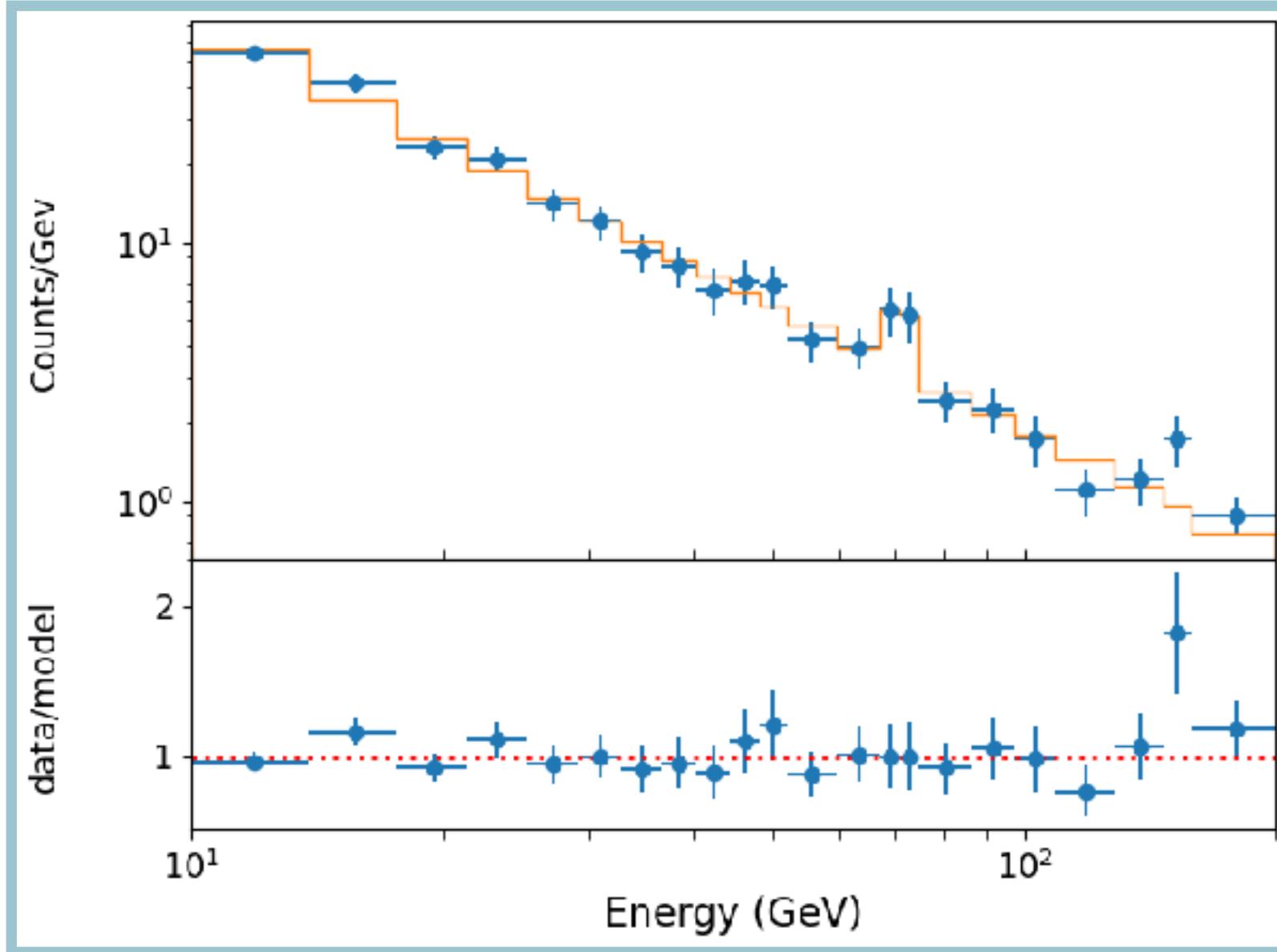


# Sampling + Model Comparison, Part II



# What about physics?



Data:

X-ray spectrum

Model:

Power law + Gaussian line

Likelihood:

Poisson

Parameters:

$A_{\text{PL}}, \Gamma, A_{\text{Line}}, E_{\text{Line}}, \sigma_{\text{Line}}$

Priors:

???

What priors should we choose?

$$p(A_{\text{PL}}, \Gamma, A_{\text{Line}}, E_{\text{Line}}, \sigma_{\text{Line}} | D) = \frac{p(D | A_{\text{PL}}, \Gamma, A_{\text{Line}}, E_{\text{Line}}, \sigma_{\text{Line}}) p(A_{\text{PL}}, \Gamma, A_{\text{Line}}, E_{\text{Line}}, \sigma_{\text{Line}})}{p(D)}$$

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What assumption  
have we made here?

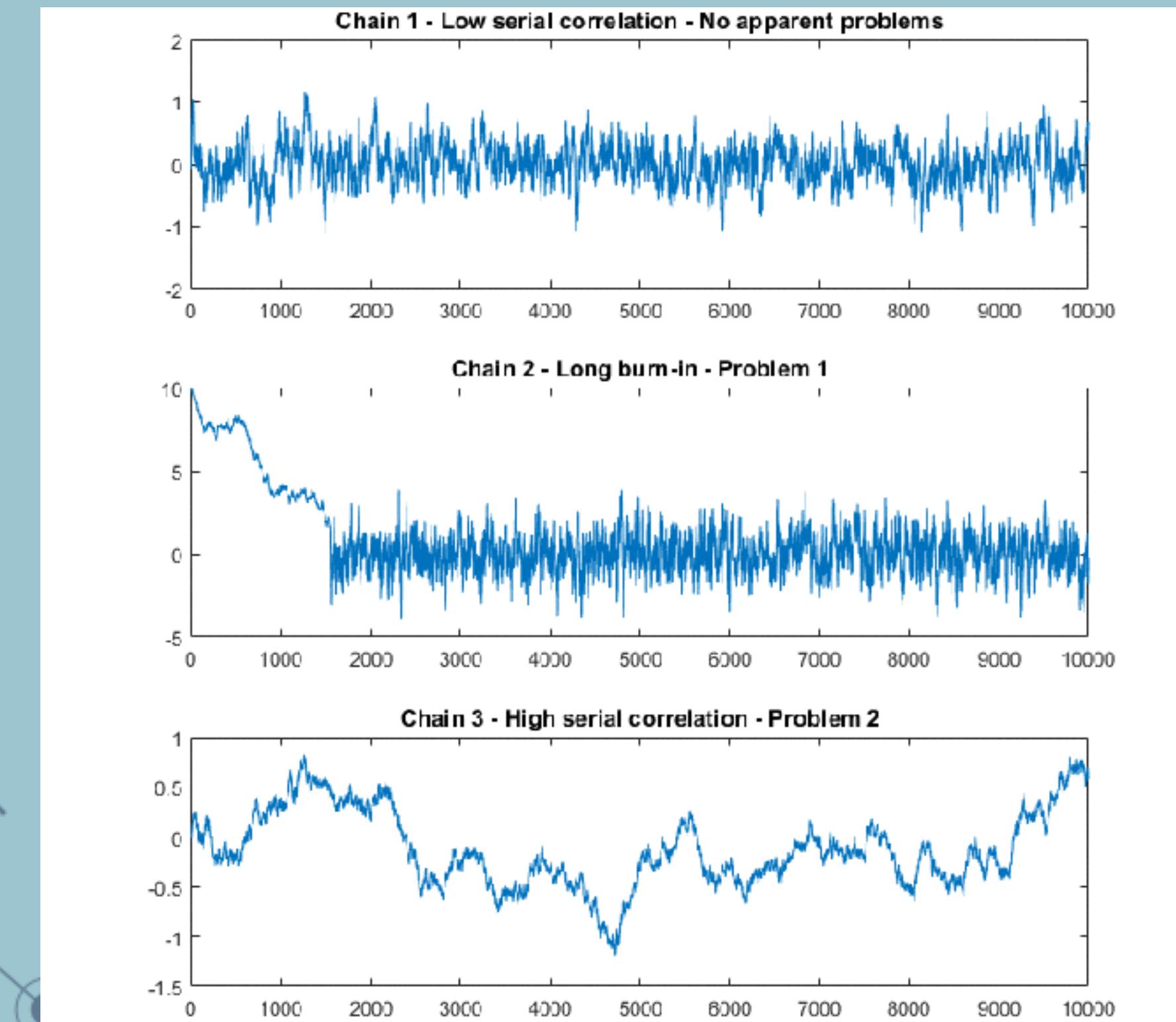
Independence of  
priors!

# **Markov Chain Monte Carlo:**

## **Assessing Convergence**

# Assessing Convergence: trace plots

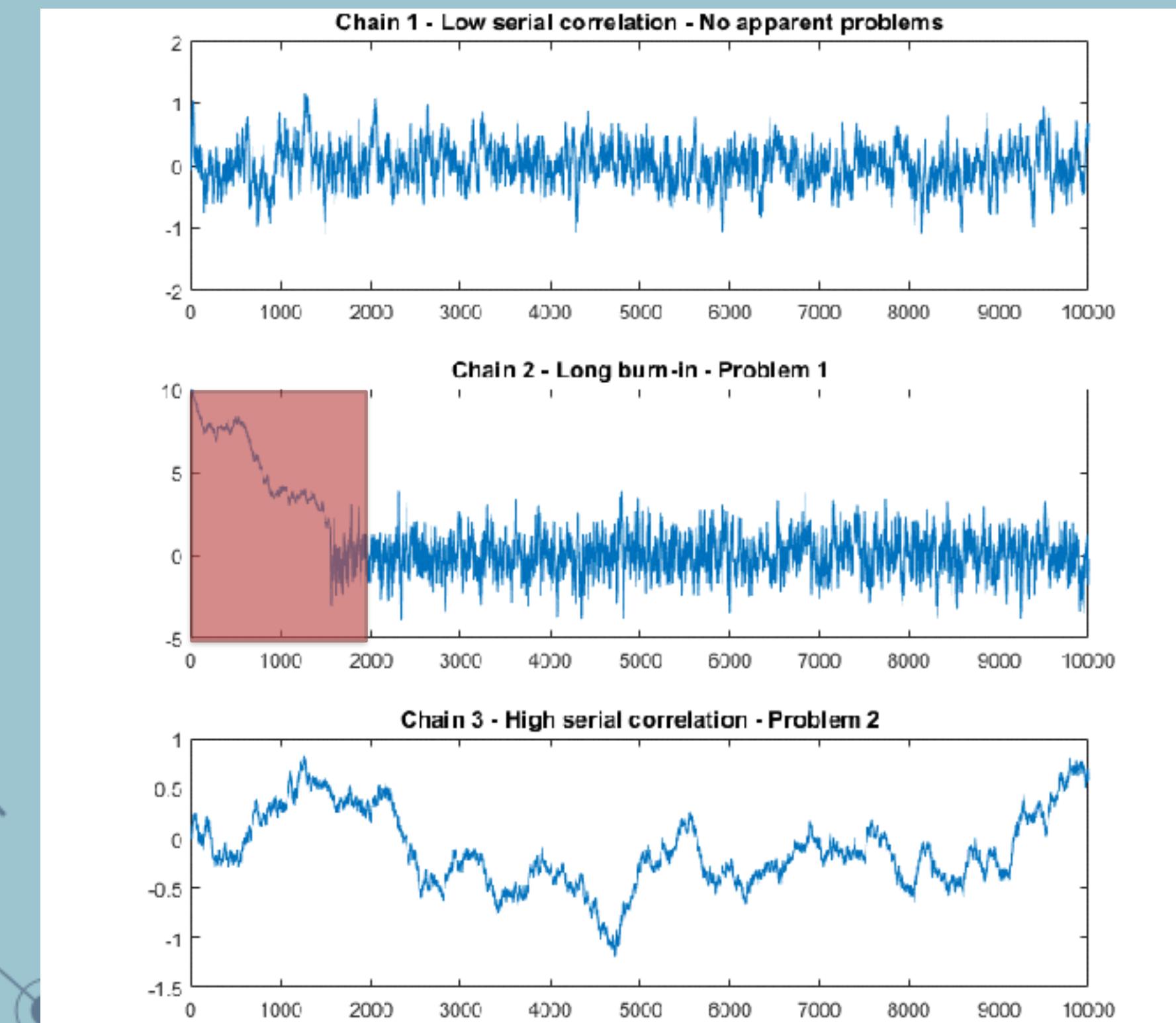
How do I know whether my Markov chain is sampling from the posterior?



