#### RANDOM PROCESS

21

HW

#### 19 efinition:

Let is be a sample space of a random experiment. A random process is a mapping x which assigns to every outcome ses a real valued function of time, x (t,s) i.e.,

$$X(S) = X(t,S)$$

The family of all such time functions is denoted by x(t,s) and is called a random process.

### Example:

If  $s = \{H, T\}$  then x assigning H into x(t, H) = sint and T into x(t, T) = cost

Classifications of Bandom process: metagogast and

-> Random processes can be classified into four types.

#### 1. Discrete Random process:

If s is discrete and T is continuous then the random process is called discrete random process.

#### Example:

het x(t) denotes the no. of telephone calls received in the interval (oit).

Here 
$$S = \{1, 2, 3, ...\}$$
  
 $Y = \{t; t \ge 0\}$ 

Here x(t) is a discrete random process.

# 2. Continuous Random process:

If both s and T are continuous then the random process is called continuous random process.

## Example:

(0,t). Here s is a continuous set (representing temperature values) and Y is a continuous set (representing time values).

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X(t) is a continuous random process.

3. Discrete Random Sequence: THU

If both s and I are discrete, the random process is Translation of the called a discrete random sequence.

Example:

fre and not be need to be and in this Let Xh denote the outcome of the nth toss of a fair die.

\* ... y = 15.17

T = 
$$\{1,2,3,---\}$$

So {xn, n=1,2,3, ... } is a discrete random sequence.

# mobiled a follow a con con 4. Continuous Random Sequence:

If s is a continuous and Y is discrete then the random process is called continuous random sequence.

#### Example:

Let Xn = Temperature at the end of the nth hour of a day. Then s is a continuous set (all possible values of temperature).

: (Xn, n=1, 2,3, -- 24) is a continuous random sequence.

Thought a short short short in the first of the south

is the flatter of the street of the

floto 5 34.0 3... }

r stepanie i

#### Mean:

The mean of the random process, x(t) is the expected value of the random variable x(t) at any time 't' i.e.

$$H_X(t) = E[X(t)]$$

# -Auto Correlation:

The auto correlation of the process {x(t): ter} defined by Rx(t,,t2) or R(t,,t2) 9s the expected value of the product of x(t1) & x(t2) of the process.

$$B(t_1,t_2)=E[x(t_1).x(t_2)]$$

# Auto covariance:

The auto covariance of the process & x(t) ; tery denoted by (x(tit2) or c(tit2)) Pardefined as, . ((t1,t2) = R(t1,t2) = M(t1). M(t2)

2M

2M

Auto Correlation Coefficient: 10000 A - 101x)

The auto correlation coefficient of the process {x(t); tet} denoted by,

Px(t1,t2) or P(t1,t2) is defined as;

$$P(t_1,t_2) = \frac{C(t_1,t_2)}{\sqrt{C(t_1,t_2)(t_2,t_2)}}$$

# Stationary process: The said of the party and the same

A random process is said to be stationary if its mean, variance are constants. Other process is called non-stationary.

# Problems:

7. TM A random process X(t) is given by X(t) = ycos27t, t>0 & cohere y' is a random variable with E(4) = 1 is the process is stationary.

Soln:

Given,

x(t) = ycosants t>0 ; E(y)=1

E(x(t)) = E[ycosant]

E(y) cosant

= 1. cosant

E[x(t)] = cosant +constant

. It is non-stationary process

2. If the random process x(t) takes the value [-1] with probability - 1/3 e takes the value [1] with probability - 2/3.

Find x(t) is a stationary or not.

Solo:

The random process x(t) with probability,

x(t)=n	-1	· · · · L
Po(+)	1/3	2/3

$$E[X(t)] = \sum_{n} P_n(t)$$

$$= (-1)(\frac{1}{3}) + 1(\frac{2}{3})$$

$$= -\frac{1}{3} + \frac{2}{3}$$

$$= \frac{1}{3}$$

$$E[x(t)] = \frac{1}{3} = \text{constant solution} \quad \text{adjustant of } \quad \text{constant solution} \quad \text{adjustant of } \quad \text{constant} \quad \text{consta$$

: Variance = 
$$E(x^2(t) - [E[x(t)]^2] = 0$$
 or only interest in the house of the part of th

· pik with - mo

. It is a stationary process.

