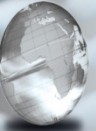


# Chapter 9

## One-Sample Estimation Problems

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# Section 9.3

## Classical Methods of Estimation

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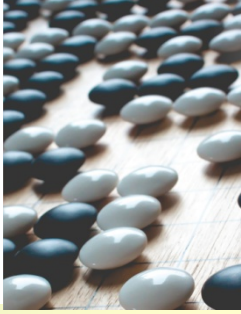
# Definition 9.1



A statistic  $\hat{\Theta}$  is said to be an **unbiased estimator** of the parameter  $\theta$  if

$$\mu_{\hat{\Theta}} = E(\hat{\Theta}) = \theta.$$

# Definition 9.1



- Sample mean

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

is an unbiased estimator of  $\mu$

# Definition 9.1



- Sample variance

$$S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$$

is an unbiased estimator of  $\sigma^2$



# Section 9.4

## Single Sample: Estimating the Mean

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# Confidence interval



- Motivation
  - Suppose we draw  $n$  random samples
  - assume we know variance  $\sigma$  but don't know mean  $\mu$
  - what can be say about true mean  $\mu$ , by using our sample mean?
  - assume  $n$  is sufficiently big, then

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \sim N(0, 1)$$

# Confidence interval



---

## Confidence Interval on $\mu$ , $\sigma^2$ Known

If  $\bar{x}$  is the mean of a random sample of size  $n$  from a population with known variance  $\sigma^2$ , a  $100(1 - \alpha)\%$  confidence interval for  $\mu$  is given by

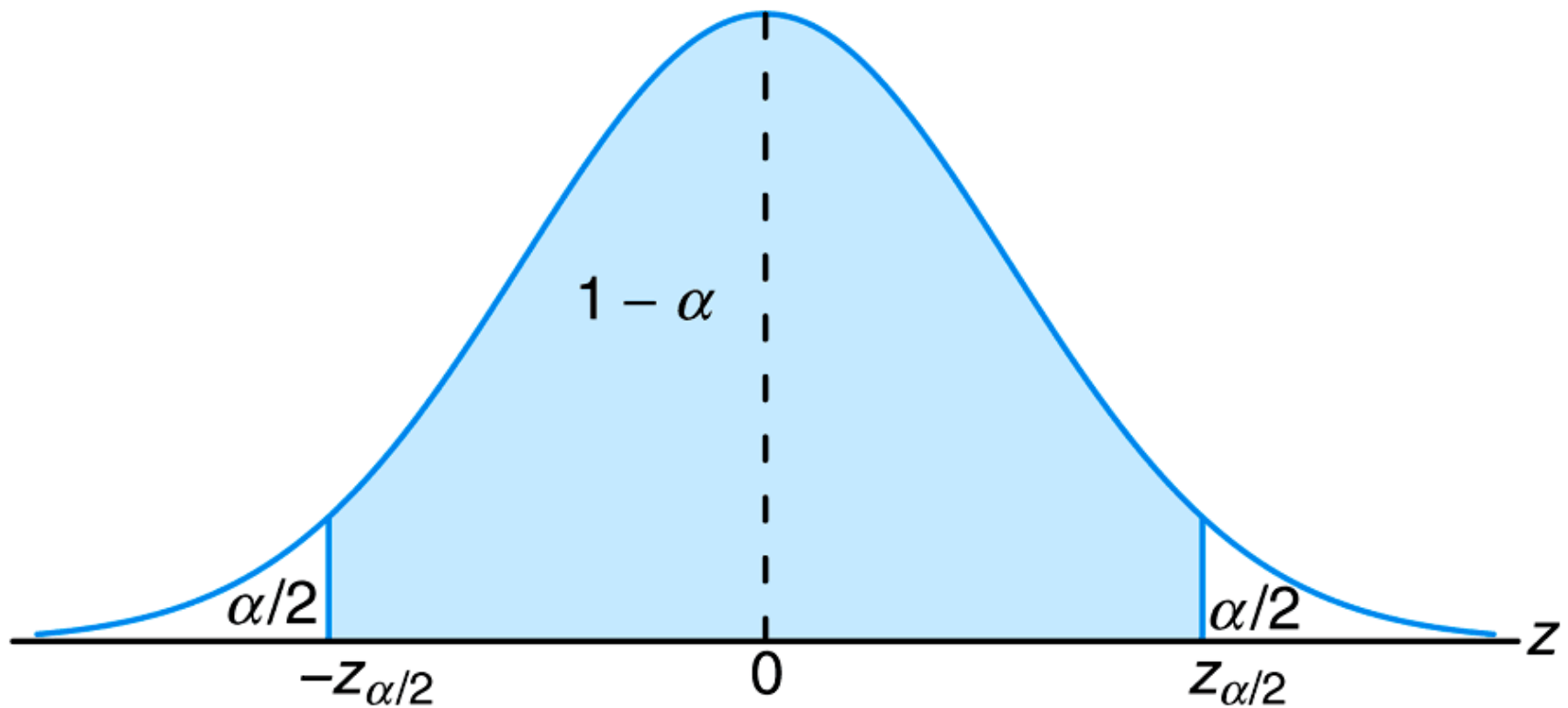
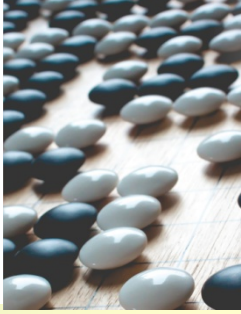
$$\bar{x} - z_{\alpha/2} \frac{\sigma}{\sqrt{n}} < \mu < \bar{x} + z_{\alpha/2} \frac{\sigma}{\sqrt{n}},$$