# Assessing Convergence: trace plots



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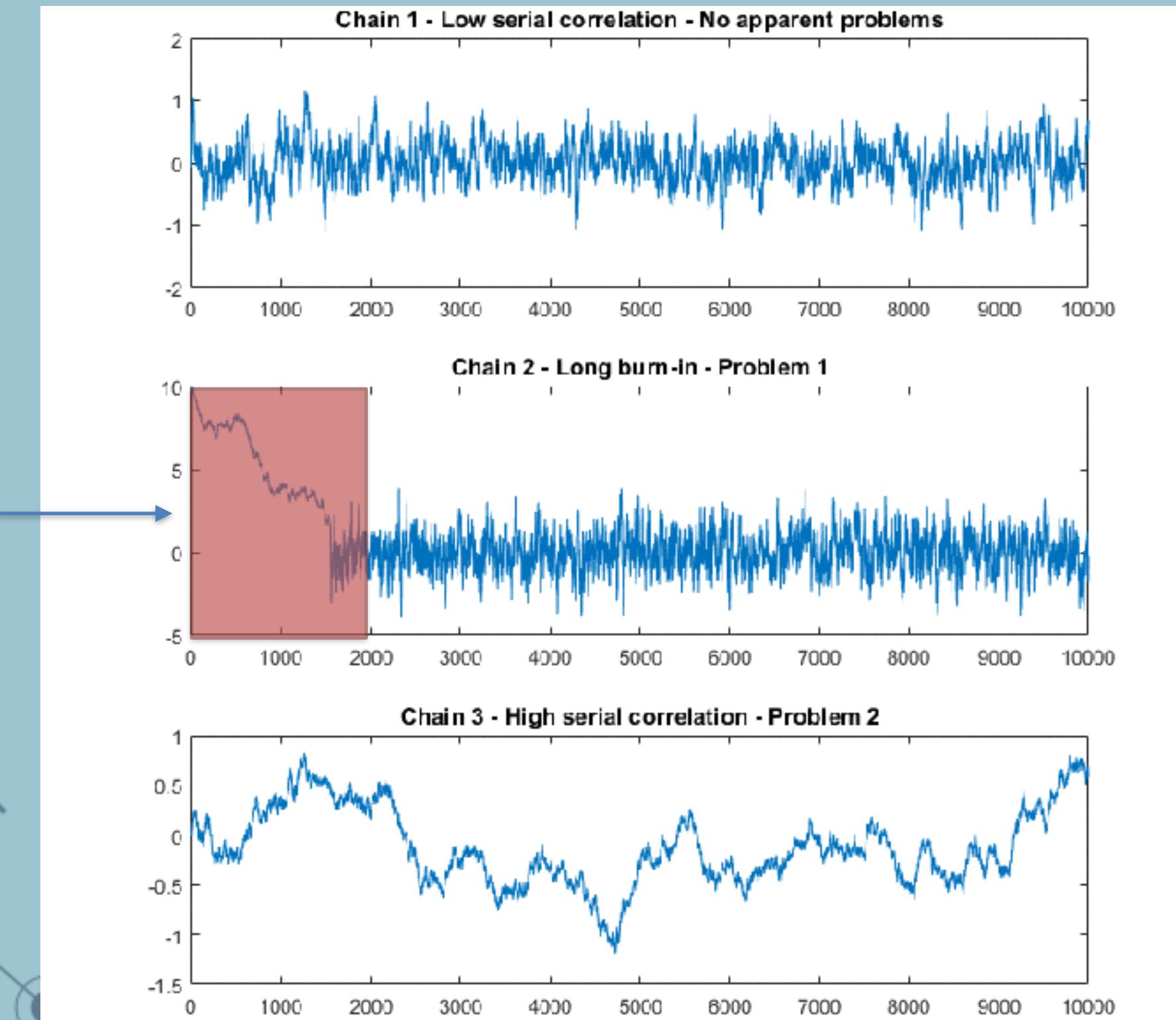


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Cut out  
this part

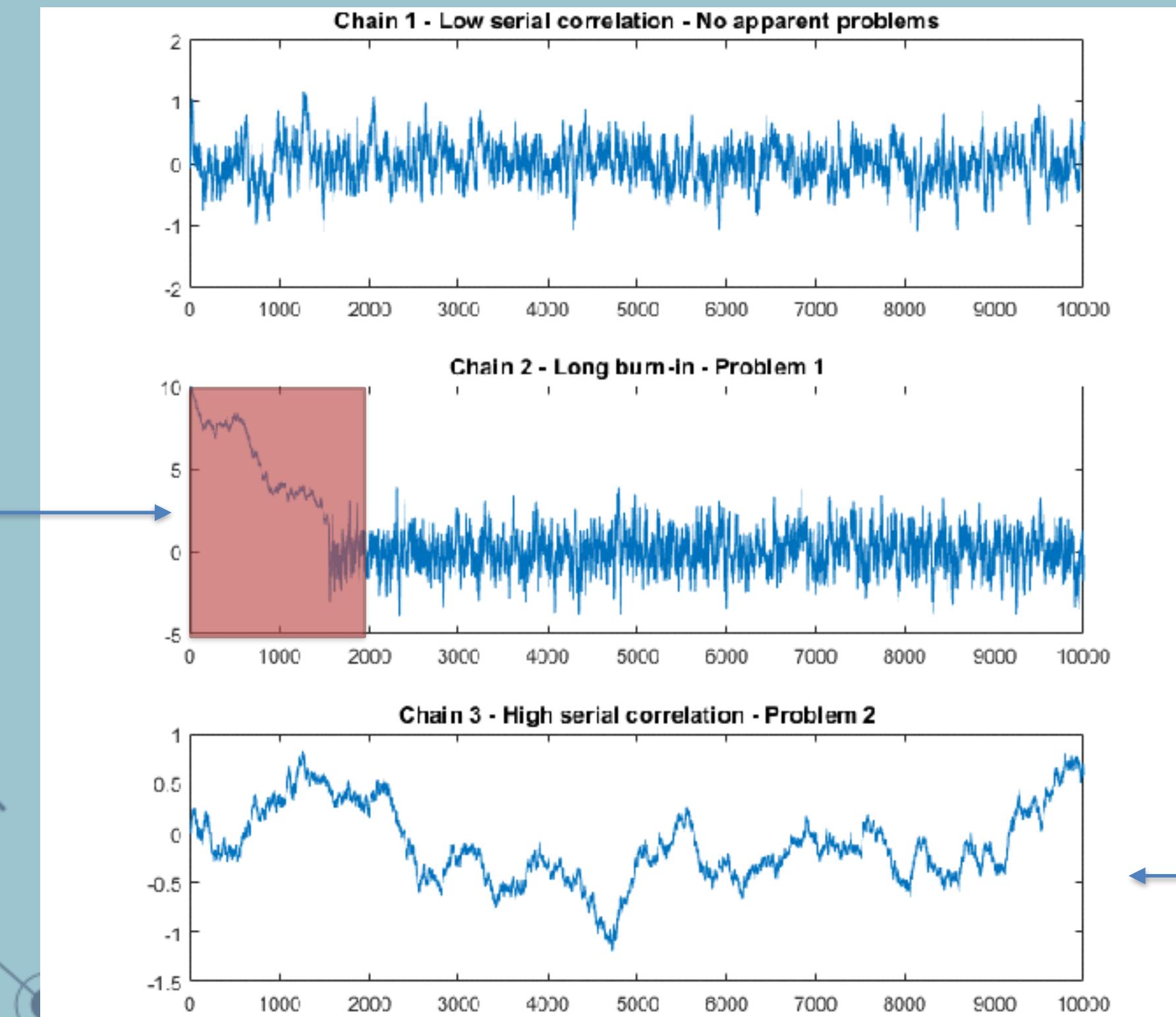


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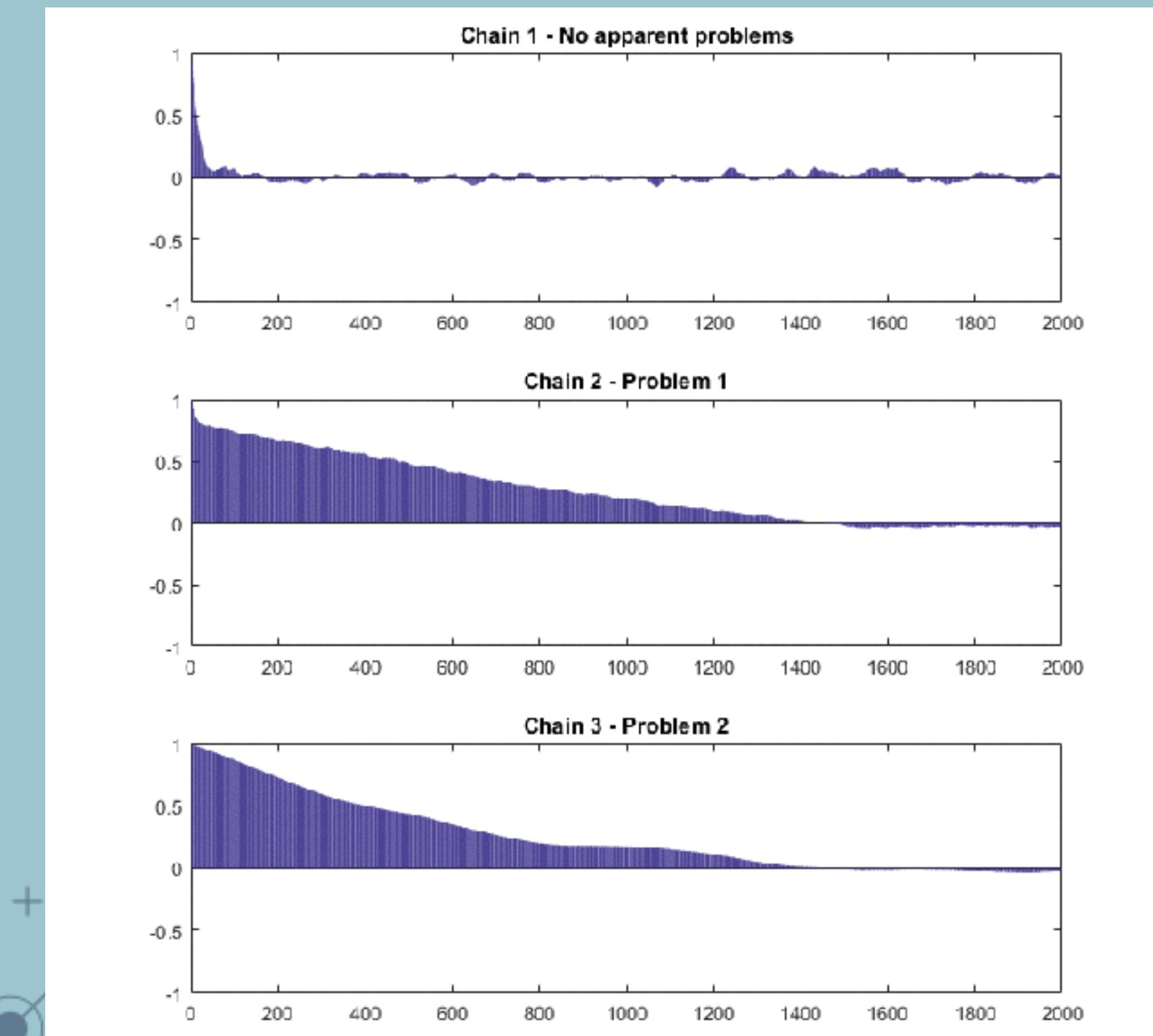
Cut out  
this part



# Assessing Convergence: autocorrelation



“How long do I have to sample until I get to an independent sample?