3. The process 
$$\{x(t)\}$$
 whose probability distribution is given by,  $p[x(t)=n] = \begin{cases} \frac{(at)^{n-1}}{(1+at)^{n+1}}, n=1,2,3 \\ \frac{at}{1+at}, n=0 \end{cases}$ 

The probability distribution of x(t) is, Soln!

P	٦		to be to	- 1
x(t)=0	0	2	2	3
Pn(+)	ot	(1+at)2	at (1+at)3	(1+at)4
Temmera	1+at	J. 105.12	1000	12

Mean 
$$E[x(t)] = \frac{2}{n} \sin n(t)$$

$$= 0 + \frac{1}{(1+at)^{2}} + 2\left(\frac{at}{(1+at)^{3}}\right) + 3\left(\frac{(at)^{2}}{(1+at)^{4}}\right)$$

$$= \frac{1}{(1+at)^{2}} \left(\frac{at}{(1+at)}\right) + 3\left(\frac{at}{(1+at)^{4}}\right)$$

$$= \frac{1}{(1+at)^{2}} \left(\frac{1+at}{(1+at)}\right)^{2} \left(\frac{(1+x)^{2}}{(1+at)^{2}}\right)$$

$$= \frac{1}{(1+at)^{2}} \left(\frac{(1+at)^{2}}{(1+at)^{2}}\right)$$

$$= \frac{1}{(1+at)^{2}} \left(\frac{(1+at)^{2}}{(1+at)^{2}}\right)$$

$$= \frac{1}{(1+at)^{2}} \left(\frac{(1+at)^{2}}{(1+at)^{2}}\right)$$

$$E[x(t)] = 1$$

$$E[x^{2}(t)] = \sum_{n=0}^{\infty} n^{2} \cdot f_{n}(t)$$

$$= \sum_{n=0}^{\infty} [n(n+1) - n] \cdot f_{n}(t)$$

$$= \sum_{n=0}^{\infty} n(n+1) \cdot f_{n}(t) \cdot \sum_{n=0}^{\infty} n \cdot f_{n}(t)$$

$$= \sum_{n=0}^{\infty} n(n+1) \cdot f_{n}(t) - 1$$

$$= \left( 0 + 2 \left( \frac{1}{(1+at)^{2}} \right) + 6 \left( \frac{at}{(1+at)^{3}} \right) + 12 \left( \frac{(at)^{2}}{(1+at)^{4}} \right) - 1$$

$$= \left( \frac{2}{(1+at)^{2}} \left( 1 + 3 \left( \frac{at}{1+at} \right) + 6 \left( \frac{(at)}{(1+at)^{3}} \right) \right) - 1$$

$$= \frac{2}{(1+at)^{2}} \left( 1 - \frac{at}{1+at} \right)^{\frac{3}{2}} - 1$$

$$= \frac{2}{(1+at)^{2}} \left( \frac{1}{1+at} - at \right)^{\frac{3}{2}} - 1$$

$$= \frac{2}{(1+at)^{2}} \left( \frac{1}{1+at} \right)^{\frac{3}{2}} - 1$$

$$= \frac{2}{(1+at)^{2}} \left( \frac{1}{1+at} \right)^{\frac{3}{2}} - 1$$

$$= 2 \left( \frac{1}{1+at} \right)^{\frac{3}{2}} - 1$$

$$= \frac{2}{1+at} - 1$$

$$= \frac{2}{1+at$$

: It is a non-stationary process.

If the random process  $x(t) = \cos(t+\phi)$  where  $\phi'$  is I random variable with density function f(0) = 1/1, where (-7 < 0 2 1/2). Check whether the process is stationary or not.