where  $z_{\alpha/2}$  is the  $z$ -value leaving an area of  $\alpha/2$  to the right.

---



# Figure 9.2 $P(-z_{\alpha/2} < Z < z_{\alpha/2}) = 1 - \alpha$



# Theorem 9.1



If  $\bar{x}$  is used as an estimate of  $\mu$ , we can be  $100(1 - \alpha)\%$  confident that the error will not exceed  $z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$ .

# example 9.2



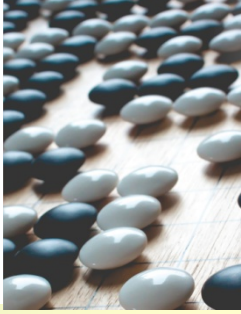
**Example 9.2:** The average zinc concentration recovered from a sample of measurements taken in 36 different locations in a river is found to be 2.6 grams per milliliter. Find the 95% and 99% confidence intervals for the mean zinc concentration in the river. Assume that the population standard deviation is 0.3 gram per milliliter.

the 95% confidence interval is

$$2.6 - (1.96) \left( \frac{0.3}{\sqrt{36}} \right) < \mu < 2.6 + (1.96) \left( \frac{0.3}{\sqrt{36}} \right),$$

which reduces to  $2.50 < \mu < 2.70$ .

# Theorem 9.2



If  $\bar{x}$  is used as an estimate of  $\mu$ , we can be  $100(1 - \alpha)\%$  confident that the error will not exceed a specified amount  $e$  when the sample size is

$$n = \left( \frac{z_{\alpha/2}\sigma}{e} \right)^2 .$$

# example 9.3



---

**Example 9.3:** How large a sample is required if we want to be 95% confident that our estimate of  $\mu$  in Example 9.2 is off by less than 0.05?

**Solution:** The population standard deviation is  $\sigma = 0.3$ . Then, by Theorem 9.2,

$$n = \left[ \frac{(1.96)(0.3)}{0.05} \right]^2 = 138.3.$$

# Stdev unknown for normal samples



---

Confidence Interval on  $\mu$ ,  $\sigma^2$  Unknown  
If  $\bar{x}$  and  $s$  are the mean and standard deviation of a random sample from a normal population with unknown variance  $\sigma^2$ , a  $100(1-\alpha)\%$  confidence interval for  $\mu$  is

$$\bar{x} - t_{\alpha/2} \frac{s}{\sqrt{n}} < \mu < \bar{x} + t_{\alpha/2} \frac{s}{\sqrt{n}},$$

