(data)  
likelihood  $\times$   
prior  
(prior knowledge)

## Sequential Bayesian analysis

Bayesian analysis in wireless sensor networks

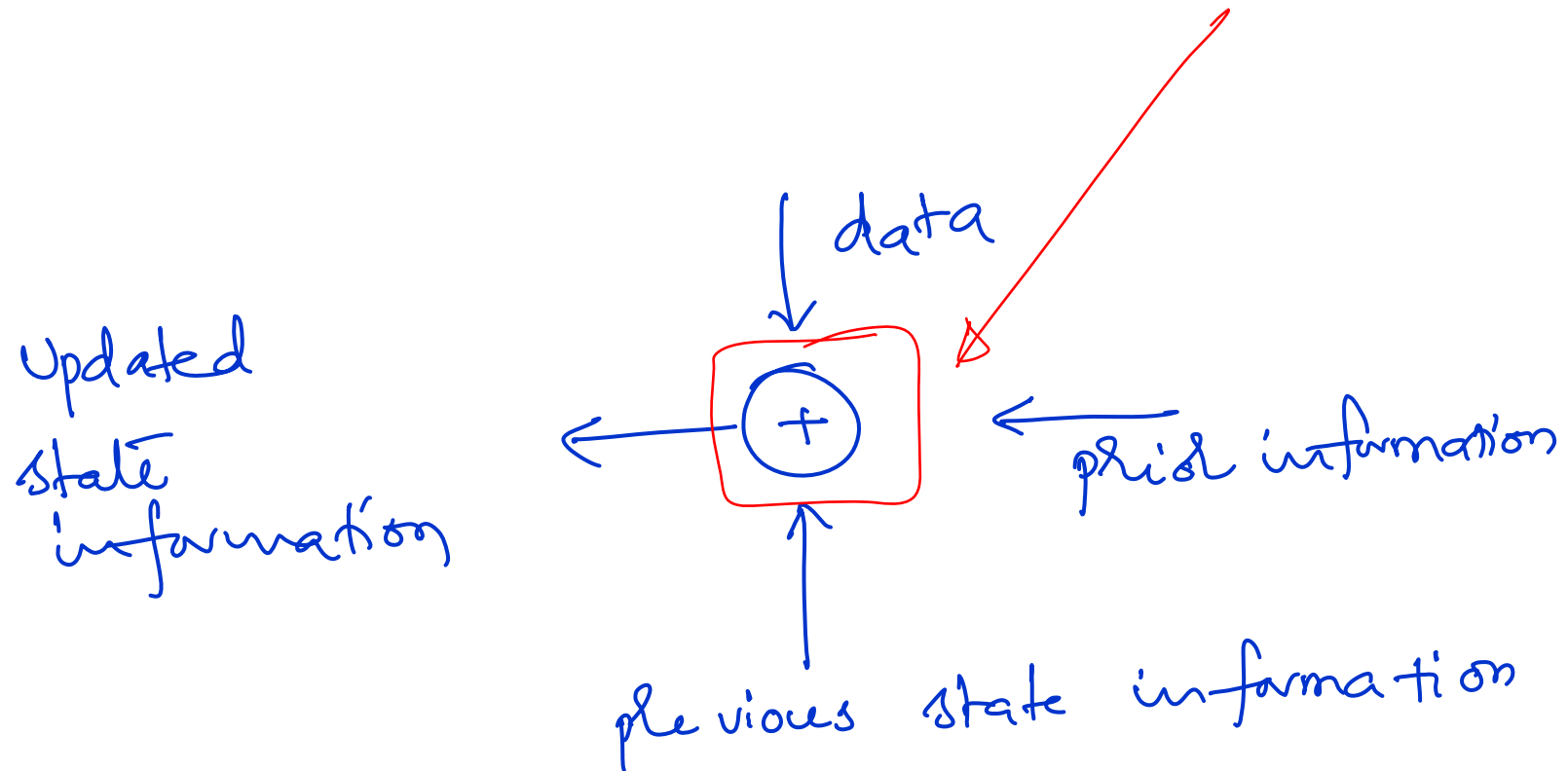




## Sequential Bayesian analysis

in general

