

Advanced Econometrics

Lecture 8: Restricted Estimation (Hansen Chapter 8)

有约束的估计

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只讲前面四节.

Introduction

- ▶ The linear regression model:

$$Y_i = \mathbf{X}'_i \boldsymbol{\beta} + e_i$$
$$\mathbb{E}(\mathbf{X}_i e_i) = \mathbf{0}.$$

Partition $\mathbf{X}'_i = (\mathbf{X}'_{1i}, \mathbf{X}'_{2i})$ and $\boldsymbol{\beta}' = (\boldsymbol{\beta}'_1, \boldsymbol{\beta}'_2)$.

- ▶ “Exclusion restriction”: $\boldsymbol{\beta}_2 = \mathbf{0}$.
- ▶ The constrained model:

$$Y_i = \mathbf{X}'_{1i} \boldsymbol{\beta}_1 + e_i$$
$$\mathbb{E}(\mathbf{X}_i e_i) = \mathbf{0}.$$

- ▶ In the constrained model, e_i is uncorrelated with the entire regressor vector $\mathbf{X}'_i = (\mathbf{X}'_{1i}, \mathbf{X}'_{2i})$.

Introduction

- ▶ A set of q linear constraints on β takes the form

$$\mathbf{R}'\beta = \mathbf{c}.$$

This is thought of as some restriction that the parameter satisfies we believe to be true, based on prior knowledge.

- ▶ \mathbf{R} is $k \times q$ and $\text{rank}(\mathbf{R}) = q < k$, \mathbf{c} is $q \times 1$. The constraints are linearly independent.
- ▶ The restricted parameter space:

$$\mathbf{B}_R = \{\beta : \mathbf{R}'\beta = \mathbf{c}\}.$$

$\mathbf{R}'\beta - \mathbf{c} = 0$ 是一个约束.

- ▶ The exclusion restriction $\beta_2 = 0$ is a special case:

$$\mathbf{R} = \begin{pmatrix} \mathbf{0} \\ \mathbf{I} \end{pmatrix} \text{ and } \mathbf{c} = \mathbf{0}.$$

Constrained Least Squares

- ▶ We estimate the constrained linear regression by minimizing the least-squares criterion subject to the constraint $\mathbf{R}'\boldsymbol{\beta} = \mathbf{c}$.
- ▶ The constrained LS estimator:

$$\tilde{\boldsymbol{\beta}}_{\text{cls}} = \underset{\mathbf{R}'\boldsymbol{\beta}=\mathbf{c}}{\operatorname{argmin}} SSE(\boldsymbol{\beta})$$

$$SSE(\boldsymbol{\beta}) = \sum_{i=1}^n (Y_i - \mathbf{X}'_i \boldsymbol{\beta})^2 = \mathbf{Y}'\mathbf{Y} - 2\mathbf{Y}'\mathbf{X}\boldsymbol{\beta} + \boldsymbol{\beta}'\mathbf{X}'\mathbf{X}\boldsymbol{\beta}.$$

- ▶ We find solution by using Lagrange multipliers. The Lagrangian function:

$$\mathcal{L}(\boldsymbol{\beta}, \boldsymbol{\lambda}) = \frac{1}{2} SSE(\boldsymbol{\beta}) + \boldsymbol{\lambda}'(\mathbf{R}'\boldsymbol{\beta} - \mathbf{c}).$$

- ▶ The first-order conditions:

$$\frac{\partial}{\partial \boldsymbol{\beta}} \mathcal{L}(\tilde{\boldsymbol{\beta}}_{\text{cls}}, \tilde{\boldsymbol{\lambda}}_{\text{cls}}) = -\mathbf{X}'\mathbf{Y} + \mathbf{X}'\mathbf{X}\tilde{\boldsymbol{\beta}}_{\text{cls}} + \mathbf{R}\tilde{\boldsymbol{\lambda}}_{\text{cls}} = \mathbf{0}$$

$$\frac{\partial}{\partial \boldsymbol{\lambda}} \mathcal{L}(\tilde{\boldsymbol{\beta}}_{\text{cls}}, \tilde{\boldsymbol{\lambda}}_{\text{cls}}) = \mathbf{R}'\tilde{\boldsymbol{\beta}}_{\text{cls}} - \mathbf{c} = \mathbf{0}.$$

$\min_{\mathbf{b} \in B_n} \sum_{i=1}^n (Y_i - \mathbf{X}_i' \mathbf{b})^2$
↑
有约束的最小二乘估计

Constrained Least Squares

- Premultiplying by $R' (X'X)^{-1}$,

$$-R'\hat{\beta} + R'\tilde{\beta}_{\text{cls}} + R' (X'X)^{-1} R\tilde{\lambda}_{\text{cls}} = 0,$$

where $\hat{\beta} = (X'X)^{-1} X'Y$ is the unrestricted LS.

