

SSIE 660: Stochastic Systems  
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Note 17  
Chapter 8. Queueing Theory

Introduction: Queueing System

- Customers arrive in some random manner.
- An arrival will be served by the server (if the server is free).
- They are made to wait in queue (if the server is busy).
- Once served, they are generally assumed to leave the system.

Notation:

$L$  : average number of customers in the system.

$L_Q$  : average number of customers waiting in queue.

$W$  : average amount of time a customer spends in the system.

$W_Q$  : average amount of time a customer spends waiting in queue.

$\lambda_a$  : average arrival rate of entering customers.

$$\lambda_a = \lim_{t \rightarrow \infty} \frac{N(t)}{t}$$

*elementary renewal theorem*

$$\frac{m(t) = E[N(t)]}{t} = \frac{1}{\mu}$$

Imagine that entering customers are forced to pay money (according to some rule) to the system.

Basic cost identity :

average rate at which the system earns =  $\lambda_a \times$  average amount an entering customer pays

Little's formula :

↓  
# of  
customers

$$L = \lambda_a W$$

time

Also,

$$L_Q = \lambda_a W_Q$$

Let us define:

$E[S]$  : average amount of time a customer spends in service.

Then,

$$\lambda_a E[S] = \text{average number of customers in service}$$

### Steady-State Probabilities

- $X(t)$  : the number of customers in the system at time  $t$
- $P_n = \lim_{t \rightarrow \infty} P\{X(t) = n\}, n \geq 0.$

Steady-probability that there will be exactly  $n$  customers in the system

= long-run proportion of time that the system contains exactly  $n$  customers

- $a_n, n \geq 0$  : proportion of customers that find  $n$  in the system when they arrive.
- $d_n, n \geq 0$  : proportion of customers leaving behind  $n$  in the system when they depart.

**Example 1.** Consider a queueing model in which all customers have service times equal to 1, and where the times between successive customers are always greater than 1.

1. Find  $a_0, d_0$ .
2. Is  $a_0 = P_0$ ?

Since every arrival finds the system empty and every departure leaves it empty, we have:

$$a_0 = 1, d_0 = 1$$

However,

$$P_0 \neq 1$$

as the system is not always empty of customers.

**Proposition 2.** *In any system in which customers arrive and depart one at a time,*

rate at which arrivals find  $n$  = rate at which departures leave  $n$

*That is,*

$$a_n = d_n$$

*Proof.*     • arrival sees  $n$ : systems goes from  $n$  to  $n + 1$ .

• departure sees  $n$ : systems goes from  $n + 1$  to  $n$ .

• In any interval of time  $T$ , # of transitions from  $n$  to  $n + 1$  = # of transitions from  $n + 1$  to  $n$ .

• Now,

$$a_n = \frac{\text{rate at which arrival finds } n}{\text{overall arrival rate}}$$

$$d_n = \frac{\text{rate at which departure leaves } n}{\text{overall departure rate}}$$

• If the overall arrival rate is equal to the overall departure rate, then  $a_n = d_n$ .

• If the overall arrival rate is greater than the overall departure rate, then  $a_n = d_n = 0$ .

□

**Proposition 3.** *Poisson Arrivals always See Time Averages (PASTA principle). In particular, for Poisson arrivals,*

$$P_n = a_n$$

• The total time the system is in state  $n$  by time  $T = P_n T$ .

• The number of arrivals in  $[0, T]$  that find the system is in state  $n$  is  $\lambda P_n T$ .

• The long-run rate at which the arrival finds the system is in state  $n = \lambda P_n$ .

• The above divided by overall arrival rate  $= \lambda P_n / \lambda = P_n \rightarrow$  proportion of arrivals that find the system is in state  $n$ .

## Exponential Models

### Single Server Exponential Queueing System

- Arrival follows Poisson process with rate  $\lambda$ .
- If the server is free, the arrival will be served by the server.
- If the server is busy, then the arrival will be waiting in the queue.
- Service time is assumed to be independent, exponentially distributed with mean  $1/\mu$ .
- M/M/1     Markovian Markovian : interarrival & service distributions are exponential  $\rightarrow$  Memoryless.

