

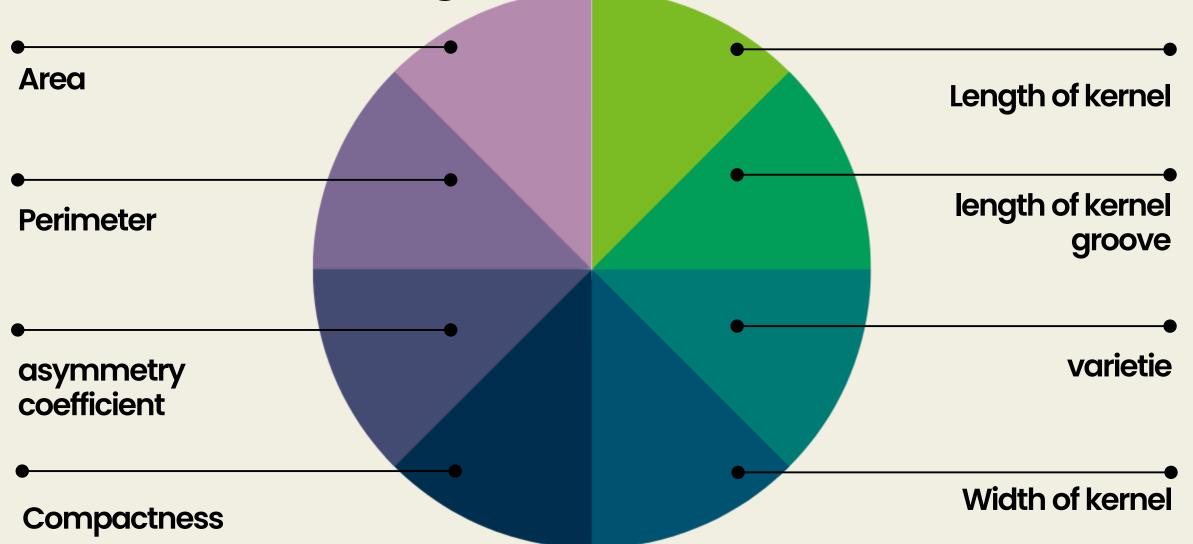
Plan

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02	Importing data
03	Data preprocessing
04	Univariate analysis
05	Bivariate analysis
06	Linear regression
07	Generalized linear regression

1 Introduction

To build our database, 8 geometric parameters

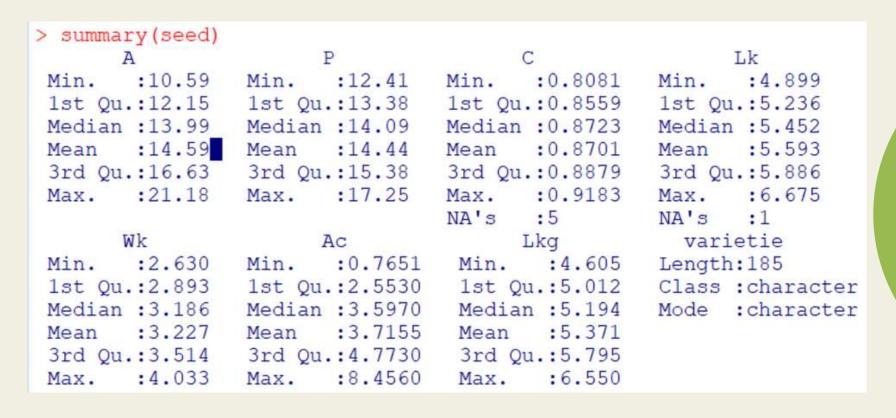
wheat grains were measured:

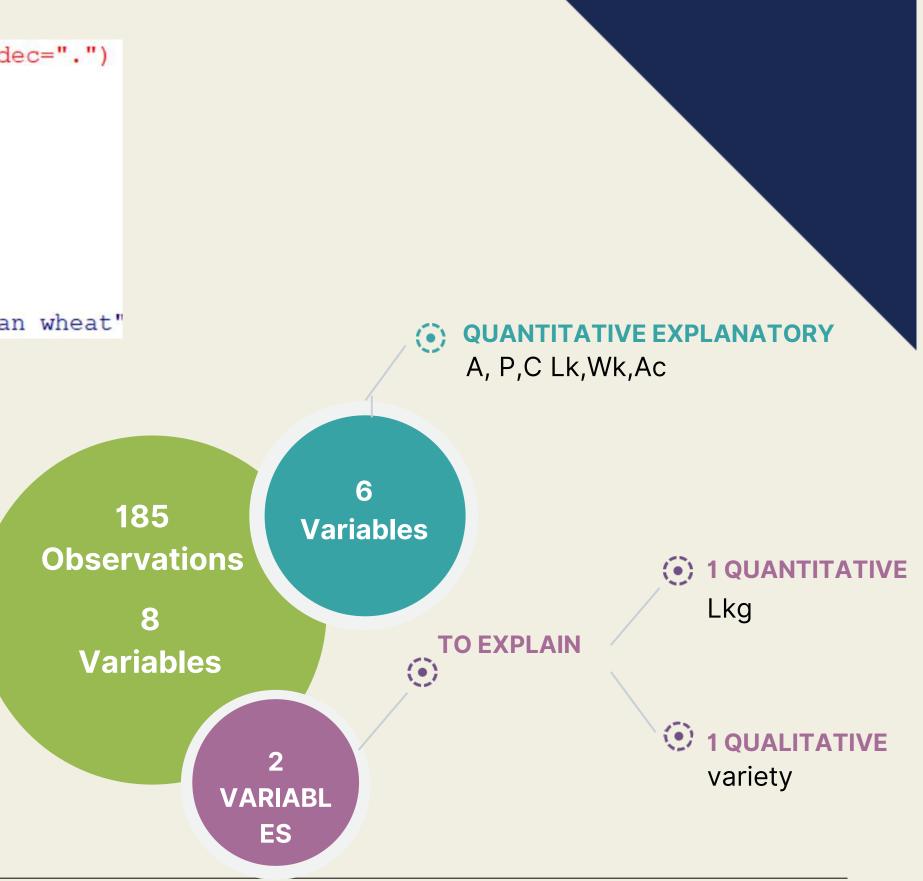


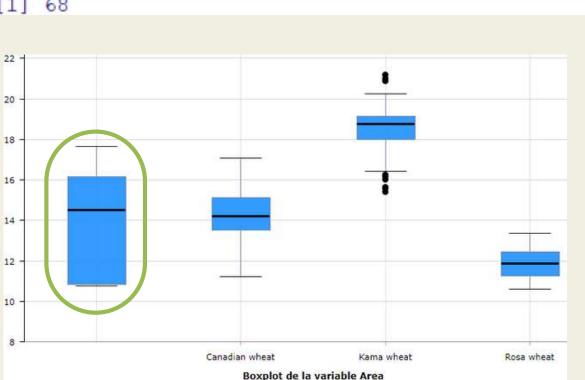
02

Importing data

```
> seed = read.table(file=file.choose(), header=TRUE, sep=", ", dec=".")
> str(seed)
'data.frame':
                185 obs. of 8 variables:
 SA
           : num 15.3 14.9 14.3 13.8 16.1 ...
 $ P
                  14.8 14.6 14.1 13.9 15 ...
           : num
 $ C
                 0.871 0.881 0.905 0.895 0.903 ...
           : num
 $ Lk
                 5.76 5.55 5.29 5.32 5.66 ...
           : num
 $ Wk
                 3.31 3.33 3.34 3.38 3.56 ...
           : num
                 2.22 1.02 2.7 2.26 1.35 ...
 $ Ac
 $ Lkq
           : num 5.22 4.96 4.83 4.8 5.17 ...
 $ varietie: chr "Canadian wheat" "Canadian wheat" "Canadian wheat"
```





