

On estimating the structure factor of a point process, with applications to hyperuniformity

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- 1 Definitions and Motivations
- 2 Structure factor estimators
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 - Tests
 - Comparison of the estimators
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Definitions and Motivations

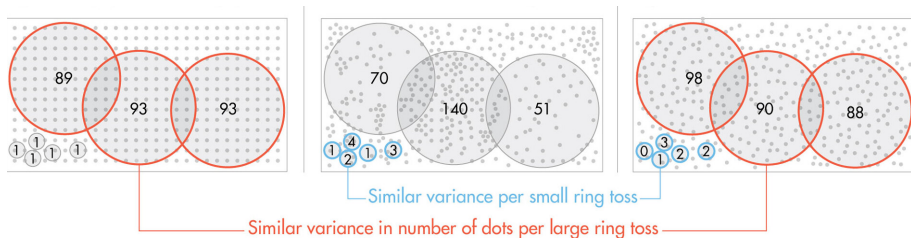
Hyperuniform or Superhomogeneous

Definitions and Motivations

Let \mathcal{X} be a stationary point process of \mathbb{R}^d of intensity ρ , \mathcal{X} is hyperuniform iff

■ Variance:

$$\lim_{R \rightarrow \infty} \frac{\text{Var}(\text{Card}(\mathcal{X} \cap B(0, R)))}{|B(0, R)|} = 0.$$

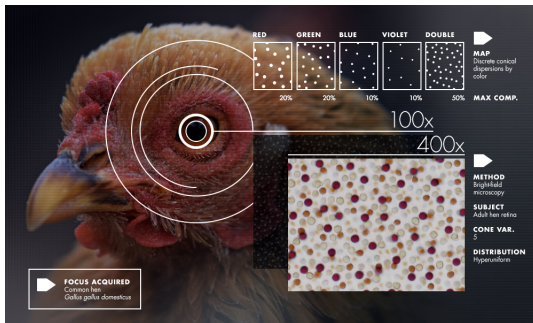


S. Torquato, *Hyperuniform States of Matter*, 2018.

S. Coste, *Order, Fluctuations, Rigidities*, 2021.

Bird's-Eye

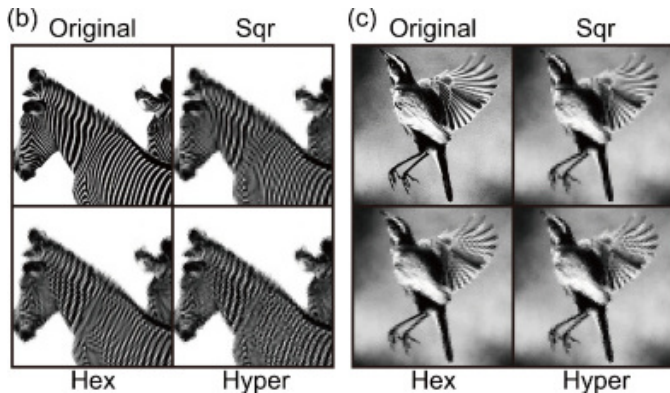
Definitions and Motivations



Y. Jiao, T. Lau, H. Hatzikirou, M Meyer-Hermann, J. C. Corbo, and S. Torquato, *Avian photoreceptor patterns represent a disordered hyperuniform solution to a multiscale packing problem*, 2014.

Image reconstruction

Definitions and Motivations



Ming-Jie Sun, Xin-Yu Zhao, and Li-Jing Li, *Imaging using hyperuniform sampling with a single-pixel camera*, 2018.

