

SSIE 660: Stochastic Systems
 Dr. Sung H. Chung
 Note 16
 Chapter 7. Renewal Theory - Cont'd

Relationship between $m(t)$, mean number of renewals by time t , and $E[S_{N(t)+1}]$, the expected time of the first renewal after t .

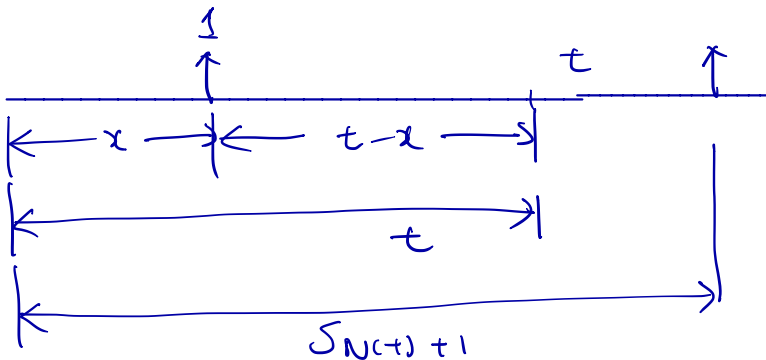
- Let $g(t) = E[S_{N(t)+1}]$.
- The first renewal occurs by time x . That is,

$$X_1 = x$$

- Conditioning on the time of the first renewal,

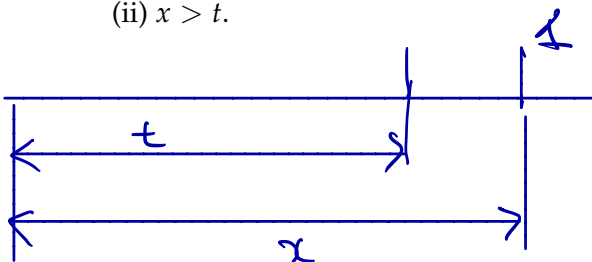
$$g(t) = \int_0^{\infty} E[S_{N(t)+1} | X_1 = x] f_{X_1}(x) dx$$

(i) $x < t$.



$$\begin{aligned} E[S_{N(t)+1} | X_1 = x] \\ &= x + E[S_{N(t-x)+1}] \\ &= x + g(t-x) \end{aligned}$$

(ii) $x > t$.



$$E[S_{N(t)+1} | X_1 = x] = x$$

$$g(t) = \int_0^t [x + g(t-x)] f_X(x) dx + \int_t^\infty x f_X(x) dx = \int_0^\infty x f_X(x) dx + \int_0^t g(t-x) f_X(x) dx$$

$$= E[X] + \int_0^t g(t-x) f_X(x) dx$$

Renewal equation:

$$g(t-x) = \mu g_1(t-x) + 1 \quad m(t) = F_X(t) + \int_0^t m(t-x) f_X(x) dx$$

$$\text{Let } g_1(t) = \frac{g(t)}{\mu} - 1 = \frac{E[S_{N(t)+1}]}{\mu} - 1 = \frac{1}{\mu} \int_0^t g(t-x) f_X(x) dx$$

$$= \int_0^t [1 + g_1(t-x)] f_X(x) dx$$

$$\therefore g_1(t) = F_X(t) + \int_0^t g_1(t-x) f_X(x) dx$$

g_1 satisfies renewal equation uniqueness

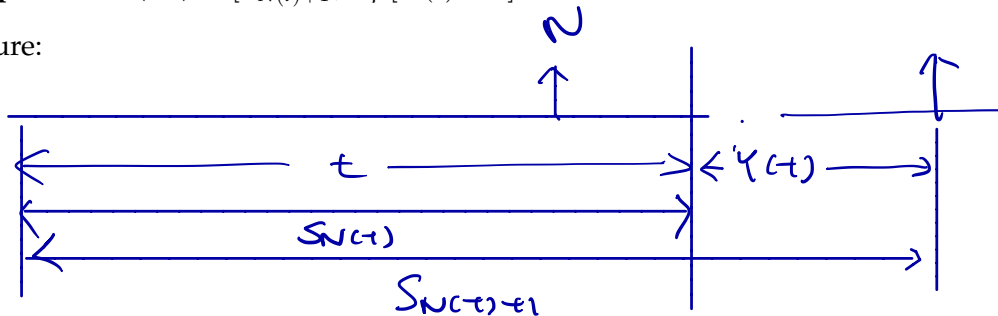
$$\therefore g_1(t) = m(t)$$

$$\underline{E[S_{N(t)+1}] = \mu [m(t) + 1] = g(t)}$$

\hookrightarrow proposition 7.2

Proposition 1 (7.2). $E[S_{N(t)+1}] = \mu[m(t) + 1]$

Figure:



$Y(t) =$ time from t until the next arrival
 $S_{N(t)+1} = t + Y(t)$

Taking expectations:

$$g(t) = E[S_{N(t)+1}] = t + E[Y(t)] = \mu[m(t) + 1]$$

divide by
 nt

$$\frac{m(t)}{t} + \frac{1}{t} = \frac{1}{\mu} + \frac{E[Y(t)]}{t} \cdot \frac{1}{\mu}$$

if $\frac{E[Y(t)]}{t} \rightarrow 0$ as $t \rightarrow \infty$ $\frac{m(t)}{t} \rightarrow \frac{1}{\mu}$ as $t \rightarrow \infty$

↑
 show this to
 prove the

elementary
renewal theorem

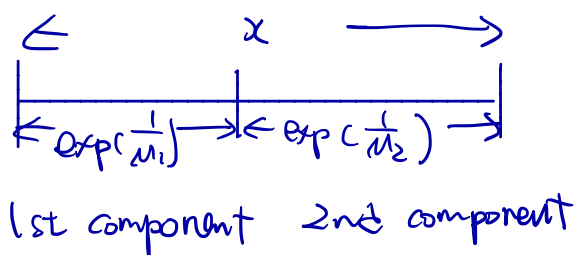
Example 2. Consider the renewal process whose inter-arrival distribution is the convolution of two exponentials. Find $m(t)$.

do this by determining $E[Y(t)]$: $Y(t)$: remaining life
 $F_i(t) = 1 - e^{-\mu_i t}$ $i=1, 2$

$$F = F_1 * F_2$$

$$F(t) = \int_0^t F_1(z) F_2(t-z) dz$$

Figure:

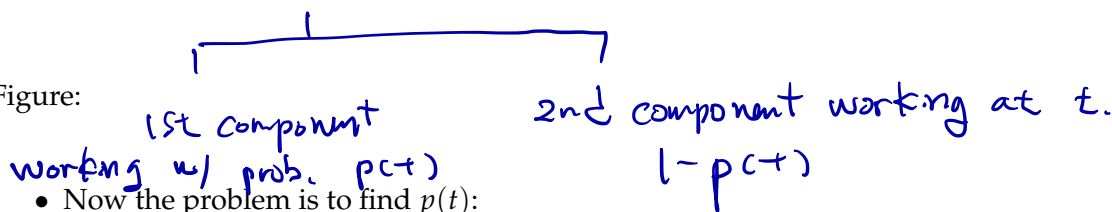


renewal = new machine
 each machine has two compo.
 initially 1st compo is used.
 then 2nd. after 2nd fails,
 another machine will be
 put. (renewal)

$$EX] = \mu = \frac{1}{\mu_1} + \frac{1}{\mu_2}$$

$$E[Y(t)] = E[\text{remaining life of the machine part}]$$

Figure:



- Now the problem is to find $p(t)$:

Prob $[1^{\text{st}} \text{ component is employed at } t]$

- Let us define a stochastic process $\{X(t), t \geq 0\}$ such that

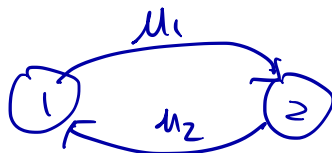
$$X(t) = \begin{cases} 1 & \text{if component 1 is employed at } t \\ 2 & \text{if component 2 is employed at } t \end{cases}$$

if 1st is still in use: remaining life is $\frac{1}{\mu_1} + \frac{1}{\mu_2}$

if 2nd is being used: $\frac{1}{\mu_2}$

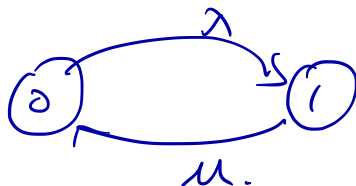
$$\therefore E[Y(t)] = \left(\frac{1}{\mu_1} + \frac{1}{\mu_2}\right) p(t) + \frac{1}{\mu_2} (1-p(t))$$