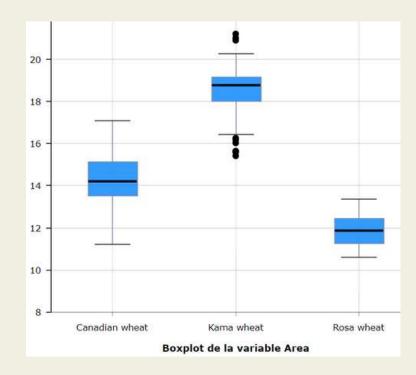


> Percent_MV_varieties = (nrow(subset(seed,varietie == ""))/nrow(seed))*100
> Percent_MV_varieties
[1] 3.243243



Visualization of outlier data by variety

Delete the lines where there are missing values in varietie.



3 modalities

1. Visualization of Outlier values:

Example Area:

```
summary(seed$A[seed$varietie=="Kama wheat"])
Min. 1st Qu. Median Mean 3rd Qu. Max.
15.38 17.99 18.76 18.43 19.14 21.18
```

```
Q1=17.99
Q3=19.14
Vmin=Q1-1.5*(Q3-Q1)
Vmax=Q3+1.5*(Q3-Q1)
IC=(Vmax-Vmin)*0.05
min_outliers=which(seed$A[seed$varietie=="Kama wheat"]<Vmin)
max_outliers=which(seed$A[seed$varietie=="Kama wheat"]>Vmax)
Area_outliers=c(min_outliers,max_outliers)
Area_outliers
for (a in Area_outliers) {
    seed$A[seed$varietie == "Kama wheat"][a]=NA
```

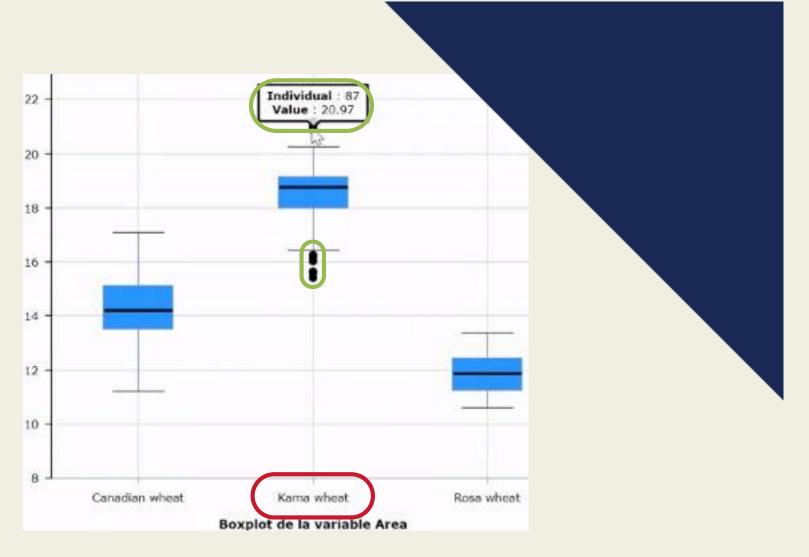
3. Check the existence of Area outliers for the Kama modality by grubbs test:

HO: There is no outlier in our Area column

H1: There is an outlier in our Area column



p-value = 0.2918 > 0.05



2. We replace the outliers with Nan:

```
> grubbs.test(seed$A[seed$varietie=="Kama wheat"], type = 11,two.sided = FALSE)

Grubbs test for two opposite outliers

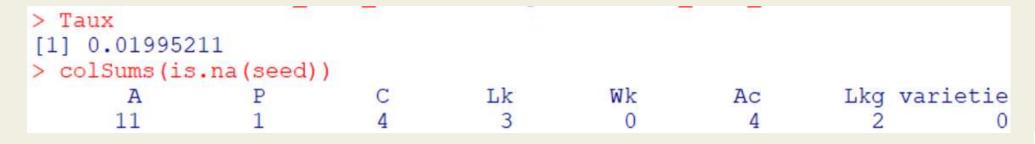
data: seed$A[seed$varietie == "Kama wheat"]

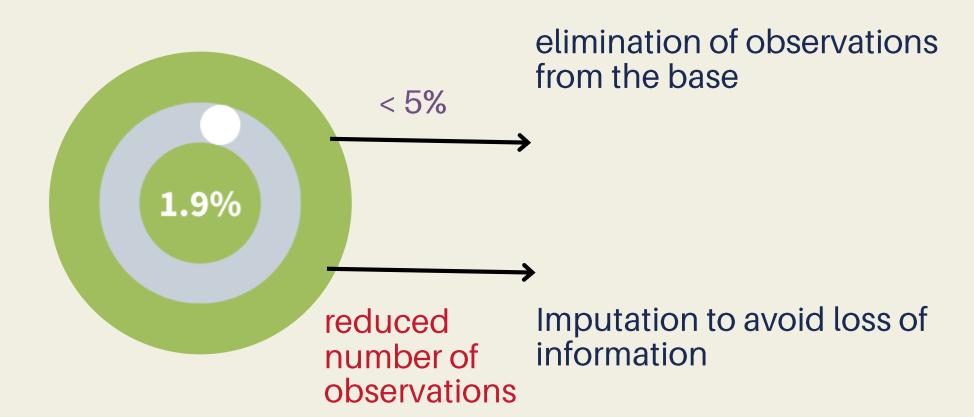
G = 4.76987, U = 0.70795, p-value = 0.2918
alternative hypothesis: 16.41 and 20.24 are outliers
```

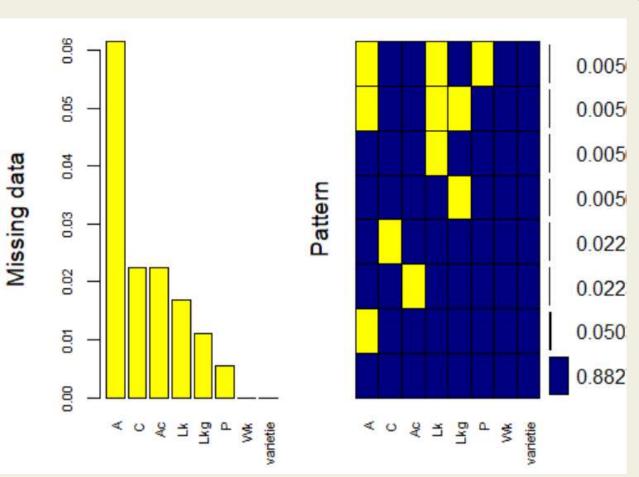




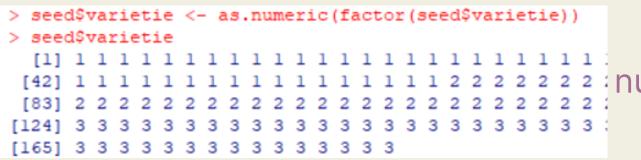
1. Visualization of Outlier values:

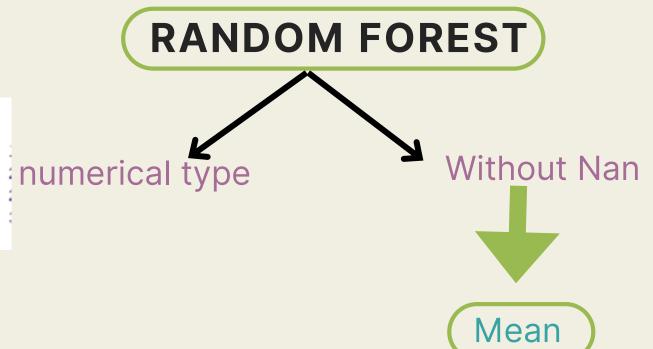






2. Processing Missing values :



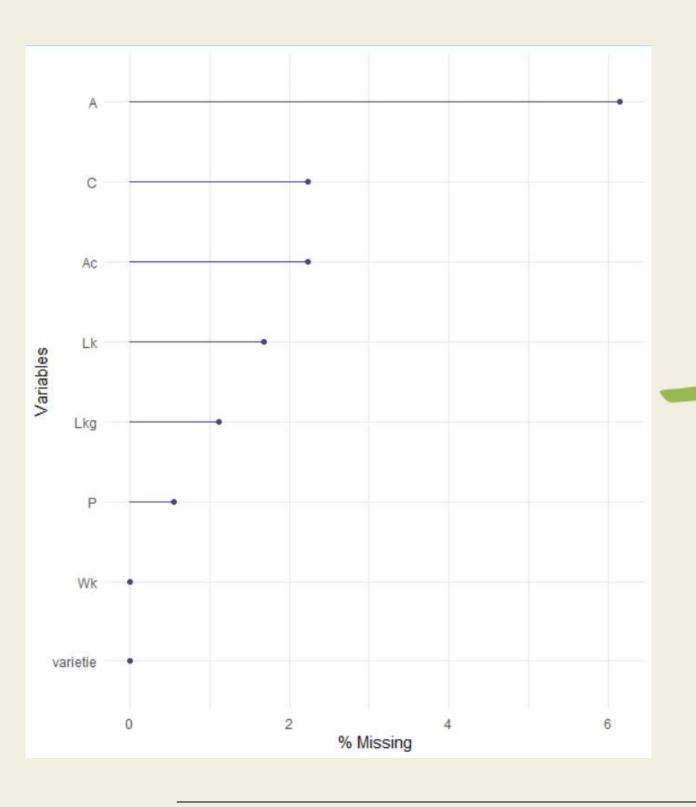


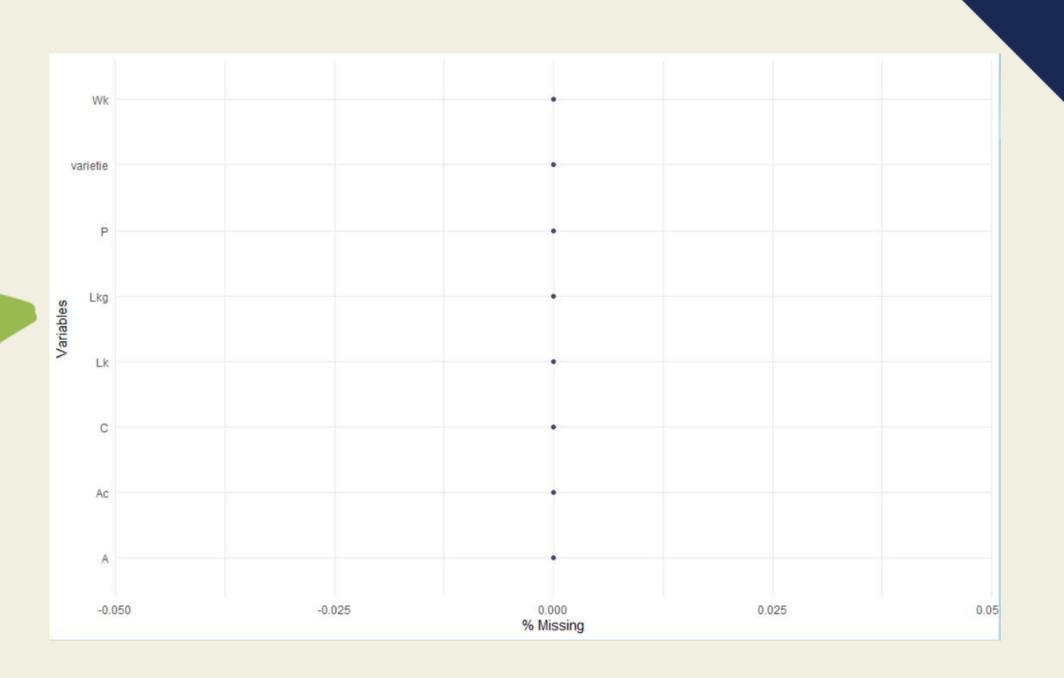


Score	Age		Score	Age	Train the Random Forest	Score	Age
98	10		98	10	Training model on the data	98	10
94	?	10 + 6 + 8	94	8	Predict Random	94	10
57	6	3	57	6	Training Forest	57	6
78	?	Impute missing	78	8	Predict Model	78	7
74	8	values using mean	64	7	Training	64	7
		to generate missing valu					

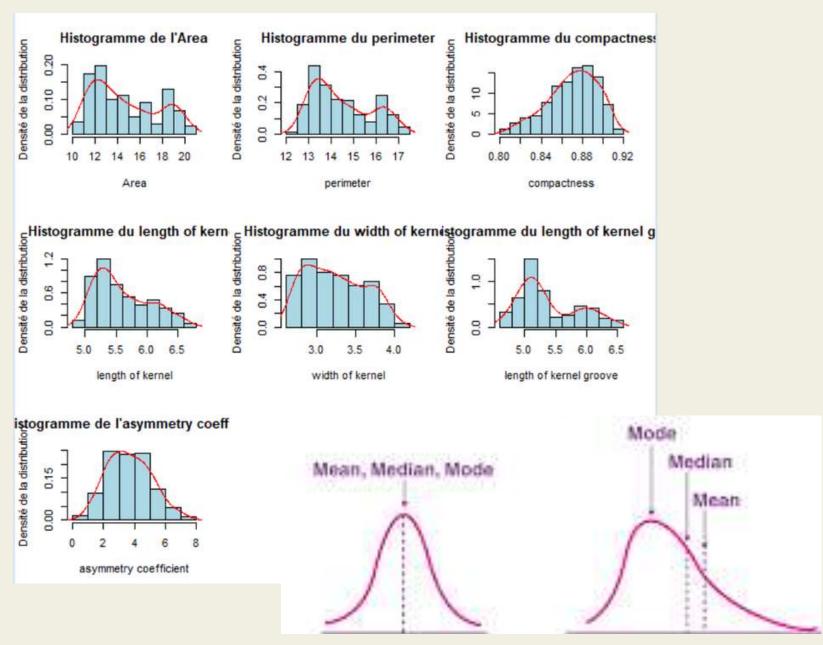
- > library(missForest)
- > seed_mf <- missForest(seed)