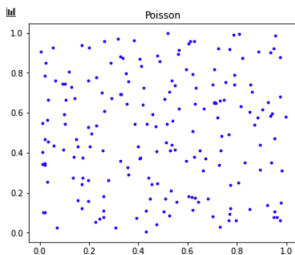
Numerical integration

Definitions and Motivations

■ Monte Carlo integration:

$$\int f(x)\mu(dx) \approx \sum_{i=1}^N w_i f(\mathbf{x}_i)$$

■ Rate of convergence with a Poisson point process: $O(1/\sqrt{N})$



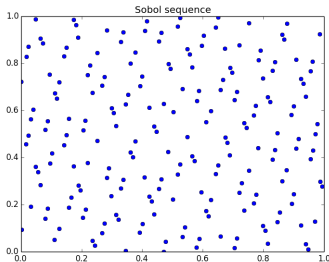
Numerical integration

Definitions and Motivations

■ Monte Carlo integration:

$$\int f(x) \mu(dx) \approx \sum_{i=1}^N w_i f(\mathbf{x}_i)$$

■ Rate of convergence with Sobol sequence: $O(\log(N)^d / N)$



Structure Factor

Definitions and Motivations

Let $\mathcal{X} \subset \mathbb{R}^d$ be a stationary point process of intensity ρ

- Structure factor

$$S(\mathbf{k}) = 1 + \rho \mathcal{F}(g - 1)(\mathbf{k})$$

- Pair correlation function

$$\mathbb{E} \left[\sum_{\mathbf{x}, \mathbf{y} \in \mathcal{X}}^{\neq} f(\mathbf{x}, \mathbf{y}) \right] = \int_{\mathbb{R}^d \times \mathbb{R}^d} f(\mathbf{x} + \mathbf{y}, \mathbf{y}) \rho^2 g(\mathbf{x}) d\mathbf{x} d\mathbf{y}$$

S. Coste, *Order, Fluctuations, Rigidities*, 2021.

S. Torquato, *Hyperuniform States of Matter*, 2018.

Structure Factor

Definitions and Motivations

$$\mathcal{X} \text{ is hyperuniform} \iff \lim_{R \rightarrow \infty} \frac{\text{Var}(\text{Card}(\mathcal{X} \cap B(0, R)))}{|B(0, R)|} = 0$$

- Structure factor

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- \mathcal{X} is hyperuniform iff

$$S(\mathbf{0}) = 0$$

S. Coste, *Order, Fluctuations, Rigidities*, 2021.

S. Torquato, *Hyperuniform States of Matter*, 2018.

Hyperuniformity class

Definitions and Motivations

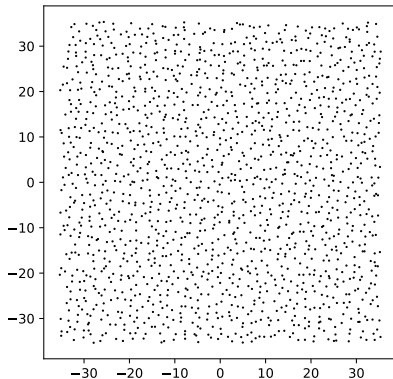
$$\mathcal{X} \text{ is hyperuniform} \iff \lim_{\mathbf{k} \rightarrow 0} \underbrace{1 + \rho \mathcal{F}(g-1)(\mathbf{k})}_{S(\mathbf{k})} = 0$$

- \mathcal{X} is hyperuniform with $|S(\mathbf{k})| \sim c \|\mathbf{k}\|_2^\alpha$ in the neighborhood of 0 then,

α	$\text{Var} [\text{Card}(\mathcal{X} \cap B(0, R))]$	class
> 1	$O(R^{d-1})$	I
1	$O(R^{d-1} \log(R))$	II
$]0, 1[$	$O(R^{d-\alpha})$	III

Ginibre ensemble

Definitions and Motivations



Ginibre ensemble

Definitions and Motivations

- Intensity: $\rho_{\text{Ginibre}} = 1/\pi$
- Pair correlation function: $g_{\text{Ginibre}}(r) = 1 - \exp(-r^2)$
- Structure factor: $S_{\text{Ginibre}}(k) = 1 - \exp(-k^2/4)$
- Power decay: $\alpha_{\text{Ginibre}} = 2$
- Hyperuniform class: *I*