Given,  $x(t) = \cos(t+\phi)$ 

 $f(\emptyset) = \frac{1}{\pi}$ ,  $-\frac{\pi}{3} + \emptyset < \frac{\pi}{3}$ ,  $+(\phi^{\dagger}) = \frac{1}{3}$ 

Soln!

$$E[x(t)] = E[\cos(t+\phi)]$$

$$= \int_{-\infty}^{\infty} \cos(t+\phi) \cdot d\phi d\phi$$

$$= \frac{1}{\pi} \int_{-\infty}^{\infty} \cos(t+\phi) \cdot d\phi$$

$$= \frac{1}{\pi} \int_{-\infty}^{\infty} \cos(t+\phi) \cdot d\phi$$

$$= \frac{1}{\pi} \left[ \sin(t+\phi) - \sin(t-\phi) \right]$$

$$= \frac{1}{\pi} \left[ \sin(x+\phi) - \sin(t-\phi) \right]$$

$$= \frac{1}{\pi} \left[ \sin(x+\phi) - \sin(x-\phi) \right]$$

$$= \frac{1}{\pi} \left[ \sin(x+\phi) - \sin(x-\phi) \right]$$

$$= \frac{1}{\pi} \left[ \sin(x+\phi) - \sin(x-\phi) \right]$$

$$= \frac{1}{\pi} \left[ \cos t + \cos t \right]$$

$$= \frac{1}{\pi} \left[ \cosh t = \cos t \right]$$

$$= \frac{1}{\pi} \left[ \cosh t = \cos t \right]$$

$$\therefore \text{ It isn't a constant value.}$$

$$\therefore \text{ It is a non-stationary process.}$$

7W 2.

If the random process  $x(t) = \sin(t+\phi)$  where  $\phi'$  is a random variable with density function  $f(\phi) = \frac{1}{n}$ , where  $(-\frac{\pi}{n} < \phi < \frac{\pi}{n})$ . Check whether the process is stationary or not.

Solo:

Given, 
$$x(t) = \sin(t+\phi)$$

$$f(\phi) = \frac{1}{\pi}, (-\frac{\pi}{3}c\phi c\frac{\pi}{3})$$

$$E[x(t)] = E[\sin(t+\phi)]$$

$$= \int_{-\pi/3}^{2} \sin(t+\phi) \cdot f(\phi) d\phi$$

$$= \int_{-\pi/3}^{2} \sin(t+\phi) \cdot \frac{1}{\pi} d\phi$$

$$= \int_{-\pi/3}^{2} \sin(t+\phi) \cdot \frac{1}{\pi} d\phi$$

$$= \frac{1}{\pi} \int_{-\pi/3}^{2} \sin(t+\phi) d\phi$$

$$= \frac{1}{\pi} \left[-\cos(t+\phi)\right]_{-\pi/3}^{\pi/3}$$

$$= \frac{1}{\pi} \left[-\cos(t+\phi)\right]_{-\pi/3}^{\pi/3}$$

$$= \frac{1}{\pi} \left[-\cos(t+\phi)\right]_{-\pi/3}^{\pi/3}$$

 $= \frac{1}{\pi} \left[ -\cos(t+\eta_0) + \cos(t-\eta_0) \right]$ 

$$= \frac{1}{\pi} \left[ \cos(\sqrt{2} + t) + \cos(\sqrt{2} - t) \right]$$

$$= \frac{1}{\pi} \left[ \sinh + \sinh \right]$$

$$= \frac{1}{\pi} \left[ \cosh t \right]$$

$$= \left[ x(t) \right] = \frac{2}{\pi} \sinh t$$

- : It isn't a constant value,
- .. It is a non-stationary process.

# Wide Sense Stationary Process:

A random process x(t) is said to be wide sense stationary ( wss), if the following conditions are satisfied.

