

StatMod2 - Hierarchical Models - Exercises 4

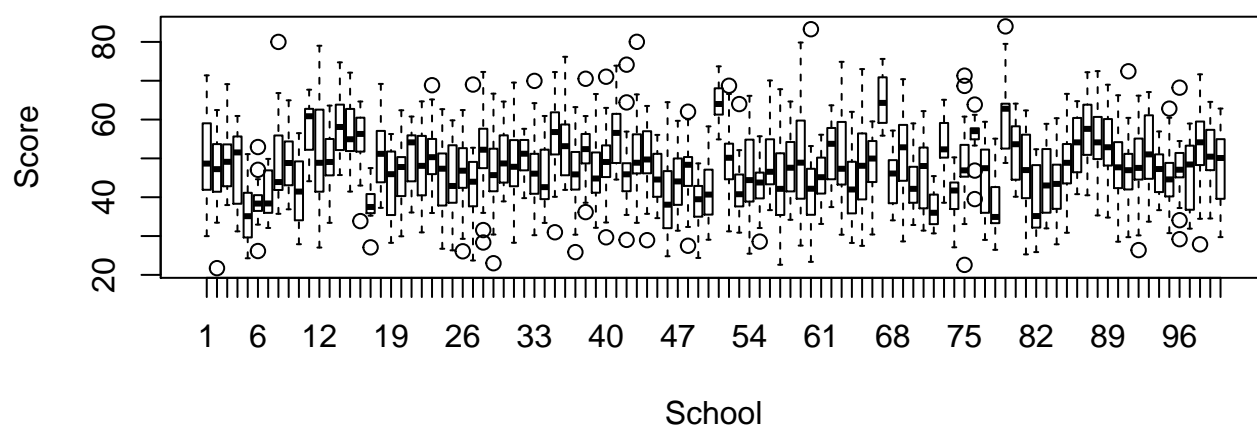
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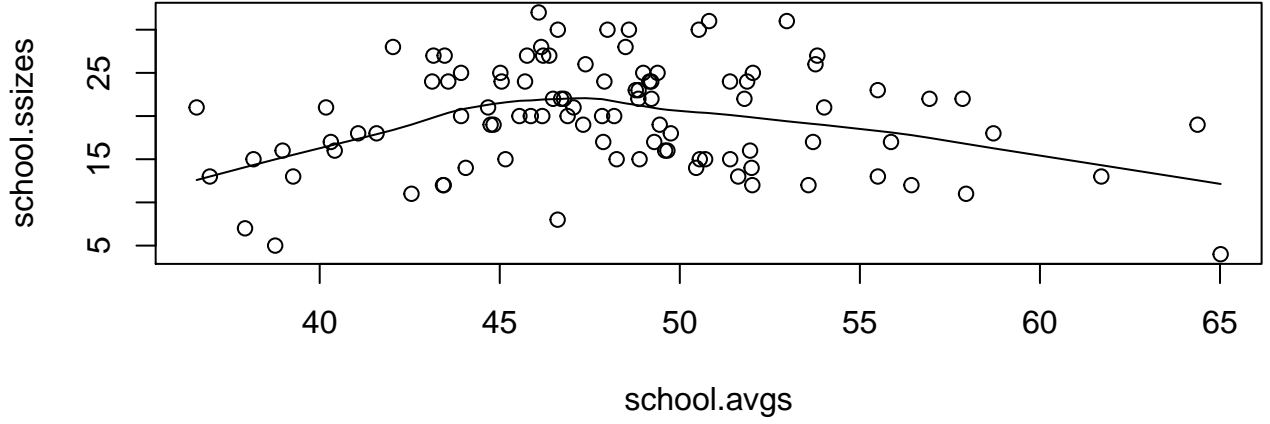
1 School Averages and Sample Size

Larger samples tend to "smooth" out extreme scores, so small samples are more likely to be extreme.

```
data <- read.csv("mathtest.csv")
attach(data)
boxplot(mathscore ~ school, xlab="School", ylab="Score")
```



```
school.avgs <- aggregate(data, list(school=school), mean)[,3]
school.sizes <- aggregate(data, list(school=school), length)[,3]
plot(cbind(school.avgs, school.sizes))
fit <- lowess(cbind(school.avgs, school.sizes))
lines(fit)
```



