



# Linear Regression: Frequentist vs. Bayesian approaches

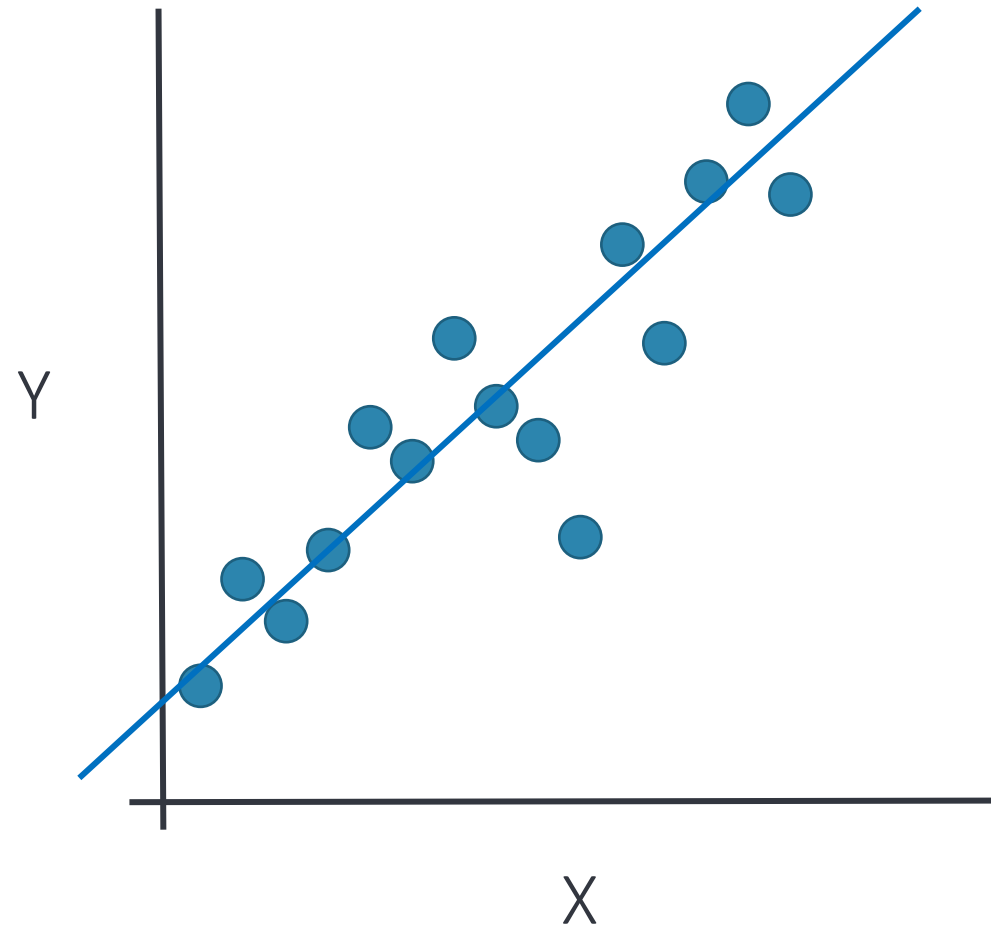
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# Terminology used in statistics and for regression

- X values can be called...
  - Covariates
  - Predictors
  - Explanatory variables
  - Independent variables
  - Features (computer science/machine learning)
- Y values can be called
  - Response variable
  - Dependent variable
  - Labels (ML)