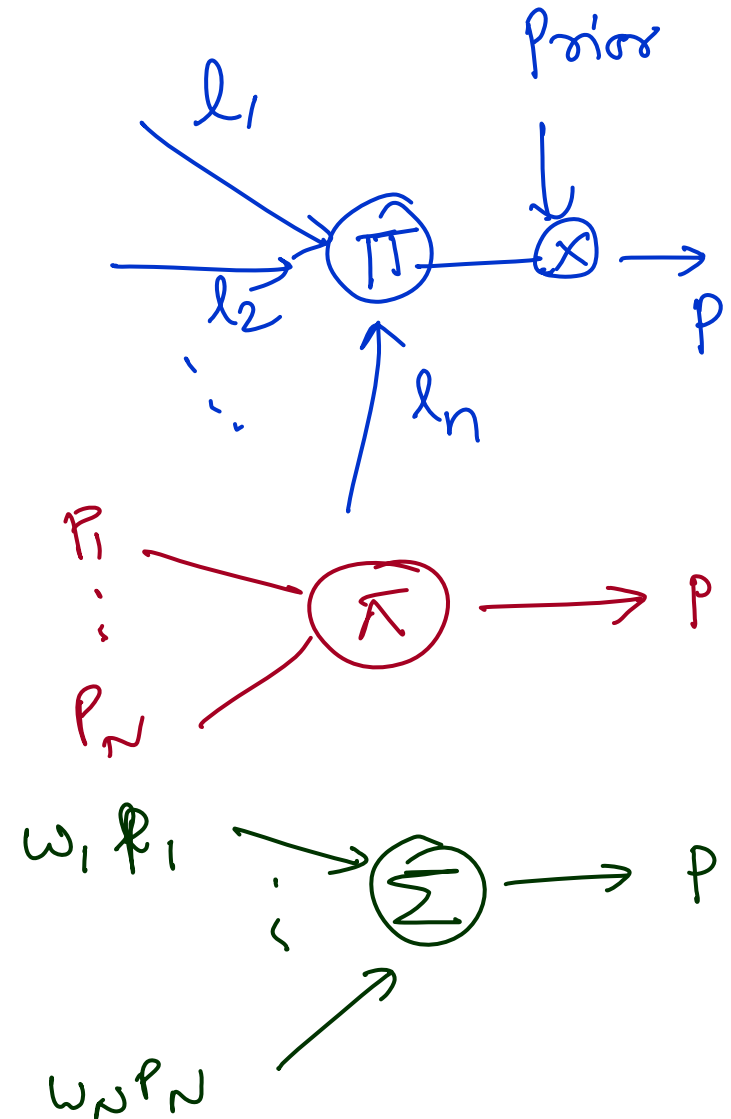
Probabilistic  
fusion center



## Sequential Bayesian analysis

Types of fusion:

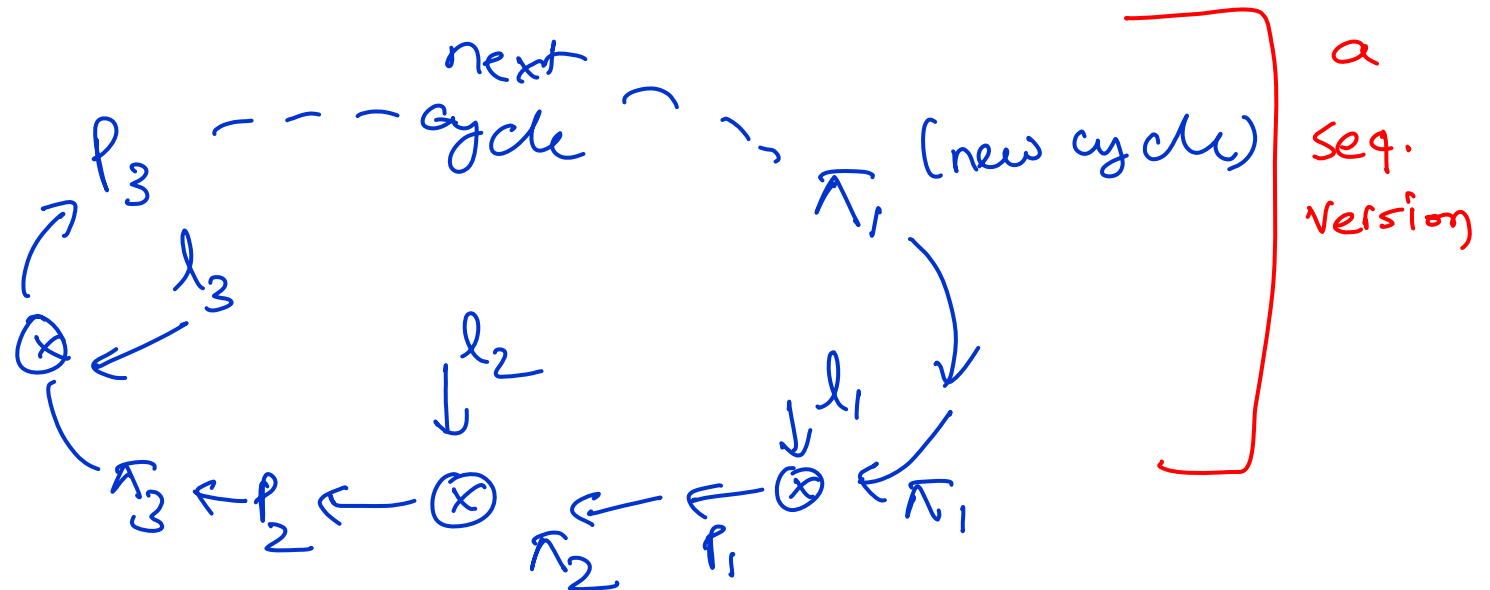
- Independent likelihood pool
- Independent opinion pool
- Linear opinion pool