Structure factor estimators

$$S(\mathbf{k}) = 1 + \rho \mathcal{F}(g - 1)(\mathbf{k})$$

- **Given:** $\mathcal{X}_N = \{\mathbf{x}_i\}_1^N$ a realization of a **stationary** point process \mathcal{X} of intensity ρ in $W = [-L/2, L/2]^d$
- **Need:** Use \mathcal{X}_N to approximate $S(\mathbf{k}) = 1 + \rho \int_{\mathbb{R}^d} (g(\mathbf{r}) - 1) e^{-i\langle \mathbf{k}, \mathbf{r} \rangle} d\mathbf{r}$

$$S(\mathbf{k}) = 1 + \rho \mathcal{F}(g - 1)(\mathbf{k})$$

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- **Idea:**

- 1 Use $\alpha_t(\mathbf{r}, W) = \int_{\mathbb{R}^d} t(\mathbf{r} + \mathbf{y}, W) t(\mathbf{y}, W) d\mathbf{y}$ s.t. $\lim_{L \rightarrow \infty} \alpha_t(\mathbf{r}, W) = 1$ and $\|t\|_2 = 1$:

$$\begin{aligned} S(\mathbf{k}) &= 1 + \rho \int_{\mathbb{R}^d} \lim_{L \rightarrow \infty} (g(\mathbf{r}) - 1) \alpha_t(\mathbf{r}, W) e^{-i\langle \mathbf{k}, \mathbf{r} \rangle} d\mathbf{r} \\ &= \lim_{L \rightarrow \infty} \mathbb{E}[\widehat{S}(t, \mathbf{k})] - \underbrace{\rho \mathcal{F}(\alpha_t)(\mathbf{k}, W)}_{\epsilon_t(\mathbf{k}, \mathbf{L})} \end{aligned}$$

- 2 Reduce bias: Consider the zeros of $\epsilon_t(\mathbf{k}, \mathbf{L})$ as the set of allowed wavevectors of \widehat{S} , or remove the bias term.

Scattering intensity

Structure factor estimators Estimators

- **Taper:** $t_0(\mathbf{x}, W) = \frac{1}{\sqrt{|W|}} \mathbb{1}_W(\mathbf{x})$
- $S(\mathbf{k}) = \lim_{L \rightarrow \infty} \underbrace{\mathbb{E} \left[\frac{1}{\rho |W|} \left| \sum_{\mathbf{x} \in \mathcal{X} \cap W} e^{-i \langle \mathbf{k}, \mathbf{x} \rangle} \right|^2 \right]}_{\hat{S}_{\text{SI}}(\mathbf{k})} - \underbrace{\rho \left(\prod_{j=1}^d \frac{\sin(k_j L/2)}{k_j \sqrt{L}/2} \right)^2}_{\epsilon_0(\mathbf{k}, \mathbf{L})}$
- $\epsilon_0(\mathbf{k}, \mathbf{L}) = \begin{cases} 0 & \text{if } \mathbf{k} \in \mathbb{A}_{\mathbf{L}} \\ \rho L^d & \text{as } \|\mathbf{k}\|_2 \rightarrow 0 \\ 2^{2d} \prod_{j=1}^d \frac{1}{L k_j^2} & \text{as } \|\mathbf{k}\|_2 \rightarrow \infty \end{cases}$
- **Allowed wavevectors:**
 $\mathbb{A}_{\mathbf{L}} = \{(k_1, \dots, k_d) \in (\mathbb{R}^d)^*, \exists j \in \{1, \dots, d\}, n \in \mathbb{Z}^* \text{ s.t. } k_j = \frac{2\pi n}{L}\}$

Scattering intensity estimator

1 Estimator:

$$\hat{S}_{\text{SI}}(\mathbf{k}) = \frac{1}{\rho|W|} \left| \sum_{\mathbf{x} \in \mathcal{X} \cap W} e^{-i\langle \mathbf{k}, \mathbf{x} \rangle} \right|^2$$