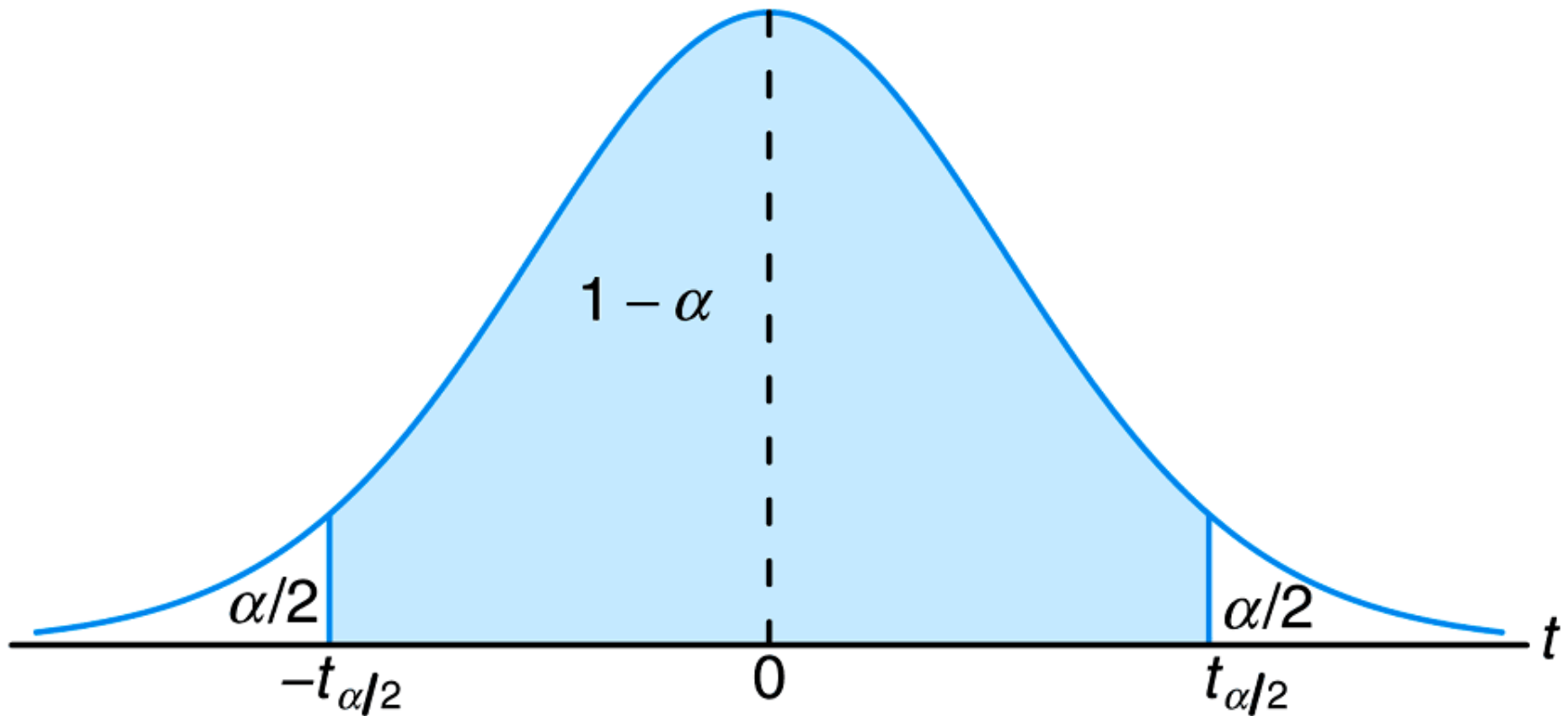
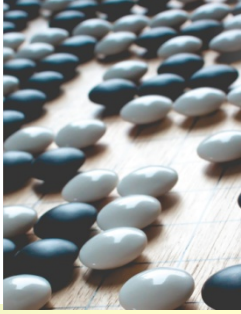
where  $t_{\alpha/2}$  is the  $t$ -value with  $v = n - 1$  degrees of freedom, leaving an area of  $\alpha/2$  to the right.

---



# Figure 9.5

$$P(-t_{\alpha/2} < T < t_{\alpha/2}) = 1 - \alpha$$



# example 9.5



**Example 9.5:** The contents of seven similar containers of sulfuric acid are 9.8, 10.2, 10.4, 9.8, 10.0, 10.2, and 9.6 liters. Find a 95% confidence interval for the mean contents of all such containers, assuming an approximately normal distribution.

The sample mean and standard deviation for the given data are

$$\bar{x} = 10.0 \quad \text{and} \quad s = 0.283.$$

95% confidence interval for  $\mu$  is

$$10.0 - (2.447) \left( \frac{0.283}{\sqrt{7}} \right) < \mu < 10.0 + (2.447) \left( \frac{0.283}{\sqrt{7}} \right),$$

which reduces to  $9.74 < \mu < 10.26$ .

# Chapter 10

## One- and Two-Sample Tests of Hypotheses

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# Section 10.2

## Testing a Statistical Hypothesis

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# Definition 10.1



A **statistical hypothesis** is an assertion or conjecture concerning one or more populations.

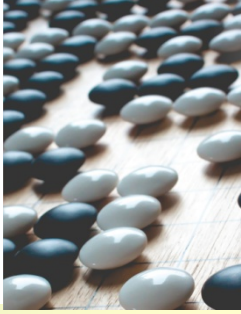
# Definition 10.1



- Null Hypothesis,  $H_0$ 
  - Hypothesis that the population is in default or unchanged state (status quo)
- Suppose in certain area, the size of apples was reduced by 10% after measuring random samples. On that year, there was a long draught period.
- **Q: Did the reduction of size happen by chance?**



# Definition 10.1



- Null Hypothesis,  $H_0$ 
  - **There was no change in the size of apples** (population maintains the default state)
  - So the observation of reduction in the sample mean happen **by chance** (b/c it is RV)
- How likely this happened by chance?
  - affected by the actual value of sample mean, sample size, variance, etc.