## Sequential Bayesian analysis

### Independent Likelihood pool

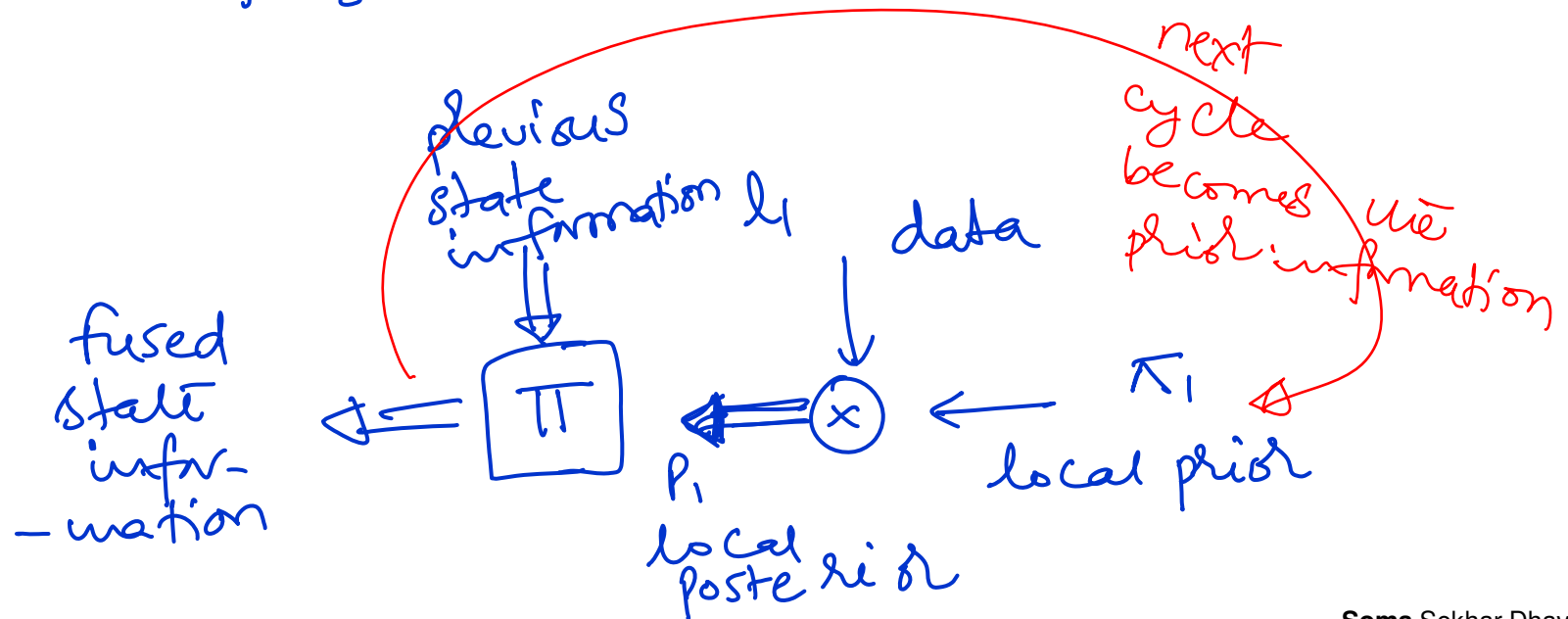
- Suitable if all the sensors have the same prior
- and can independently estimate the state information