3. Visualization after imputation of Missing values:

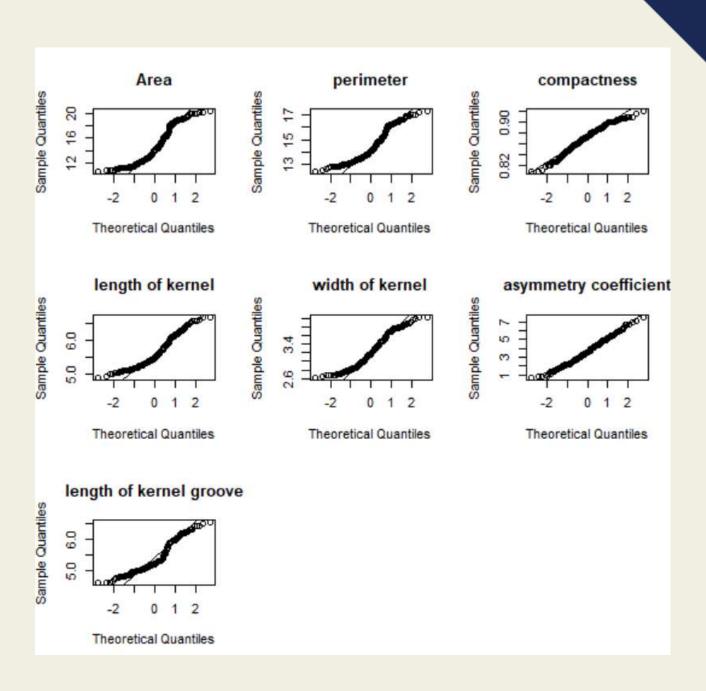




4.1- Study of normality a- Graphical representation:



Histogram

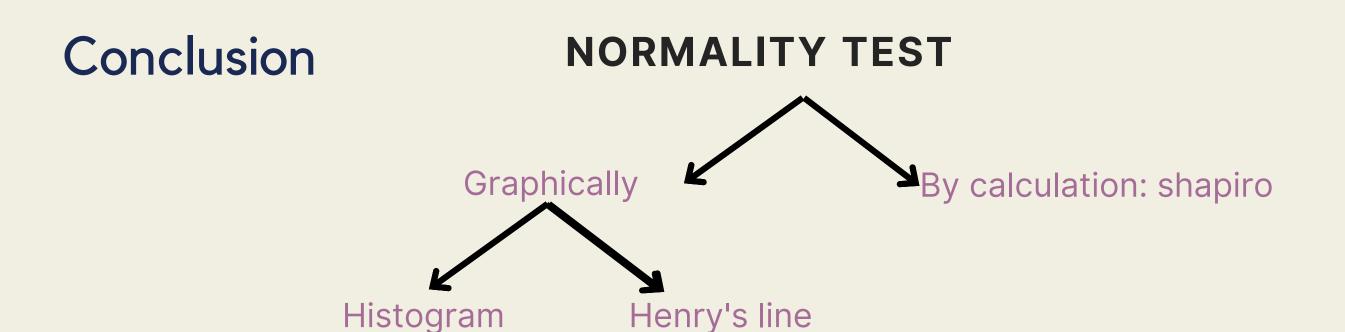


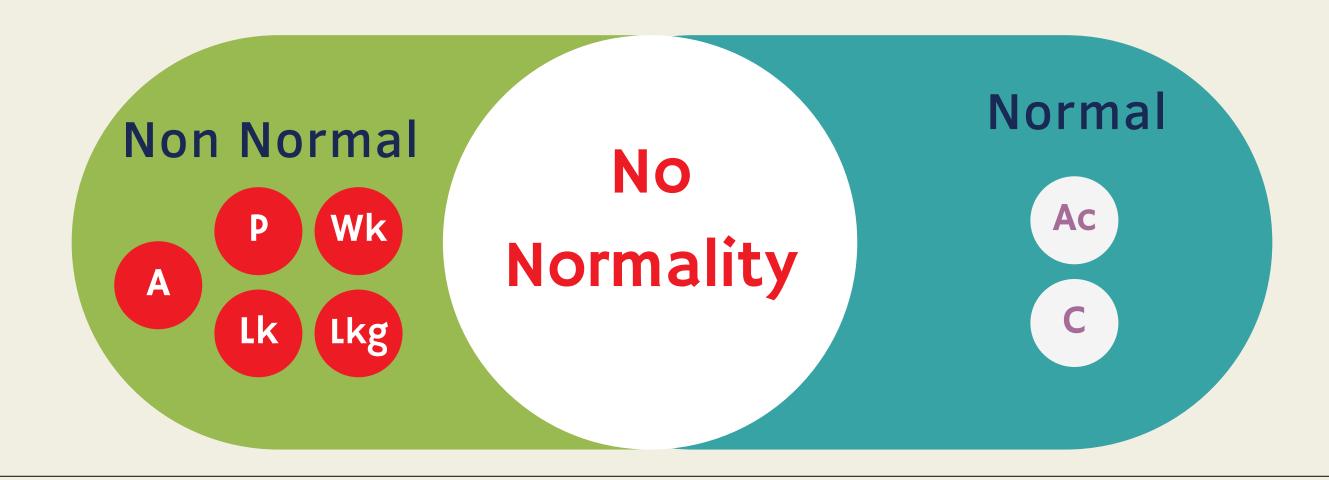
Henry's line

b- By Calculation:

Shapiro test:

```
> #-----Pour la colonne width of kernel------
> #H0 : Wk suit une normale
> #Hl : Wk ne suit pas une normale
> shapiro.test(seed imputed$Wk)
       Shapiro-Wilk normality test
data: seed imputed$Wk
W = 0.95014, p-value = 6.197e-06
> #p_value<0.05 : on rejette H0 et on accepte H1
> #La variable Wk ne suit pas la loi normale
> #-----Pour la colonne asymmetry coefficient------
> #H0 : Ac suit une normale
> #H1 : Ac ne suit pas une normale
> shapiro.test(seed imputed$Ac)
       Shapiro-Wilk normality test
data: seed imputed$Ac
W = 0.98745, p-value = 0.1122
> #p_value>0.05 : on accepte H0 et on rejette H1
> #La variable Ac suit la loi normale
```



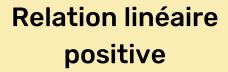


4.2- Study of the modality of the variable variety:

```
> unique(seed imputed$varietie)
[1] 1 2 3
> table(seed imputed$varietie)
1 2 3
58 53 68
> by (seed imputed, seed$varietie, nrow)
seedSvarietie: 1
[1] 58
seedSvarietie: 2
[1] 53
seedSvarietie: 3
[1] 68
> #on a trois modalité :
> #"Canadian wheat" avec 58 observations
> #"Kama wheat" avec 53 observations
> #"Rosa wheat" avec 68 observations
> #On a une database balanced : les modalités sont réparties de facon équitable
```

5.1. Study of the dependence relationship between quantitative variables.

a- Graphical representation:



A-P

A-Lk

A-Wk

A-Lkg

P-Lk

P-Wk

P-Lkg

Lk-Lkg

Wk-Lkg

Liaison monotone positive non linéaire

A-C

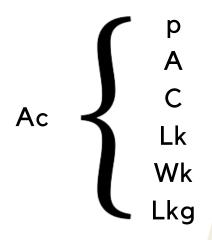
P-C

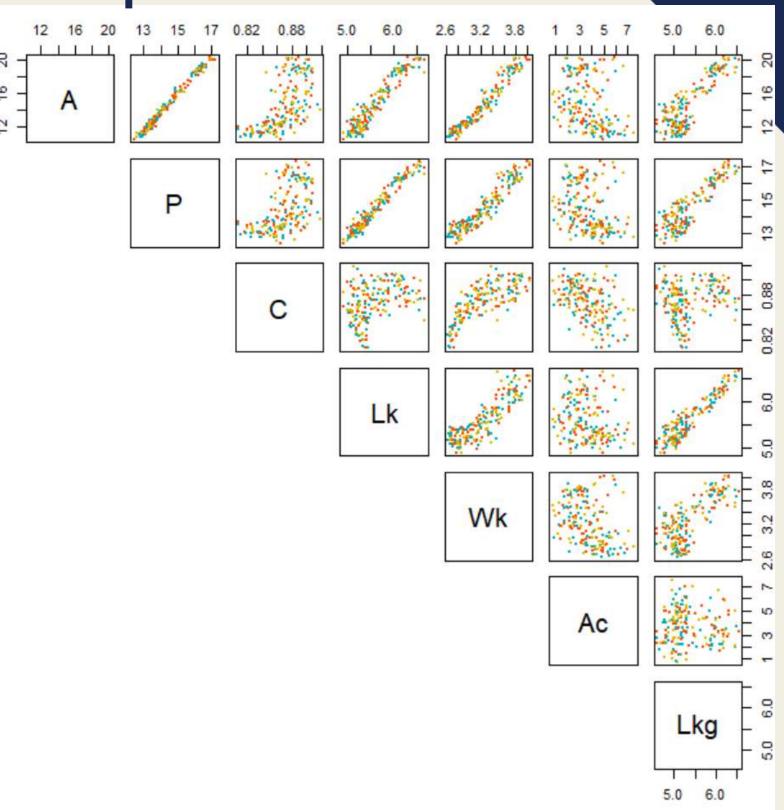
C-Lk

C-Wk

Lk-Wk

Absence de liaison

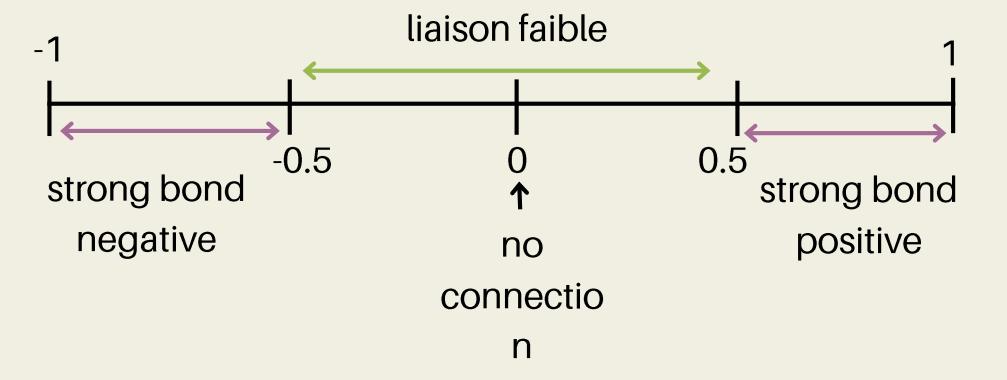




b- Calculation of statistical indicators

Spearman's correlation coefficient:

> RAC=cor(seed_imputed\$A, seed_imputed\$C, method="spearman")





> corrplot(cor(seed_imputed[,1:7], method="s"), type = "lower", method = "number")

c - Correlation test

```
#H0:r=0 => absence de dépendance
#H1:r<>0 => présence de dépendance

cor.test(seed_imputed$A, seed_imputed$P, method="spearman")

#p-value=2.2e-16<<<<<0.05 => p-value est négligeable=>accepter H1
#=>A et P sont fortement corrélées

cor.test(seed_imputed$Ac, seed_imputed$Lk, method="spearman")

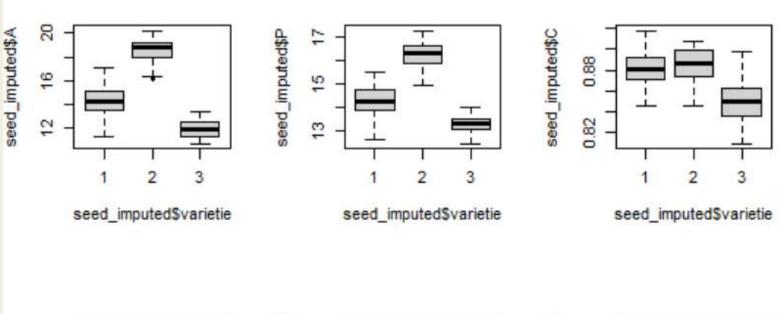
#p-value= 0.0003664 <0.05 =>accepter H1
#=>Ac et Lk sont faiblement corrélées car p-value est proche de 0.05
```

5.2- 5.2- Study of the dependence relationship between the quantitative variables and the qualitative variable:

a- Graphical representation

By changing the modality of the variable varietie, we notice a change in the reference values of the other variables.

There is an effect of the target variable varietie on all quantitative variables.



