

Week 6 - Lecture 28

$$\Pr \{ \hat{\mu} \geq \mu + \varepsilon \} \leq \sqrt{\frac{\sigma^2}{2\pi n \varepsilon^2}} \exp \left\{ \frac{-n \varepsilon^2}{2\sigma^2} \right\}$$

For sufficiently large n
[using CLT]

$$\Pr \{ \hat{\mu} \geq \mu + \varepsilon \} \leq \frac{\sigma^2}{n \varepsilon^2} \quad [\text{using Chebyshev}]$$

SubGaussian Random Variables

A r.v. X is σ -subgaussian if $\forall \lambda \in \mathbb{R}$
it holds that

$$\mathbb{E} [e^{\lambda X}] \leq \exp(\lambda^2 \sigma^2 / 2) \quad \text{--- (1)}$$

Let $X \sim \mathcal{N}(\mu, \sigma^2)$

$$\mathbb{E} [e^{\lambda X}] = \exp \left\{ \mu \lambda + \lambda^2 \sigma^2 / 2 \right\}$$

If $\mu = 0$, then :

$$\mathbb{E} [e^{\lambda X}] = \exp \left\{ \lambda^2 \sigma^2 / 2 \right\}$$

Thus $\mathcal{N}(0, \sigma^2)$ is σ -subgaussian

Also,

$$\log \mathbb{E}[e^{\lambda x}] \leq \lambda^2 \sigma^2 / 2$$

If $X \sim \exp(\mu)$

$$\mathbb{E}[e^{\lambda x}] = \mu \int_0^{\infty} e^{\lambda x} e^{-\mu x} dx$$

$$= \left(\frac{\mu}{\lambda - \mu} \right) e^{(\lambda - \mu)x} \Big|_0^{\infty}$$

If $\lambda > \mu$ then this quantity $\sim \infty$

If $\lambda < \mu$ then this quantity $= 0$

$\therefore X \sim \exp(\mu)$ is not σ -subgaussian r.v.

R.v. X is called heavy tailed if

$$\log \mathbb{E}[e^{\lambda x}] = \infty \quad \forall \lambda > 0$$

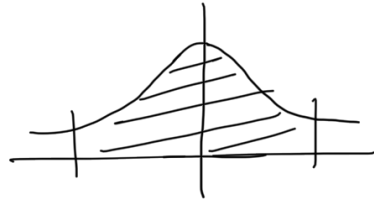
otherwise it is called light-tailed.

Focus will be on light-tailed r.v.

$X \sim \mathcal{N}(0, \sigma^2)$ is σ -subgaussian r.v.

support of this is $-\infty$ to $+\infty$ (unbounded support)

should be light-tailed which means:



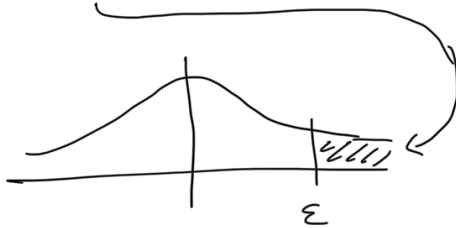
As we go towards the tails the probability reduces, probability content is less in the tails

Thus subgaussian r.v. can have unbounded support but should be light-tailed.

Theorem: If X is σ -subgaussian

then $\forall \epsilon \geq 0$

$$\Pr\{X \geq \epsilon\} \leq \exp\left\{-\frac{\epsilon^2}{2\sigma^2}\right\} \quad \text{--- (2)}$$



This probability is upper bounded

Proof:

$$\Pr\{X \geq \epsilon\} = \Pr\{e^X \geq e^\epsilon\}$$

\uparrow \uparrow
 X can be \uparrow is +ve r.v.
 +ve or -ve

$$= \Pr\{e^{\lambda X} \geq e^{\lambda \epsilon}\}$$

$$\leq \frac{\mathbb{E}[e^{\lambda X}]}{e^{\lambda \epsilon}} \quad \left[\text{Markov inequality} \right]$$

Now if X is σ -subgaussian, then

$$\leq \exp \left\{ \lambda^2 \sigma^2 / 2 - \lambda \varepsilon \right\} \quad \text{is true for any } \lambda > 0$$

[from eq ①
defn of subgaussian]

Optimizing for λ and plugging in the value

$$\leq \exp \left\{ -\varepsilon^2 / 2\sigma^2 \right\} \quad \left[\begin{array}{l} \text{optimal value} \\ \text{of } \lambda = \varepsilon / \sigma^2 \end{array} \right]$$

Hence proved

Theorem :

$$\Pr \{ X \leq -\varepsilon \} \leq \exp \left\{ -\varepsilon^2 / 2\sigma^2 \right\} \quad \text{--- (2a)}$$

$$\therefore \Pr \{ |X| \geq \varepsilon \} = \Pr \{ X \geq \varepsilon \} + \Pr \{ X \leq -\varepsilon \}$$

$$\Pr \{ |X| \geq \varepsilon \} = 2 \exp \left\{ -\varepsilon^2 / 2\sigma^2 \right\} \quad \text{--- (3)}$$

Let $\varepsilon = \sqrt{2\sigma^2 \log 1/\delta}$, then :