## Sequential Bayesian analysis

### Independent opinion pool

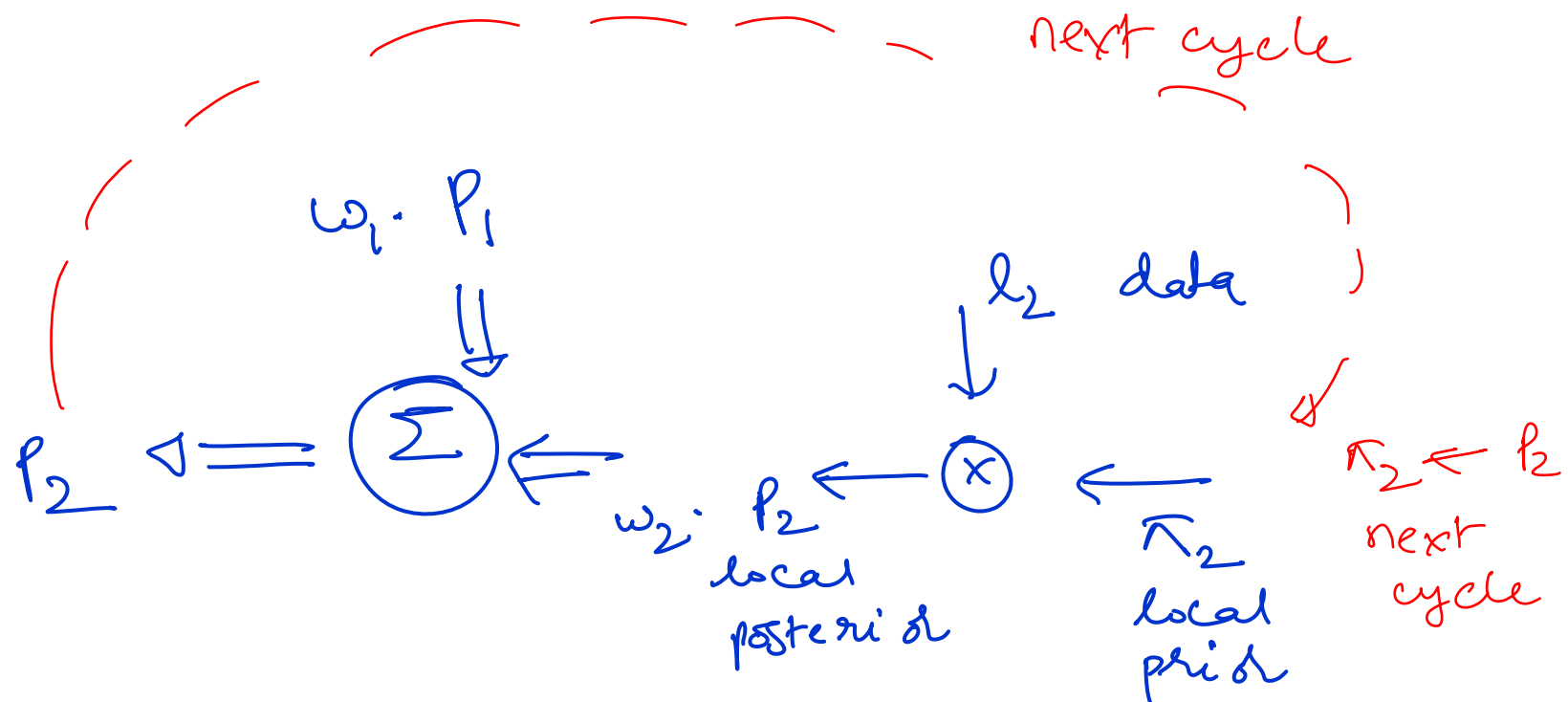
- Suitable when each sensor has different prior information
- and can estimate the state information reliably by each sensor



## Sequential Bayesian analysis

linear opinion pool.

- sensor estimates have strong dependency (possibly linear)



## Sequential Bayesian analysis

- prior is normal  $\rightarrow \mathcal{N}(\mu, \Sigma)$
- posterior is approximated as normal
- likelihood =  $C \mathcal{N}\left(\frac{A}{d^2}, \sigma^2\right)$

posterior calculations are based on

- Monte-carlo integrations &
- Gauss-Hermite numerical approximations

## Sequential Bayesian analysis

Efficient calculation of Posterior density:

$$E_{post}(g(\theta_1, \theta_2)) = \int \int g(\theta_1, \theta_2) L(\theta_1, \theta_2) \Pi(\theta_1, \theta_2) d\theta_1 d\theta_2$$

Upon expanding, we get,

$$= \int \frac{\exp(-\frac{(\theta_2 - \mu_2)^2}{2\sigma_2^2})}{(2\pi\sigma_2^2)^{0.5}} \int g(\theta_1, \theta_2) \frac{\exp(-\frac{(\theta_1 - \mu_1 - \rho \frac{\sigma_1}{\sigma_2}(\theta_2 - \mu_2))^2}{2\sigma_1^2(1-\rho^2)})}{(2\pi\sigma_1^2(1-\rho^2))^{0.5}} d\theta_1 d\theta_2$$

A numerical approximation to the above equation is sought using the Hermite polynomials [Ref2] which leads us to the following:

$$E_{post}[g(\theta_1, \theta_2)] \approx \sum_{p=1}^{N1} m_{1,p} \sum_{q=1}^{N2} m_{2,q} g(z_{1,p}, z_{2,q}) L(z_{1,p}, z_{2,q})$$