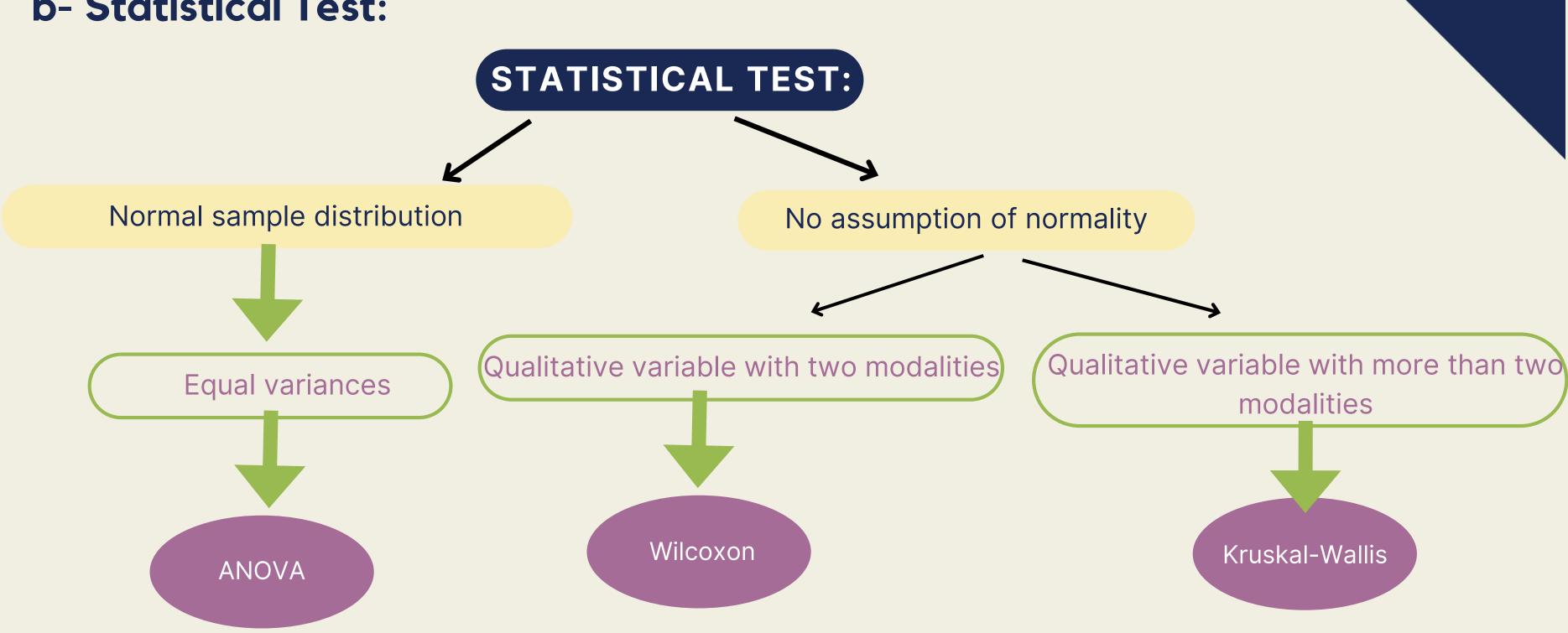
seed_imputed\$V



05

Bivariate analysis:

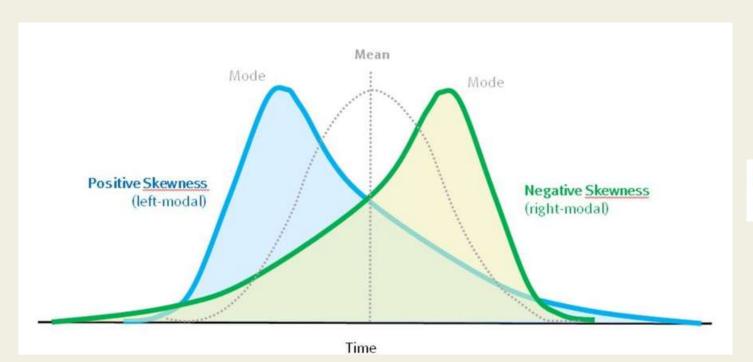
b- Statistical Test:

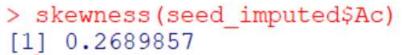


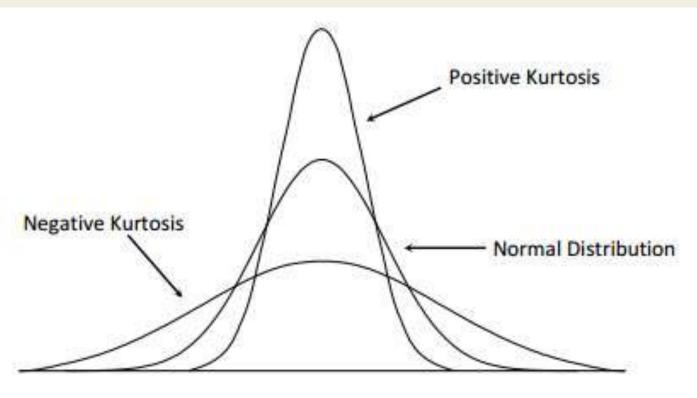


```
> #On est dans le cas de deux variables une quantitative
> #A/P/C/Lk/Wk/Ac/Lkg et une qualitative varietie
> #Le premier test possible à appliquer est l'ANOVA
> #Vérifiant alors ces conditions d'application:
> tapply(seed imputed$A, seed imputed$varietie, shapiro.test)
        Shapiro-Wilk normality test
data: X[[i]]
W = 0.99097, p-value = 0.944
$`2`
        Shapiro-Wilk normality test
data: X[[i]]
W = 0.90785, p-value = 0.0006049
$`3`
        Shapiro-Wilk normality test
data: X[[i]]
W = 0.96808, p-value = 0.0785
> #pvalue = 0.944 > 0.05, alors on a la normalité pour le groupe "canadian wheat".
> #pvalue = 0.0006571 < 0.05, alors n'on pas la normalité pour le groupe "kama wheat".
> #pvalue = 0.0785 > 0.05, alors a la normalité pour le groupe "Rosa wheat".
> #==> on n'a pas la normalité
```

```
> tapply(seed imputed$Ac, seed imputed$varietie, shapiro.test)
$'1'
       Shapiro-Wilk normality test
data: X[[i]]
W = 0.98231, p-value = 0.5567
$.2.
       Shapiro-Wilk normality test
data: X[[i]]
W = 0.95733, p-value = 0.05606
$,3,
        Shapiro-Wilk normality test
data: X[[i]]
W = 0.98957, p-value = 0.8446
> #on a la normalité
```







> kurtosis(seed_imputed\$Ac)
[1] 2.539139

Normality Shapiro.test

Equal variances Bartlet



```
> bartlett.test(seed imputed$Ac ~ seed imputed$varietie)
        Bartlett test of homogeneity of variances
data: seed imputed$Ac by seed imputed$varietie
Bartlett's K-squared = 1.381, df = 2, p-value = 0.5013
> #p-value = 0.5013>0.05 =>On accepte H0
> #==> On peut appliquer le test statistique ANOVA
> #HO: variables indépendantes
> #H1: présence de dépendance entre les deux varibles
> fit <- aov(seed imputed$Ac ~ seed imputed$varietie)
> summary(fit)
                      Df Sum Sq Mean Sq F value Pr(>F)
                      1 140.8 140.77 112.6 <2e-16 ***
seed imputed$varietie
Residuals
                      177 221.2
                                 1.25
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
> #p-value<2e-16 <0.001
> # On accepte H1
> # au moins un niveau(groupe) avec une moyenne significativement différente
  # Il existe un effet de varietie sur Ac
```

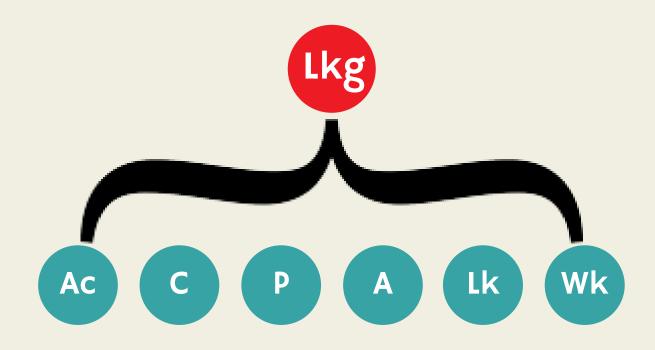








6.1- Regression of the quantitative target variable Lkg according to the others:



(R squared) = Sum of Squares Explained/ Sum of Squares Total

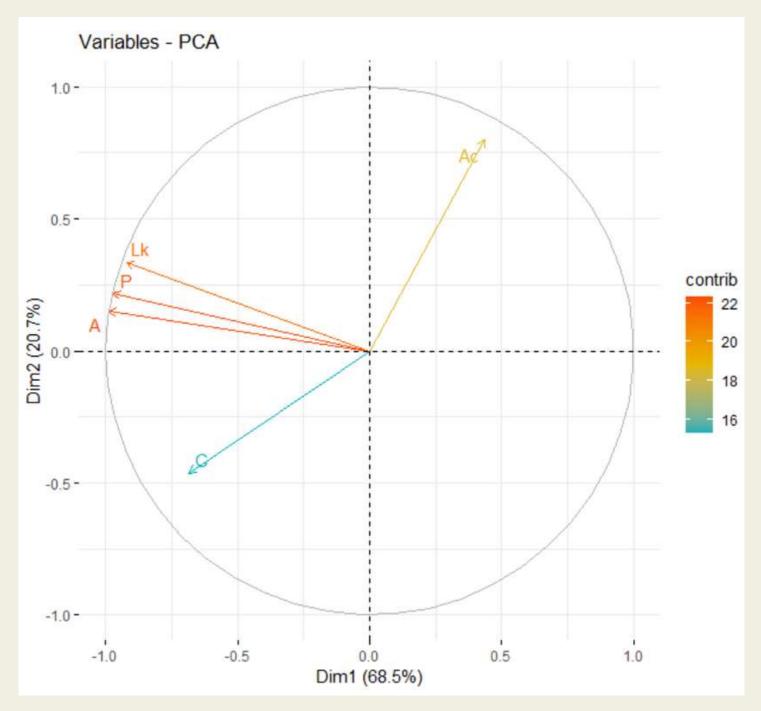
```
> D=lm(Lkg ~ A+P+C+Lk+Wk+Ac , data=seed imputed)
> summary(D)
Call:
lm(formula = Lkg ~ A + P + C + Lk + Wk + Ac, data = seed imputed)
Residuals:
    Min
              1Q Median
                                        Max
-0.41102 -0.08040 0.00099 0.09716 0.29221
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 11.903466
                       2.172037 5.480 1.49e-07 ***
                       0.057306 7.299 1.02e-11 ***
            0.418297
                       0.145726 -4.300 2.86e-05 ***
           -0.626575
                       1.965942 -3.698 0.000292 ***
           -7.269379
                       0.156694 4.294 2.92e-05 ***
Lk
            0.672920
Wk
           -0.372342
                       0.295683 -1.259 0.209642
                       0.008604 5.450 1.72e-07 ***
            0.046892
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.136 on 172 degrees of freedom
Multiple R-squared: (0.9236,)
                               Adjusted R-squared: 0.9209
F-statistic: 346.6 on 6 and 172 DF, p-value: < 2.2e-16
```

6.2- Improvement of the performance of the regression model:

```
> DSWk=lm(Lkg ~ A+P+C+Lk+Ac ,data=seed imputed)
> summary (DSWk)
Call:
lm(formula = Lkg ~ A + P + C + Lk + Ac, data = seed imputed)
Residuals:
    Min
              10 Median
-0.41296 -0.08306 -0.00036 0.09821 0.29515
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 13.552810 1.735628 7.809 5.31e-13 ***
            0.412312 0.057205 7.208 1.69e-11 ***
           -0.711640 0.129342 -5.502 1.33e-07 ***
           -9.288077 1.139932 -8.148 7.13e-14 ***
          0.715176 0.153318 4.665 6.16e-06 ***
           0.042825 0.007988 5.361 2.62e-07 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.1362 on 173 degrees of freedom
Multiple R-squared: 0.9229,
                             Adjusted R-squared: 0.9207
F-statistic: 414.2 on 5 and 173 DF, p-value: < 2.2e-16
> #Le modèle ne perd pas de qualité
> #=> on a toujours R2=0.9229
> #le même taux d'information expliquée de la variabilité de Lkg.
> #la variable éliminée n'a pas d'importance
```

```
> DSLk=lm(Lkg ~ A+P+C+Ac , data=seed imputed)
> summary(DSLk)
Call:
lm(formula = Lkg ~ A + P + C + Ac, data = seed imputed)
Residuals:
    Min
              10 Median
-0.37028 -0.09825 -0.00697 0.09316 0.36179
Coefficients:
             Estimate Std. Error t value Pr(>|t|)
(Intercept) 16.325692 1.725188 9.463 < 2e-16 ***
                        0.060487 6.966 6.44e-11 ***
             0.421362
            -0.467177 0.125103 -3.734 0.000255 ***
           -12.081246 1.026227 -11.772 < 2e-16 ***
             0.041996 0.008449 4.971 1.59e-06 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 '' 1
Residual standard error: 0.1441 on 174 degrees of freedom
Multiple R-squared: (0.9132, Adjusted R-squared: 0.9112
F-statistic: 457.7 on 4 and 174 DF, p-value: < 2.2e-16
> # R2: 0.9132 =>1e modèle perd de qualité
> #=>la varible Lk est significative
> #=>On va considérer le modèle précédant DSWk comme modèle optimal
```

6.3- principal component analysis:



Wk: insignificant variable (According to linear regression)

Lkg: the target quantitative variable

varietie: target qualitative variable

To eliminate variables to be projected by the PCA

```
> results$rotation

PC1 PC2 PC3 PC4 PC5

A -0.5329520 0.1479898 -0.02333715 0.460678135 0.69375226

P -0.5256544 0.2132081 0.07541820 0.402892332 -0.71429657

C -0.3697605 -0.4569437 -0.77633620 -0.216961097 -0.06862654

Lk -0.4975099 0.3273642 0.25137922 -0.760494534 0.06142566

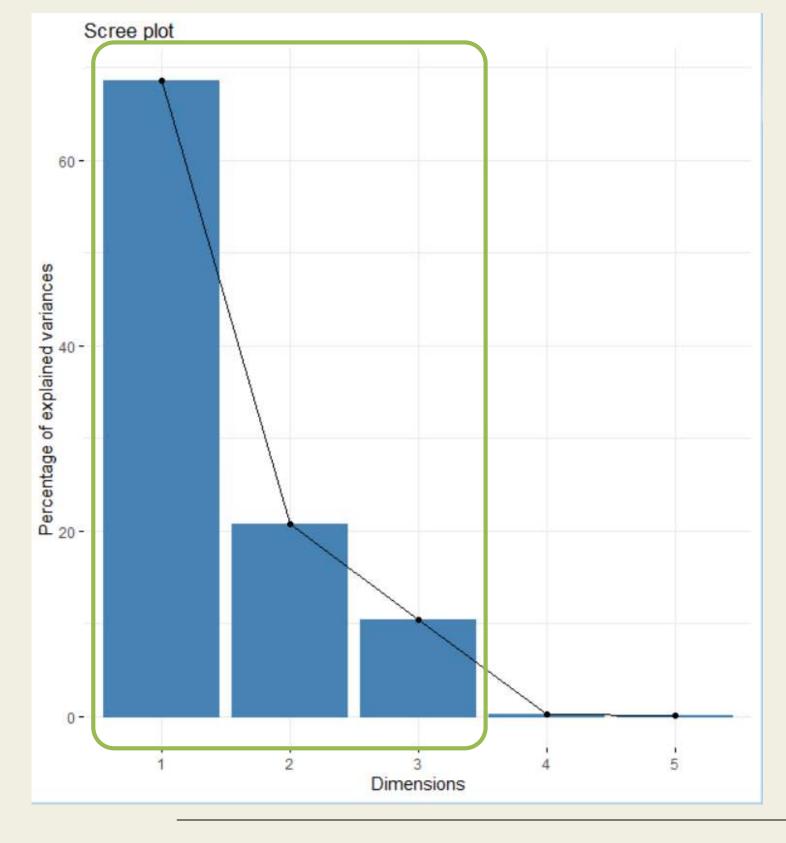
Ac 0.2353946 0.7852875 -0.57260638 -0.005419556 -0.00234499
```



The first 5 principal components



6.3- Principal component analysis:



This graph represents the percentages of variances explained by each main axis.

