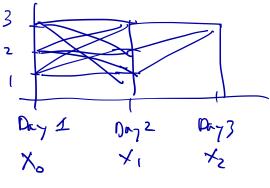
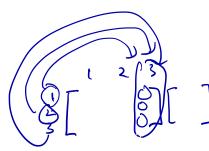
SSIE 660: Applied Stochastic Processes Dr. Sung H. Chung Note 8 Chapter 4. Discrete Markov Chains

Unconditional Probabilities

Example: What is the probability that the inventory at the beginning of the 3rd day is 3? That is $P[X_2 = 3]$?





Beginning step is '0', not '1'

$$P^2 = \left[\begin{array}{ccc} 0.23 & 0.28 & 0.49 \\ 0.3 & 0.21 & 0.49 \\ 0.30 & 0.28 & 0.42 \end{array} \right]$$

We cannot find the unconditional probability $P[X_2 = 3]$, unless we know

$$P[X_0=i]$$
 $i=1,2.3$
(ef 5 J , d , d) dende

$$P[X_2 = 3] = \sum_{i=1}^{3} P[X_2 = 3 | x_0 = i] P[X_0 = i]$$

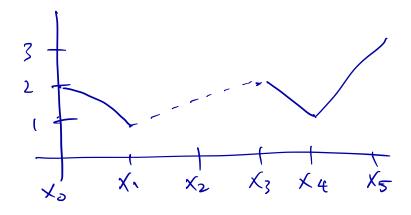
Let
$$\alpha_1 = 0.3$$
, $\alpha_2 = 0.6$, $\alpha_3 = 0.1$. Then, $P[X_2 = 3] = \begin{pmatrix} .49 \\ .49 \\ .42 \end{pmatrix}^T \begin{pmatrix} .3 \\ .6 \\ .1 \end{pmatrix} = .483$.

In general,

$$P[X_n = j] = \sum_i P[X_i = j | X_0 = i] P[X_0 = i]$$

Other Probabilities

$$P[X_5 = 3, X_4 = 1, X_3 = 2, X_1 = 1 | X_0 = 2] =$$



$$= P_{21} * P_{12}^{(2)} * P_{21} * P_{13} = .4 * .28 * .4 * .7 = .0314$$
 \rightarrow Conditional probability

$$P[X_4 = 3, X_3 = 1, X_2 = 1] =$$

Recall: $(\alpha_1 = 0.3, \alpha_2 = 0.6, \alpha_3 = 0.1)$

$$\sum_{i=1}^{3} P[X_4 = 3, X_3 = 1, X_2 = 1] X_6 = x] P[X_6 = x]$$

$$= P_{11}^{2} \cdot P_{13} \cdot Q_1 + P_{21}^{2} \cdot P_{11} \cdot P_{13} \cdot Q_2 + P_{31}^{2} P_{11} \cdot P_{13} \cdot Q_3$$

$$= 0.586$$

Classification of States:

(A) Transition probability matrix in our inventory example.

$$P = \begin{bmatrix} 0.3 & 0 & 0.7 \\ 0.4 & 0.3 & 0.3 \\ 0.2 & 0.4 & 0.4 \end{bmatrix}, P^4 = \begin{bmatrix} 0.284 & 0.26 & 0.456 \\ 0.279 & 0.265 & 0.456 \\ 0.279 & 0.260 & 0.461 \end{bmatrix}$$

$$P^8 = \begin{bmatrix} 0.280 & 0.262 & 0.458 \\ 0.280 & 0.262 & 0.458 \\ 0.280 & 0.262 & 0.458 \end{bmatrix}, P^{12} = \begin{bmatrix} 0.280 & 0.262 & 0.458 \\ 0.280 & 0.262 & 0.458 \\ 0.280 & 0.262 & 0.458 \end{bmatrix}$$
(B)
$$P = \begin{bmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \end{bmatrix}, P^2 = \begin{bmatrix} 0 & 0 & 1 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \end{bmatrix}, P^{12} = \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}$$

$$P^{13} = \begin{bmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \end{bmatrix}, P^{14} = \begin{bmatrix} 0 & 0 & 1 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \end{bmatrix}$$
(C)
$$P = \begin{bmatrix} 0.3 & 0.3 & 0.4 \\ 0.4 & 0.2 & 0.4 \\ 0 & 0 & 1 \end{bmatrix}, P^{2} = \begin{bmatrix} 0.21 & 0.15 & 0.64 \\ 0.20 & 0.16 & 0.64 \\ 0 & 0 & 1 \end{bmatrix}$$