2 Allowed wavevectors:

$$\mathbb{A}_{\mathbf{L}} = \left\{ (k_1, \dots, k_d) \in (\mathbb{R}^d)^*, \exists j \in \{1, \dots, d\}, n \in \mathbb{Z}^* \text{ s.t. } k_j = \frac{2\pi n}{L} \right\}.$$

Formulation in the literature:

1 Estimator:

$$\hat{S}_{\text{SI},s}(\mathbf{k}) \triangleq \frac{1}{N} \left| \sum_{j=1}^N e^{-i\langle \mathbf{k}, \mathbf{x}_j \rangle} \right|^2$$

2 Allowed wavevectors:

$$\mathbb{A}_{\mathbf{L}}^{\text{res}} = \left\{ \left(\frac{2\pi n_1}{L}, \dots, \frac{2\pi n_d}{L} \right) \text{ with, } \mathbf{n} = (n_1, \dots, n_d) \in \mathbb{Z}^d \setminus \{\mathbf{0}\} \right\}.$$

Preprint: D. Hawat, G. Gautier, R. Bardenet, R. Lachièze-Rey *On estimating the structure factor of a point process, with applications to hyperuniformity, 2022.*

Tapered estimator

General case

- Tapered estimator:

$$S(\mathbf{k}) = \lim_{L \rightarrow \infty} \mathbb{E} \left[\underbrace{\frac{1}{\rho} \left| \sum_{j=1}^N t(\mathbf{x}_j, W) e^{-i\langle \mathbf{k}, \mathbf{x}_j \rangle} \right|^2}_{\hat{S}_T(t, \mathbf{k})} \right] - \underbrace{\rho |\mathcal{F}(t)(\mathbf{k}, W)|^2}_{\epsilon_t(\mathbf{k}, L)}.$$

- Debiased tapered estimator:

- 1 Directly debiased:

$$\hat{S}_{\text{DDT}}(t, \mathbf{k}) \triangleq \frac{1}{\rho} \left| \sum_{j=1}^N t(\mathbf{x}_j, W) e^{-i\langle \mathbf{k}, \mathbf{x}_j \rangle} - \rho \mathcal{F}(t)(\mathbf{k}, W) \right|^2$$

- 2 Undirectly debiased:

$$\hat{S}_{\text{UDT}}(t, \mathbf{k}) \triangleq \frac{1}{\rho} \left| \sum_{j=1}^N t(\mathbf{x}_j, W) e^{-i\langle \mathbf{k}, \mathbf{x}_j \rangle} \right|^2 - \rho |\mathcal{F}(t)(\mathbf{k}, W)|^2$$

Preprint: T. Rajala and S. C. Olhede and D. John Murrell *Spectral estimation for spatial point patterns*, 2020.

Multitapered estimator

$$\hat{S}_T(t, \mathbf{k}) = \frac{1}{\rho} \left| \sum_{j=1}^N t(\mathbf{x}_j, W) e^{-i\langle \mathbf{k}, \mathbf{x}_j \rangle} \right|^2$$

More generally,

- Family of orthogonal tapers: $(t_q)_{q=1}^P$
- Multitapered estimator:

$$\hat{S}_{MT}((t_q)_{q=1}^P, \mathbf{k}) = \frac{1}{P} \sum_{q=1}^P \hat{S}_T(t_q, \mathbf{k}).$$

Estimators assuming stationarity and isotropy

Structure factor estimators Estimators

- Structure factor: $S(\mathbf{k}) = 1 + \rho \mathcal{F}(g - 1)(\mathbf{k})$
- Isotropic case:

$$S(k) = 1 + \rho \frac{(2\pi)^{d/2}}{k^{d/2-1}} \int_0^\infty (g(r) - 1) r^{d/2} J_{d/2-1}(kr) dr.$$