$$\mu'_1 = \mu_1 + \rho \frac{\sigma_1}{\sigma_2} (\theta_2 - \mu_2)$$

$$\sigma'_1 = \sigma_1 (1 - \rho^2)^{0.5}$$

$$m_{1,p} = w_{1,p} \exp(t_{1,p}) \sqrt{2} \sigma_2$$

$$z_{1,p} = \mu_2 + \sqrt{2} \sigma_2 t_{1,p}$$

$$m_{2,q} = w_{2,q} \exp(t_{2,q}) \sqrt{2} \sigma'_1$$

$$z_{2,q} = \mu'_1 + \sqrt{2} \sigma'_1 t_{2,q}$$

$$\begin{aligned} \hat{\theta}_1 &= M_{11} \\ \hat{\theta}_2 &= M_{21} \\ \hat{\sigma}_1^2 &= M_{12} - \hat{\theta}_1^2 \\ \hat{\sigma}_2^2 &= M_{22} - \hat{\theta}_2^2 \\ \hat{\rho} &= \frac{M_{co} - \hat{\theta}_1 \hat{\theta}_2}{\sigma_1 \sigma_2} \end{aligned}$$

$$\begin{aligned} M_{11} &= E_{post}[\theta_1] \\ M_{21} &= E_{post}[\theta_2] \\ M_{12} &= E_{post}[\theta_1^2] \\ M_{22} &= E_{post}[\theta_2^2] \\ M_{co} &= E_{post}[\theta_1 \theta_2] \end{aligned}$$

(using  
Gauss-Hermite  
numerical  
approximation)

15 x 15 grid  
is chosen

## Sequential Bayesian analysis

Some approximations:

$$p_{\text{new}} \propto \underbrace{p_j^k(\theta)}_{N_{\text{sig}}} \cdot \underbrace{p_{j-1}^k(\theta)}_{N(\mu_{j-1}^k, \Sigma_{j-1}^k)} \quad \left[ \begin{array}{l} \text{Indep.} \\ \text{opinion} \\ \text{pool} \end{array} \right]$$

$$p_{\text{new}} \approx \sim N[\mu_{\text{new}}, \Sigma_{\text{new}}]$$

$$\mu_{\text{new}} = \Sigma_{\text{new}} \left( \Sigma_j^k \mu_j^k + \Sigma_{j-1}^k \mu_{j-1}^k \right)$$

$$\Sigma_{\text{new}} = \left( \Sigma_{j-1}^k + \Sigma_0^k \right)^{-1}$$

(mean: a weighed average of the  
local state & previous state)



## Sequential Bayesian analysis

Some approximations:

$$p_{\text{new}} \propto \underbrace{w_1 p_1}_{\mathcal{N}(\mu_1, \Sigma_1)} + \underbrace{w_2 p_2}_{\mathcal{N}(\mu_2, \Sigma_2)} \quad \left\{ \begin{array}{l} \text{linear} \\ \text{opinion} \\ \text{pool} \end{array} \right\}$$

$$\cdot \quad p_{\text{new}} \approx \mathcal{N}(\mu, \Sigma) \quad \left\{ \begin{array}{l} \text{normal approx is} \\ \text{sought} \end{array} \right\}$$

when

$$\Sigma_{\text{new}} = w_1 \left( \Sigma_1 + \mu_1 \mu_1^T \right) + w_2 \left( \Sigma_2 + \mu_2 \mu_2^T \right)$$

$$- \mu_{\text{new}} \mu_{\text{new}}^T$$

$$\mu_{\text{new}}^T = w_1 \mu_1 + w_2 \mu_2$$

## Sequential Bayesian analysis

Some approximations:

$$p_{\text{new}} \propto \underbrace{w_1 p_1}_{\mathcal{N}(\mu_1, \Sigma_1)} + \underbrace{w_2 p_2}_{\mathcal{N}(\mu_2, \Sigma_2)} \quad \left\{ \begin{array}{l} \text{linear} \\ \text{opinion} \\ \text{pool} \end{array} \right.$$