# Definition 10.1



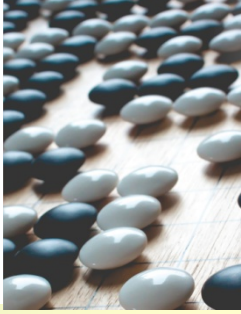
- We consider two hypotheses
  - Null Hypothesis,  $H_0$
  - Alternative Hypothesis,  $H_1$
  - $H_1$  means  $H_0$  is rejected
    - that is, the change in the default state is significant, and did not happen by chance
- $H_1$  typically stands for question to be answered or theory to be tested

# Definition 10.1



- Look at the sample mean, and consider the probability that the sample mean obtained by chance
  - if the probability is significantly low, reject  $H_0$
  - otherwise do not reject  $H_0$
- Careful: NOT rejecting  $H_0 \neq$  ACCEPT  $H_0$ !
- Not rejecting  $H_0$ 
  - don't have sufficient evidence to rule out  $H_0$

# Definition 10.1



- Typical null and alternative hypothesis
  - $H_0$ : defendant is innocent
  - $H_1$ : defendant is guilty
- By default, defendant is innocent, unless there is ‘undoubtable’ evidence (not by chance) against defendant
- Of course, if an evidence happened by chance, it means that we cannot reject  $H_0$  – but not saying that defendant is innocent

# Definition 10.2



Rejection of the null hypothesis when it is true is called a **type I error**.

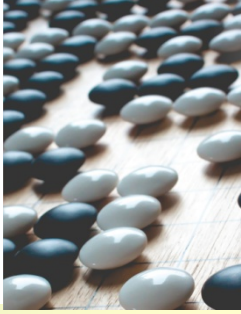
# Definition 10.3



Nonrejection of the null hypothesis when it is false is called a **type II error**.



# Table 10.1 Possible Situations for Testing a Statistical Hypothesis



	$H_0$ is true	$H_0$ is false
Do not reject $H_0$	Correct decision	Type II error
Reject $H_0$	Type I error	Correct decision

# Example

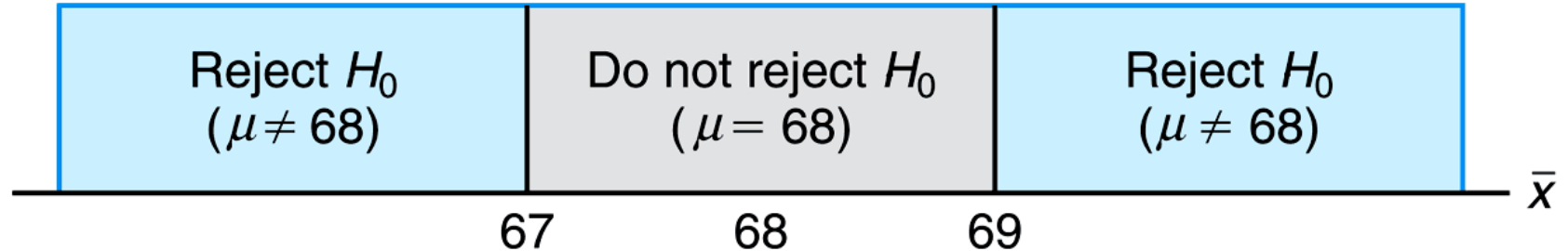


**Consider the null hypothesis that the average weight of male students in a certain college is 68 kilograms against the alternative hypothesis that it is unequal to 68.**

