



EM Variants

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Review of the EM Algorithms

- To use EM algorithms, the key steps below are required

1) Computing the posteriori distribution

$$p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)})$$

2) Evaluating the expectation of $\log p(\mathbf{x}, \mathbf{z}; \boldsymbol{\theta})$ w.r.t. the posteriori $p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)})$, i.e.,

$$Q(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)}) = \mathbb{E}_{p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)})} [\log p(\mathbf{x}, \mathbf{z}; \boldsymbol{\theta})]$$

3) Maximizing

$$\boldsymbol{\theta}^{(t+1)} = \arg \max_{\boldsymbol{\theta}} Q(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)})$$

However, not all of them are always achievable

Two Issues in the EM

- Issue one

The maximization is not achievable

$$\boldsymbol{\theta}^{(t+1)} = \arg \max_{\boldsymbol{\theta}} Q(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)})$$

- Issue two

- 1) The posteriori $p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)})$ cannot be derived analytically
- 2) Even if $p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)})$ can be obtained, we still cannot derive the close-form expression for the expectation

$$\mathbb{E}_{p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)})} [\log p(\mathbf{x}, \mathbf{z}; \boldsymbol{\theta})]$$

Outline

- Addressing Issue One
- Addressing Issue Two

Generalized EM

- It is quite often in training LVMs that the optimization $\max_{\boldsymbol{\theta}} \mathcal{Q}(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)})$ cannot be solved

How to address this issue?

- Maximizing $\mathcal{Q}(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)})$ is not necessary. Increasing $\mathcal{Q}(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)})$ is sufficient to guarantee the EM algorithm to work
- That is, if we adopt SGD to update the parameter as

$$\boldsymbol{\theta}^{(t+1)} = \boldsymbol{\theta}^{(t)} + \gamma \cdot \left. \frac{\partial \mathcal{Q}(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)})}{\partial \boldsymbol{\theta}} \right|_{\boldsymbol{\theta} = \boldsymbol{\theta}^{(t)}}$$

we can also guarantee the monotonic increase of log-likelihood

$$\log p(\mathbf{x}; \boldsymbol{\theta}^{(t+1)}) \geq \log p(\mathbf{x}; \boldsymbol{\theta}^{(t)})$$

- Sketch of proof

➤ First, after the SGD update, it can be easily seen that

$$Q(\boldsymbol{\theta}^{(t+1)}; \boldsymbol{\theta}^{(t)}) \geq Q(\boldsymbol{\theta}^{(t)}; \boldsymbol{\theta}^{(t)})$$

➤ From $\mathcal{L}(p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}), \boldsymbol{\theta}) = \int p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}) \log \frac{p(\mathbf{x}, \mathbf{z}; \boldsymbol{\theta})}{p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)})} d\mathbf{z} = Q(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)}) - \int p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}) \log p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}) d\mathbf{z}$, we further have

$$\mathcal{L}(p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}); \boldsymbol{\theta}^{(t+1)}) \geq \underbrace{\mathcal{L}(p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}); \boldsymbol{\theta}^{(t)})}_{=\log p(\mathbf{x}; \boldsymbol{\theta}^{(t)})}$$

➤ Due to

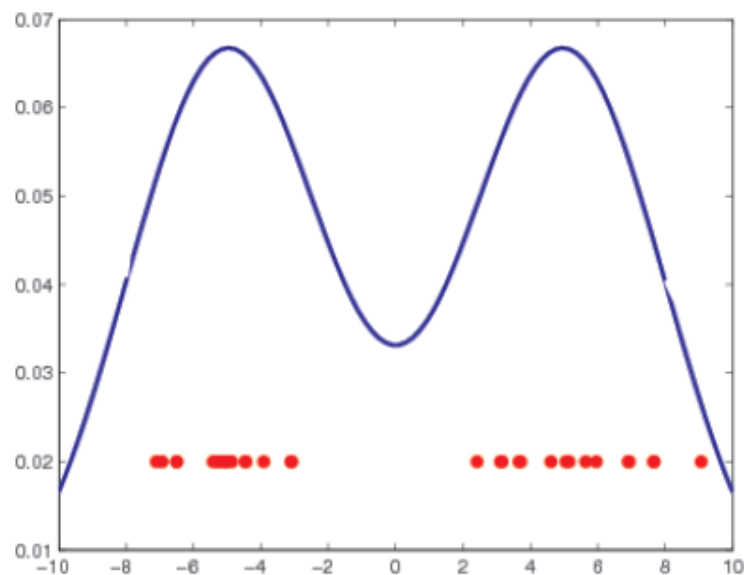
$$\begin{aligned} \log p(\mathbf{x}; \boldsymbol{\theta}^{(t+1)}) &= \underbrace{\mathcal{L}(p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}), \boldsymbol{\theta}^{(t+1)})}_{\geq \log p(\mathbf{x}; \boldsymbol{\theta}^{(t)})} + \underbrace{KL(p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}) || p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t+1)}))}_{\geq 0} \\ \Rightarrow \log p(\mathbf{x}; \boldsymbol{\theta}^{(t+1)}) &\geq \log p(\mathbf{x}; \boldsymbol{\theta}^{(t)}) \end{aligned}$$