# Assessing Convergence:

## Gelman-Rubin Statistic

Compare the variance within chains to the variance between chains

$$\bar{x}_j = \frac{1}{L} \sum_{i=1}^L x_i^{(j)} \text{ Mean value of chain j}$$



# Assessing Convergence:

## Gelman-Rubin Statistic

Compare the variance within chains to the variance between chains

$$\bar{x}_j = \frac{1}{L} \sum_{i=1}^L x_i^{(j)} \text{ Mean value of chain } j$$

$$\bar{x}_* = \frac{1}{J} \sum_{j=1}^J \bar{x}_j \text{ Mean of the means of all chains}$$



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$$B = \frac{L}{J-1} \sum_{j=1}^J (\bar{x}_j - \bar{x}_*)^2 \text{ Variance of the means of the chains}$$



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$$W = \frac{1}{J} \sum_{j=1}^J \left( \frac{1}{L-1} \sum_{i=1}^L (x_i^{(j)} - \bar{x}_j)^2 \right) \text{ Averaged variances of the individual chains across all chains}$$



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An estimate of the Gelman-Rubin statistic  $R$  then results as<sup>[1]</sup>

$$R = \frac{\frac{L-1}{L} W + \frac{1}{L} B}{W}.$$

When  $L$  tends to infinity and  $B$  tends to zero,  $R$  tends to 1.



# Summarising Posterior Probability Distributions

You have sampled a posterior PDF, which has produced samples (vectors of parameters). What do you do now?



# Summarising Posterior Probability Distributions:

Marginalised  
Distributions +  
Credible Intervals



## Reminder: Marginalisation

Summarising  
Posterior Probability  
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Reminder: Marginalisation

$$p(x) = \int_{-\infty}^{\infty} p(x, y) dy = \int_{-\infty}^{\infty} p(x | y) p(y) dy$$

# Summarising Posterior Probability Distributions: Marginalised Distributions + Credible Intervals



## Reminder: Marginalisation

$$p(x) = \int_{-\infty}^{\infty} p(x, y) dy = \int_{-\infty}^{\infty} p(x | y)p(y) dy$$

$$p(\theta_1 | D) = \int_{-\infty}^{\infty} p(\theta_1, \theta_2 | D)p(\theta_1)p(\theta_2)d\theta_2$$

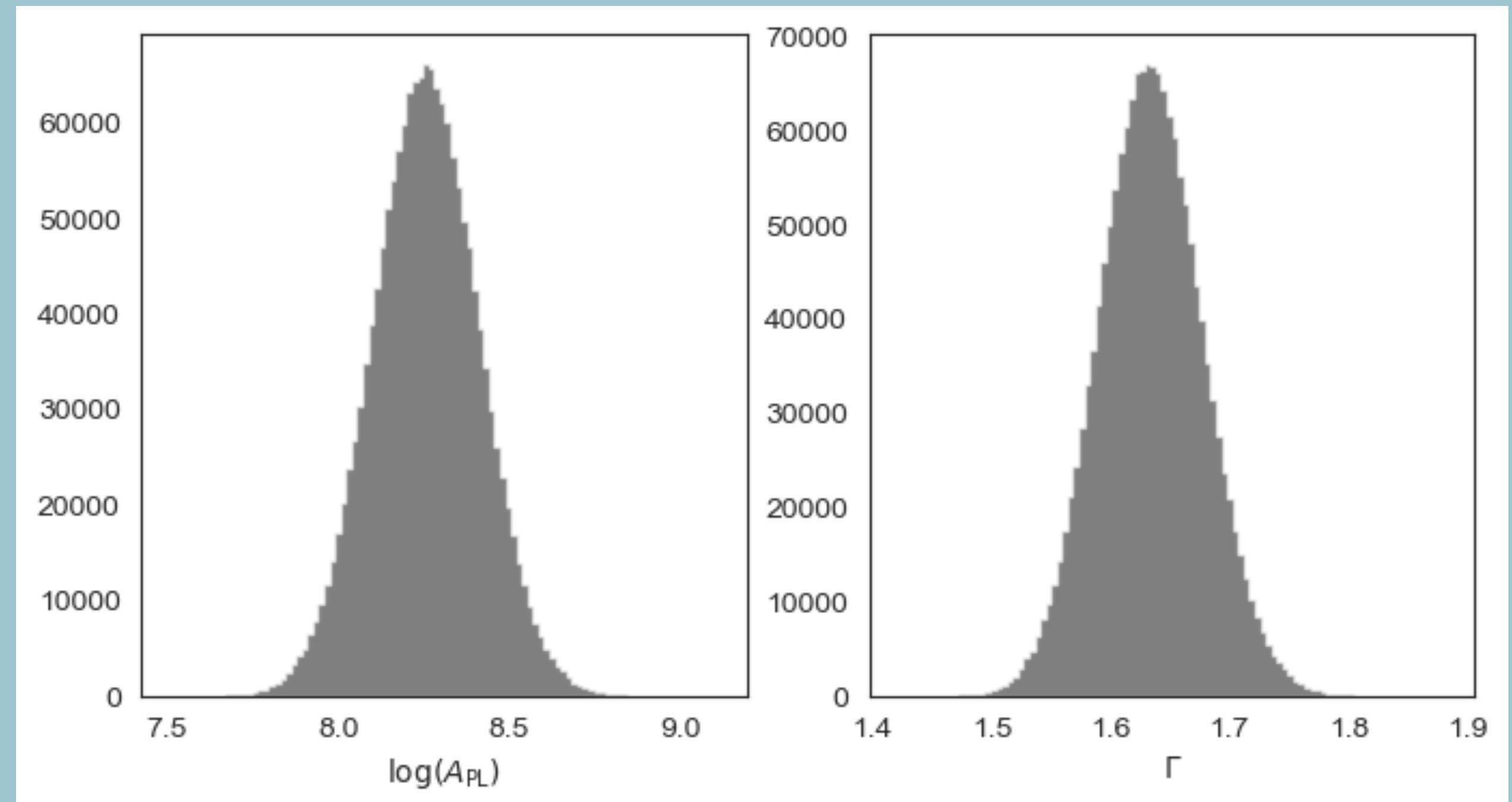
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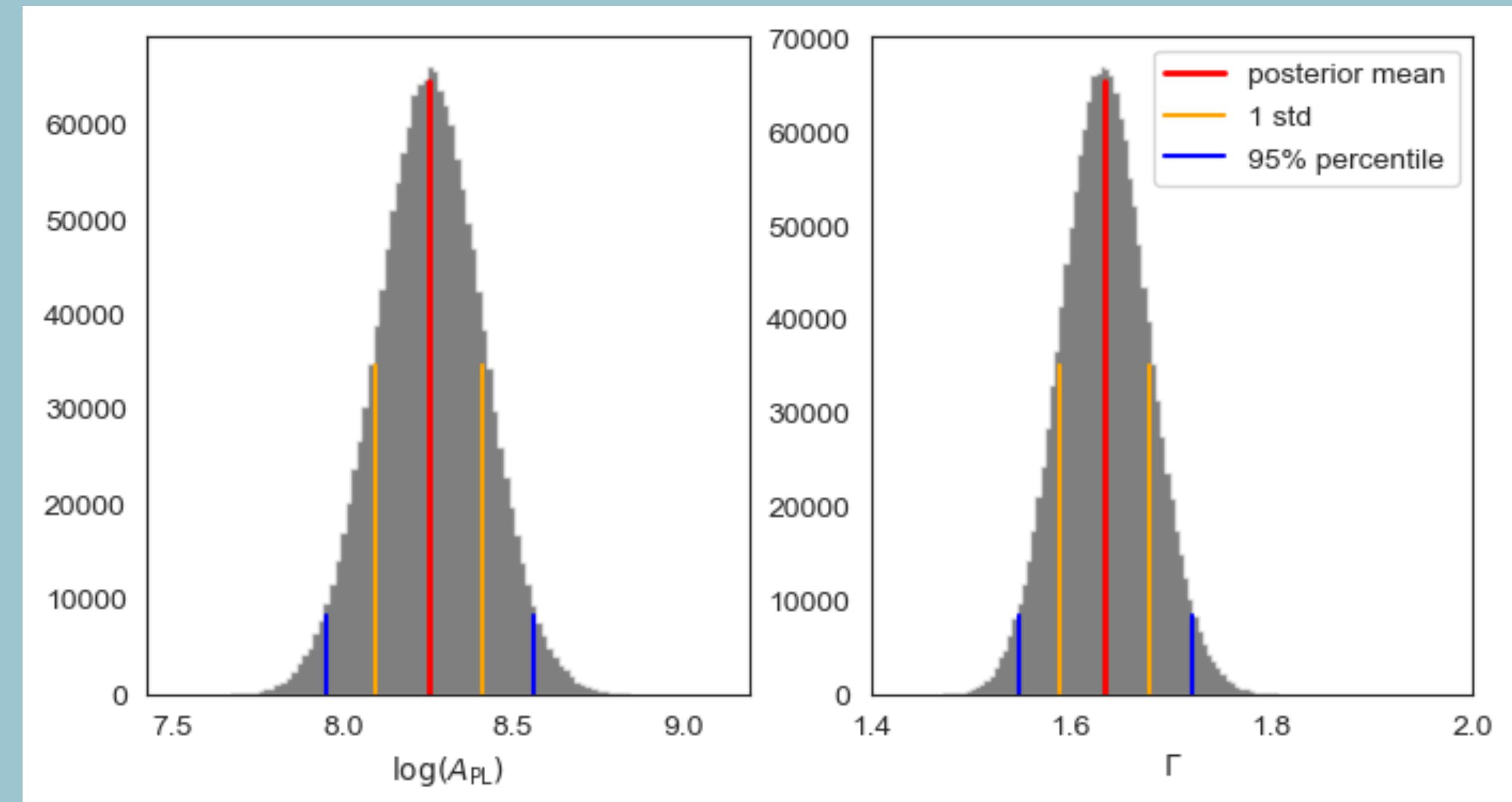
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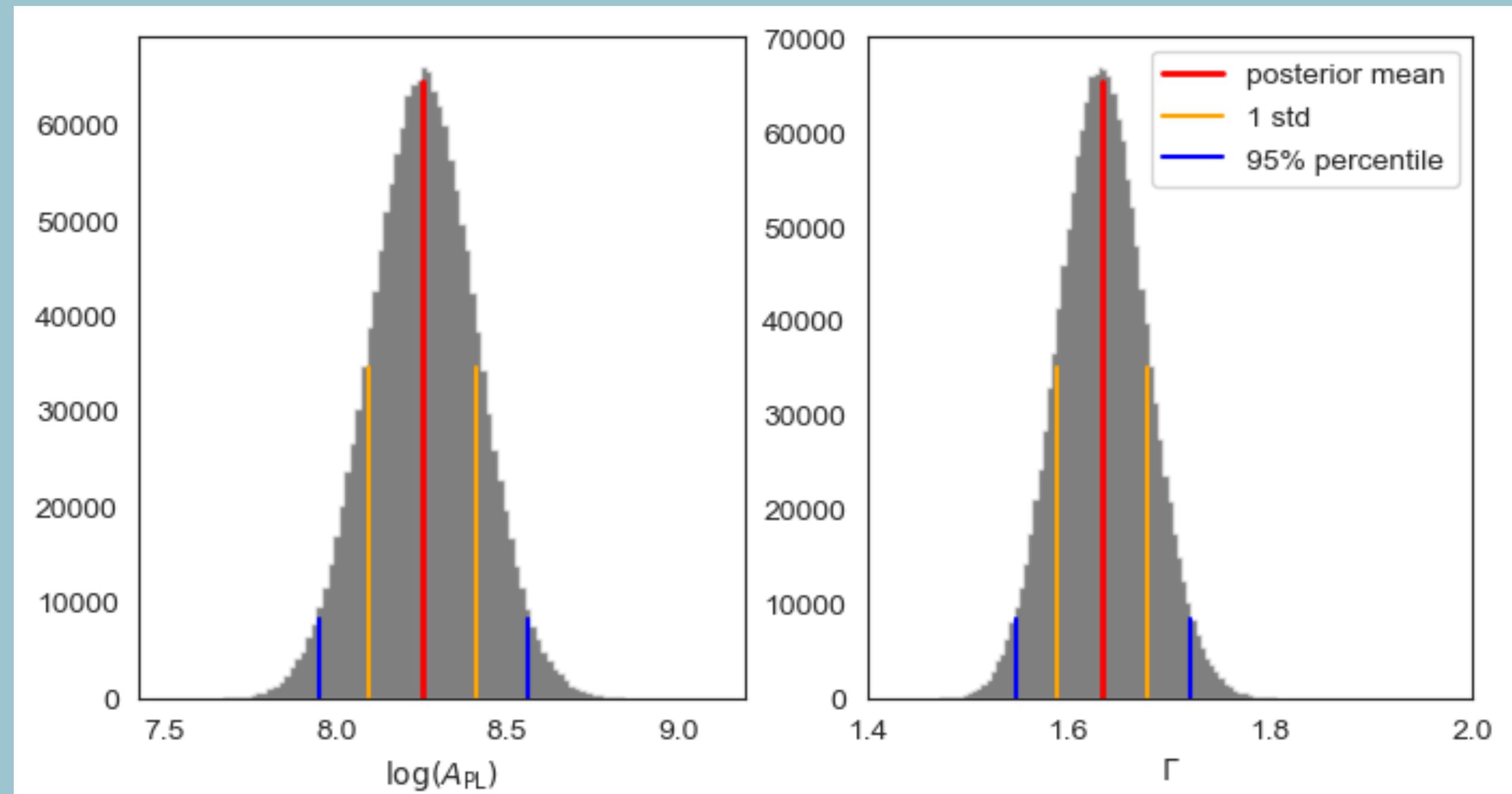
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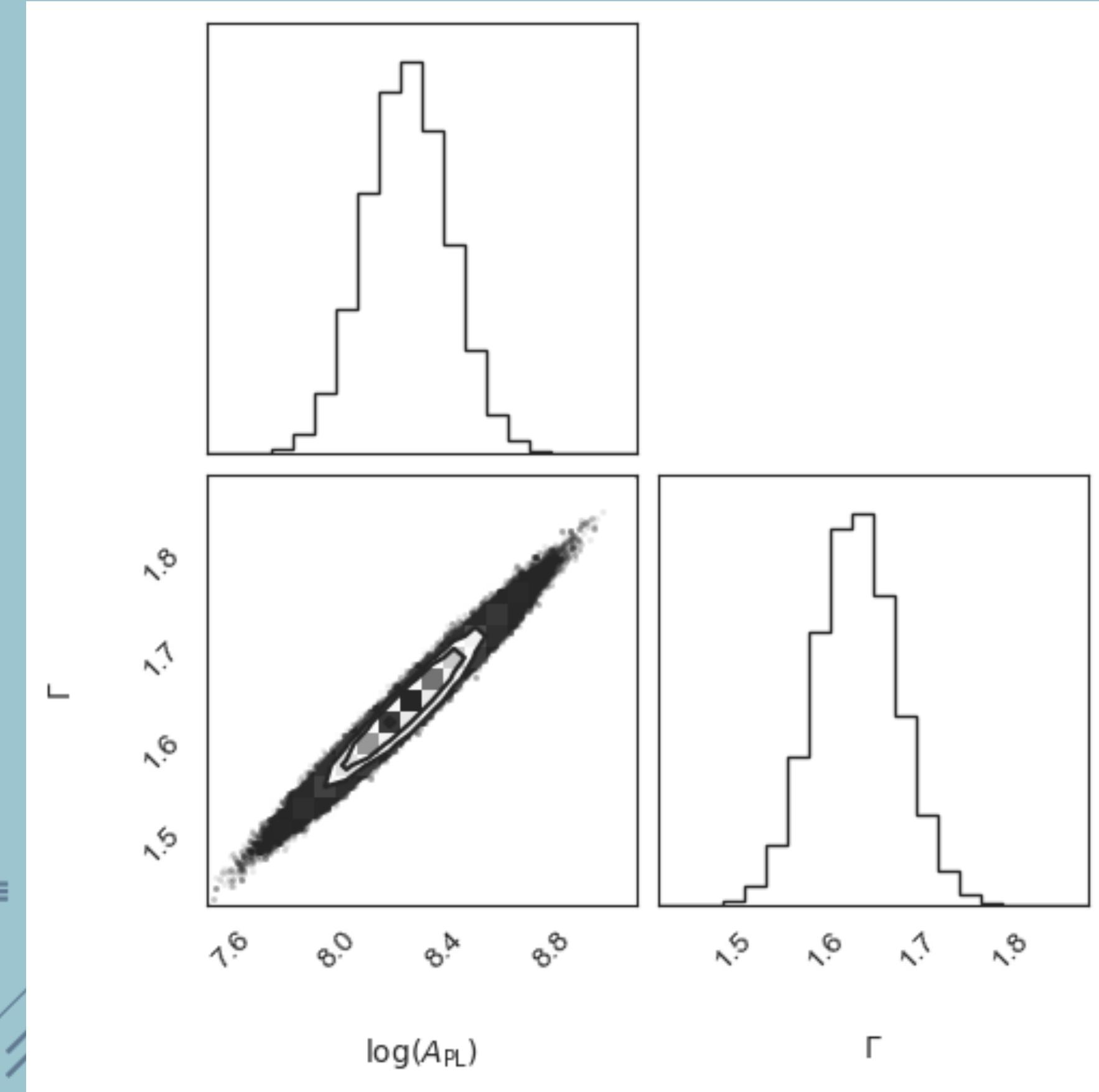
Marginalised Distributions + Credible Intervals