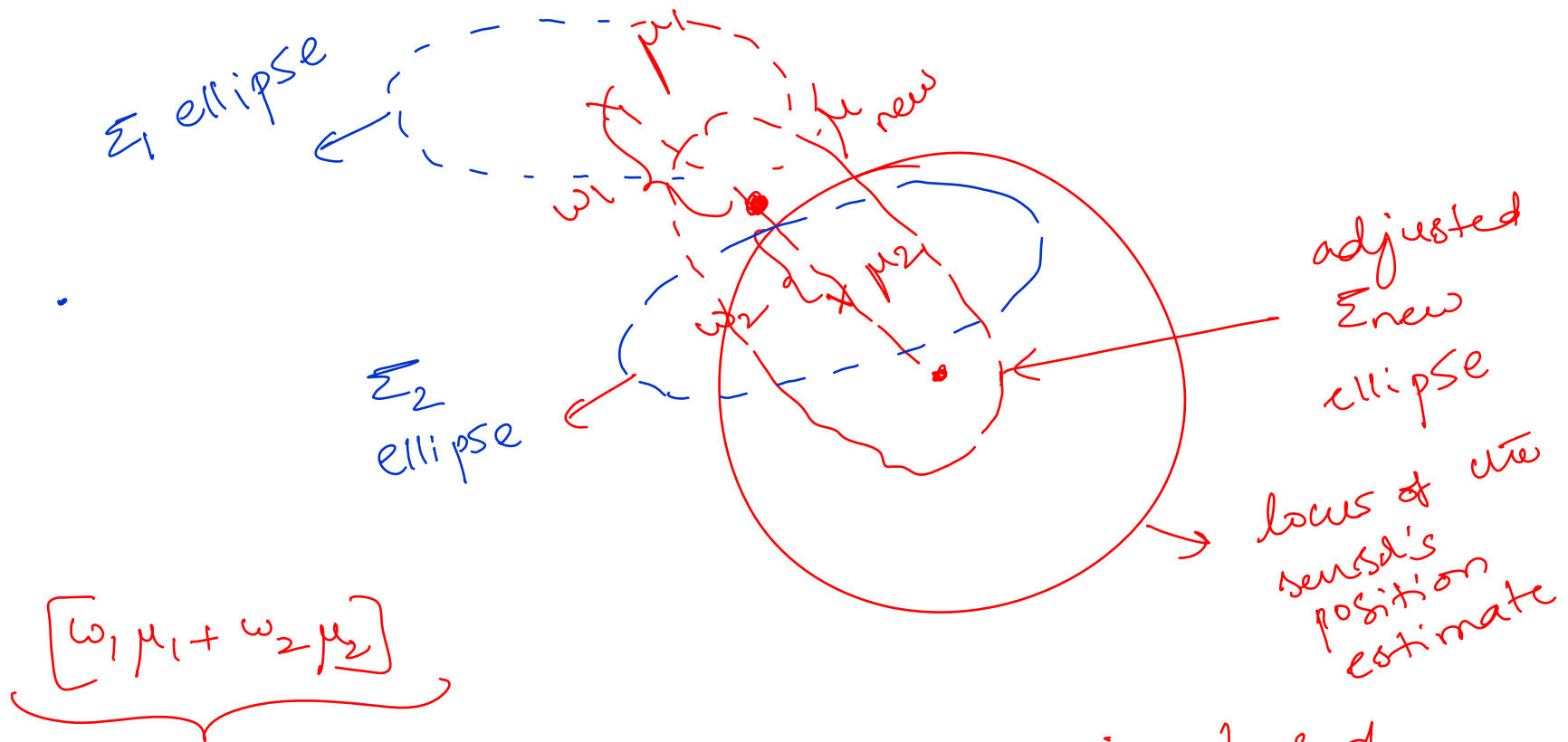
•  $p_{\text{new}} \approx \mathcal{N}(\mu_{\text{new}}, \Sigma_{\text{new}})$  { normal approx is }  
when sought

if  $\Sigma_1 \rightarrow 0$  (absolutely certain about it)

$$\mu_{\text{new}} = w_1 \mu_1 + w_2 \mu_2$$

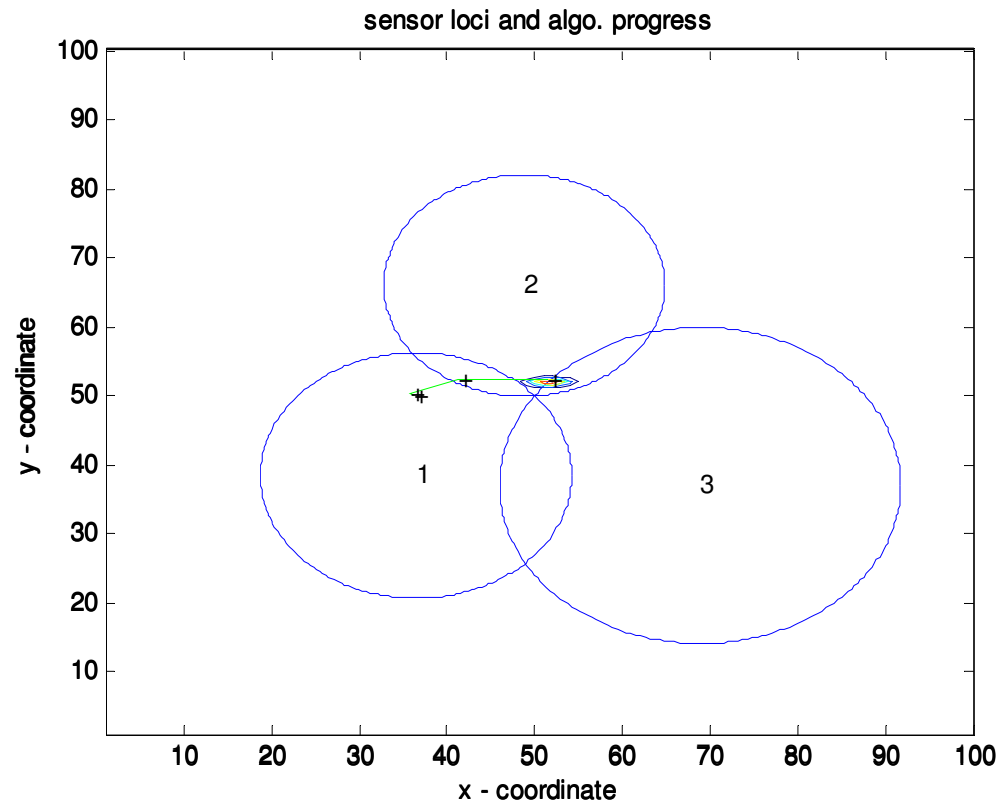
$$\Sigma_{\text{new}} = w_2 \Sigma_2 + w_1 w_2 (\mu_1 - \mu_2)(\mu_1 - \mu_2)^T$$

