

# MCMC HW 3

Maurice Diesendruck

February 27, 2015

(a)

$$Pr(y_i = 1|X_i) = \Phi(X_i'\beta). \quad (1)$$

$$y_i = I(z_i > 0) \quad (2)$$

$$z_i \sim N(X_i'\beta, 1) \quad (3)$$

Show that (1)  $\equiv$  (2) and (3):

First, " $\Leftarrow$ " (2) and (3) imply

$$y_i = \begin{cases} 1 : & z_i > 0, \text{ or when } z_i = X_i'\beta + \epsilon_i > 0 \Rightarrow \epsilon_i > -X_i'\beta \\ 0 : & \text{otherwise} \end{cases}$$

Where  $z_i = X_i'\beta + \epsilon_i$ ,  $\epsilon \sim N(0, 1)$ .

So,  $P(y_i = 1|X_i) = P(z_i > 0) = P(\epsilon_i > -X_i'\beta) = P(\epsilon_i < X_i'\beta) = \Phi(X_i'\beta) \equiv (1)$ .

The inequality in the penultimate step is allowed to switch signs by the symmetry of the Normal distribution.

(b) Find conditional posterior  $P(z_i|\beta, y)$  and  $P(\beta|z, y)$ , using prior on  $\beta$ ,  $P(\beta) = 1$ .

The conditional posterior of  $z_i$  is identical to (3), except that the given value of  $y_i$  informs the value of  $z_i$ , by definition.

$$z_i|\beta, y \sim \begin{cases} N(X_i'\beta, 1), \text{ left truncated at } 0, \text{ if } y_i = 1 \\ N(X_i'\beta, 1), \text{ right truncated at } 0, \text{ if } y_i = 0 \end{cases}$$

In summary, if  $y_i = 1$  then the distribution is the right side of the normal distribution centered at  $X_i'\beta$ , whereas if  $y_i = 0$  then the distribution is the left side.

The conditional posterior of  $\beta$ , given a flat, uninformative prior, is simply the standard least squares result on the Normal  $z$ :

$$\beta|z, y \sim N(\hat{\beta}_z, 1 \cdot (X'X)^{-1}) \quad (4)$$

$$\text{Where } \hat{\beta}_z = (X'X)^{-1}X'z \quad (5)$$

(c) Gibbs sampler proposition:

Initialize value for  $\beta$  as MLE, and call it  $\beta_0$ . Then set  $\beta_0 = \beta_k$ .

1. Sample  $z_k$  from  $P(z_k|\beta_k, y)$ .
2. Sample  $\beta_{k+1}$  from  $P(\beta_{k+1}|z_k, y)$ .
3. Repeat until convergence and good mixing.

After discarding values from the beginning of the chain (the burn in period), the resulting values of  $\beta$  and  $z$  will be an estimate of the exact joint posterior distribution  $P(\beta, z|y)$ .

(d) Show the histogram of simulated  $\beta$  values as an estimate of  $p(\beta_j|y)$ .