$$H_0: \mu = 68,$$

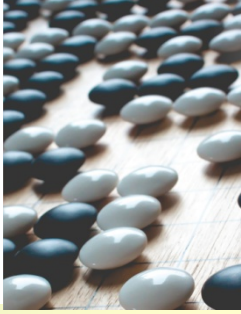
$$H_1: \mu \neq 68.$$

# Figure 10.4 Critical region (in blue)

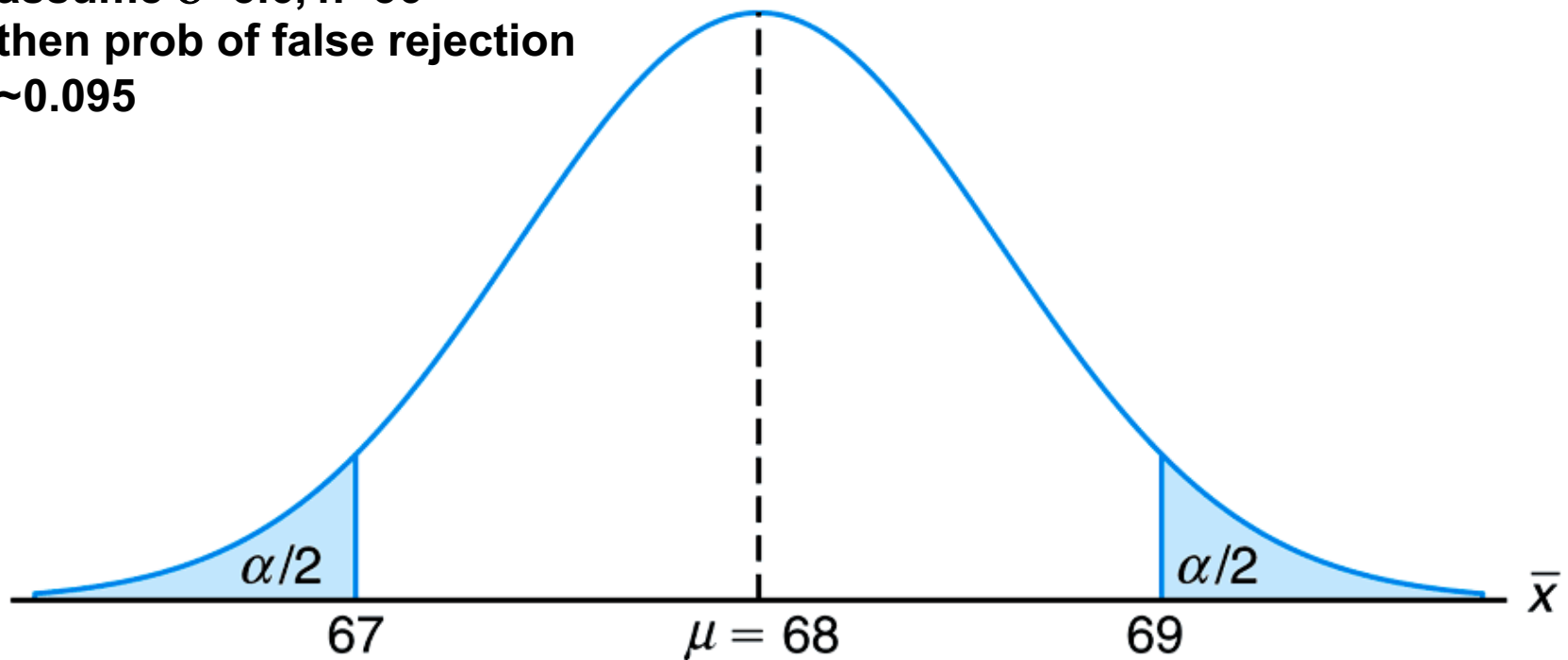


**Suppose we reject  $H_0$  if sample mean  $< 67$  or  $> 69$**

# Figure 10.5 Critical Region for testing $\mu = 68$ versus $\mu \neq 68$



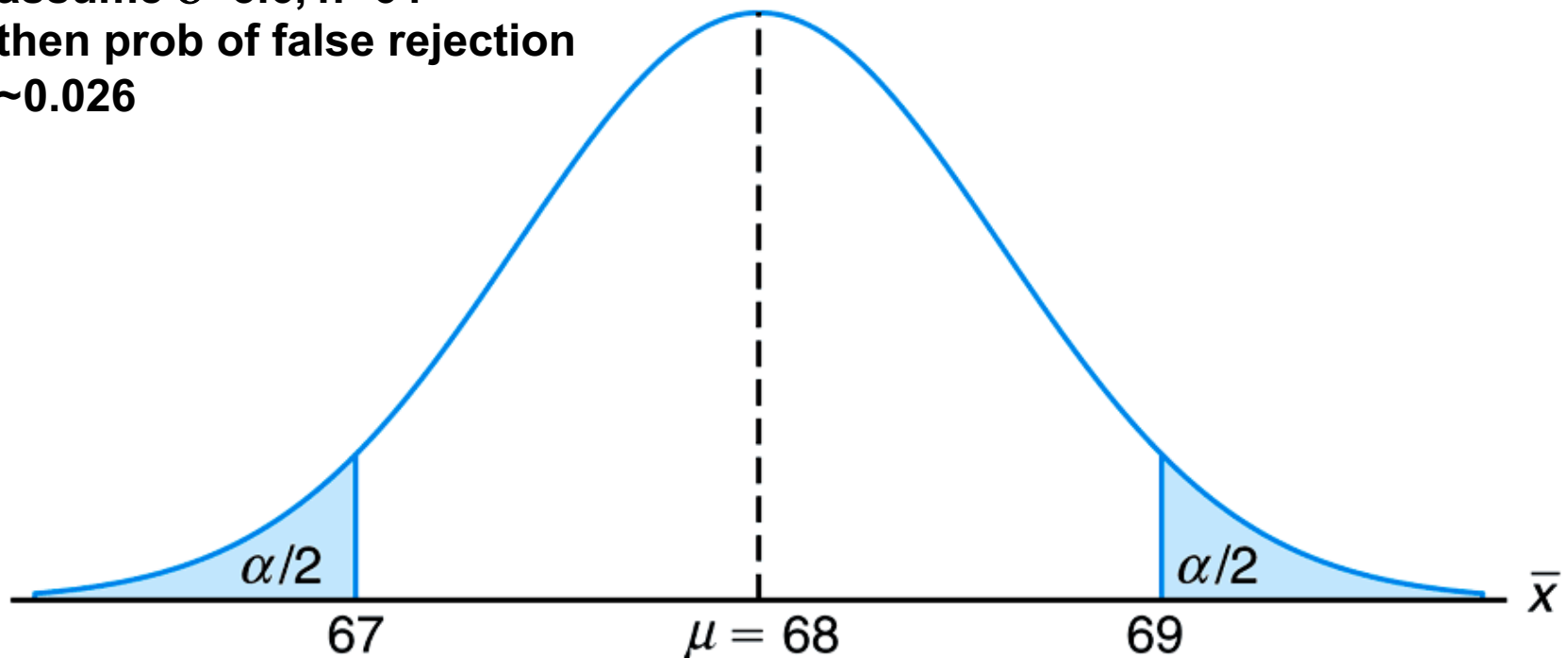
assume  $\sigma=3.6$ ,  $n=36$   
then prob of false rejection  
~0.095



# Figure 10.5 Critical Region for testing $\mu = 68$ versus $\mu \neq 68$



assume  $\sigma=3.6$ ,  $n=64$   
then prob of false rejection  
 $\sim 0.026$





- one-sided test

$$H_0: \theta = \theta_0,$$

$$H_1: \theta > \theta_0$$

**or**

$$H_0: \theta = \theta_0,$$

$$H_1: \theta < \theta_0,$$

- two-sided test

$$H_0: \theta = \theta_0,$$

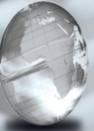
$$H_1: \theta \neq \theta_0,$$



# Section 10.3

## The Use of $P$ -Values for Decision Making in Testing Hypotheses

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# Definition 10.5



A ***P*-value** is the lowest level (of significance) at which the observed value of the test statistic is significant.

# P-value



- For given test statistic, probability of committing false rejection if the statistic corresponded to the critical value.

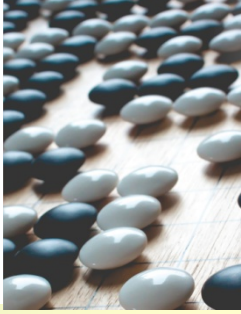
$$\begin{aligned}H_0: \mu \text{ is } 10 \\ H_1: \mu \text{ is not } 10\end{aligned}$$

- suppose  $z=1.87$ . If  $z$  were critical value, prob. of false rejection is

$$P = 2 P(z > 1.87) = 0.0614$$

This is called P-value

# P-value



- significance level of testing
- If the probability of random chance (false rejection) is below this level, we think the given statistic represent 'rare' event, and considers significant
- Typical significance level: 5%, 1%

# P-value

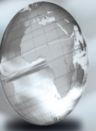


- Typical significance level: 5%, 1%
- If P value is less than the significance level, we reject  $H_0$ .
- In the example,  $P=0.614 > 5\%$ , so we cannot reject  $H_0$ 
  - Probability of random chance is too high!
- P-value is widely used