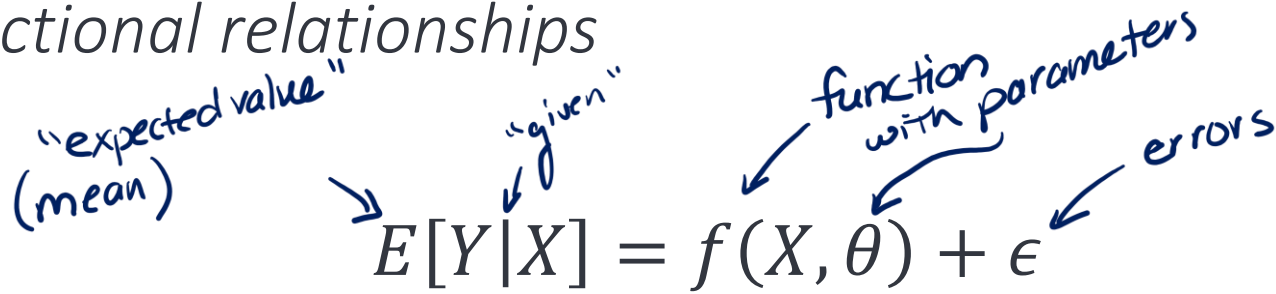


# Regression

- Estimate functional relationships
- Many types!
  - Linear regression, generalized linear models (GLMs), lasso regression, Poisson regression, logistic regression, ...
- “standard” regression assumes  $X$  is fixed without error
- **Linear regression does not imply we are fitting a line**
  - E.g. “linear” regression means *linear in the parameters*

# Concept of Regression

- *Estimating functional relationships*



The diagram shows the regression equation  $E[Y|X] = f(X, \theta) + \epsilon$  with handwritten annotations. An arrow points from the text "expected value (mean)" to  $E[Y|X]$ . Another arrow points from the text "given" to  $X$ . A third arrow points from the text "function with parameters" to  $f(X, \theta)$ . A fourth arrow points from the text "errors" to  $\epsilon$ .

$$E[Y|X] = f(X, \theta) + \epsilon$$

- Note the asymmetry in most regression analysis. This is not a fit to the joint distribution of  $(Y, X)$
- *Homoscedastic errors*:  $\epsilon$  is an  $n$ -vector with  $\sigma^2$
- *Heteroscedastic errors*:  $\epsilon$  is an  $n$ -vector with  $\sigma_i^2$  (*known or unknown*)
- Errors-in-Variables models assume  $X$  has error as well

# Are these models linear or non-linear?

Example

Linear or Non-Linear?

- $Y = \underline{\beta_0} + \underline{\beta_1}X + \underline{\beta_2}e^X + \epsilon$