Estimators assuming stationarity and isotropy

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- Ball window : $W = B(0, R)$
- Radial Taper: $t_0(\mathbf{x}, W) = \frac{1}{\sqrt{|W|}} \mathbb{1}_W(\mathbf{x})$

Estimators assuming stationarity and isotropy

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- Ball window : $W = B(0, R)$
- Radial Taper: $t_0(\mathbf{x}, W) = \frac{1}{\sqrt{|W|}} \mathbb{1}_W(\mathbf{x})$
- Bartlett's isotropic estimator

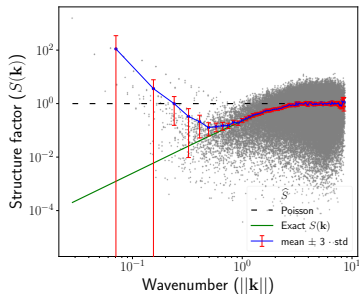
$$\hat{S}_{\text{BI}}(k) = 1 + \frac{(2\pi)^{\frac{d}{2}}}{\rho |W| \omega_{d-1}} \sum_{\substack{i,j=1 \\ i \neq j}}^N \frac{J_{d/2-1}(k \|\mathbf{x}_i - \mathbf{x}_j\|_2)}{(k \|\mathbf{x}_i - \mathbf{x}_j\|_2)^{d/2-1}}$$

- Allowed wavenumbers: $\mathbb{A}_R = \left\{ \frac{x}{R} \in \mathbb{R}, \text{ s.t. } J_{d/2}(x) = 0 \right\}$

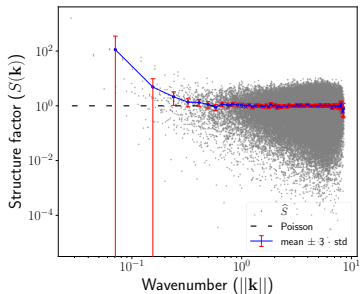
Scattering intensity

Structure factor estimators Tests

$$\hat{S}_{\text{SI}}(\mathbf{k}) = \frac{1}{\rho|W|} \left| \sum_{\mathbf{x} \in \mathcal{X} \cap W} e^{-i\langle \mathbf{k}, \mathbf{x} \rangle} \right|^2$$



Ginibre



Poisson

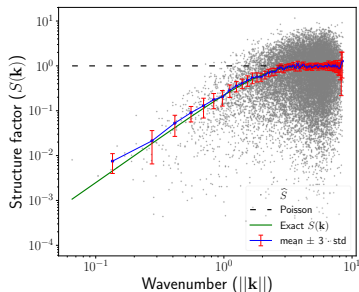
Figure: On arbitrary wavevectors \mathbf{k}

<https://github.com/For-a-few-DPPs-more/structure-factor>

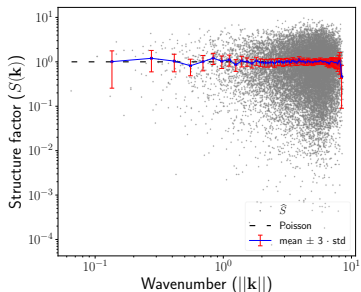
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Ginibre



Poisson

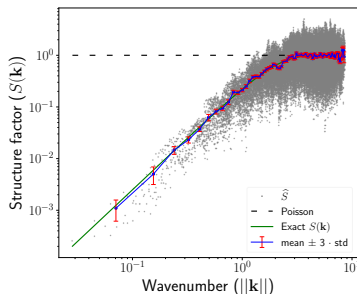
Figure: On allowed wavevectors \mathbf{k}

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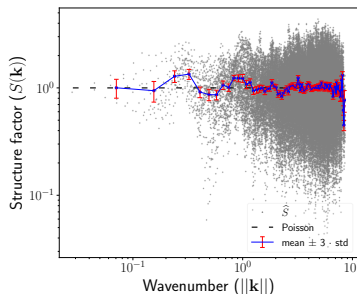
Multitapered estimator

