A2_DiTong

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1 Assignment 2

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- 1. Imputing age and gender
- (a) I will first use the SurveyIncome data to fit a simple linear regression model and a logistic regression model listed below:

```
age_i = \beta_0 + \beta_1 * totinc_i + \beta_2 * wgt_i

log(P_f/P_m) = \beta_0 + \beta_1 * totinc_i + \beta_2 * wgt
```

Then I will create a total income variable for the BestIncome data by adding up labor income and capital income:

```
best_tot_inc = lab_inc + cap_inc
```

Then, I will plug in the total income variable and the weight variable of the BestIncome dataset to the two regression models above in order to predict the values of age and gender. Finally, I will impute the predicted values of age and gender into the BestIncome dataset.

(b) Here is where I'll use my proposed method from part (a) to impute variables.

```
reg_age = linear_model.LinearRegression()
         reg_age.fit(survey_x, age)
         # Logistic regression of gender on tot_inc and wgt
         reg gender = linear model.LogisticRegression()
         reg_gender.fit(survey_x, gender)
         # Create new variables and datasets needed for prediction and imputation
         best_tot_inc = BestInc.lab_inc + BestInc.cap_inc
         best_x = np.column_stack((best_tot_inc, BestInc.wgt))
         #Impute age and gender using information from the SurvInc
         imp_age = reg_age.predict(best_x)
         imp_gender = reg_gender.predict(best_x)
         imp_BestInc = np.column_stack((BestInc, imp_age, imp_gender))
         # Convert numpy array into panda dataframe
         new BestInc = pd.DataFrame(imp BestInc)
         new_BestInc.columns = ['lab_inc', 'cap_inc', 'hgt', 'wgt', 'age', 'gender']
  c) Here is where I'll report the descriptive statistics for my new imputed variables.
In [20]: # Get descriptive statistics of age
         new_BestInc.age.describe()
Out[20]: count
                  10000.000000
                    44.890828
         mean
         std
                      0.219150
         min
                     43.976495
         25%
                     44.743776
         50%
                     44.886944
         75%
                     45.038991
         max
                     45.703819
         Name: age, dtype: float64
In [21]: # Get descriptive statistics of gender
         new_BestInc.gender.describe()
Out[21]: count
                  10000.000000
                      0.471700
         mean
         std
                      0.499223
         min
                      0.000000
         25%
                      0.000000
         50%
                      0.000000
         75%
                      1.000000
                      1.000000
         max
         Name: gender, dtype: float64
```

Simple linear regression of age on tot_inc and wgt

(d) Correlation matrix for the now six variables

Out[22]: <pandas.io.formats.style.Styler at 0x1e72e98d128>

- 2. Stationarity and data drift
- (a) Estimate by OLS and report coefficients and standard errors on those coefficients.

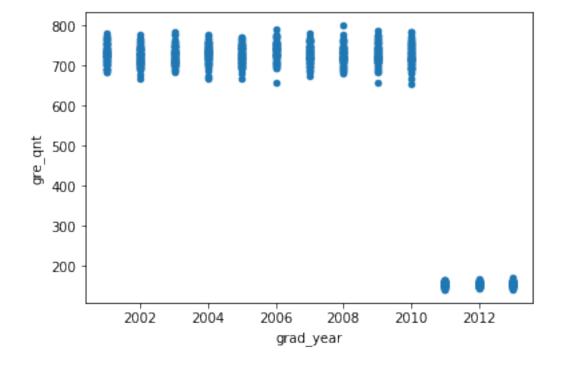
OLS Regression Results

Dep. Variable:		salary_p4		R-squared:			0.263	
Model:		OLS		Adj.	R-squared:		0.262	
Method:		Least Squares		F-sta	tistic:		356.3	
Date:		Tue, 16 Oct 2018		Prob	(F-statisti	c):	3.43e-68	
Time:		23:31:00		Log-L	ikelihood:		-10673.	
No. Observations:		1	000	AIC:			2.135e+04	
Df Residuals:			998	BIC:			2.136e+04	
Df Model:			1					
Covariance Type:		nonrobust						
=======					=======			
	coet	f std err		t	P> t	[0.025	0.975]	
gre_qnt	-25.7632	1.365	-18	3.875	0.000	-28.442	-23.085	
const	8.954e+04	4 878.764	101	.895	0.000	8.78e+04	9.13e+04	

Omnibus:	9.118	Durbin-Watson:	1.424				
<pre>Prob(Omnibus):</pre>	0.010	Jarque-Bera (JB):	9.100				
Skew:	0.230	Prob(JB):	0.0106				
Kurtosis:	3.077	Cond. No.	1.71e+03				

Warnings:

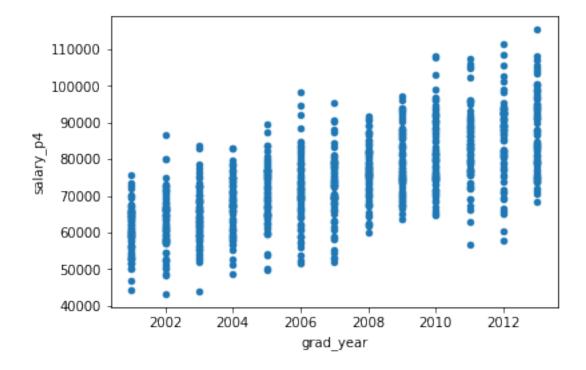
- [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
- [2] The condition number is large, 1.71e+03. This might indicate that there are strong multicollinearity or other numerical problems.
 - (b) Create a scatterplot of GRE score and graduation year.