linear

- $Y = \left(\frac{X}{\beta_0}\right)^{-\beta_1} + \epsilon$

non-linear

- $Y = \beta_0 + \beta_1 \sin(X) + \beta_2 \cos(X) + \epsilon$

linear

- $Y = \beta_0 e^{-\beta_1 X} + \epsilon$

$$y = \beta_0 x + e^{-\beta_1 x}$$

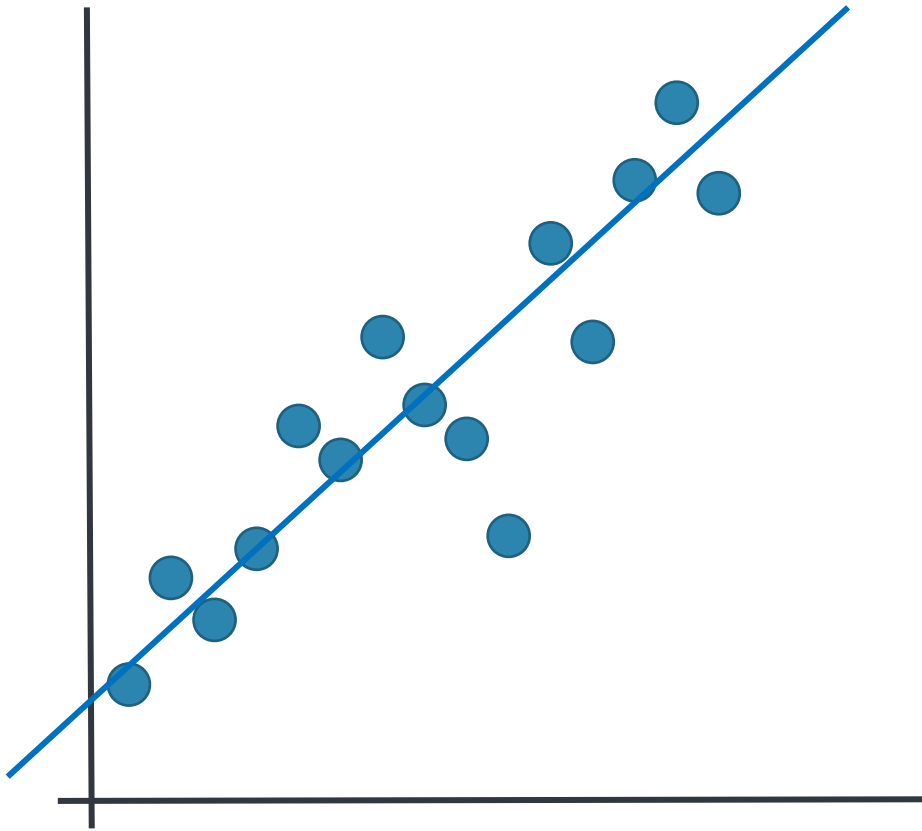
non-linear

# Conditions for linear regression

- Nearly normal residuals
  - Residuals should be normally distributed about 0
  - No trends in residuals
  - No major outlier(s) or "influential points" (we'll get to this!)
- Constant variability
  - Variability above/below least squares line shouldn't change as  $x$  changes
- Independent observations
  - e.g., typically don't apply to time series data

# Frequentist:

## Fitting a line using Ordinary Least Squares (OLS) Regression



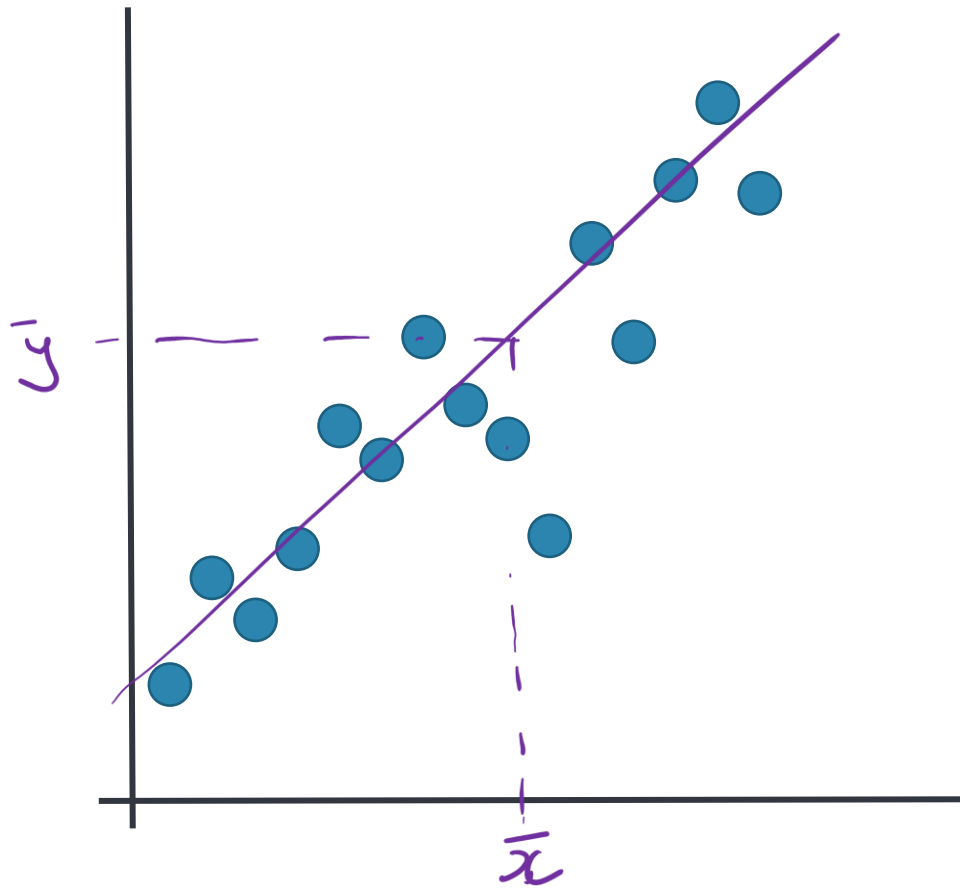
- Useful when the relationship between two quantities can be summarized by a straight line
- Correlation describes the *strength* of the correlation between x and y, whereas the regression line is used to describe the relationship between x and y
- The regression line is a model which follows the equation:

$$y = \underset{\substack{\uparrow \\ \text{parameters}}}{\beta_0} + \overset{\text{intercept}}{\underset{\substack{\uparrow \\ \text{slope}}}{\beta_1}} x + \varepsilon$$

$$\hat{y} = b_0 + b_1 x \quad (\text{fitted line, } \hat{y} \text{ predicted expected values})$$