\* Mx = E(x(t)) + Mean is constant with ita)

\* E[x(t)+(++1)] = Bxx(T)

Then the auto correlation process function only depends on the time difference ? - to no

6. 4M

\*

If the process x(t)= vcost + (v+1) sint then E(u) = E(v) = 0  $e E(u^2) = E(v^2) = 1$ . Find the process is wide sense stationary procession not + 01300 200 (arda 1200 to ]

Soln:

1 ( Sust 0+ HO + 8 | Ho ) 100 ( ) x(t) = vcost + (v+1) sint

E(u) = E(v) = 0  $E(u^2) = E(v^2) = 1$ 

E[x(t)] = E[vcost+(v+1)sint]

= E (ucost + vsint + sint)

= E(u) cost + E(v) sint + sint

11. DIF WELLO + O + STOTE CONTINUE I SO

E[x(t)] = sint

- .. It isn't constant value
- : It is not a wide sense stationary process.

of (vouses) A

IaM

Show that the random process x(t) = Acos (cot+0) is a wide sense stationary, where A & w dre constants & o' is uniformly distributed on the interval [0,25]

Soln!

Given, x(t) = Acos (wt+0)

The probability density function of  $f(0) = \frac{1}{b-a}$ 

$$f(0) = \frac{1}{2\pi - 0}$$

 $E[X(t)] = E[A\cos(\omega t + \theta)]$ = JAcos (wtte). Han do literation of the second server

 $= \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \left( \frac{1}{2\pi} \right) \right) \right) \right)^{2\pi} = \frac{A}{2\pi} \left( \frac{1}{2\pi} \left($ 

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= A (sincet-sinut) with suit out to should

E[x(t)] = 0

E(x(t)+(t+1)) = E[ncos(wt+0). A,cos(w(t+1)+0)].

= E[ 1200 (wt+0) (cos (wt+0+w2)),

= 1 A2 cos (w+0), cos (w+0+w1)]. 27. do

 $=\frac{\Lambda^2}{2\pi}\int_{0}^{2\pi}\left[\cos(\omega t+\theta+\omega t+\theta+\omega t)\right]d\theta$ 

 $=\frac{A^2}{4\pi}\int \left[\cos\left(2\omega t+20+\omega T\right)+\cos\omega T\right]\cdot d\theta$ 

 $= \frac{A^2}{4\pi} \left[ \frac{\sin(2\omega t + 20 + \omega t)}{2} + \cos(\omega t) \right]^{2\pi}$ 

 $=\frac{A^2}{4\pi}\left(\frac{\sin 2\omega t + 4\pi + \omega \tau}{2} + 2\pi \cos \omega \tau\right) - \frac{\sin 2\omega t + \omega \tau}{2}$ 

 $=\frac{A^2}{4\pi}\left[\frac{\sin(2\omega t+\omega t)}{2}+\frac{2\pi\cos(2\omega t+\omega t)}{2}\right]$ 

 $= \frac{A^2}{2} \left( \frac{p \times \cos(w)}{2} \right)^{\frac{1}{2}}$   $= \frac{A^2}{2} \left( \cos(w) \right)_{\mu}$ 

8.

Show that the random process X(t) = Asin (witter) is a cuide sense stationary, where A & w are constants & e' is uniformly distributed on the interval [0,27]

Solo:

The probability density function of  $f(\theta) = \frac{1}{b-a}$   $f(\theta) = \frac{1}{an-0}$ 

$$E[X(t)] = E[Asin(\omega t + \theta)]$$

$$= \int_{0}^{A} Asin(\omega t + \theta) \cdot \frac{1}{2\pi} d\theta$$

$$= \frac{A}{2\pi} \int_{0}^{2\pi} sin(\omega t + \theta) d\theta$$

$$= \frac{A}{2\pi} \left[ -cos(\omega t + 2\pi) + cos(\omega t) \right]$$

$$= \frac{A}{2\pi} \left[ -cos(\omega t + 2\pi) + cos(\omega t) \right]$$

$$= \frac{A}{2\pi} \left[ -cos(\omega t + 2\pi) + cos(\omega t) \right]$$

$$\mathbb{E}[x(t)] = 0$$

$$E[X(t)+(t+t)] = E[Asin(\omega t+0). Asin(\omega (t+t)+0)]$$

$$= E[A^{2}sin(\omega t+0). sin(\omega t+0+\omega t)]$$

$$= \int_{A^{2}}^{A^{2}}sin(\omega t+0). sin(\omega t+0+\omega t)] \cdot \frac{1}{2n} d\theta$$