# Section 10.4

## Single Sample: Tests Concerning a Single Mean

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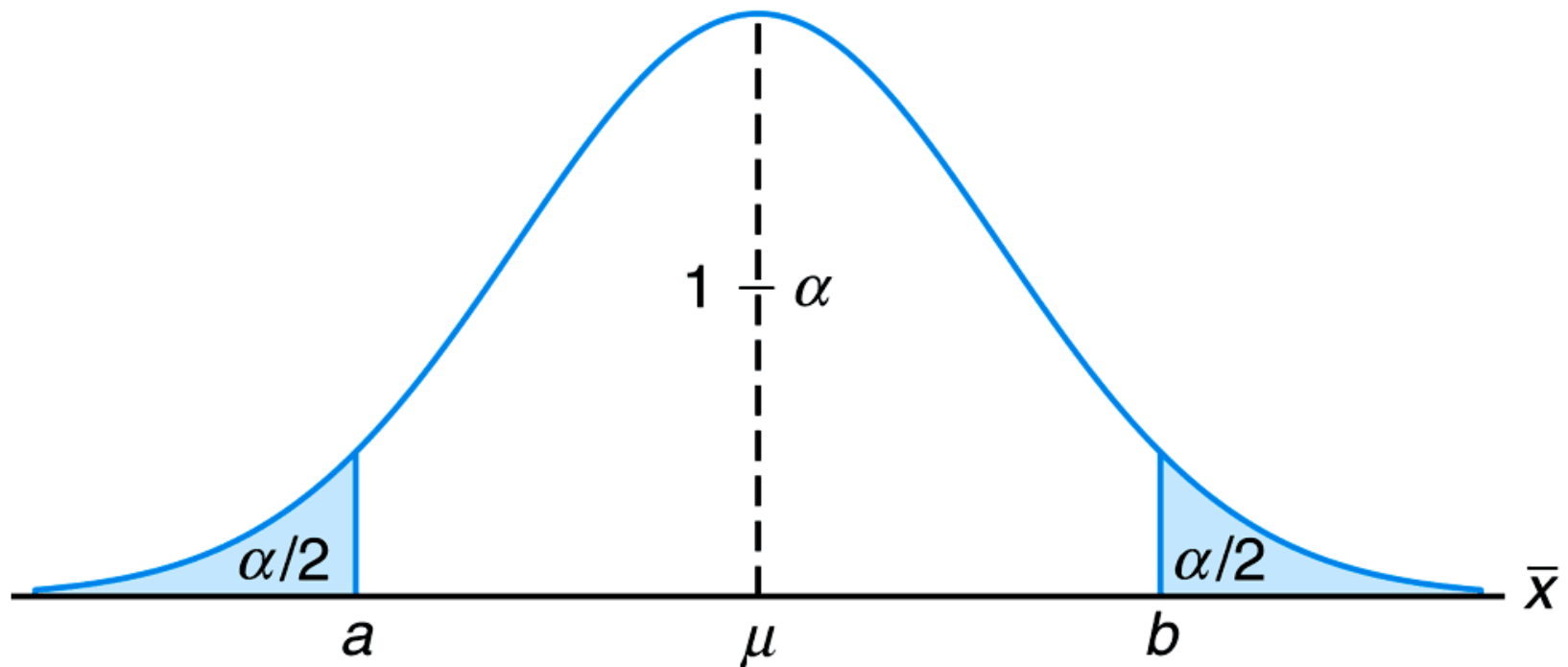


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**Figure 10.9** Critical region for the alternative hypothesis  $\mu \neq \mu_o$





# example 10.3




**Example 10.3:** A random sample of 100 recorded deaths in the United States during the past year showed an average life span of 71.8 years. Assuming a population standard deviation of 8.9 years, does this seem to indicate that the mean life span today is greater than 70 years? Use a 0.05 level of significance.

- Solution:**
1.  $H_0: \mu = 70$  years.
  2.  $H_1: \mu > 70$  years.
  3.  $\alpha = 0.05$ .
  4. Critical region:  $z > 1.645$ , where  $z = \frac{\bar{x} - \mu_0}{\sigma / \sqrt{n}}$ .
  5. Computations:  $\bar{x} = 71.8$  years,  $\sigma = 8.9$  years, and hence  $z = \frac{71.8 - 70}{8.9 / \sqrt{100}} = 2.02$ .
  6. Decision: Reject  $H_0$  and conclude that the mean life span today is greater than 70 years.

The  $P$ -value corresponding to  $z = 2.02$  is given by the area of the shaded region in Figure 10.10.