- What 'stochastic process' does $X(t)$ follow? Continuous time Markov Chain



$$P_{11}(t) = \frac{\mu_1}{\mu_1 + \mu_2} e^{-(\mu_1 + \mu_2)t} + \frac{\mu_2}{\mu_1 + \mu_2}$$

- Consider the following problem:



- We found $P_{00}(t)$ and $P_{10}(t)$ (We could also find $P_{01}(t)$ and $P_{11}(t)$)

$$P_{00}(t) = \frac{\mu}{\mu + \lambda} + \frac{\lambda}{\mu + \lambda} e^{-(\mu + \lambda)t}, \quad P_{10}(t) = \frac{\mu}{\mu + \lambda} - \frac{\mu}{\mu + \lambda} e^{-(\mu + \lambda)t}$$

$$E[Y(t)] = \frac{1}{\mu_2} + \frac{1}{\mu_1} \left[\frac{\mu_2}{\mu_1 + \mu_2} + \frac{\mu_1}{\mu_1 + \mu_2} e^{-(\mu_1 + \mu_2)t} \right]$$

$$= \frac{1}{\mu_2} + \frac{\mu_2}{\mu_1(\mu_1 + \mu_2)} + \frac{1}{\mu_1 + \mu_2} e^{-(\mu_1 + \mu_2)t}$$

$$\mu[m(t) + 1] = t + E[Y(t)]$$

$$m(t) = \frac{t}{\mu} - 1 + \frac{E[Y(t)]}{\mu}$$

$$\mu = E[X] = \frac{1}{\mu_1} + \frac{1}{\mu_2} = \frac{\mu_1 + \mu_2}{\mu_1 \mu_2}$$

$$m(t) = t \frac{\mu_1 \mu_2}{\mu_1 + \mu_2} - 1 + \frac{\mu_1 \mu_2}{(\mu_1 + \mu_2)^2} + \frac{\mu_2}{\mu_1(\mu_1 + \mu_2)} \frac{\mu_1 \mu_2}{(\mu_1 + \mu_2)}$$

$$+ \frac{\mu_1 \mu_2}{(\mu_1 + \mu_2)^2} e^{-(\mu_1 + \mu_2)t}$$

$$= t \frac{\mu_1 \mu_2}{\mu_1 + \mu_2} - \frac{\mu_1 \mu_2}{(\mu_1 + \mu_2)^2} [1 - e^{-(\mu_1 + \mu_2)t}]$$

Renewal Reward Process

- A reward is received each time a renewal occurs. Specifically, a reward of R_n is received at the time of the n^{th} renewal.
- $R_n, n \geq 1$ are independent and identically distributed.
- But R_n may depend on X_n , the length of the n^{th} renewal interval.
- The total reward earned by time t is $R(t)$.

$$R(t) = \sum_{n=1}^{N(t)} R_n$$

- Let $E[R] = E[R_n]$ and $E[X] = E[X_n]$.

Proposition 3. If $E[R] < \infty$ and $E[X] < \infty$, then

1. with probability 1, $\lim_{t \rightarrow \infty} \frac{R(t)}{t} = \frac{E[R]}{E[X]}$

2. $\lim_{t \rightarrow \infty} \frac{E[R(t)]}{t} = \frac{E[R]}{E[X]}$

Proof.
$$\frac{R(t)}{t} = \frac{\sum_{n=1}^{N(t)} R_n}{t} = \left(\frac{\sum_{n=1}^{N(t)} R_n}{N(t)} \right) \left(\frac{N(t)}{t} \right)$$

↓ $E[R]$ by strong law of large numbers as $t \rightarrow \infty$

and by proposition 7.1

$$\frac{N(t)}{t} \rightarrow \frac{1}{\mu} = \frac{1}{E[X]} \text{ as } t \rightarrow \infty$$

Remark:

1. If we say that a cycle is completed every time a renewal occurs, then Proposition 3 states that the long-run average reward per unit time is equal to $\frac{\text{the expected reward earned during a cycle}}{\text{the expected length of a cycle}}$
2. This result is valid when the reward is earned gradually throughout the renewal cycle, also.