- Solving for $\tilde{\lambda}_{\text{cls}}$:

$$\tilde{\lambda}_{\text{cls}} = \left[R' (X'X)^{-1} R \right]^{-1} (R'\hat{\beta} - c).$$

- Solving for $\tilde{\beta}_{\text{cls}}$:

$$\tilde{\beta}_{\text{cls}} = \hat{\beta} - (X'X)^{-1} R \left[R' (X'X)^{-1} R \right]^{-1} (R'\hat{\beta} - c). \quad \text{if } R'\hat{\beta} - c = 0, \tilde{\beta}_{\text{cls}} = \hat{\beta}$$

如果无约束的极值点满足约束, 那么无约束与有约束是一样的.

Exclusion Restriction

- ▶ The unconstrained model:

$$Y_i = \mathbf{X}'_{1i}\boldsymbol{\beta}_1 + \mathbf{X}'_{2i}\boldsymbol{\beta}_2 + e_i.$$

- ▶ The exclusion restriction: $\boldsymbol{\beta}_2 = \mathbf{0}$. The constrained equation is

$$Y_i = \mathbf{X}'_{1i}\boldsymbol{\beta}_1 + e_i.$$

- ▶ The constrained LS is

$$\tilde{\boldsymbol{\beta}}_{\text{cls}} = \hat{\boldsymbol{\beta}} - (\mathbf{X}'\mathbf{X})^{-1} \begin{pmatrix} \mathbf{0} \\ \mathbf{I} \end{pmatrix} \left[(\mathbf{0} \quad \mathbf{I}) (\mathbf{X}'\mathbf{X})^{-1} \begin{pmatrix} \mathbf{0} \\ \mathbf{I} \end{pmatrix} \right]^{-1} (\mathbf{0} \quad \mathbf{I}) \hat{\boldsymbol{\beta}}.$$

- ▶ Easy to check:

$$\tilde{\boldsymbol{\beta}}_{\text{cls}} = \begin{pmatrix} (\sum_{i=1}^n \mathbf{X}_{1i}\mathbf{X}'_{1i})^{-1} (\sum_{i=1}^n \mathbf{X}_{1i}Y_i) \\ \mathbf{0} \end{pmatrix}.$$

无约束最小二乘估计

$$Y_i = X_i' \beta + e_i$$

$$H_0: \beta_1 = 0, \beta_2 = 0$$

$$H_1: \beta_1 \neq 0 \text{ 或 } \beta_2 \neq 0$$

用 F test (有限样本)

Wald test (大样本)

$$\tilde{\beta} = \hat{\beta} - (X'X)^{-1}R'(R(X'X)^{-1}R')^{-1}(R\hat{\beta} - c)$$

$$\tilde{e} = Y - X\tilde{\beta} = \hat{e} + (X'X)^{-1}R'(R(X'X)^{-1}R')^{-1}(R\hat{\beta} - c)$$

$$\tilde{e}'\tilde{e} = \hat{e}'\hat{e} + (R\hat{\beta} - c)'(R(X'X)^{-1}R')^{-1}(R\hat{\beta} - c)$$

$$\underbrace{\tilde{e}'\tilde{e}}_{SSR_r} = \underbrace{\hat{e}'\hat{e}}_{SSR_{ur}} + 2\hat{e}'(X'X)^{-1}R'(R(X'X)^{-1}R')^{-1}(R\hat{\beta} - c)$$

$$= \underbrace{\hat{e}'\hat{e}}_{SSR_{ur}} + \underbrace{(R\hat{\beta} - c)'(R(X'X)^{-1}R')^{-1}(R\hat{\beta} - c)}_{> 0 \Rightarrow \chi^2} \Rightarrow \boxed{SSR_r > SSR_{ur}}$$

$$F = \frac{(SSR_r - SSR_{ur})/p}{SSR_{ur}/(n-k)}$$

$$\frac{SSR_{ur}}{n-k} = s^2$$

$$= \frac{(R\hat{\beta} - c)'(R(X'X)^{-1}R')^{-1}(R\hat{\beta} - c)}{p}$$

F.p = Wald (同方差下)

$R s^2 (X'X)^{-1} R'$ — 同方差假设下, $R \hat{V} \hat{\beta} R'$.