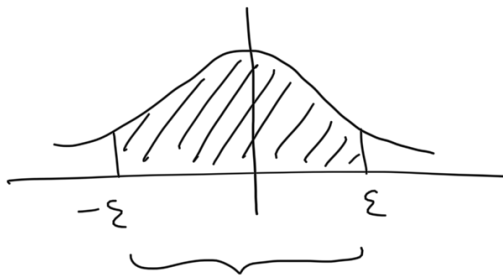
$$\Pr \left\{ X \geq \sqrt{2\sigma^2 \log 1/\delta} \right\} \leq \exp \left\{ \frac{-2\sigma^2 \log 1/\delta}{2\sigma^2} \right\}$$

$$\leq \delta$$

∴ For any given δ

$$X \in \left(-\sqrt{2\sigma^2 \log 1/\delta}, \sqrt{2\sigma^2 \log 1/\delta} \right)$$

will happen with probability $1 - 2\delta$



$$X \in (-\varepsilon, \varepsilon)$$

For small δ ,
this probability will
be large.

For small δ , value of ε
 $= \sqrt{2\sigma^2 \log 1/\delta}$ will be
 bigger and hence the
 interval $-\varepsilon, +\varepsilon$ will be larger, will
 be far apart, hence makes sense
 that probability will be large.

We are finally interested in knowing about the
estimators.

So if r.v. are subgaussian then what will be:

$$\left[\Pr\{\hat{\mu} \geq \mu + \varepsilon\} \quad \text{where } \hat{\mu} = \frac{1}{n} \sum_{i=1}^n X_i \right]$$

Properties of subgaussian r.v. :

Lemma: Suppose X is σ -subgaussian and X_1 & X_2 are σ_1 and σ_2 -subgaussian independent and respectively, then:

$$1) \mathbb{E}[X] = 0, \text{Var}[X] \leq \sigma^2$$

$$2) cX \text{ is } |c|\sigma\text{-subgaussian } \forall c \in \mathbb{R}$$

$$3) X_1 + X_2 \text{ is } \sqrt{\sigma_1^2 + \sigma_2^2}\text{-subgaussian}$$

So now coming back to

$$\Pr\{\hat{\mu} \geq \mu + \varepsilon\} \quad \text{where } \hat{\mu} = \frac{1}{n} \sum_{i=1}^n X_i$$

$$\hat{\mu} - \mu = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)$$

Theorem: Assume $X_i - \mu$ are independent and σ -subgaussian r.v.s. Then $\forall \varepsilon \geq 0$

$$\Pr\{\hat{\mu} \geq \mu + \varepsilon\} \leq \exp\left\{-n\varepsilon^2/2\sigma^2\right\}$$

$$\Pr\{\hat{\mu} \leq \mu - \varepsilon\} \leq \exp\left\{-n\varepsilon^2/2\sigma^2\right\}$$

$$\therefore \Pr\{\hat{\mu} \geq \mu + \varepsilon\}$$

$$= \Pr \left\{ \frac{1}{n} \sum_{i=1}^n (x_i - \mu) \geq \varepsilon \right\}$$

$$= \Pr \left\{ \sum_{i=1}^n \frac{(x_i - \mu)}{n} \geq \varepsilon \right\}$$

$$\therefore x_i - \mu \sim \frac{\sigma}{n} \text{ subgaussian}$$

We are hence taking sum of n r.v.s. with $\frac{\sigma}{n}$ - subgaussian

$$\therefore \sum_{i=1}^n \frac{(x_i - \mu)}{n} \sim \sqrt{\sum_{i=1}^n \left(\frac{\sigma}{n}\right)^2} = \sigma/\sqrt{n}$$

$$\therefore \sum_{i=1}^n \frac{(x_i - \mu)}{n} \text{ is } \sigma/\sqrt{n} \text{ - subgaussian}$$

Putting this into the defn of subgaussian r.v. we get

$$\Pr \{ \hat{\mu} \geq \mu + \varepsilon \} \leq \exp \left\{ \frac{-\varepsilon^2}{2(\sigma/\sqrt{n})^2} \right\}$$

$$\Pr \{ \hat{\mu} \geq \mu + \varepsilon \} \leq \exp \left\{ \frac{-n\varepsilon^2}{2\sigma^2} \right\}$$

Now we have a bound which is exponentially decaying in n , and this holds for any n .

When we had proved with CLT n had to be sufficiently large. but with subgaussian assumption

the bound holds for any n .

$$\Pr\{\hat{\mu} \geq \mu + \varepsilon\} \leq \frac{\sigma^2}{n\varepsilon^2} \quad [\text{Chebyshev}]$$

$$\leq \exp\left\{-\frac{\varepsilon^2 n}{2\sigma^2}\right\} \quad [\text{Subgaussian}]$$

↑

this is tighter than Chebyshev

$$\text{Also for } \forall x > 0, e^{-x} < \frac{1}{e^x}$$

$$\therefore \Pr\{\hat{\mu} \geq \mu + \varepsilon\} \leq \frac{2}{e^{\varepsilon^2 n}}$$

$$\leq \frac{\sigma^2}{\varepsilon^2 n}$$