Les trois premières dimensions sont les variances expliquées les plus importantes.



6.3- Principal component analysis:

```
> NEW DATA=results$x[, 1:3]
> NEW MODEL=lm(seed imputed$Lkg ~ NEW DATA )
> summary (NEW MODEL)
Call:
lm(formula = seed imputed$Lkg ~ NEW DATA)
Residuals:
     Min
              1Q Median
-0.46235 -0.09641 0.00799 0.09683 0.35030
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 5.372542 0.011340 473.761 < 2e-16 ***
NEW DATAPC1 -0.207583 0.006143 -33.790 < 2e-16 ***
NEW DATAPC2 0.232842 0.011184 20.818 < 2e-16 ***
NEW DATAPC3 0.120737 0.015716 7.682 1.07e-12 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.1517 on 175 degrees of freedom
Multiple R-squared: 0.9033, Adjusted R-squared: 0.9016
F-statistic: 344.7 on 3 and 175 DF, p-value: < 2.2e-16
> #R-squared: 0.9032
> #90.32% de la variation de Lkg est expliquée par les variables résultantes
> #de la PCA
> #Malgré la diminution de R2, il reste un bon modèle
>
```



Generalized linear regression:

GLM is a flexible generalization of linear regression.

The dependent variable is linearly related to the variables via a precise link function.

Why we use GLMs



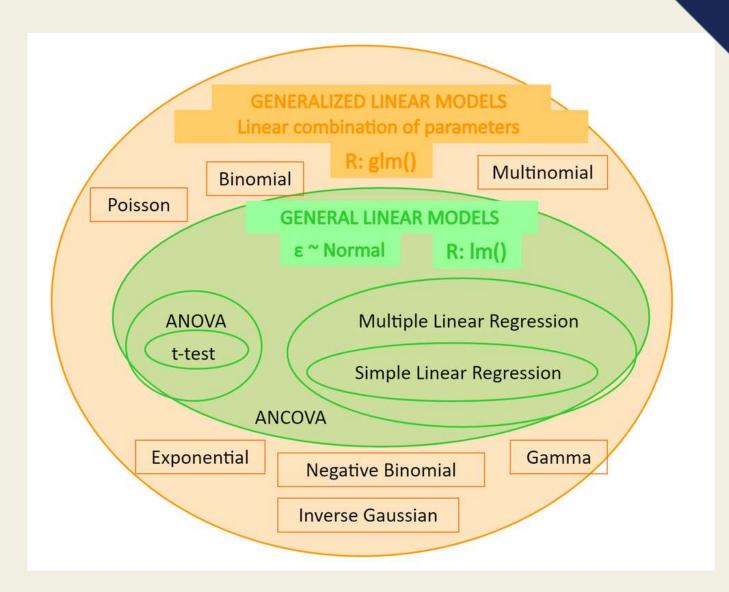
They provide a way to model dependent variables:

Non-normal distributions

no linear relationship with the explanatory variables

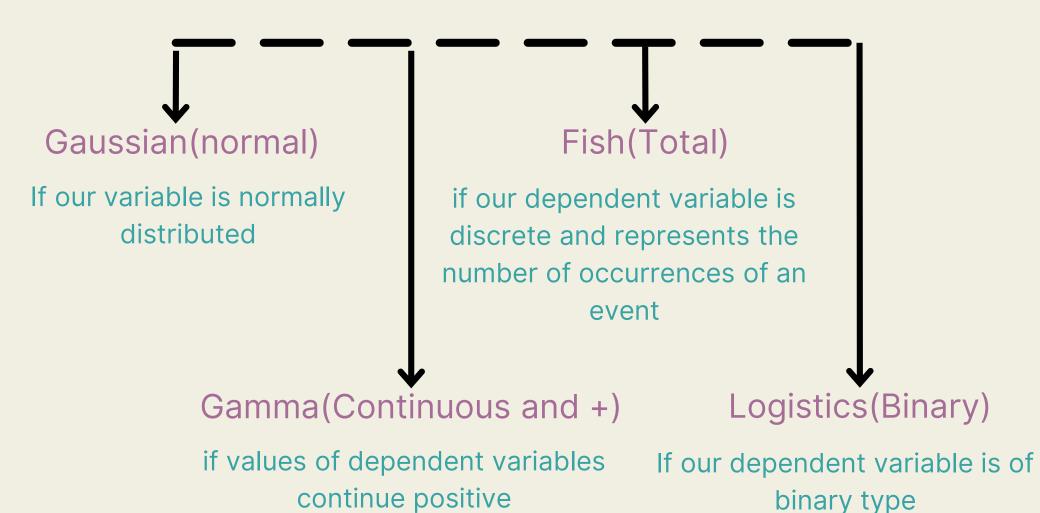
No need to check:

- The normality of the residuals
- Homoscedasticity
- That the dependent variables are continuous



Generalized linear regression:

TYPE OF GENERALIZED REGRESSION



Input-output Error (residual) Link function Examples Meaning of the of Y coefficients relationship distribution and inverse x₁, x₂, ..., x_n Left Identity Gaussian Differences Ventricular Mass, LVM Identity Risk of a Binomial Odds Ratios Binary Event Rates of a Poisson Rate Ratios Count Event

Generalized linear regression:

```
> GENERALIZED MODEL <- glm(Lkg ~ A+P+C+Lk+Wk+Ac, data=seed imputed, family=Gamma())
> #gamma : valeurs des variables dépendantes continue positifs
> summary (GENERALIZED MODEL)
Call:
glm(formula = Lkg ~ A + P + C + Lk + Wk + Ac, family = Gamma(),
   data = seed imputed)
Deviance Residuals:
     Min
                      Median
                                                Max
-0.074866 -0.016248 0.001226 0.017230 0.058137
                                                            > AIC (GENERALIZED MODEL)
Coefficients:
                                                            [1] -193.1547
             Estimate Std. Error t value Pr(>|t|)
                                                            > AIC (NEW MODEL)
(Intercept) -0.0280443 0.0732571 -0.383 0.702326
                                                          [1] -161.5193
           -0.0128009 0.0018663 -6.859 1.20e-10 ***
                                                          > #AIC (GENERALIZED MODEL) < AIC (NEW MODEL)
          0.0190367 0.0048757 3.904 0.000135 ***
                                                            > #==> Le modèle linéaire généralisé a une qualité d'ajustement meillsseure
          0.2558217 0.0676622 3.781 0.000215 ***
                                                            > ‡que le modèle obtenu par la modélisation linéaire des données obtenues
         -0.0211096 0.0052306 -4.036 8.18e-05 ***
                                                            > #de la PCA
          0.0088303 0.0101111 0.873 0.383702
           -0.0015049 0.0003017 -4.987 1.49e-06 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
(Dispersion parameter for Gamma family taken to be 0.000658218)
   Null deviance: 1.39491 on 178 degrees of freedom
Residual deviance: 0.11374 on 172 degrees of freedom
AIC: -193.15
Number of Fisher Scoring iterations: 3
```