Considering just one dimension of your MCMC sample automatically marginalises over the other dimensions! Magic!

# Summarising Posterior Probability Distributions:

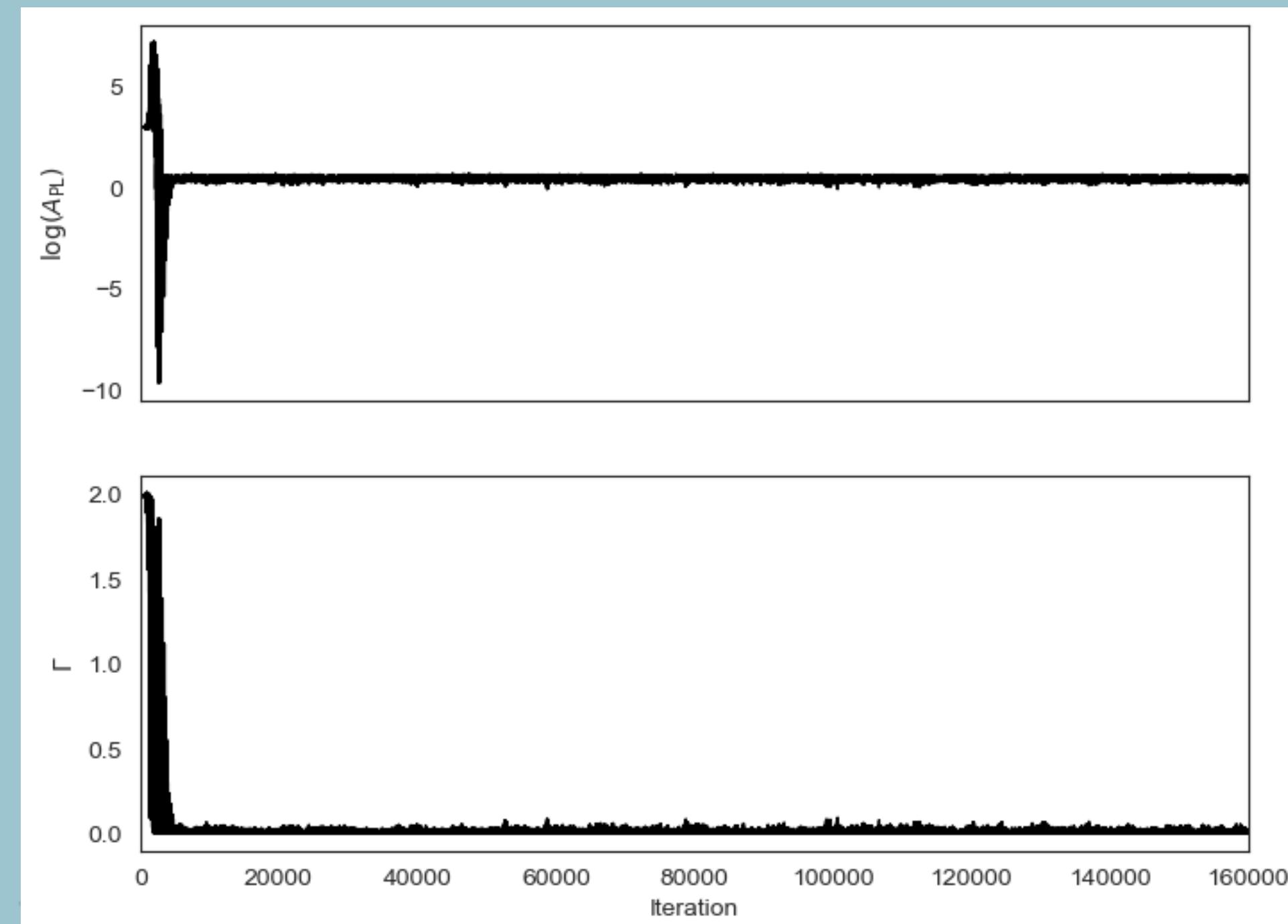
Corner Plots or  
Pair Plots



How do you know if/  
when things go  
wrong?