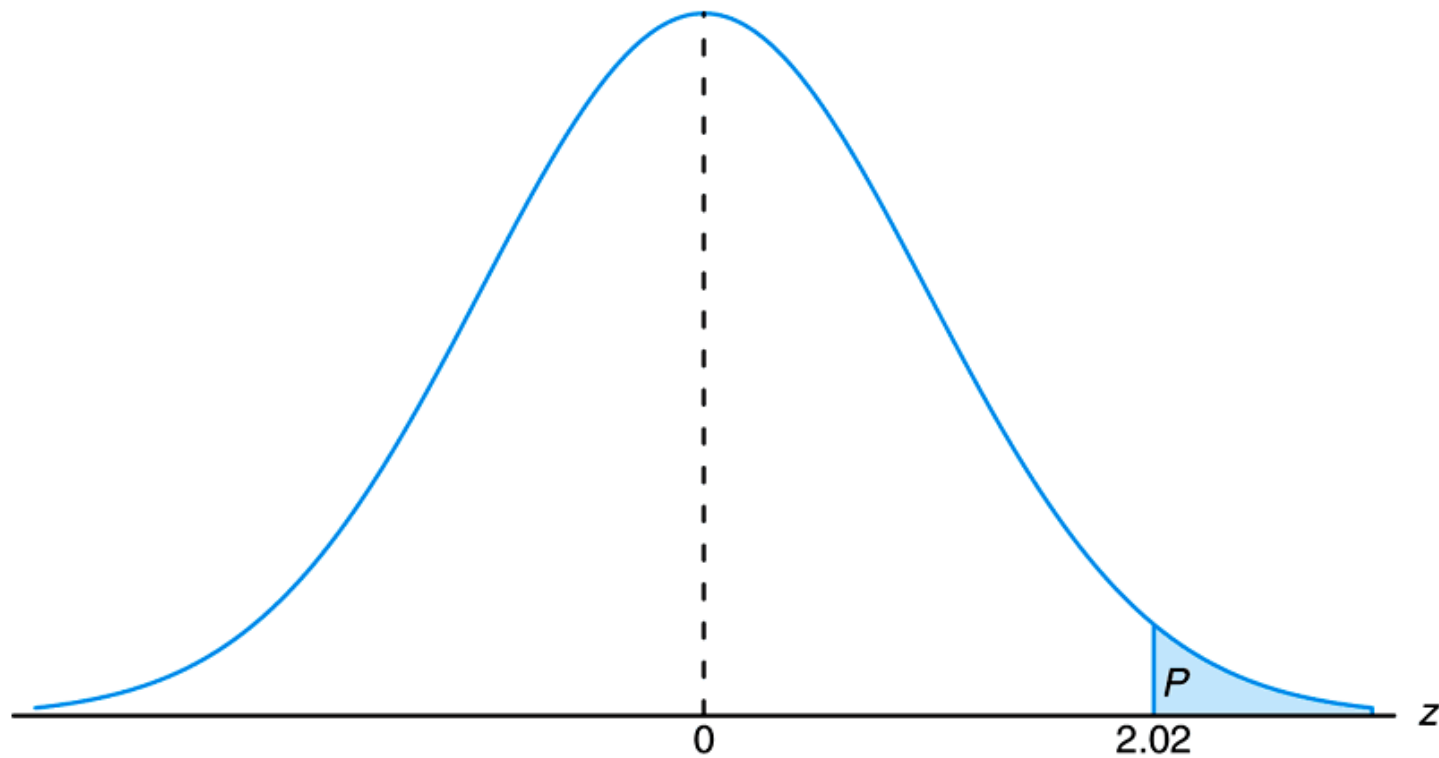
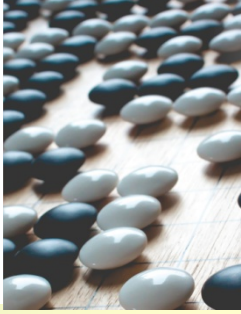
Using Table A.3, we have

$$P = P(Z > 2.02) = 0.0217.$$

As a result, the evidence in favor of  $H_1$  is even stronger than that suggested by a 0.05 level of significance. 



# Figure 10.10 P-value for Example 10.3



# example 10.4



**Example 10.4:** A manufacturer of sports equipment has developed a new synthetic fishing line that the company claims has a mean breaking strength of 8 kilograms with a standard deviation of 0.5 kilogram. Test the hypothesis that  $\mu = 8$  kilograms against the alternative that  $\mu \neq 8$  kilograms if a random sample of 50 lines is tested and found to have a mean breaking strength of 7.8 kilograms. Use a 0.01 level of significance.

- Solution:**
1.  $H_0: \mu = 8$  kilograms.
  2.  $H_1: \mu \neq 8$  kilograms.
  3.  $\alpha = 0.01$ .
  4. Critical region:  $z < -2.575$  and  $z > 2.575$ , where  $z = \frac{\bar{x} - \mu_0}{\sigma / \sqrt{n}}$ .
  5. Computations:  $\bar{x} = 7.8$  kilograms,  $n = 50$ , and hence  $z = \frac{7.8 - 8}{0.5 / \sqrt{50}} = -2.83$ .
  6. Decision: Reject  $H_0$  and conclude that the average breaking strength is not equal to 8 but is, in fact, less than 8 kilograms.

Since the test in this example is two tailed, the desired  $P$ -value is twice the area of the shaded region in Figure 10.11 to the left of  $z = -2.83$ . Therefore, using Table A.3, we have

$$P = P(|Z| > 2.83) = 2P(Z < -2.83) = 0.0046,$$

which allows us to reject the null hypothesis that  $\mu = 8$  kilograms at a level of significance smaller than 0.01. ■

# Figure 10.11 P-value for Example 10.4