Example 4 (A Car Buying Model). The life time of a car is continuous random variable having a distribution H and probability density h . Mr. Brown has a policy that he buys a new car as soon as his old one either breaks down or reaches age of T years. Suppose that a new car costs C_1 dollars and also that an additional cost of C_2 dollars is incurred whenever Mr. Brown's car breaks down. Under the assumption that a used car has no resale value, what is Mr. Brown's long-run average cost?

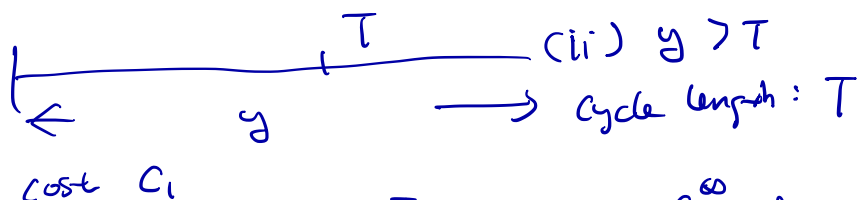
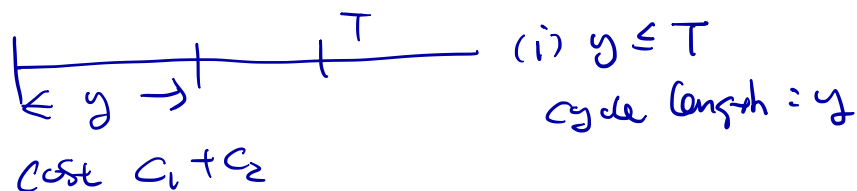
Find optimal T to minimize the long-run average cost.

Renewal = Brown buying a new car

a cycle completes everytime Brown buying a new car

$$\frac{\text{long-run average cost}}{\text{unit time}} = \frac{E[\text{cost per cycle}]}{E[\text{length of a cycle}]}$$

$$y \equiv \text{life of the car} \sim h_Y(y)$$



$$E[\text{length of a cycle}] = \int_0^T y \cdot h_Y(y) dy + \int_T^\infty T h_Y(y) dy$$

$$\begin{aligned} \text{length} & \begin{cases} y & \text{if } y \leq T \\ T & \text{if } y > T \end{cases} \\ & = \int_0^T y \cdot h_Y(y) dy + T(1 - H_Y(T)) \end{aligned}$$

$$\begin{aligned} E[\text{cost for cycle}] &= C_1 P[y > T] + (C_1 + C_2) P[y \leq T] \\ &= C_1 (P[y > T] + P[y \leq T]) + C_2 P[y \leq T] \end{aligned}$$

$$= C_1 + C_2 H(T)$$

\therefore Mr Brown's long-run average cost will be

$$\frac{C_1 + C_2 H(T)}{\int_0^T y h(y) dy + T[1 - H(T)]}$$

Now, suppose that $h(y) \sim \text{uniform}(0, 10)$ $h(y) = \frac{1}{10}$
 $C_1 = 3K$ $C_2 = \frac{1}{2}K$ $H(T) = \int_0^T \frac{1}{10} dy$

if $T \leq 10$, then $3 + \frac{1}{2}(T/10) = \frac{T}{10}$

$$\int_0^T \left(\frac{y}{10}\right) dy + T\left(1 - \frac{T}{10}\right)$$

$$= \frac{3 + T/20}{T^2/20 + (10T - T^2)/10} = \frac{60 + T}{20T - T^2}$$

$$\text{let } g(T) = \frac{60 + T}{20T - T^2} \quad g'(T) = \frac{(20T - T^2) - (60 + T)(20 - 2T)}{(20T - T^2)^2} = 0$$

$$\rightarrow 20T - T^2 = (60 + T)(20 - 2T)$$

$$\rightarrow T^2 + 120T - 1200 = 0$$

$$T = 9.25 \text{ or } -129.25$$