

3 Comparison

This is an analysis of the LASSO, BLASSO, and Horseshoe priors on diabetes data.

1. LASSO

The LASSO is an L1-penalized least squares estimate of β , with \tilde{y} as the mean-centered outcome variable, where the following is optimized:

$$\min_{\beta} (\tilde{y} - XB)'(\tilde{y} - XB) + \lambda \sum_{j=1}^p |\beta_j|$$

(a) Strengths

- i. Easy to implement.

(b) Weaknesses

- i. Proposed standard error estimators are not considered satisfactory.

2. BLASSO

The BLASSO interprets the L1-penalty as a prior on β and σ^2 . In this case, y is normal, the prior $\pi(\beta|\sigma^2)$ is Laplace (i.e. double exponential), and the prior $\pi(\sigma^2)$ is non-informative and defined as such to ensure a unimodal posterior:

$$\begin{aligned} y|\mu, X, \beta, \sigma^2 &\sim N(\mu 1_n + X\beta, \sigma^2 I_n) \\ \pi(\beta|\sigma^2) &= \prod_{j=1}^p \frac{\lambda}{2\sqrt{\sigma^2}} e^{-\lambda|\beta_j|/\sqrt{\sigma^2}} \quad (\text{Laplace distribution}) \\ \pi(\sigma^2) &= \frac{1}{\sigma^2} \end{aligned}$$

(a) Strengths

- i. Easy to implement
- ii. Automatically provides interval estimates for all parameters, including the error variance.

(b) Weaknesses

- i. More computationally intensive than LASSO.

3. Horseshoe

The Horseshoe (HS) prior assumes $y|\theta \sim N(\theta, \sigma^2 I)$, and aims to (1) estimate θ , and (2) predict future realizations of y . It is especially useful in cases where most covariates are nearly zero (i.e. sparse θ). The model is as follows:

$$\begin{aligned} \theta_i|\lambda_i &\sim N(0, \lambda_i^2) \\ \lambda_i|\tau &\sim C^+(0, \tau) \\ \tau &\sim C^+(0, \sigma) \\ E(\theta_i|y) &= \int_0^1 (1 - \kappa_i) y_i p(\kappa_i|y) d\kappa_i = (1 - E(\kappa_i|y)) y_i \end{aligned}$$

Where $\kappa_i = 1/(1 + \lambda_i^2)$, assuming fixed values $\sigma^2 = \tau^2 = 1$. Given $\lambda_i \geq 0$, this implies that λ_i 's are indirectly related to amount of shrinkage (e.g. high λ_i means low κ_i , and less shrinkage).

- (a) Strengths
 - i. Easy to implement.
 - ii. Has fewer hyperparameters.
 - iii. Robust and flexible to high or low sparsity situations.
 - iv. Converges efficiently.
- (b) Weaknesses
 - i.
 - ii.

4 Perform Gibbs Sampling for Bayesian Lasso

Reference: Park and Casella (2008) (Section 2).

1. Model and Prior for BLASSO

See 3.2, above.

2. Complete Conditional for each (set of) Parameters

To get complete conditionals, the prior $\pi(\beta|\sigma^2)$ is represented differently, as a scaled mixture of normals (with exponential mixing density). Again, the Laplace can be re-written as the integrated product of a Normal and Exponential distribution:

$$\begin{aligned}\pi(\beta|\sigma^2) &= \prod_{j=1}^p \frac{\lambda}{2\sqrt{\sigma^2}} e^{-\lambda|\beta_j|/\sqrt{\sigma^2}} && \text{(Laplace distribution)} \\ &= \prod_{j=1}^p \int_0^\infty \text{Normal} * \text{Exponential}\end{aligned}$$

This way, full conditionals become:

$$\begin{aligned}y|\mu, X, \beta, \sigma^2 &\sim N(\mu 1_n + X\beta, \sigma^2 I_n), \\ \beta|\sigma^2, \tau_1^2, \dots, \tau_p^2 &\sim N_p(0_p, \sigma^2 D_\tau), && \text{(Normal part of Laplace)} \\ \text{where } D_\tau &= \text{diag}(\tau_1^2, \dots, \tau_p^2), \\ \sigma^2, \tau_1^2, \dots, \tau_p^2 &\sim \pi(\sigma^2) d\sigma^2 \prod_{j=1}^p \frac{\lambda^2}{2} e^{\lambda^2 \tau_j^2 / 2} d\tau_j^2, && \text{(Exponential part of Laplace)} \\ \sigma^2, \tau_1^2, \dots, \tau_p^2 &> 0.\end{aligned}$$

Results and Discussion

Our model is run with the following command:

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> # R code here!
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