$$P^{12} = \begin{bmatrix} 0.001 & 0.0009 & 0.998 \\ 0.001 & 0.0009 & 0.998 \\ 0 & 0 & 0.3 & 0.7 \\ 0 & 0 & 0.3 & 0.7 \\ 0 & 0 & 0.5 & 0.5 \end{bmatrix}, P^2 = \begin{bmatrix} 0.36 & 0.64 & 0 & 0 \\ 0.32 & 0.68 & 0 & 0 \\ 0 & 0 & 0.44 & 0.56 \\ 0 & 0 & 0.44 & 0.56 \\ 0 & 0 & 0.44 & 0.583 \\ 0 & 0 & 0.447 & 0.583 \\ 0 & 0.00 & 0.447 & 0.583 \\ 0 & 0.00 & 0.447 & 0.583 \\ 0 & 0.00 & 0.447 & 0.583 \\ 0 & 0.00 & 0.447 & 0.583 \\ 0 & 0.00 & 0.447 & 0.583 \\ 0 & 0.00 & 0.447 & 0.583 \\ 0 & 0.00 & 0.447 & 0.583 \\ 0 & 0.00 & 0.447 & 0.583 \\ 0 & 0.00 & 0.447 & 0.583 \\ 0 & 0.00 & 0.447 & 0.583 \\ 0 & 0.00 & 0.447 & 0.583$$

• It can be seen that in some Markov Chains, multiplying *P* by itself a large number of times will result in identical rows (A and C) - independent of starting states.

- This means that n-step transition probability does not depend on starting state.
- These probabilities are called steady state or stationary probabilities.)
- In B and D, the rows are not identical; this means that only certain type of Markov Chains have steady state probabilities. This depends upon the types of states a Markov Chain has, which we will study next.

Classification of States:

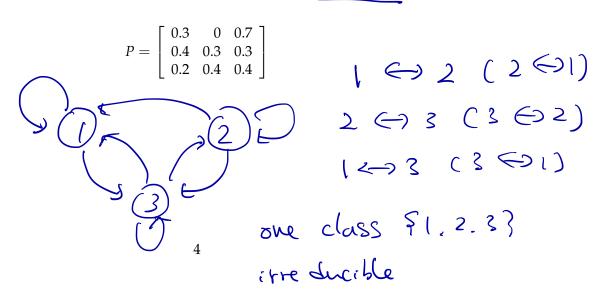
1. State *j* is "accessible" from state *i*, if

$$P_{ij}^n > 0$$
, for some $n \ge 0$.

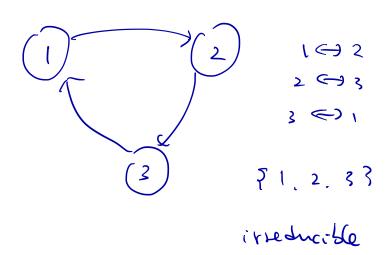
That is, process or system will enter state j, at some time in the future (some n), starting in state i now.

- 2. Two states, i and j, that are accessible to each other are said to "communicate each other"
 - (a) If $i \leftrightarrow j$, then $j \leftrightarrow i$.
 - (b) If $i \leftrightarrow j$, and $j \leftrightarrow k$, then $i \leftrightarrow k$.
 - (c) A state communicates with itself if $i \leftrightarrow i$.
- 3. If $i \leftrightarrow j$, then i, j belong to the same class.
- 4. A Markov Chain could have more than one class of states.
- 5. If a Markov Chain has only one class, then it is said to be irreducible.

Example 1.

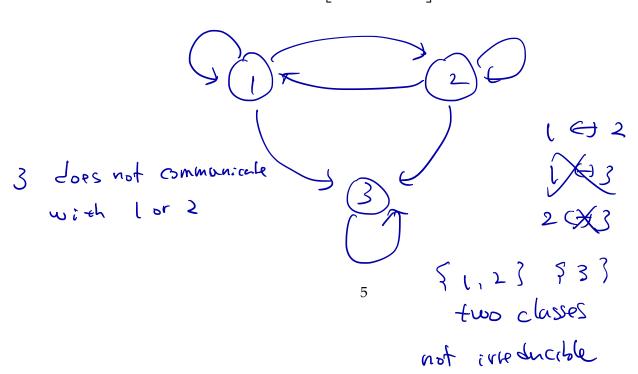


$$P = \left[\begin{array}{ccc} 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \end{array} \right]$$



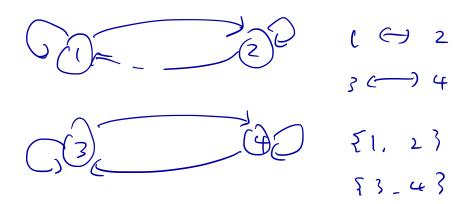
Example 3.

$$P = \left[\begin{array}{ccc} 0.3 & 0.3 & 0.4 \\ 0.4 & 0.2 & 0.4 \\ 0 & 0 & 1 \end{array} \right]$$



Example 4.

$$P = \left[\begin{array}{cccc} 0.2 & 0.8 & 0 & 0 \\ 0.4 & 0.6 & 0 & 0 \\ 0 & 0 & 0.3 & 0.7 \\ 0 & 0 & 0.5 & 0.5 \end{array} \right]$$



not meducible



Closed Sets:

- 1. No state outside the set can be reached from it.
- 2. If only one state forms a closed set, then it is called an *absorbing state*. That is if j is an absorbing state, then $P_{jj} =$

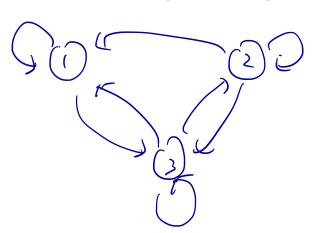
Definition:

Probability [the system will *ever* revisit state i, after it leaves it] = f_i

- A state is said to be *recurrent* if $f_i = 1$.
- A state is said to be *transient* if $f_i < 1$.