2 Normal Hierarchical Model with Gibbs Sampling

2.1 Model

For school $i = 1, \dots, p$; student $j = 1, \dots, n_i$; and $\sum n_i = n$; let $a = b = c = d = 1$, and note that for this data, $p = 100$ and $n = 1993$:

$$\begin{aligned}
 y_{ij} &\sim N(\theta_i, \sigma^2) \\
 \theta_i &\sim N(\mu, \tau^2) \\
 \mu &\sim Unif(40, 60) \\
 \sigma^2 &\sim InvGa(a, b) \\
 \tau^2 &\sim InvGa(c, d)
 \end{aligned}$$

Or alternatively,

$$\begin{aligned}
 \bar{y}_i &\sim N(\theta_i, \frac{\sigma^2}{n_i}) \\
 \theta_i &\sim N(\mu, \tau^2) \\
 \mu &\sim Unif(40, 60) \\
 \sigma^2 &\sim InvGa(a, b) \\
 \tau^2 &\sim InvGa(c, d)
 \end{aligned}$$

$$N(\bar{\mathbf{y}}|\boldsymbol{\theta}, \sigma^2 \text{diag}(1/n_i)) N(\boldsymbol{\theta}|\boldsymbol{\mu}, \tau^2 I_p) Unif(\mu|40, 60) Ga(\frac{1}{\sigma^2}|a, b) Ga(\frac{1}{\tau^2}|c, d)$$

2.2 Joint and Posterior Distributions

Let $\boldsymbol{\nu} = (\boldsymbol{\theta}, \boldsymbol{\mu}, \sigma^2, \tau^2)$; $\boldsymbol{\theta} = \theta_1, \dots, \theta_p$; $\boldsymbol{\sigma}_*^2 = \sigma^2 \text{diag}(1/n_i)$; $\bar{\mathbf{y}}$ be the $p \times 1$ vector of mean math scores for all schools; and note that $\boldsymbol{\mu}$ is a $p \times 1$ vector with a single value, μ .

$$\begin{aligned}
P(\boldsymbol{\nu}|\mathbf{y}) &\propto P(\mathbf{y}|\boldsymbol{\nu})P(\boldsymbol{\nu}) \\
&= N(\bar{\mathbf{y}}|\boldsymbol{\theta}, \boldsymbol{\sigma}_*^2)N(\boldsymbol{\theta}|\boldsymbol{\mu}, \tau^2 I_p) \text{Unif}(\mu|40, 60)Ga\left(\frac{1}{\sigma^2}|a, b\right)Ga\left(\frac{1}{\tau^2}|c, d\right) \\
&= \left(\frac{1}{\boldsymbol{\sigma}_*^2}\right)^{p/2} \exp\left(-\frac{1}{2} \frac{(\bar{\mathbf{y}} - \boldsymbol{\theta})'(\bar{\mathbf{y}} - \boldsymbol{\theta})}{\boldsymbol{\sigma}_*^2}\right) \\
&\quad \times \left(\frac{1}{\tau^2}\right)^{p/2} \exp\left(-\frac{1}{2} \left(\frac{(\boldsymbol{\theta} - \boldsymbol{\mu})'(\boldsymbol{\theta} - \boldsymbol{\mu})}{\tau^2}\right)\right) \\
&\quad \times \left(\frac{1}{\sigma^2}\right)^{a-1} \exp\left(-b\left(\frac{1}{\sigma^2}\right)\right) \\
&\quad \times \left(\frac{1}{\tau^2}\right)^{c-1} \exp\left(-d\left(\frac{1}{\tau^2}\right)\right)
\end{aligned}$$

Conditional distributions are as follows:

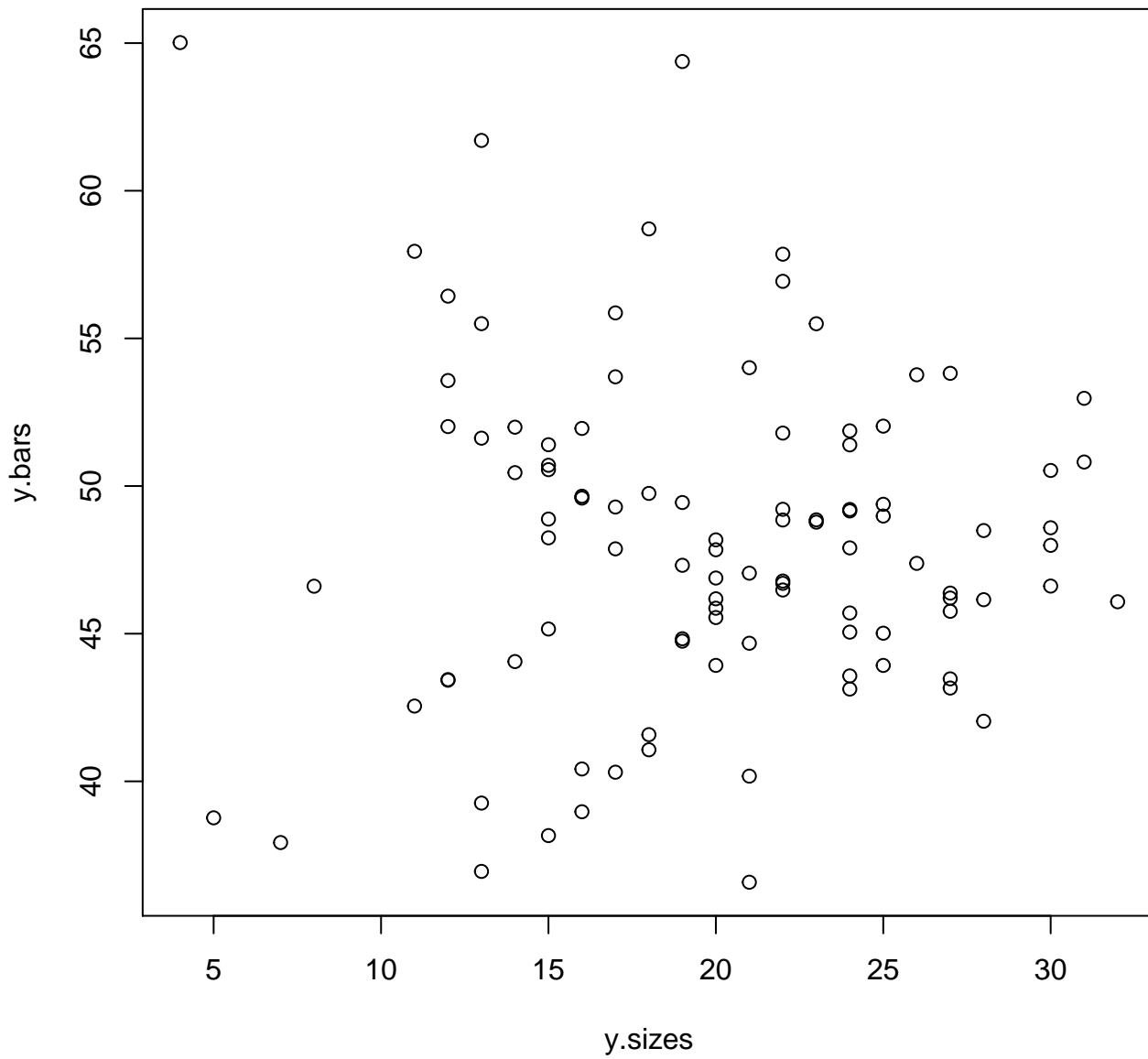
$$\begin{aligned}
P(\boldsymbol{\theta}|\bar{\mathbf{y}}, \boldsymbol{\mu}, \sigma^2, \tau^2) &\propto \exp\left(-\frac{1}{2}\left(\frac{(\bar{\mathbf{y}} - \boldsymbol{\theta})'(\bar{\mathbf{y}} - \boldsymbol{\theta})}{\boldsymbol{\sigma}_*^2} + \frac{(\boldsymbol{\theta} - \boldsymbol{\mu})'(\boldsymbol{\theta} - \boldsymbol{\mu})}{\tau^2}\right)\right) \\
&= \exp\left(-\frac{1}{2}\left(\frac{\bar{\mathbf{y}}'\bar{\mathbf{y}}}{\boldsymbol{\sigma}_*^2} - \frac{\bar{\mathbf{y}}'\boldsymbol{\theta}}{\boldsymbol{\sigma}_*^2} - \frac{\boldsymbol{\theta}'\bar{\mathbf{y}}}{\boldsymbol{\sigma}_*^2} + \frac{\boldsymbol{\theta}'\boldsymbol{\theta}}{\boldsymbol{\sigma}_*^2}\right) + \left(\frac{\boldsymbol{\theta}'\boldsymbol{\theta}}{\tau^2} - \frac{\boldsymbol{\theta}'\boldsymbol{\mu}}{\tau^2} - \frac{\boldsymbol{\mu}'\boldsymbol{\theta}}{\tau^2} + \frac{\boldsymbol{\mu}'\boldsymbol{\mu}}{\sigma^2}\right)\right) \\
&\sim N\left(\left(\frac{I_n}{\boldsymbol{\sigma}_*^2} + \frac{I_n}{\tau^2}\right)^{-1}\left(\frac{\bar{\mathbf{y}}}{\boldsymbol{\sigma}_*^2} - \frac{\boldsymbol{\mu}}{\tau^2}\right), \left(\frac{I_n}{\boldsymbol{\sigma}_*^2} + \frac{I_n}{\tau^2}\right)^{-1}\right)
\end{aligned}$$