Structure factor estimators Tests

$$\hat{S}_{\text{DDT}}(t, \mathbf{k}) \triangleq \frac{1}{\rho} \left| \sum_{j=1}^N t(\mathbf{x}_j, W) e^{-i\langle \mathbf{k}, \mathbf{x}_j \rangle} - \rho \mathcal{F}(t)(\mathbf{k}, W) \right|^2$$



Ginibre



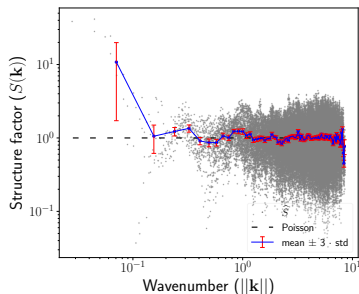
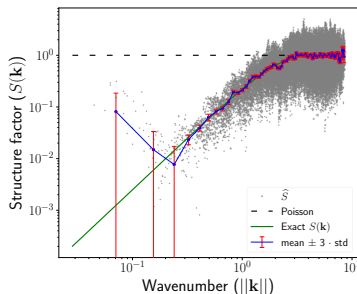
Poisson

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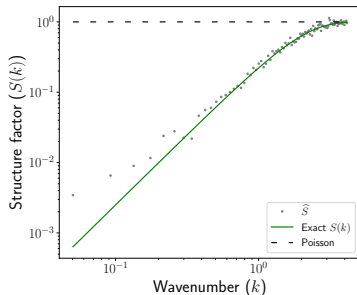


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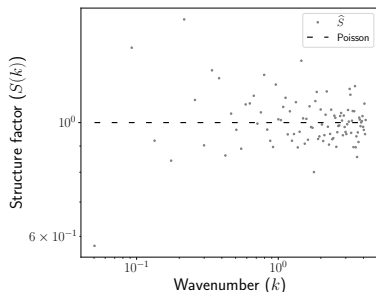
Bartlett's isotropic estimator

Structure factor estimators Tests

$$\hat{S}_{\text{BI}}(k) \triangleq 1 + \frac{(2\pi)^{\frac{d}{2}}}{\rho |W| \omega_{d-1}} \sum_{\substack{i,j=1 \\ i \neq j}}^N \frac{J_{d/2-1}(k \|\mathbf{x}_i - \mathbf{x}_j\|_2)}{(k \|\mathbf{x}_i - \mathbf{x}_j\|_2)^{d/2-1}}$$



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Poisson

Figure: On allowed wavenumbers k

<https://github.com/For-a-few-DPPs-more/structure-factor>

Integrated mean square error

Sample integrated variance and MSE across 50 samples

Estimators	\widehat{iVar}	$CI[\widehat{iMSE}]$	\widehat{iVar}	$CI[\widehat{iMSE}]$
$\widehat{S}_{SI}(2\pi\mathbf{n}/L)$	0.32	0.32 ± 0.02	1.31	1.34 ± 0.06
$\widehat{S}_{DDMT}((t_q)_1^4)$	0.08	0.08 ± 0.007	0.37	0.38 ± 0.02
$\widehat{S}_{BI}(\frac{x}{R})$	3.9×10^{-3}	$4.0 \times 10^{-3} \pm 3 \times 10^{-4}$	0.057	$0.058 \pm 9 \times 10^{-3}$
	Ginibre		Poisson	

Hyperuniformity diagnostics

Effective hyperuniformity

$$\mathcal{X} \text{ is effectively hyperuniform} \iff H = \frac{\hat{S}(0)}{\hat{S}(k_{peak})} \leq 10^{-3},$$

- $\hat{S}(0)$ is a linear extrapolation of the estimated structure factor \hat{S} in $k = 0$.
- k_{peak} is the location of the first dominant peak value of \hat{S} .