$$= \frac{A^{2}}{2n} \int_{a}^{1} \left[cos(\omega t+0-\omega t-\phi-\omega t) - cos(\omega t+0+\omega t+0+\omega t)\right] d\theta$$

$$= \int_{an}^{2} \left[cos(\omega t+0-\omega t-\phi-\omega t) - cos(\omega t+0+\omega t+0+\omega t)\right] d\theta$$

$$= \int_{an}^{2} \left[cos(\omega t+0-\omega t-\phi-\omega t) - cos(\omega t+0+\omega t+0+\omega t)\right] d\theta$$

$$= \frac{A^{2}}{2\pi} \int_{0}^{1} \left[ \cos wt - \cos (2wt + 20 + wt) \right] d\theta$$

$$= \frac{A^{2}}{4\pi} \left[ -\frac{\sin (2wt + 20 + wt)}{3} + \cos wt \right]_{0}^{2\pi}$$

$$= \frac{A^2}{4\pi} \left[ \left( \cos \omega v (2\pi) - \frac{\sin (2\omega t + 4\pi + \omega t)}{2} \right) - \left( -\frac{\sin (2\omega t + \omega t)}{2} \right) \right]$$

$$= \frac{A^2}{4\pi} \left[ 2\pi \cos \omega v - \frac{\sin (2\omega t + \omega t)}{2} + \frac{\sin (2\omega t + \omega t)}{2} \right]$$

$$= \frac{A^2}{4\pi} \left( \frac{\pi}{\pi} \cos \omega t \right)$$

$$= \frac{\hbar^2}{2} (\cos \omega t)$$

Poisson's Distribution Process:

If x(t) represents the no. of occurance of a certain even in (oit) then the discrete random process {x(t)} is called poisson process.

$$p[x(t)=n] = \frac{e^{\lambda t} (\lambda t)^{n}}{n!}$$

- Properties of poisson process:
  - + The poisson process & a markovprocess.
  - → The sum of two independent poisson processes is again a poisson process.
  - The difference of two independent poisson processes is not a poisson process.

(Burson / sattunos - ]

Theorem 1: \*

Proove that the difference of two independent poisson 9. processes isn't a poisson process. 1 (10+10+4) = + [AGIN(WELS). AS (WEL+WITE)

Proof:

Let x1(t) & x2(t) is two poisson process with mean hie ha respectively.

$$E[x(t)] = E[x(t) - x_2(t)]$$

$$E[x(t)] = E[x(t) - x_2(t)]$$

$$= E[x_1(t)] - E[x_2(t)]$$

$$= \lambda_1 t - \lambda_2 t$$

$$E[x(t)] = (\lambda_1 - \lambda_2) t$$

$$= \mathbb{E} \left[ \mathbf{x}(\mathbf{t}) \right] = \mathbb{E} \left[ \mathbf{x}(\mathbf{t}) - \mathbf{x}_{2}(\mathbf{t}) \right]$$

$$= \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] - \mathbb{E} \left[ \mathbf{x}_{2}(\mathbf{t}) \right]$$

$$= \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] - \mathbb{E} \left[ \mathbf{x}_{2}(\mathbf{t}) \right]$$

$$= \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] + \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right]$$

$$= \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] = \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] + \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right]$$

$$= \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] = \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] + \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right]$$

$$= \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] = \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] + \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right]$$

$$= \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] = \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] + \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right]$$

$$= \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] = \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] + \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right]$$

$$= \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] + \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right] + \mathbb{E} \left[ \mathbf{x}_{1}(\mathbf{t}) \right]$$