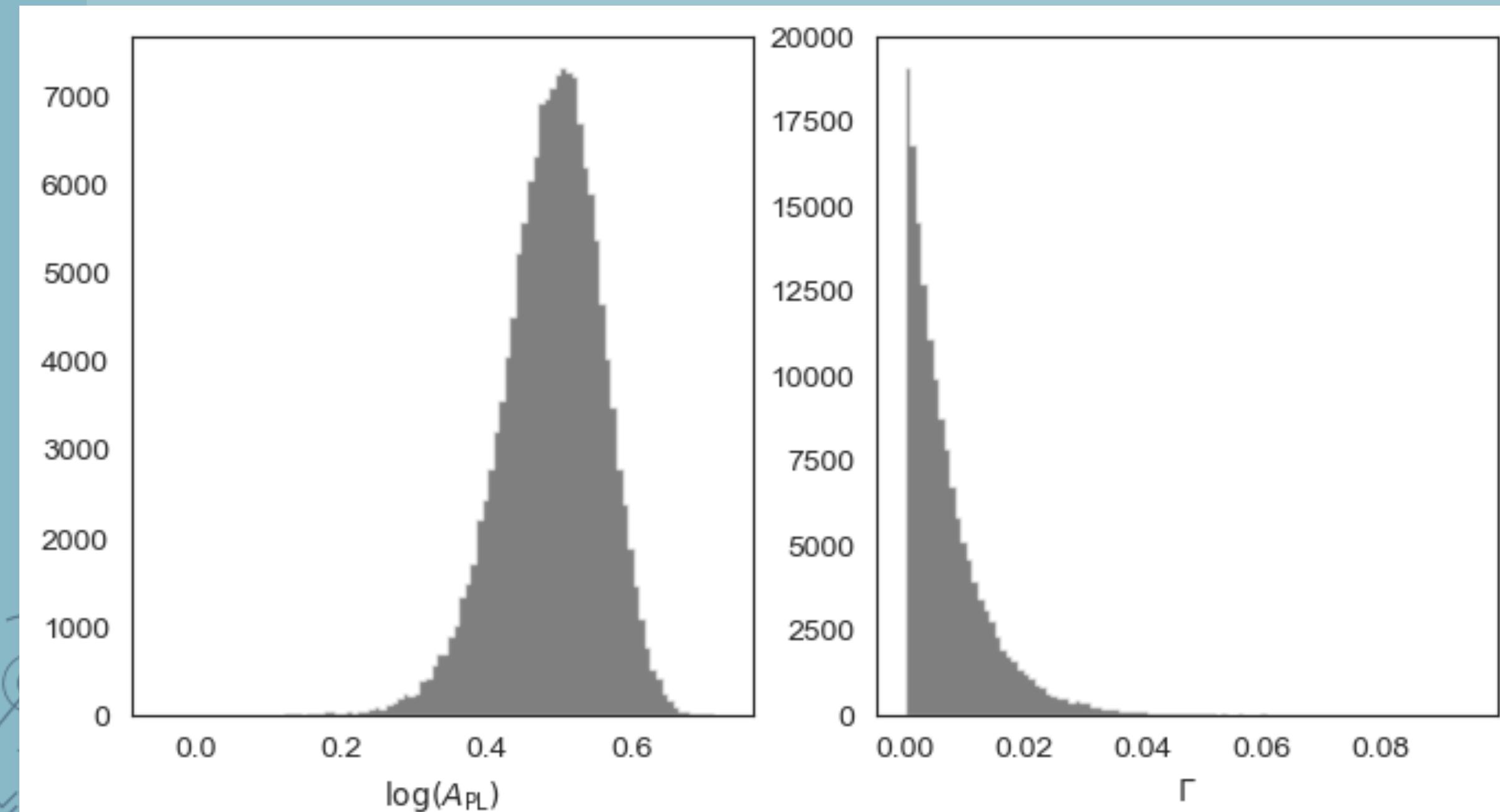


Earlier this week ...



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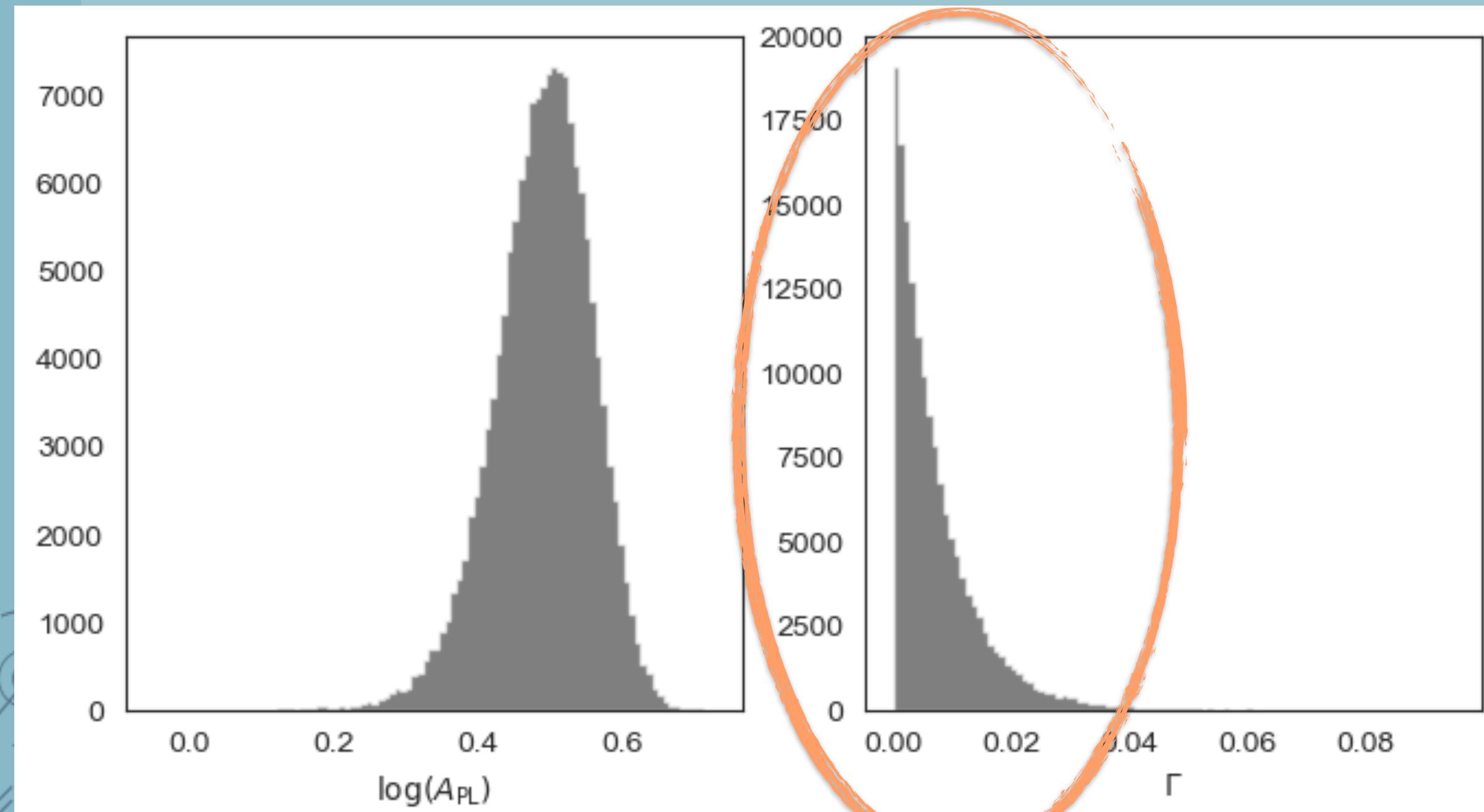
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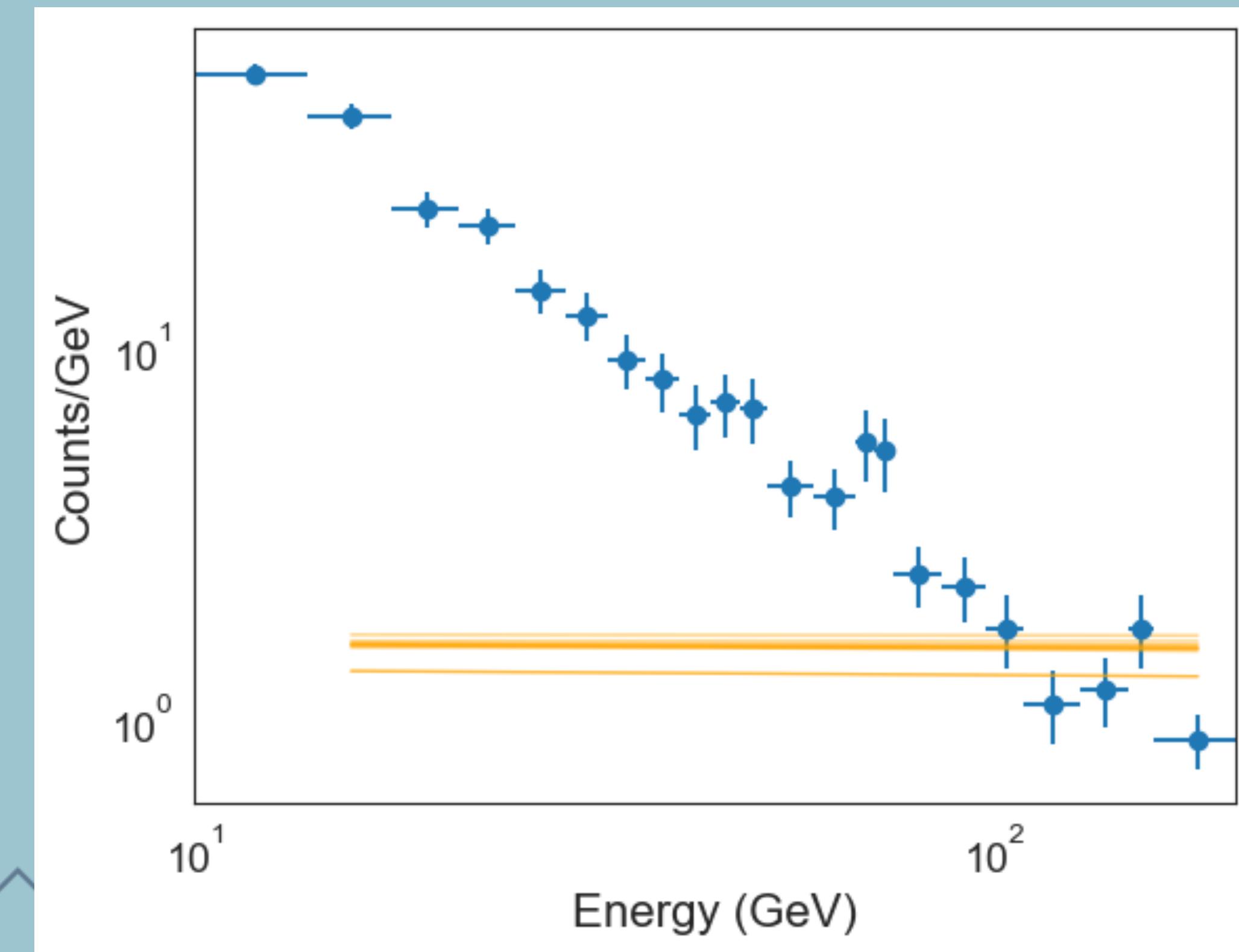
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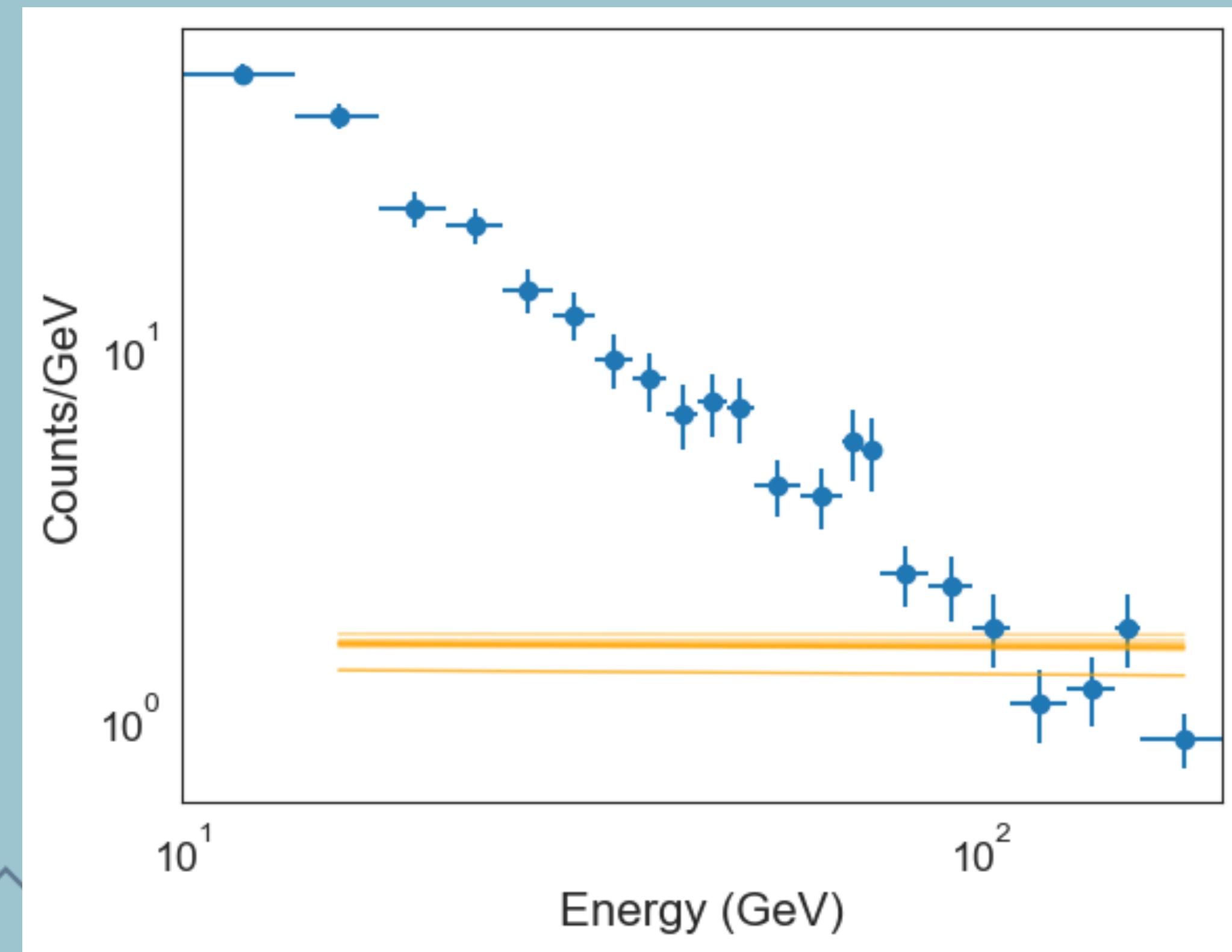
???