- Need: Check if $S(\mathbf{0}) = 0$
- Problem: We don't have an unbiased estimator of $S(\mathbf{0})$
- Idea: Use biased estimators to construct an unbiased estimator

Coupled sum estimator

Need: Check if $S(\mathbf{0}) = 0$

- $(Y_m)_m$ a sequence of approximations of a r.v. Y

$$Z_m = \sum_{j=1}^{m \wedge M} \frac{Y_j - Y_{j-1}}{\mathbb{P}(M \geq j)}, \quad m \geq 1$$

- M is an \mathbb{N} -r.v. such that $\mathbb{P}(M \geq j) > 0$ for all j , and $Y_0 = 0$

C. Rhee and P.W. Glynn, *Unbiased estimation with square root convergence for SDE models*, 2015

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- M is an \mathbb{N} -r.v. such that $\mathbb{P}(M \geq j) > 0$ for all j , and $Y_0 = 0$
- $\mathbb{E}[Z_m] = \mathbb{E}[Y_m]$ and $Z_m \xrightarrow[m \rightarrow \infty]{\text{a.s.}} Z := \sum_{j=1}^M \frac{Y_j - Y_{j-1}}{\mathbb{P}(M \geq j)}$.
- $Y_m \xrightarrow[m \rightarrow \infty]{L^2} Y$ + some hypotheses $\implies \mathbb{E}[Z] = \mathbb{E}[Y]$

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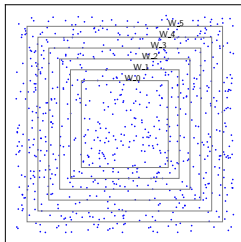
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- Consider an increasing sequence of sets $(\mathcal{X} \cap W_m)_{m \geq 1}$, with $W_\infty = \mathbb{R}^d$



Coupled sum estimator

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- Consider an increasing sequence of sets $(\mathcal{X} \cap W_m)_{m \geq 1}$, with $W_\infty = \mathbb{R}^d$
- Take $Y_m = 1 \wedge \hat{S}_m(\mathbf{k}_m^{\min})$ and $\mathbf{k}_m^{\min} \xrightarrow{m \rightarrow \infty} \mathbf{0}$

Multiscale hyperuniformity test

Hyperuniformity diagnostics Statistical test

Need: Check if $S(\mathbf{0}) = 0$

- Take $Y_m = 1 \wedge \hat{S}_m(\mathbf{k}_m^{\min})$, $\mathbf{k}_m^{\min} \xrightarrow{m \rightarrow \infty} \mathbf{0}$, $\{W_m\}_m \uparrow$, and $W_\infty = \mathbb{R}^d$
- $Z = \sum_{j=1}^M \frac{Y_j - Y_{j-1}}{\mathbb{P}(M \geq j)}$
- M is an \mathbb{N} -r.v. such that $\mathbb{P}(M \geq j) > 0$ for all j , and $Y_0 = 0$

Proposition

Assume that $M \in L^p$ for some $p \geq 1$. Then $Z \in L^p$ and $Z_m \rightarrow Z$ in L^p . Moreover,

- 1 If \mathcal{X} is hyperuniform, then $\mathbb{E}[Z] = 0$.
- 2 If \mathcal{X} is not hyperuniform and $\sup_m \mathbb{E}[\hat{S}_m^2(\mathbf{k}_m^{\min})] < \infty$, then $\mathbb{E}[Z] \neq 0$.

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Multiscale hyperuniformity test

Hyperuniformity diagnostics Statistical test

Need: Check if $S(\mathbf{0}) = 0 \iff \mathbb{E}[Z] = 0$, with $Z = \sum_{j=1}^M \frac{Y_j - Y_{j-1}}{\mathbb{P}(M \geq j)}$

Test:

- M a Poisson r.v. of parameter λ
- i.i.d. pairs $(\mathcal{X}_a, M_a)_{a=1}^A$ of realizations of (\mathcal{X}, M)
- Asymptotic confidence interval $CI[\mathbb{E}[Z]]$ of level ζ

$$CI[\mathbb{E}[Z]] = \left[\bar{Z}_A - z \bar{\sigma}_A A^{-1/2}, \bar{Z}_A + z \bar{\sigma}_A A^{-1/2} \right]$$