$$\begin{aligned}
P(\sigma^2|\mathbf{y}, \boldsymbol{\theta}, \tau^2) &\propto \left(\frac{1}{\sigma^2}\right)^{n/2} \exp\left(-\frac{1}{2}\left(\frac{(\bar{\mathbf{y}} - \boldsymbol{\theta})'(\bar{\mathbf{y}} - \boldsymbol{\theta})}{\boldsymbol{\sigma}_*^2}\right)\right) \\
&\quad \times \left(\frac{1}{\sigma^2}\right)^{a-1} \exp\left(-b\left(\frac{1}{\sigma^2}\right)\right) \\
&= \left(\frac{1}{\sigma^2}\right)^{n/2+a-1} \exp\left(-\left(\frac{(\bar{\mathbf{y}} - \boldsymbol{\theta})'(\bar{\mathbf{y}} - \boldsymbol{\theta})}{2\text{diag}(1/n_i)} + b\right)\left(\frac{1}{\sigma^2}\right)\right) \\
&\sim \text{Ga}\left(n/2 + a, \left(\frac{(\bar{\mathbf{y}} - \boldsymbol{\theta})'(\bar{\mathbf{y}} - \boldsymbol{\theta})}{2\text{diag}(1/n_i)} + b\right)\right)
\end{aligned}$$

$$\begin{aligned}
P(\tau^2|\mathbf{y}, \boldsymbol{\theta}, \sigma^2) &\propto \left(\frac{1}{\tau^2}\right)^{p/2} \exp\left(-\frac{1}{2}\left(\frac{(\boldsymbol{\theta} - \boldsymbol{\mu})'(\boldsymbol{\theta} - \boldsymbol{\mu})}{\tau^2}\right)\right) \\
&\quad \times \left(\frac{1}{\tau^2}\right)^{c-1} \exp\left(-d\left(\frac{1}{\tau^2}\right)\right) \\
&= \left(\frac{1}{\tau^2}\right)^{p/2+c-1} \exp\left(-\left(\frac{(\boldsymbol{\theta} - \boldsymbol{\mu})'(\boldsymbol{\theta} - \boldsymbol{\mu})}{2} + d\right)\left(\frac{1}{\tau^2}\right)\right) \\
&\sim \text{Ga}\left(p/2 + c, \left(\frac{(\boldsymbol{\theta} - \boldsymbol{\mu})'(\boldsymbol{\theta} - \boldsymbol{\mu})}{2} + d\right)\right)
\end{aligned}$$

$$\begin{aligned}
P(\boldsymbol{\mu}|\mathbf{y}, \boldsymbol{\theta}, \sigma^2, \tau^2) &\propto \exp\left(-\frac{1}{2}\left(\frac{(\boldsymbol{\theta} - \boldsymbol{\mu})'(\boldsymbol{\theta} - \boldsymbol{\mu})}{\tau^2}\right)\right) \\
&\sim N(\bar{\boldsymbol{\theta}}, \tau^2)
\end{aligned}$$

2.3 Gibbs Sampler



```
## Error in eval(expr, envir, enclos): object 'ybars' not found
```

2.4 Shrinkage

In general, the smaller the sample size, the more extreme the sample mean, and the larger the shrinkage factor.