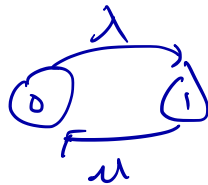
Limiting probabilities:  $P_n, n = 0, 1, 2, \dots$

rate at which the process enters state  $n$  = rate at which the process leaves state  $n$

Figures:



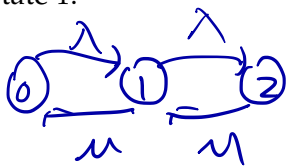
State 0:



the rate at which the process leaves state 0 is  $\lambda P_0$

on the other hand, the rate at which the process enters state 0 is  $\mu P_1$ ,  $\therefore \lambda P_0 = \mu P_1$

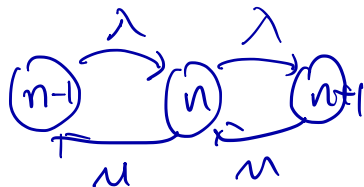
State 1:



rate out of  $(\lambda + \mu) P_1$   
rate in  $\lambda P_0 + \mu P_2$

$$\therefore (\lambda + \mu) P_1 = \lambda P_0 + \mu P_2$$

State n:



$$(\lambda + \mu) P_n = \lambda P_{n-1} + \mu P_{n+1}$$

$$\lambda P_0 = \mu P_1$$

$$P_{n+1} = \frac{\lambda}{\mu} P_n + \left( P_n - \frac{\lambda}{\mu} P_{n-1} \right), n \geq 1$$

Solving in terms of  $P_0$  :

$$P_0 = P_0$$

$$P_1 = \frac{\lambda}{\mu} P_0$$

$$P_2 = \frac{\lambda}{\mu} P_1 + \left( P_1 - \frac{\lambda}{\mu} P_0 \right) = \frac{\lambda}{\mu} P_1 = \left( \frac{\lambda}{\mu} \right)^2 P_0$$

$\vdots$

$$P_n = \left( \frac{\lambda}{\mu} \right)^n P_0$$

Also,

$$1 = \sum_{n=0}^{\infty} P_n = \sum_{n=0}^{\infty} \left( \frac{\lambda}{\mu} \right)^n P_0 = \frac{P_0}{1 - \lambda/\mu} \Rightarrow P_0 = 1 - \frac{\lambda}{\mu}$$

Therefore,

$$P_n = \left( \frac{\lambda}{\mu} \right)^n \left( 1 - \frac{\lambda}{\mu} \right) \quad n \geq 1$$

Note: we assumed that  $\frac{\lambda}{\mu} < 1 \rightarrow$  the mean service time is less than the mean time between successive arrivals.  $\lambda < \mu$

otherwise, if  $\lambda > \mu$  then  $P_n = 0$   $\forall n$

avg. # of customers in the system

Now,

$$L = \sum_{n=0}^{\infty} n P_n$$

$$= \sum_{n=0}^{\infty} n \left(\frac{\lambda}{\mu}\right)^n (1 - \frac{\lambda}{\mu}) = \frac{\lambda}{\mu - \lambda}$$

(Note:  $\sum_{n=1}^{\infty} nx^n = \frac{x}{(1-x)^2}$ )

Let's know

$$W = \frac{L}{\lambda} = \frac{1}{\mu - \lambda}$$

$$W_Q = W - E[S] = W - \frac{1}{\mu} = \frac{\lambda}{\mu(\mu - \lambda)}$$

$$L_Q = \lambda W_Q = \frac{\lambda^2}{\mu(\mu - \lambda)}$$

**Example 4.** Suppose that customers arrive at a Poisson rate of one per every 12 minutes, and that the service time is exponential at a rate of one service per 8 minutes. What are  $L$  and  $W$ ?

$$\lambda = \frac{1}{12}$$

$$\mu = \frac{1}{8}$$

$$L = \frac{\lambda}{\mu - \lambda} = 2$$

$$W = \frac{1}{\mu - \lambda} = 24$$

Now suppose that the arrival rate increases 20 percent to  $\lambda = 1/10$ . What is the corresponding  $L$  and  $W$ ?

$$L = 4, \quad W = 40$$

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**Example 5.** For an  $M/M/1$  queue in steady state, what is the probability that the next arrival finds  $n$  in the system?

- If  $t$  is the current time, then the time from  $t$  until the next arrival is exponentially distributed with rate  $\lambda$ .
- Also, the time from the last arrival before  $t$  until the time  $t$  is exponentially distributed with rate  $\lambda$ .
- Although, the times between successive arrivals of a Poisson process are exponential with rate  $\lambda$ , the time between the previous arrival before  $t$  and the first arrival after  $t$  is distributed as the sum of two independent exponentials. (the inspection paradox)
- The length of an inter arrival interval that contains a specified time tends to be longer than an ordinary inter arrival interval.
- $N_a$ : the number found by the next arrival.
- $X$ : the number currently in the system.