Plot posterior  
draws:

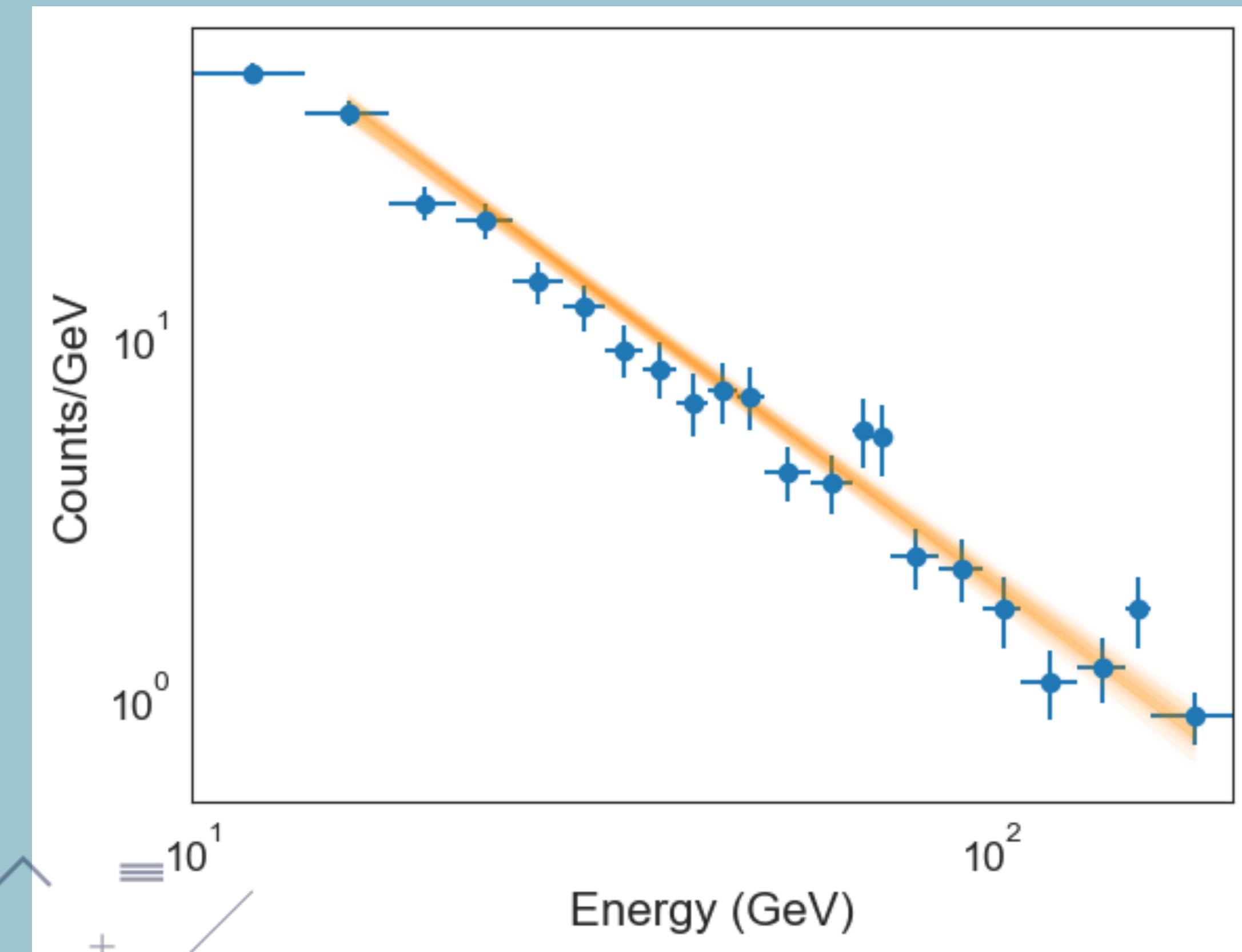


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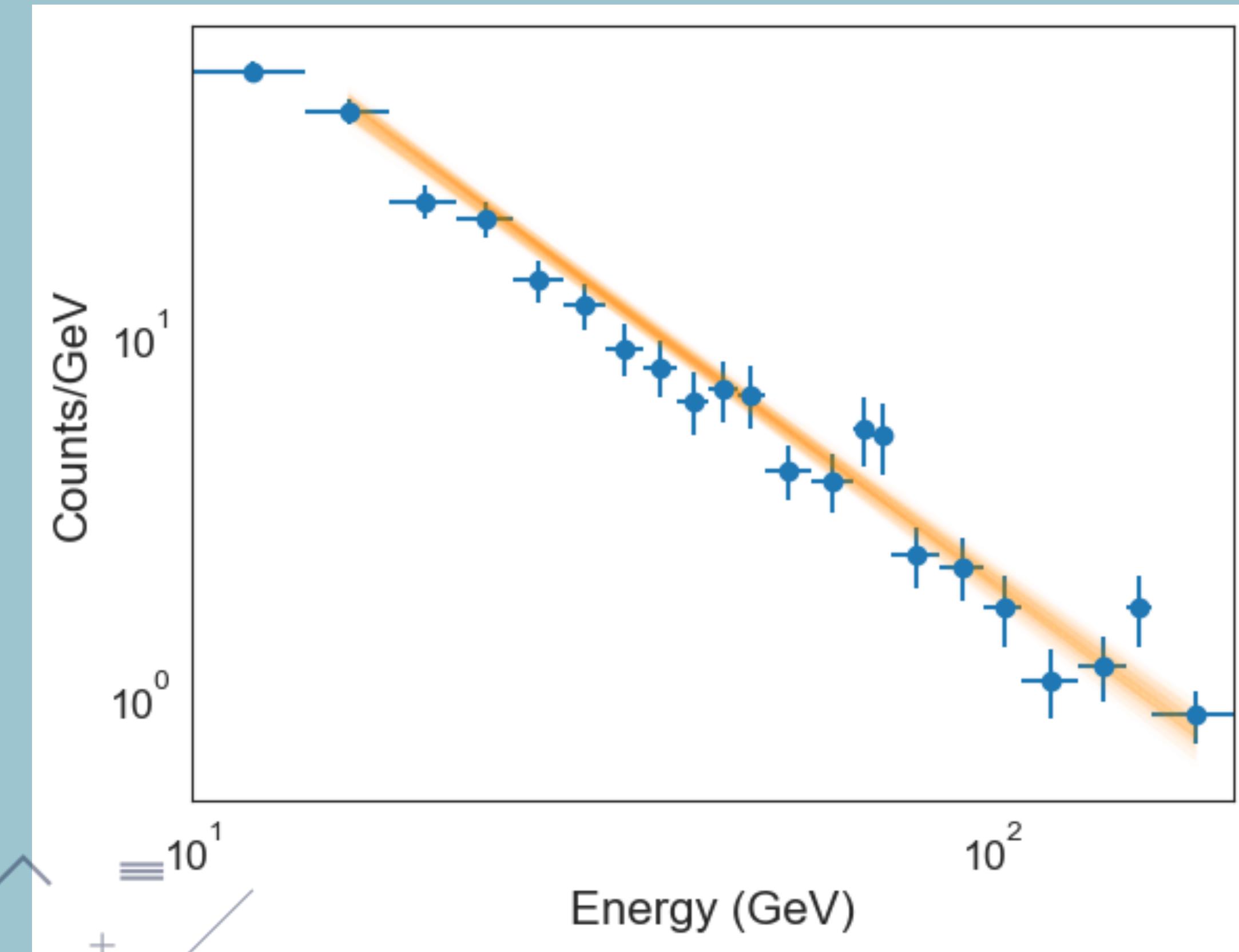


$\Gamma$  was defined with the wrong sign!

Fixing  $\Gamma$ :



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Plotting posterior draws allows you to see the range of models the posterior produces

# MCMC Notes and Pitfalls

Further reading:

<https://arxiv.org/pdf/1710.06068>



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Sampling is not a one-size-fits-all solution! Different problems may do better with different sampling algorithms! And there are many (see also: Hamiltonian Monte Carlo, Importance Sampling, Slice Sampling, Gibbs Sampling, Nested Sampling, AI-supported sampling algorithms, ...)