Some approximations: geometrical interpretation

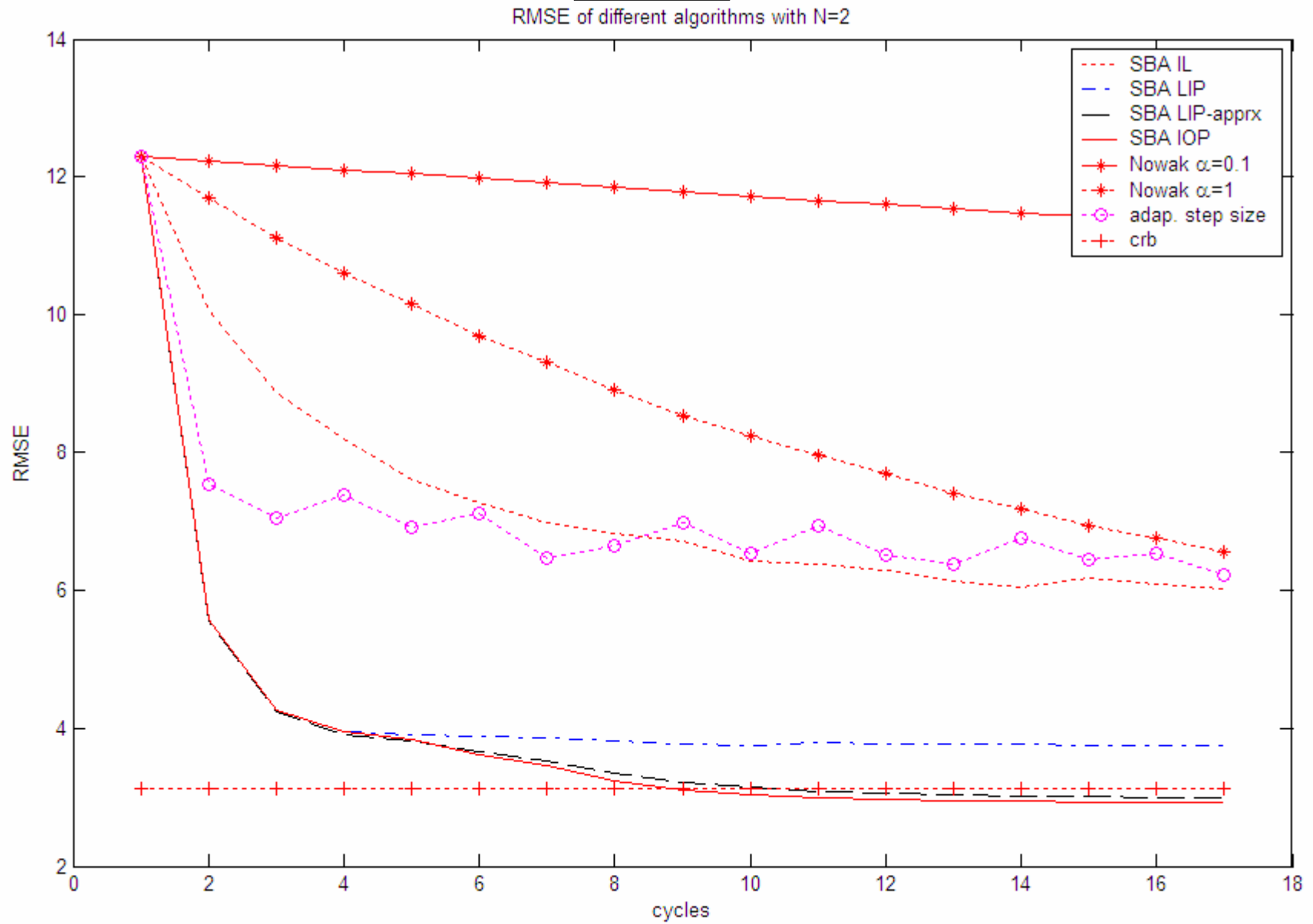


Covariance matrix is scaled  
in the direction of the error

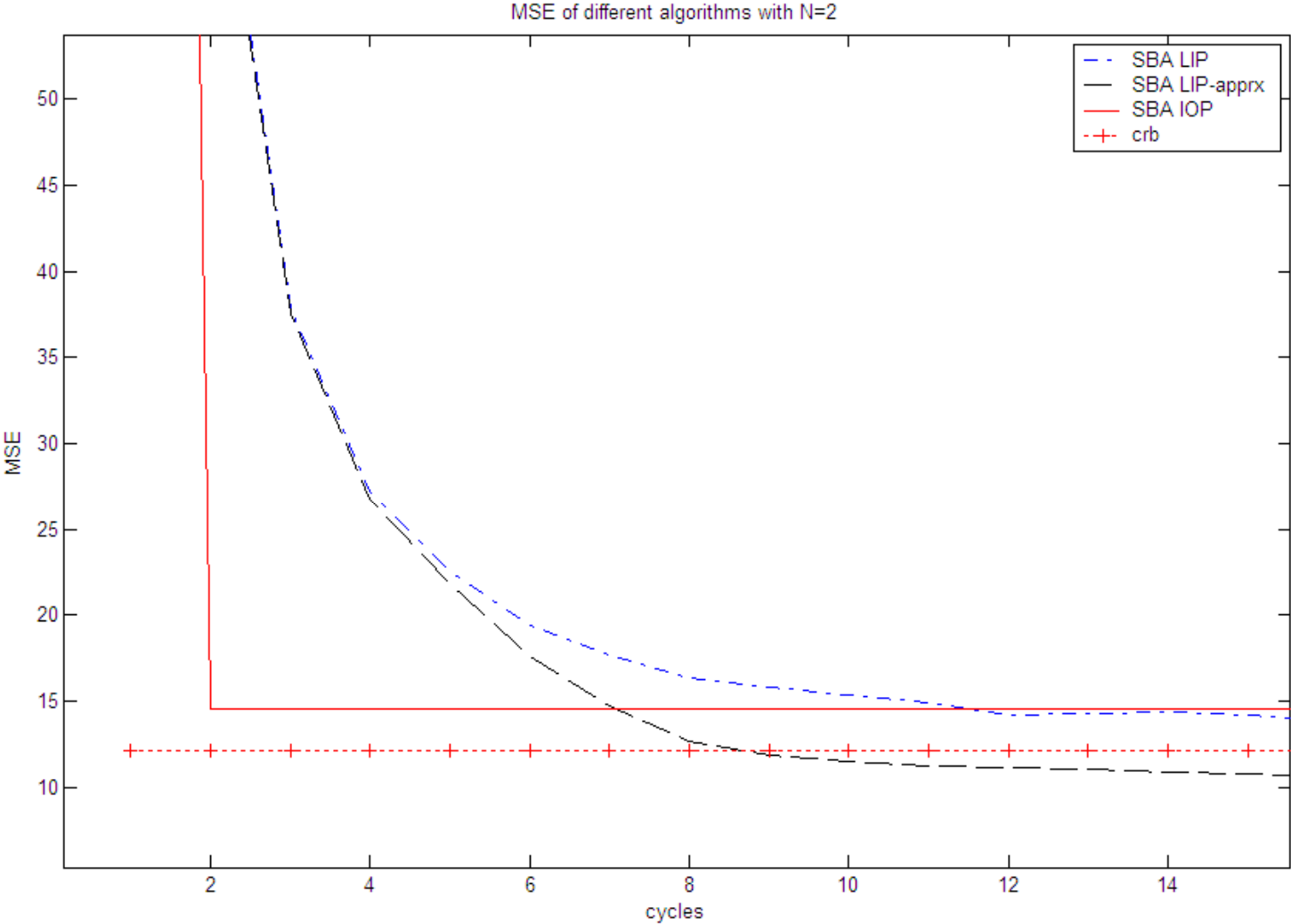
## Sequential Bayesian analysis



## Results



Results



## Conclusions/ Future work

Who leads the race?

Computational/ Communication cost  
Modified sub-gradient/ Baye's

Oversimplified assumptions?

Too many assumptions  
Noise variance unknown  
Source energy r.v  
Multiple sources  
Reverberation (indoor applications)

What else can be done?

Moving source  
Bounds on convergence  
Advanced Baye's filters  
De-simplifying the assumptions to be  
more realistic

## References

**Rabbat** and **Nowak**, “Decentralized source localization and tracking,” ICASSP, 2004, pp. III-921-924

**Naylor** and **Smith**, “Applications of a method for efficient computation of Posterior Distributions,” Applied Statistics, Vol.31, No.3 (1982), pp. 214-225

**Jeffrey Hightower** and **Gaetano Borriello**, "Location Sensing Techniques," UW CSE 01-07-01, University of Washington, Department of Computer Science and Engineering, Seattle, WA, July 2001

Thank you!