Conditioning on  $X$  yields:

$$\begin{aligned}
 P[N_a = n] &= \sum_{k=0}^{\infty} P\{N_a = n \mid X = k\} P\{X = k\} \\
 &= \sum_{k=0}^{\infty} P\{N_a = n \mid X = k\} \left(\frac{\lambda}{\mu}\right)^k (1 - \lambda/\mu) \\
 &= \sum_{k=n}^{\infty} P\{N_a = n \mid X = k\} \left(\frac{\lambda}{\mu}\right)^k (1 - \lambda/\mu) \\
 &= \sum_{i=0}^{\infty} P\{N_a = n \mid X = n+i\} \left(\frac{\lambda}{\mu}\right)^{n+i} (1 - \lambda/\mu)
 \end{aligned}$$

By the way

$$P\{N_a = n \mid X = n+i\} = \left(\frac{\mu}{\lambda + \mu}\right)^i \frac{\lambda}{\lambda + \mu}, \quad n \geq 0$$

Consequently  $P\{N_a = n\} = \left(\frac{\lambda}{\mu}\right)^{n+1} (1 - \lambda/\mu)$

on the other hand,  $P\{N_a = 0\} = \sum_{i=0}^{\infty} \left(\frac{\mu}{\lambda + \mu}\right)^i \left(\frac{\lambda}{\mu}\right)^{i+1} (1 - \lambda/\mu)$

$$= (1 + \lambda/\mu) (1 - \lambda/\mu)$$

### Single Server Exponential Queueing System Having Finite Capacity

Balance equations:

State	Rate at which the process leaves = Rate at which the process enters
0	$\lambda P_0 = \mu P_1$
$1 \leq n \leq N-1$	$(\lambda + \mu) P_n = \lambda P_{n-1} + \mu P_{n+1}$
N	$\mu P_N = \lambda P_{N-1}$

Also,  $\mu P_n = \lambda P_{n-1}, n = 1, 2, \dots, N$       the rate at which departures leave behind  $n-1$   
 Thus,      =      "      arrivals find  $n-1$

$$P_n = \frac{\lambda}{\mu} P_{n-1} = \dots = \left(\frac{\lambda}{\mu}\right)^n P_0$$

$$1 = \sum_{n=0}^N P_n = P_0 \sum_{n=0}^N \left(\frac{\lambda}{\mu}\right)^n$$

$$P_0 = \frac{(1 - \lambda/\mu)}{1 - (\lambda/\mu)^{N+1}}$$

$$P_n = \frac{(\lambda/\mu)^n (1 - \lambda/\mu)}{1 - (\lambda/\mu)^{N+1}} \quad n = 0, 1, \dots, N$$

Note: no need to assume  $\lambda/\mu < 1$ .



$$\begin{aligned}
 L &= \sum_{n=0}^N n P_n = \frac{(1 - \lambda/\mu)}{1 - (\lambda/\mu)^{N+1}} \sum_{n=0}^N n \left( \frac{\lambda}{\mu} \right)^n \\
 &= \frac{\lambda [1 + N(\lambda/\mu)^{N+1} - (N+1)(\lambda/\mu)^N]}{(\mu - \lambda)(1 - (\lambda/\mu)^{N+1})}
 \end{aligned}$$

In terms of  $W$ , there are two cases.

1. including those arrivals to find the system is full (thereby, not entering the system,  $W=0$ )

$$\lambda_a = \lambda$$

2. considering only actual entering arrivals ( $W > 0$ )

$$\lambda_a = \lambda(1 - P_N)$$

$$W = \frac{L}{\lambda_a}$$

**Example 6.** Suppose that it costs  $c\mu$  dollars per hour to provide service at a rate  $\mu$ . Suppose also that we incur a gross profit of  $A$  dollars for each customer served. If the system has a capacity  $N$ , what service rate  $\mu$  maximizes our total profit?

$$\begin{aligned}
 \text{profit per hour} &= \lambda(1 - P_N)A - c\mu \\
 &= \lambda A \left[ 1 - \frac{(\lambda/\mu)^N (1 - \lambda/\mu)}{1 - (\lambda/\mu)^{N+1}} \right] - c\mu \\
 &= \frac{\lambda A [1 - (\lambda/\mu)^N]}{1 - (\lambda/\mu)^{N+1}} - c\mu
 \end{aligned}$$

$$\frac{d[\text{profit}]}{d\mu} = 10 \frac{(2\mu^3 - 3\mu^2 + 1)}{(\mu^3 - 1)^2} - 1$$