$$E[(x_{1}(t))] = E[(x_{1}(t) - x_{2}(t))]^{2}$$

$$= E[(x_{1}(t))^{2} + (x_{2}(t))^{2} - 2x_{1}(t) \cdot x_{2}(t)]$$

$$= E[(x_{1}(t))^{2} + E[(x_{2}(t))^{2} - 2x_{1}(t) \cdot x_{2}(t)]$$

$$= \sum_{i=1}^{n} (x_{1}(t))^{2} + E[(x_{2}(t))^{2} - 2x_{1}(t) \cdot x_{2}(t)]$$

$$= \sum_{i=1}^{n} (x_{1}(t))^{2} + \sum_{i=1}^{n} (x_{2}(t))^{2} - 2x_{1}(t) \cdot x_{2}(t)$$

$$= \sum_{i=1}^{n} (x_{1}(t))^{2} + \sum$$

.. The difference of two independent poisson processes isn't a poisson process.

16 the poisson process is stationary ? ) - ( ) x ( ) x ( ) in Let X(t) be a poisson process, 15 . (a) 5 . (D) 1 .

E[x(t)] = xt + constant

: It is a not a stationary process

- Suppose the customer arraive at a bank, according to a 11. possson process with mean rate of a per min. Find the probability that during a time interval of a min.
  - di Exactly 4 customers arrawed.
  - (ii) Greater than 4 customers arraived.
  - ciii) Fewer than a customers arraived.

soln!

Let x(t) be the customer arraived at a bank, in the time interval 't'. x(t) follows the poisson's process.

$$\lim_{t \to \infty} P[x(t) = n] = \underbrace{e^{\lambda t} (\lambda t)^{\lambda 1}}_{n!} \quad [\lambda = 3, t = 2]$$

$$\begin{array}{rcl}
\mathbf{n} = \mathbf{H}; &= \frac{e^{2(2)}[3(2)]^{4}}{4!} &= \frac{e^{2(2)}[3(2)]^{4}}{e^{2(2)}[3(2)]^{4}} &= \frac{e^{2(2)}[$$

$$= (0.0024)(54)$$

$$= 0.13 \text{ (ii)} \quad P[X(\pm)>4] \Rightarrow [1-P(X(\pm)\pm4]] \qquad (iii) \quad P[X(\pm)>4] \Rightarrow \frac{1}{2!} + \frac{\bar{c}^6(6)^2}{2!} + \frac{\bar{c}^6(6)^4}{4!} = 1 - \left[\frac{\bar{c}^{3(2)}(6)^0}{0!} + \frac{\bar{c}^6(6)^1}{1!} + \frac{\bar{c}^6(6)^2}{2!} + \frac{\bar{c}^6(6)^4}{3!} + \frac{\bar{c}^6(6)^4}{4!}\right]$$

$$= 1 - \left( \frac{\bar{c}^{6}(1)}{1} + \frac{\bar{c}^{6}(6)}{1} + \frac{\bar{c}^{6}(36)}{2} + \frac{\bar{c}^{6}(6 \times 6 \times 8)}{3} + \frac{\bar{c}^{6}(6 \times 6 \times 8 \times 8)}{3 \times 2 \times 1} + \frac{\bar{c}^{6}(6 \times 6 \times 8 \times 8)}{2 \times 2 \times 2 \times 1} \right)$$

$$= 1 - \left( \frac{\bar{e}^6}{1} + \frac{6\bar{e}^6}{1} + 18\bar{e}^6 + 36\bar{e}^6 + 54\bar{e}^6 \right)$$

$$= 1 - 115(0.002)$$

(iii) 
$$P[x(t) < 4] = \left[\frac{\bar{e}^6(6)}{0!} + \frac{\bar{e}^6(6)}{1!} + \frac{\bar{e}^6(6)^2}{3!} + \frac{\bar{e}^6(6)^3}{3!}\right]$$

$$= \left[\frac{\bar{e}^6(1)}{1!} + \frac{\bar{e}^6(6)}{1!} + \frac{\bar{e}^6(6)^4}{3!} + \frac{\bar{e}^6(6)^3}{3!} + \frac{\bar{e}^6(6)^3}{3!}\right]$$

$$= \left[\bar{e}^6 + 6\bar{e}^6 + 18\bar{e}^6 + 36\bar{e}^6\right]$$

$$= 61\bar{e}^6$$

$$= 61(0.0024)$$

$$= 0.1464 \%$$

Queries presented in a computer database are following a poisson's process a mean rate, \$16, queries per min. An experiment consist of monitoring a database for in minutes & recording in the notof queries presented in What is the probability that no queries arraive in a min arrival.

in what is the probability that exactly 6 queries arraive

(iii) What is the probability that less than 3 queries arraive in 1/2 min interval.