# Frequentist:

## Fitting a line using Ordinary Least Squares (OLS) Regression



- A regression line can tell you something about the effect of the predictor or independent variable on the response variable

$$\hat{y} = b_0 + b_1 x$$

- Slope of a regression line is related to the correlation of the points:

$$b_1 = r \frac{s_y}{s_x}$$

← sample standard deviation of y  
correlation ← " " " " x

- Intercept of a regression line is:

$$b_0 = \bar{y} - b_1 \bar{x}$$

the line passes through  $(\bar{x}, \bar{y})$



# Relationship between correlation and slope (for simple L.R.)

Correlation  $r$  between  $x$  and  $y$  :

$$r = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{(n-1) s_x s_y}$$

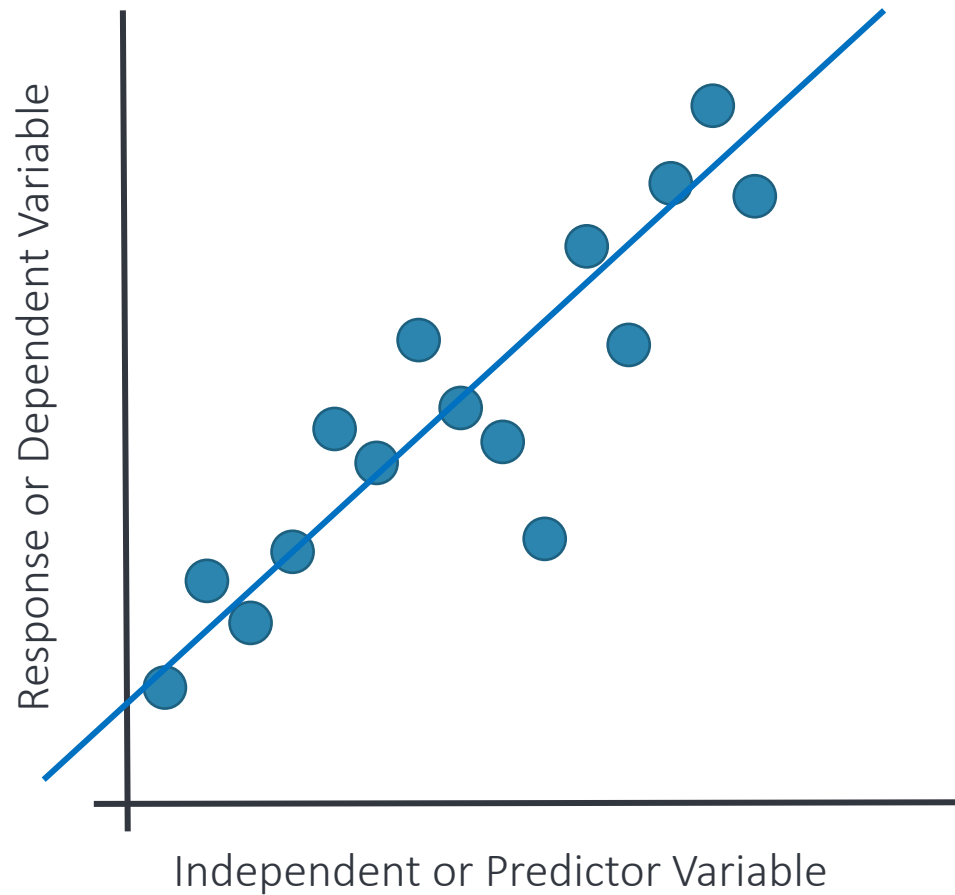
(always between -1 and 1)

- $r$  tells you about the *strength* of the relationship
- Slope tells you change in  $Y$  per unit change in  $X$