Outline

- Addressing Issue One
- Addressing Issue Two

MCMC EM

- For any probability distributions, we can always draw samples from it, e.g., using **Markov chain Monte Carlo (MCMC)** methods



- Although the exact expression of the posteriori $p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)})$ is not known, we can use samples drawn from it to approximate it

- Thus, we can draw lots of samples \mathbf{z}_s for $s = 1, \dots, S$ from the posteriori distribution $p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)})$ such that

$$\mathbf{z}_s \sim p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)})$$

- Then, the expectation $\mathbb{E}_{p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)})} [\log p(\mathbf{x}, \mathbf{z}; \boldsymbol{\theta})]$ can be approximated as

$$\mathcal{Q}(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)}) \approx \frac{1}{S} \cdot \sum_{s=1}^S \log p(\mathbf{x}, \mathbf{z}_s; \boldsymbol{\theta})$$

- We can optimize the approximate $\mathcal{Q}(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)})$ with SGD algorithm

The two sub-problems in the Issue Two are both solved. Thus, latent-variable models can always be trained with MCMC EM

- Drawing samples from a distribution is *computationally expensive*
- An alternative approach is to use a simple distribution $q(\mathbf{z}; \boldsymbol{\phi})$ to approximate the exact posterior distribution $p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)})$

How to get the approximate simple distribution $q(\mathbf{z}; \boldsymbol{\phi})$?

- Idea
 - 1) Assuming a simple form for $q(\mathbf{z}; \boldsymbol{\phi})$, e.g.,

$$q(\mathbf{z}; \boldsymbol{\phi}) = \prod \mathcal{N}(z_i; \mu_i, \sigma_i^2)$$

- 2) Finding the best $\boldsymbol{\phi}$ that minimizes the KL-divergence

$$KL(q(\mathbf{z}; \boldsymbol{\phi}) \| p(\mathbf{z}|\mathbf{x}; \boldsymbol{\theta}^{(t)}))$$

- Steps to update the model parameter θ

1) Finding the best approximate $q(\mathbf{z}; \phi)$ such that

$$\phi^{(t)} = \arg \min_{\phi} KL(q(\mathbf{z}; \phi) \| p(\mathbf{z} | \mathbf{x}; \theta^{(t)}))$$

2) Using $q(\mathbf{z}; \phi^{(t)})$ to compute expectation $\mathbb{E}_{p(\mathbf{z} | \mathbf{x}; \theta^{(t)})} [\log p(\mathbf{x}, \mathbf{z}; \theta)]$ approximately as

$$\tilde{Q}(\theta; \phi^{(t)}) = \mathbb{E}_{q(\mathbf{z}; \phi^{(t)})} [\log p(\mathbf{x}, \mathbf{z}; \theta)]$$

3) Obtaining the new value $\theta^{(t+1)}$ as

$$\theta^{(t+1)} = \arg \max_{\theta} \tilde{Q}(\theta; \phi^{(t)})$$

- The two optimization problems can be equivalently written as

$$\min_{\phi} KL(q(\mathbf{z}; \phi) \| p(\mathbf{z}|\mathbf{x}; \theta^{(t)})) \Leftrightarrow \max_{\phi} \int q(\mathbf{z}; \phi) \log \frac{p(\mathbf{z}|\mathbf{x}; \theta^{(t)})}{q(\mathbf{z}; \phi)} d\mathbf{z}$$

$$\Leftrightarrow \max_{\phi} \int q(\mathbf{z}; \phi) \log \frac{p(\mathbf{x}, \mathbf{z}; \theta^{(t)})}{q(\mathbf{z}; \phi)} d\mathbf{z}$$

$$\max_{\theta} \mathbb{E}_{q(\mathbf{z}; \phi^{(t)})} [\log p(\mathbf{x}, \mathbf{z}; \theta)] \Leftrightarrow \max_{\theta} \int q(\mathbf{z}; \phi^{(t)}) \log \frac{p(\mathbf{x}, \mathbf{z}; \theta)}{q(\mathbf{z}; \phi^{(t)})} d\mathbf{z}$$

- The algorithm to optimize θ and ϕ can be understood as solving the following optimization problem **in an alternative way**

$$\max_{\phi, \theta} \mathcal{L}(\mathbf{x}; \theta, \phi)$$

with

$$\mathcal{L}(\mathbf{x}; \theta, \phi) \triangleq \int q(\mathbf{z}; \phi) \log \frac{p(\mathbf{x}, \mathbf{z}; \theta)}{q(\mathbf{z}; \phi)} d\mathbf{z}$$