From the scatterplot above, we can see that there is a huge gap between the pre-2011 data and the post-2011 data. It is obvious that the the change in the GRE score scale rather than a

real drastic "intellengence drop" of GRE takers accounts for this gap. Hence, we will get biased coefficients using this problematic data in the regression that tests our hyphothesis. My solution to this problem is to convert the pre-2011 GRE scores according to the post-2011 scale using this equation: $g_qnt_converted = 130.0 + (g_qnt - 200) * (170 - 130) / (800-200)$

(c) Create a scatterplot of income and graduation year



From the scatterplot above, we can see that the salary data is non-stationary, namely, there is a time trend in salary. Hence, we will get biased coefficients using this problematic data in the regression that tests our hyphothesis. My proposed solution is to detrend the salary data by first

calculate the average growth rate in salaries across all 13 years using the mean salary of each year and then divide each salary by $(1 + avg_growth_rate)$ ** (grad_year - 2001). Here is the equation for calculating the growth rate for each year: growth rate = (mean salary - mean salary of the previous year) / mean salary of the previous year

(d) Re-estimate by OLS with updated variables and report coefficients and standard errors on those coefficients.

```
In [28]: #Define Outcome and Independent Variables
    outcome = 'salary_p4'
    features = ['gre_qnt']

x, y = Inc_Int[features], Inc_Int[outcome]

# Simple linear regression of salary_p4 on gre_qnt
    x = sm.add_constant(x, prepend=False)
    x.head()

m = sm.OLS(y, x)

res = m.fit()

# Report coefficients and SE's
    print(res.summary())
```

OLS Regression Results

```
______
Dep. Variable:
                    salary_p4
                            R-squared:
                                                    0.000
Model:
                        OLS Adj. R-squared:
                                                   -0.001
Method:
                Least Squares F-statistic:
                                                  0.05257
          Tue, 16 Oct 2018 Prob (F-statistic):
Date:
                                                    0.819
Time:
                    23:31:04 Log-Likelihood:
                                                   -10291.
                        1000 AIC:
                                                2.059e+04
No. Observations:
                            BIC:
Df Residuals:
                        998
                                                  2.060e+04
Df Model:
Covariance Type: nonrobust
```

	coef	std err	t	P> t	[0.025	0.975]
gre_qnt const	-9.9681 6.304e+04	43.474 7083.395	-0.229 8.900	0.819	-95.279 4.91e+04	75.343 7.69e+04
Omnibus: Prob(Omnib Skew: Kurtosis:	us):	0.		-):	2.026 0.668 0.716 5.11e+03

Warnings:

- [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
- [2] The condition number is large, 5.11e+03. This might indicate that there are strong multicollinearity or other numerical problems.

The coefficient of gre_qnt changes from -25.7632 to -9.9681, and the coefficient for the constant changes from 8.954e+04 to 6.304e+0.

In the regression in (a), GRE quantitative scores before 2011 are much higher than that after 2011, while the salary presents an increasing trend over the years. Hence, the problematic decreasing pattern of the GRE scores and increasing pattern of the salary might have magnified the negative association between them. After we adjust all GRE scores into the same range and detrend the salary data, the variance of both variables get smaller, and therefore, the negative association between them is lessened.

Since the P value for the coefficient of gre_qnt is very large, we do not have sufficient evidence to reject the null hypothesis that the GRE quantitative score has no effect on salary. Therefore, the results provide evidence against our hypothesis that higher intelligence is associated with higher income. But this regression model itself is problematic, as it does not control for a range of other variables that could affect salary. Hence, the results might not reflect the real situation in the first place.

3. Assessment of Kossinets and Watts.

See next page.

The question of Kossinets and Watts (2009)'s research is: what is the origin of homophily? The more precise version in their own words is that, "on what grounds do individuals selectively make or break some ties over others, and how do these choices shed light on the observation that similar people are more likely to become acquainted than dissimilar people?" (p. 406)

The authors use a network data set comprising interaction, affiliation, and attribute-type longitudinal data of 30,396 students, faculty, and staff in a university community during one academic year.

The data is constructed on the basis of three data sources: "(1) the logs of e-mail interactions within the university over one academic year, (2) a database of individual attributes (status, gender, age, department, number of years in the community, etc.), and (3) records of course registration, in which courses were recorded separately for each semester." (Kossinets and Watts 2009, p. 410)

For the final data set used for analysis, there are 30,396 observations who are selected as active e-mail users from the original sample of 43,553 individuals who used university e-mail to both send and receive messages during the academic year.

While the article reports analysis of only one academic year's worth of data, the full data set spans two calendar years.

A precise descriptions and definitions of all variables can be found in Appendix A.

According to the note on data cleansing and missing values in Appendix B, the error-correction strategies for errors and missing values only yield marginal improvement for the variable "field". Hence, the problem of errors and missing values in the variable field is largely unsolved. Since field is one of the measurements of similarity, this problem might diminish the authors' ability to effectively and accurately test the effect of similarity on homophily.

Due to privacy considerations, the email message content is unavailable to the researchers. Yet, email contact and its frequency per se do not necessarily represent

the existence and intensity of social ties and can not fully measure all dimensions of the concept social tie. Without message content, it is extremely difficult to interpret the meaning of any given pattern of relations.

The authors address this weakness by proposing possible supplementary approaches to construct validity. The first approach is text analysis and validation of inferred network ties through selective surveying of e-mail users. Alternatively, they can also obtain message content data for analysis through informed consent procedures under which users would be willing to provide content in exchange for benefits as well as assurances on the use of the content.