$$b_1 = r \frac{s_y}{s_x}$$

# Frequentist:

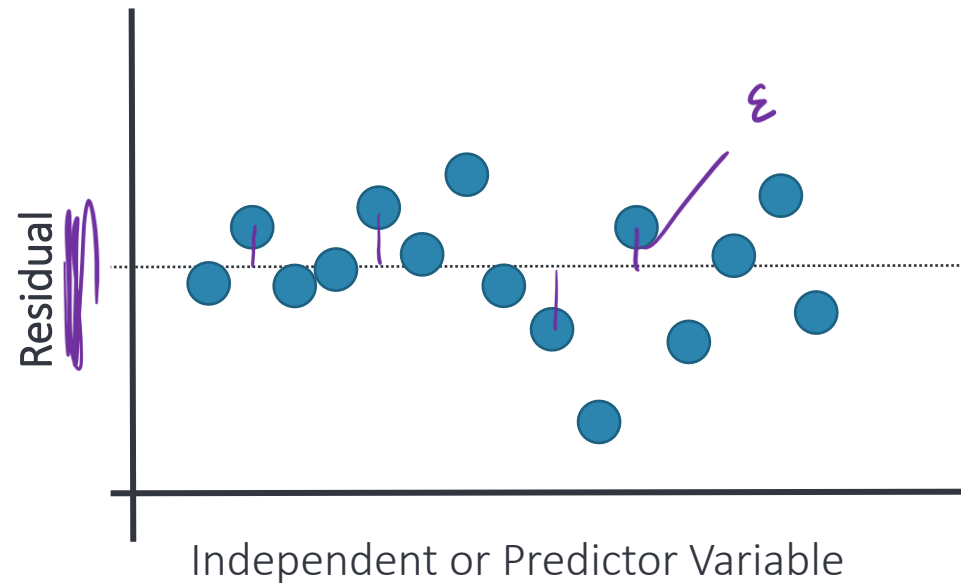
## How we find the least squares line



- The least squares line minimizes the sum of squared residuals:

minimize  $\sum_{i=1}^n \epsilon_i^2$

*n* ← n # of data points



# Linear Regression more generally

$p \rightarrow \# \text{ of parameters}$

Dimensions of each:

$$Y = X\beta + \epsilon$$

$n \times 1$     $n \times p$     $p \times 1$     $n \times 1$

It helps to visualize:

$$\begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix} = \begin{pmatrix} x_{11} & x_{12} & \dots & x_{1p} \\ \vdots & \vdots & & \vdots \\ x_{n1} & x_{n2} & \dots & x_{np} \end{pmatrix} \begin{pmatrix} \beta_1 \\ \vdots \\ \beta_p \end{pmatrix} + \begin{pmatrix} \epsilon_1 \\ \vdots \\ \epsilon_n \end{pmatrix}$$

Equivalently:

More concisely:

# Review of Linear Regression

- Model

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}$$

- $X$  often called the "design matrix"
- $X$  is an  $n \times p$  matrix of covariates/explanatory variables. These could be
  - Measurements
  - Fixed by design
  - Introduced to increase model flexibility
- In practice, the intercept  $\beta_0$  may be encoded in  $X$ :

# Least Squares estimators for $\beta$

- If the linear regression model is  $Y = \beta_0 + \beta_1 X + \epsilon$ , then the OLS estimator minimizes the *residual sums of squares* (RSS):

$$\min(\text{RSS}) = \min \left[ \sum_{i=1}^n \underbrace{\left( y_i - (\beta_0 + \beta_1 x_i) \right)}_{\epsilon_i}^2 \right]$$

- Least squares estimator of  $\beta$ :

$$\hat{\beta}_{LS} = \operatorname{argmin}_{\beta} \sum_i^n (y_i - x_i^T \beta)^2$$

# Maximum Likelihood Estimator (MLE) for $\beta$

Assuming normally-distributed, independent errors

If we assume that  $y_i \sim N(\underbrace{x_i^T \beta}_{\text{mean}}, \underbrace{\sigma^2}_{\text{variance}})$ , then we can write out the likelihood function

# Likelihood function

$$f(y_1, y_2, \dots, y_n | x_1, x_2, \dots, x_n; \beta, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(y_i - x_i^T \beta)^2}$$

$$y_i | X \sim N(X\beta, \sigma^2 I)$$

The likelihood function is

$$L(\underbrace{\beta, \sigma^2}_{\text{parameters}}; \underbrace{y}_{\text{data}}) = \frac{1}{(\sqrt{2\pi\sigma^2})^n} \exp \left\{ -\frac{1}{2} (y - X\beta)^T (y - X\beta) \right\}$$

The log-likelihood is then

$$\log \mathcal{L} = \dots$$

And now we can do maximum likelihood estimation ...