- Instead of updating θ, ϕ alternatively, we can also update them *simultaneously* with the SGD algorithm, that is,

$$\theta^{(t+1)} = \theta^{(t)} + \gamma \cdot \left. \frac{\partial \mathcal{L}(x; \theta, \phi)}{\partial \theta} \right|_{\theta=\theta^{(t)}}$$

$$\phi^{(t+1)} = \phi^{(t)} + \gamma \cdot \left. \frac{\partial \mathcal{L}(x; \theta, \phi)}{\partial \phi} \right|_{\phi=\phi^{(t)}}$$

The method is dubbed *variational Bayesian EM* (VB-EM)

- In general, we optimize $\mathcal{L}(x; \theta, \phi)$ w.r.t. the two parameters θ, ϕ simultaneously

- Actually, it can be proved that $\mathcal{L}(\mathbf{x}; \boldsymbol{\theta}, \boldsymbol{\phi})$ is a *lower bound* of the log-likelihood $\ln p(\mathbf{x}; \boldsymbol{\theta})$ for any $\boldsymbol{\theta}$ and $\boldsymbol{\phi}$, that is,

$$\ln p(\mathbf{x}; \boldsymbol{\theta}) \geq \mathcal{L}(\mathbf{x}; \boldsymbol{\theta}, \boldsymbol{\phi})$$

(Proof can be found in the next slide)

When the log-likelihood $\ln p(\mathbf{x}; \boldsymbol{\theta})$ cannot be directly maximized, we can seek to optimize its lower bound

$$\mathcal{L}(\mathbf{x}; \boldsymbol{\theta}, \boldsymbol{\phi}) \triangleq \int q(\mathbf{z}; \boldsymbol{\phi}) \log \frac{p(\mathbf{x}, \mathbf{z}; \boldsymbol{\theta})}{q(\mathbf{z}; \boldsymbol{\phi})} d\mathbf{z}$$

where $q(\mathbf{z}; \boldsymbol{\phi})$ can be set as any simple distribution forms, *e.g.*,

$$q(\mathbf{z}; \boldsymbol{\phi}) = \prod \mathcal{N}(z_i; \mu_i, \sigma_i^2)$$

Proof of $\ln p(\mathbf{x}; \boldsymbol{\theta}) \geq \mathcal{L}(\mathbf{x}; \boldsymbol{\theta}, \boldsymbol{\phi})$

$$\begin{aligned}\ln p_{\boldsymbol{\theta}}(\mathbf{x}) &= \int_{\mathbf{z}} q_{\boldsymbol{\phi}}(\mathbf{z}) \ln p_{\boldsymbol{\theta}}(\mathbf{x}) d\mathbf{z} = \int_{\mathbf{z}} q_{\boldsymbol{\phi}}(\mathbf{z}) \ln \frac{p_{\boldsymbol{\theta}}(\mathbf{x}, \mathbf{z}) q_{\boldsymbol{\phi}}(\mathbf{z})}{q_{\boldsymbol{\phi}}(\mathbf{z}) p_{\boldsymbol{\theta}}(\mathbf{z}|\mathbf{x})} d\mathbf{z} \\&= \int_{\mathbf{z}} q_{\boldsymbol{\phi}}(\mathbf{z}) \ln \frac{p_{\boldsymbol{\theta}}(\mathbf{x}, \mathbf{z})}{q_{\boldsymbol{\phi}}(\mathbf{z})} d\mathbf{z} + \int_{\mathbf{z}} q_{\boldsymbol{\phi}}(\mathbf{z}) \ln \frac{q_{\boldsymbol{\phi}}(\mathbf{z})}{p_{\boldsymbol{\theta}}(\mathbf{z}|\mathbf{x})} d\mathbf{z} \\&= \int_{\mathbf{z}} q_{\boldsymbol{\phi}}(\mathbf{z}) \ln \frac{p_{\boldsymbol{\theta}}(\mathbf{x}, \mathbf{z})}{q_{\boldsymbol{\phi}}(\mathbf{z})} d\mathbf{z} + KL(q_{\boldsymbol{\phi}}(\mathbf{z}) || p_{\boldsymbol{\theta}}(\mathbf{z}|\mathbf{x})) \\&\geq \int_{\mathbf{z}} q_{\boldsymbol{\phi}}(\mathbf{z}) \ln \frac{p_{\boldsymbol{\theta}}(\mathbf{x}, \mathbf{z})}{q_{\boldsymbol{\phi}}(\mathbf{z})} d\mathbf{z} \\&\triangleq \mathcal{L}(\mathbf{x}; \boldsymbol{\theta}, \boldsymbol{\phi})\end{aligned}$$