# Sampling Algorithms

DANIELA HUPPENKOTHEN



UNIVERSITEIT  
VAN AMSTERDAM

d.huppenkothen@uva.nl

# Gibbs Sampling



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- Can become computationally inefficient for large numbers of parameters
- Tend to get stuck in regions of parameter space for a long time, especially if high correlation in posterior and many dimensions



# Slice Sampling

Further reading:

- <https://www.cs.toronto.edu/pub/radford/slice-aos.pdf>
- arXiv:1001.0175



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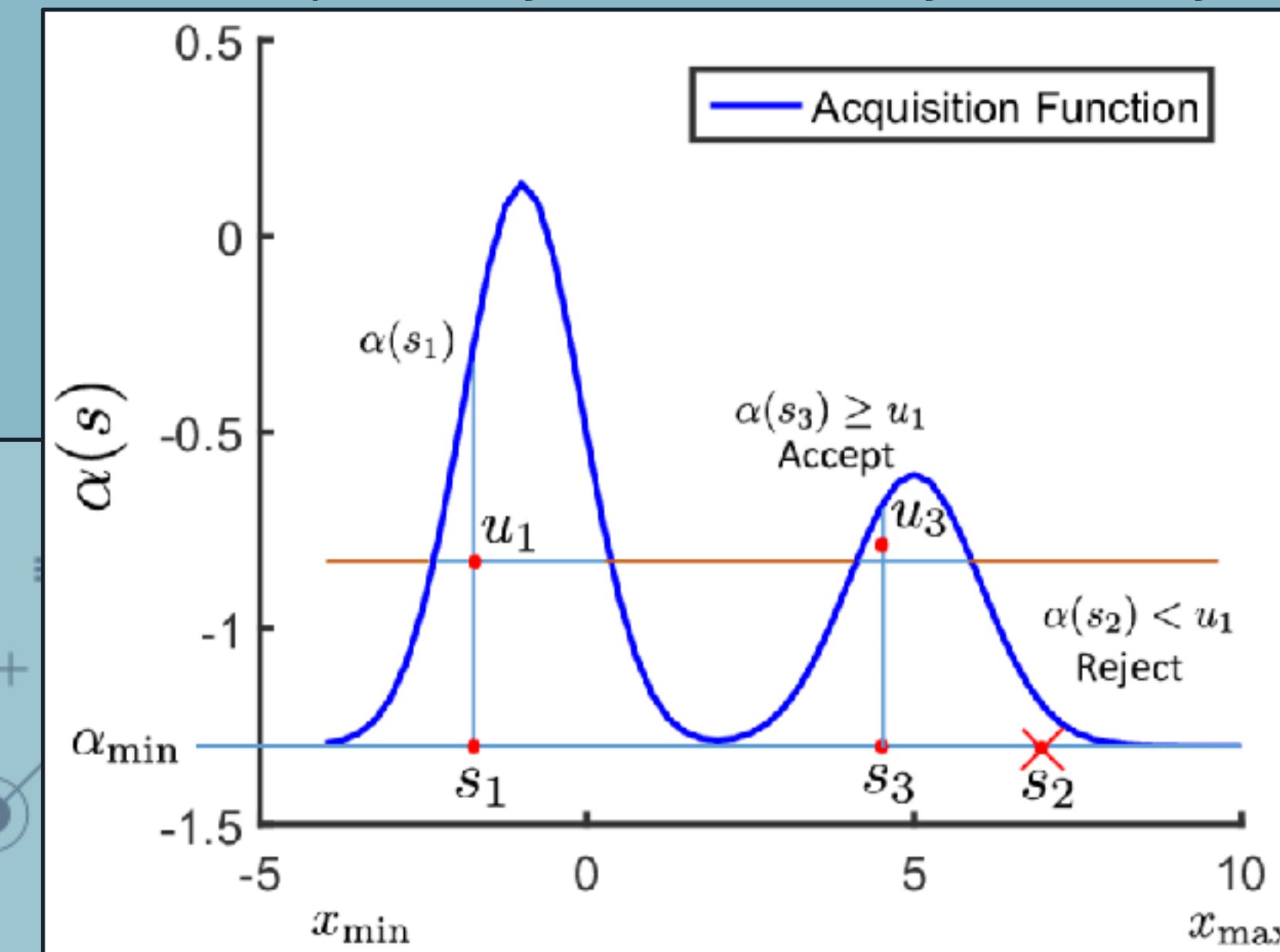
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- In 1D: easier to implement than Metropolis Hastings
  - Slice-sampling auto-adjusts the step size



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  5. Repeat from step 2 until you have as many draws as you want.
- In 1D: easier to implement than Metropolis Hastings
  - Slice-sampling auto-adjusts the step size
  - Often slower than Metropolis Hastings.



# Slice Sampling

Further reading:

- <https://www.cs.toronto.edu/pub/radford/slice-aos.pdf>
- arXiv:1001.0175

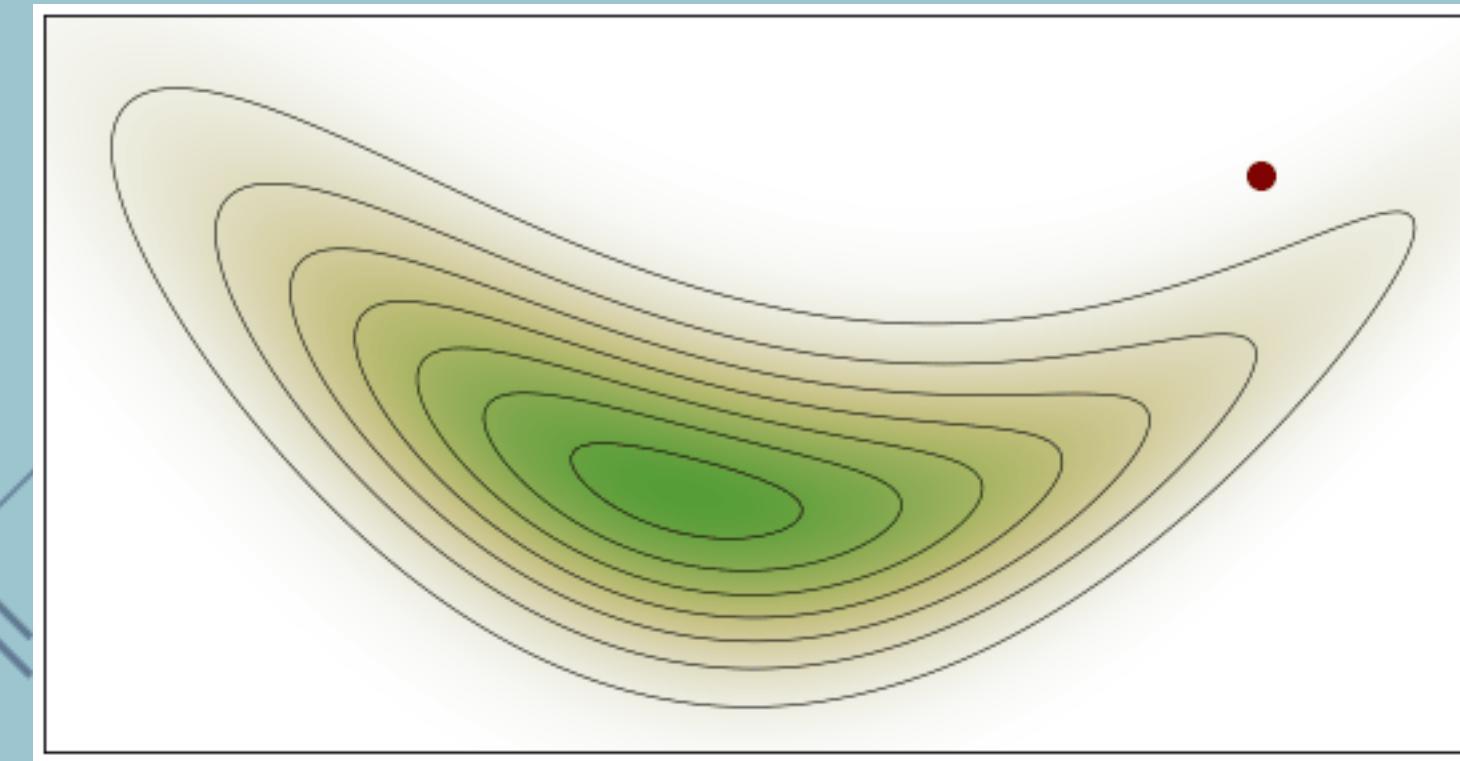
Issue with MH and Gibbs: you have to know the step size!

Slice sampling:

1. Choose a random starting value of  $x$ , let's call it  $x_1$
  2. Uniformly sample on interval  $[0, f(x_1)]$ , let's call it  $a$
  3. Imagine a horizontal line at  $y = a$ : figure out all the line segments under the curve.
  4. From all the line segments, draw a value of  $x$  uniformly.
  5. Repeat from step 2 until you have as many draws as you want.
- In 1D: easier to implement than Metropolis Hastings
  - Slice-sampling auto-adjusts the step size
  - Often slower than Metropolis Hastings.
  - Finding the roots of the intersection between horizontal line and distribution is pretty tricky.