# Maximum likelihood estimate of $\beta$

log likelihood  $\rightarrow$

$$\ell(\beta, \sigma^2; y) = -\frac{n}{2} \log \sigma^2 - \frac{1}{2\sigma^2} (y - X\beta)^T (y - X\beta)$$

Take derivative w.r.t  $\beta$  and set equal to zero:

$$\left. \frac{\partial \ell}{\partial \beta} = -\frac{1}{2\sigma^2} X^T (y - X\beta) \right|_{\hat{\beta}_{MLE}} = 0$$

$$\text{M.L.E. of } \beta: \hat{\beta}_{ML} = (X^T X)^{-1} X^T Y = \hat{\beta}_{LS}$$



# Slope and Intercept Estimators

- Under the assumption that  $\epsilon$  are i.i.d., then the slope and intercept estimators are unbiased and asymptotically distributed as:

$$\hat{\beta}_1 \sim N\left(\beta_1, \frac{\sigma^2}{S_{xx}}\right)$$

Handwritten annotations:   
 - A horizontal line above  $\hat{\beta}_1$ .   
 - An arrow from the text "true value" points to  $\beta_1$ .   
 - An arrow from the text "variance" points to  $\frac{\sigma^2}{S_{xx}}$ .

assuming  $\sigma$  is known.

- When the true variance is unknown, the distributions of the estimators are not known (because we don't know  $\beta_1$  nor  $\sigma^2$ )

# Uncertainties in the line

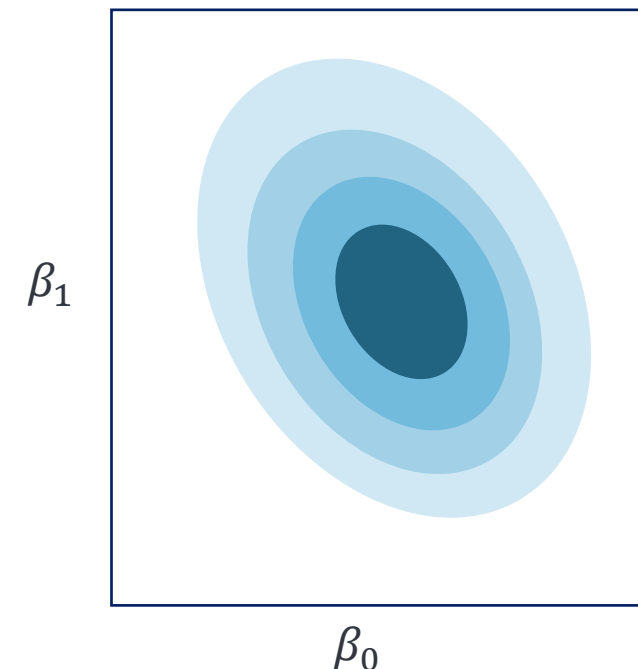
- Because we know that  $\widehat{\beta}_1 \sim N(\beta_1, \frac{\sigma^2}{s})$ , we can use this to construct confidence intervals for  $Y$  at each  $X$  using

$$Y|x = \hat{\beta}_0 + \hat{\beta}_1 x \pm t_{\alpha/2, n-2} s \sqrt{1 + \frac{1}{n} \frac{(x-\bar{x})^2}{S_{xx}}}$$

➡ do this for a range of  $x$ -values, and you will get confidence bands about the best-fit (OLS) line that are hyperbolas.

# Point Estimates vs. Bayesian Inference

- The estimators for  $\beta_0, \beta_1$ , etc. In previous slides are *point estimates*
- A maximum likelihood estimate is also a point estimate of a parameter
- In Bayesian inference...
  - We try to infer the *distribution* for a parameter
  - We get a *posterior distribution*
  - The posterior distribution encodes all the information from
    - Prior assumptions
    - Model assumptions
    - Data
  - We report the whole posterior distribution in our results
  - We can report credible intervals to express uncertainty



# Linear Regression in Bayesian context

- The model for  $y$  is

$$y \sim N(\overset{\text{mean}}{x\beta}, \overset{\text{variance}}{\sigma^2})$$

- The likelihood under this model is

$$\prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(y_i - x_i\beta)^2}$$

- We must set priors on the parameters

$$p(\beta_0, \beta_1) \rightarrow \text{joint prior distribution}$$

$$\text{or } \beta_0 \sim N(0, \sigma_0^2)$$