Example 5.

$$P = \left[\begin{array}{ccc} 0.3 & 0 & 0.7 \\ 0.4 & 0.3 & 0.3 \\ 0.2 & 0.4 & 0.4 \end{array} \right]$$



F1, 2, 3?

only one state

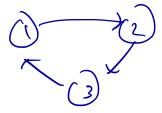
recurrent

F1, 2, 3?

trons.ent

Example 6.

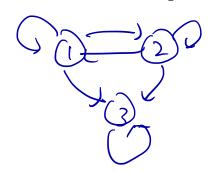
$$P = \left[\begin{array}{ccc} 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \end{array} \right]$$



recurrent \$1.2.33 transpert X

Example 7.

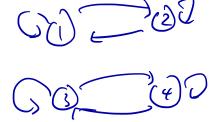
$$P = \left[\begin{array}{ccc} 0.3 & 0.3 & 0.4 \\ 0.4 & 0.2 & 0.4 \\ 0 & 0 & 1 \end{array} \right]$$

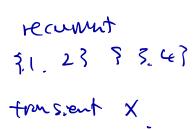


?1. 23 state
?3? State
recurrent ?3?
transport ?1.23

Example 8.

$$P = \left[\begin{array}{cccc} 0.2 & 0.8 & 0 & 0 \\ 0.4 & 0.6 & 0 & 0 \\ 0 & 0 & 0.3 & 0.7 \\ 0 & 0 & 0.5 & 0.5 \end{array} \right]$$





geome-furc

- If state i is recurrent, then the number of times the process/system visits state i is ∞
- If state i is transient, then the number of times the process/system visits state i is $\int_{-\infty}^{\infty} n \, dx \, dx$
- Let us consider a transient state *i* and derive the mean number of times the system visites *i*, before it leaves state *i* forever.
- After each visit to state *i*, there are two outcomes.
 - 1. It will come back with probability
 - 2. It will not come back with probability $1 \frac{1}{2}$
- The number of times the system visits i, before it leaves i forever

$$N = u$$
 $t_{u}(1-t_{u})$: $b(N=u) = (1-b)_{u}b$
 $N = 1$ $t_{u}(1-t_{u})$ (et $b = 1-z_{u}$)

• Mean number of times the system visits i, before it leaves i forever

 $P_{i}[N_{i}=\infty]=0$ $[new = E[N] = \frac{1}{p} = \frac{1}{1-f_{i}}$

- We can also derive an alternative expression for the expected (mean) number of visits to state *i*, before the system leaves *i* forever.
- Let us define an indicator (0-1 variable) I_n as follows.

$$I_n = \begin{cases} 1, & \text{if } \forall_N = 1 \\ 0, & \text{o.} \end{cases}$$

• Total number of visits to *i* after leaving *i*

$$\sum_{n=0}^{\infty} I_n | X_0 = i$$

• Expected number of visits to i, after leaving i

$$E\left[\sum_{n=0}^{\infty}I_{n}|X_{0}=i\right] = \sum_{n=0}^{\infty} E\left[I_{n}\left[X_{0}=i\right]\right]$$

$$= \sum_{n=0}^{\infty} P\left[Y_{n}=i\right] Y_{0}=i\right] = \sum_{n=0}^{\infty} P_{i}^{n}$$

$$= \sum_{n=0}^{\infty} P\left[Y_{n}=i\right] Y_{0}=i\right]$$

• State *i* is recurrent if
$$\sum_{n=0}^{\infty} \rho_{i\bar{r}}^{n} = 0$$

• State *i* is transient if
$$\sum_{n=0}^{\infty} \rho_{ii}^{n} < \infty$$

In summary: State i is

summary: State
$$i$$
 is

• recurrent if
$$\int_{i}^{\infty} = 1$$

$$\int_{i=0}^{\infty} \rho_{i}^{\infty} = \infty$$

• transient if
$$\int_{-\tau}^{\tau} < 1$$

$$\sum_{N=0}^{\infty} b_{i,i}^{N} < \infty$$

Also,

- re current • If i is recurrent and if $i \leftrightarrow j$, then j is
- If i is transient and if $i \leftrightarrow j$, then j is
- Recurrence and Transience are Class properties.