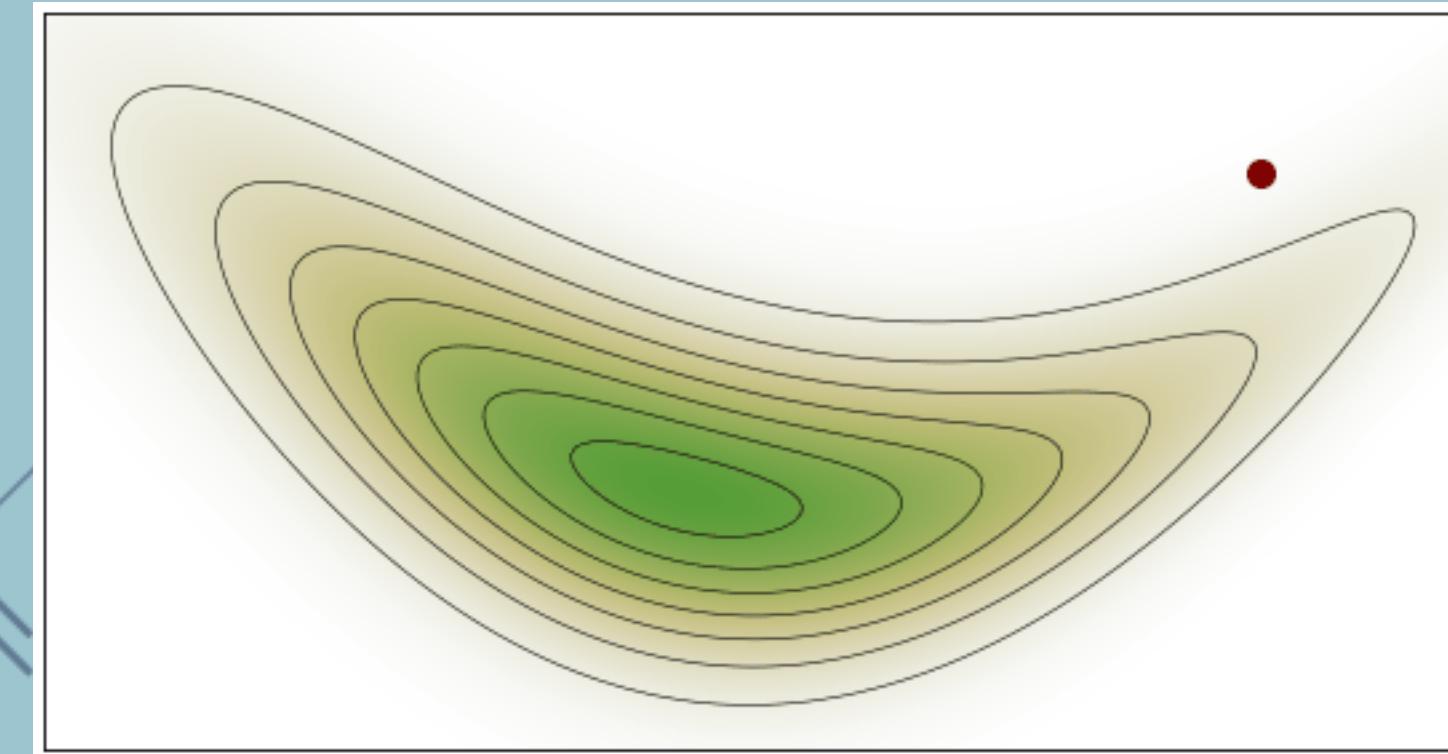


# Hamiltonian Monte Carlo



# Hamiltonian Monte Carlo

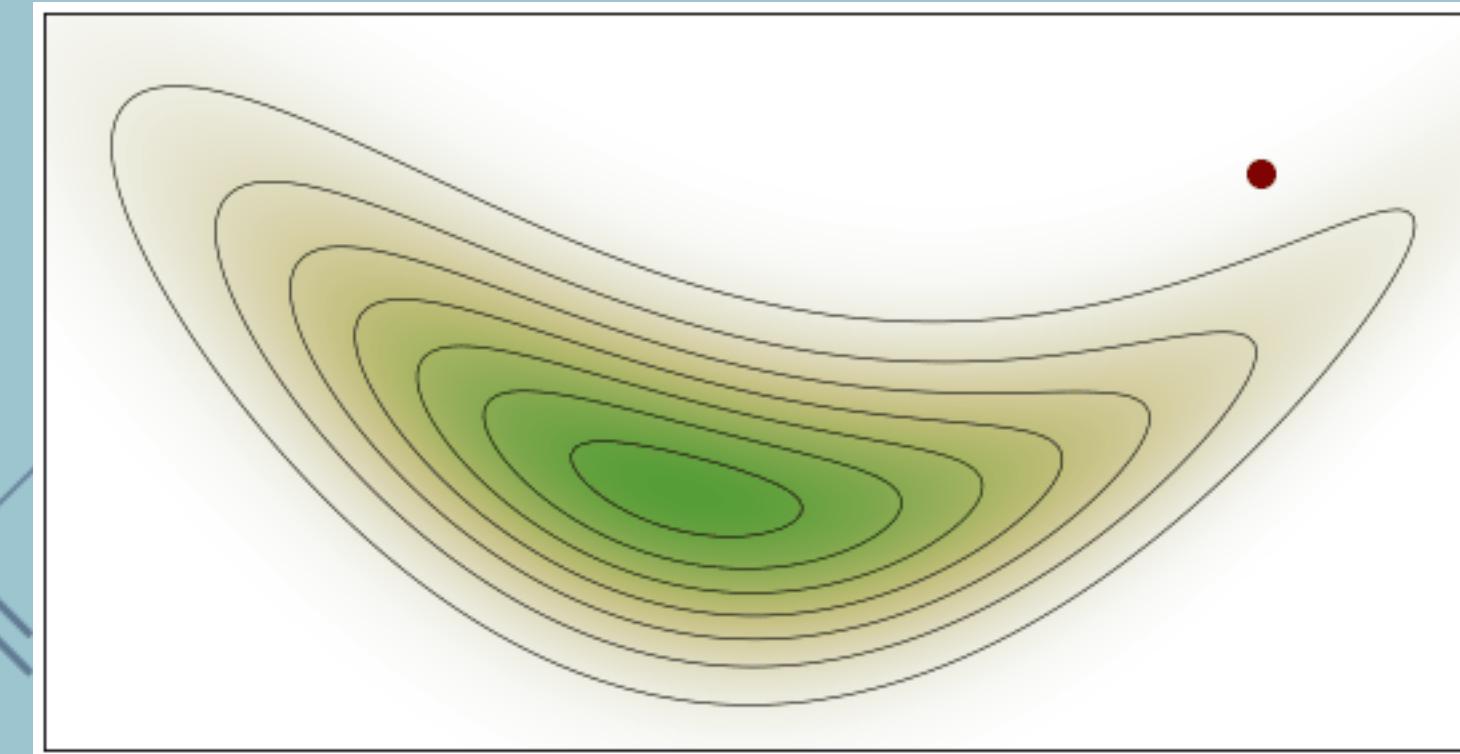
Use Hamiltonian dynamics to propose a move to a new point in state space



# Hamiltonian Monte Carlo

Use Hamiltonian dynamics to propose a move to a new point in state space

Reduces correlations between successive samples by proposing moves to distant states which maintain a high probability of acceptance

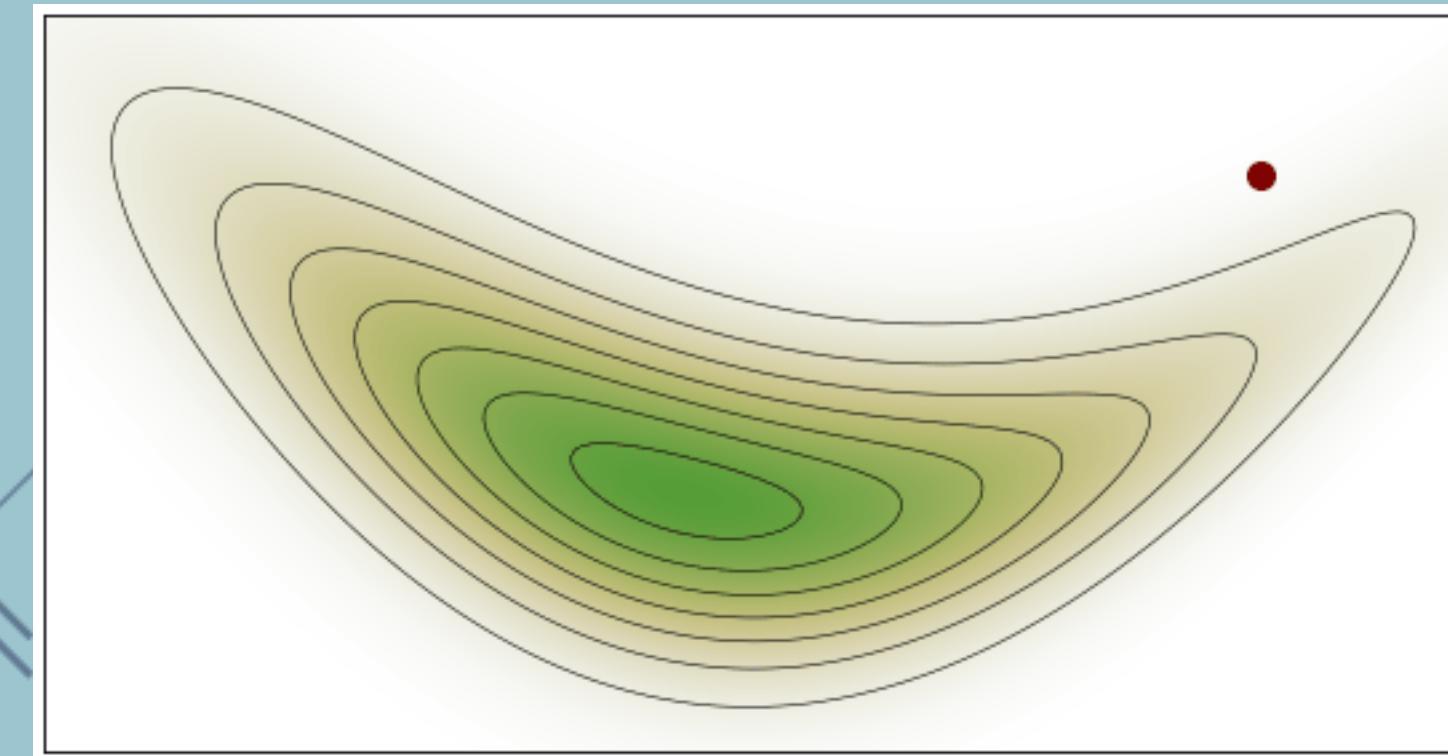


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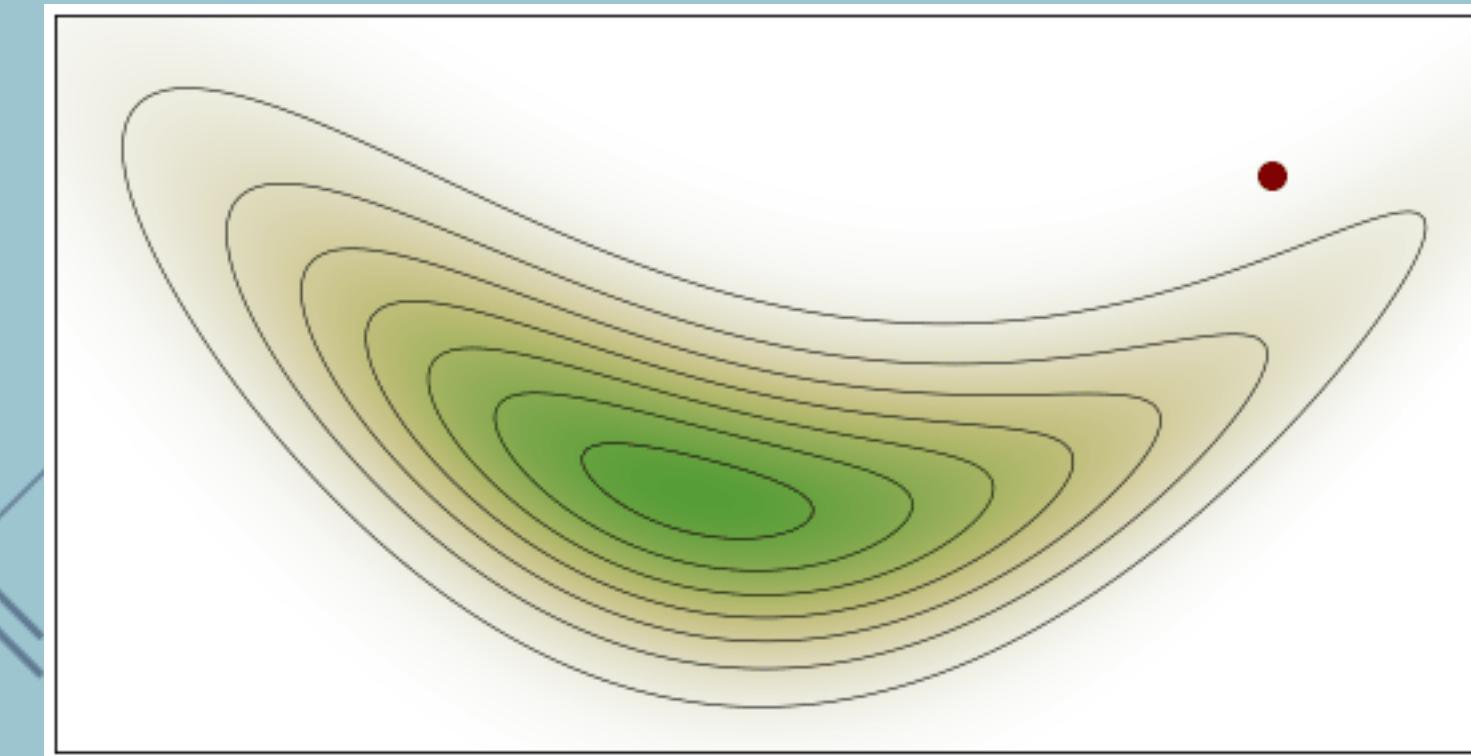
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