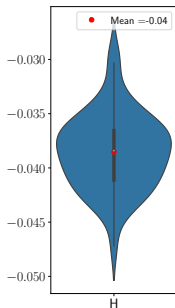
with $\mathbb{P}(-z < \mathcal{N}(0, 1) < z) = \zeta$

- Assessing whether 0 lies in $CI[\mathbb{E}[Z]]$

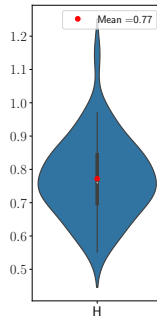
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Ginibre



Poisson

H for 50 samples from Ginibre and Poisson using Bartlett's isotropic estimator

<https://github.com/For-a-few-DPPs-more/structure-factor>

Multiscale hyperuniformity test

$$\mathcal{X} \text{ is hyperuniform} \iff \mathbb{E}[Z] = \mathbb{E}\left[\sum_{j=1}^M \frac{Y_j - Y_{j-1}}{\mathbb{P}(M \geq j)}\right] = 0$$

Table: Multiscale hyperuniformity test with 99% confidence level

	\bar{Z}_{50}	$CI[\mathbb{E}[Z]]$	\bar{Z}_{50}	$CI[\mathbb{E}[Z]]$
Ginibre	0.015	$[-0.021, 0.051]$	0.007	$[-0.003, 0.011]$
Poisson	0.832	$[0.444, 1.220]$	0.781	$[0.560, 1.001]$
\hat{S}	\hat{S}_{SI}		\hat{S}_{BI}	

Multiscale hyperuniformity test

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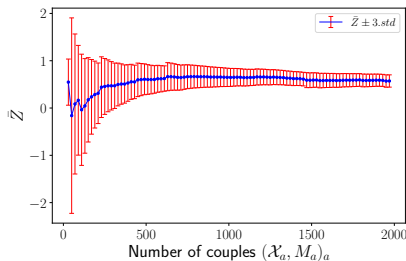
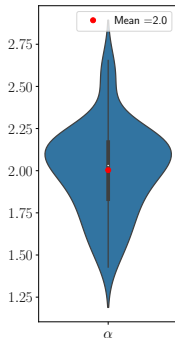


Figure: $C/[\mathbb{E}[\bar{Z}]]$ for a Poisson point process with the scattering intensity, as a function of the number of realizations of Z .

Hyperuniformity class

\mathcal{X} is hyperuniform with $|S(\mathbf{k})| \sim c \|\mathbf{k}\|_2^\alpha$ in the neighborhood of zero



α of 50 samples from Ginibre using Bartlett's isotropic estimator

Code availability

Python Package

Code availability

- 1 Open-source Python toolbox called `structure_factor`¹
- 2 Available on Github and PyPI²
- 3 Detailed documentation³
- 4 Jupyter notebook tutorial⁴

¹<https://github.com/For-a-few-DPPs-more/structure-factor>

²<https://pypi.org/project/structure-factor/>

³<https://for-a-few-dpps-more.github.io/structure-factor/>

⁴<https://github.com/For-a-few-DPPs-more/structure-factor/tree/main/notebooks>

Conclusion

Conclusion

Conclusion

- Estimators of the structure factor, properties and performances
- Available diagnostics of hyperuniformity and limitations
- The first statistical test of hyperuniformity and limitations
- Python toolbox `structure-factor`

THANK YOU

Conclusion



Github



Documentation



Preprint

Github: <https://github.com/For-a-few-DPPs-more/structure-factor>

Documentation: <https://for-a-few-dpps-more.github.io/structure-factor/>

Preprint: D. Hawat, G. Gautier, R. Bardenet, R. Lachìze-Rey *On estimating the structure factor of a point process, with applications to hyperuniformity*, 2022.