Soln:

Let X(t) be the queries presented in the time arrival t. X(t) follows the poisson's process.

(i) 
$$P[x(t)=n] = \frac{e^{\lambda t} (\lambda t)^n}{n!} (\lambda = 6, t = 1)$$
  
 $n=0; = \frac{e^{6(1)} (6(1))^0}{0!}$   
 $= \frac{e^{6} (6)^0}{0!}$   
 $= \frac{e^{6} (1)}{1} = e^{6}$   
 $= 0.0024 / 1$ 

(iii) 
$$P[x(t) < 3] = \left[\frac{\bar{c}^{3}(3)}{(3)} + \frac{\bar{c}^{3}(3)}{(3)} + \frac{\bar{c}^{3}(3)^{2}}{2!}\right] + \frac{\bar{c}^{3}(3)^{2}}{3!}$$

$$= \left[\frac{\bar{c}^{3}(1)}{(1)} + \frac{\bar{c}^{3}(3)}{(2)} + \frac{\bar{c}^{3}(3)^{2}}{2!}\right] + \frac{\bar{c}^{3}(3)^{2}}{3!}$$

$$= \left[\frac{\bar{c}^{3}(1)}{(1)} + \frac{\bar{c}^{3}(3)}{(2)} + \frac{\bar{c}^{3}(3)^{2}}{2!}\right]$$

$$= \left[\bar{c}^{3} + 3\bar{c}^{3} + \mu_{1} + 5\bar{c}^{3}\right]$$

$$= 8.5\bar{c}^{3}$$

$$= 8.5(0.05)$$

$$= 0.425/l$$

# \* Markov Process:

- → It is a random process in which future behaviour of the process depends only on the present state not on the states in the past.
- Ex: ii) The probability of raining today depends only on the previous conditions existed for the last 2 days e not on the past weather conditions.
  - (ii) Snake e ladder game.

### Markov Chain:

We define the markov chain as follows, if  $P.\{x_n = \alpha n/x_{n-1} = \alpha n-1, x_{n-2} = \alpha n-2, x_n = \alpha 0\}$   $P\{x_n = \alpha n/x_{n-1} = \alpha n-1\} \text{ for all } n, \text{ then the process } \{x_n\},$   $n = 0,1,2... \text{ is called as markov chain, where } \alpha_1,\alpha_2,...\text{ an are called state of the markov chain.}$ 

# Transition Probability Matrix [TPM]:

+ If P=[Pij] is a square matrix & is called the TPM of the chain. If the no. of states is finite say im' then the matrix ip' will be a non square matrix. Otherwise the matrix will be infinite.

\* State Chapman - Kolmogorov Theorem:  $\rightarrow \text{If 'p' is the TPM of a markov chain, then the 'n' step TPM;}$   $p^{(n)} = p^n \text{ i.e., } [Pij^{(n)}] = [Pij]^n$ 

Steady State Distribution:

(ii) 
$$n_1+n_2+n_3=1$$
; where  $n=[n_1 n_2]$  or (or)  $n_1+n_2=1$ 

12M VV2mp 13.

The 3 state markov chain is given by transition probability matrix [TPM] where,  $p = \begin{bmatrix} 0 & 2/3 & 1/3 \\ 1/2 & 0 & 1/2 \\ 0 & 1/2 & 1/2 \end{bmatrix}$ . Find the steady state distribution.

Soln:

Given, 
$$P = \begin{bmatrix} 0 & 2/3 & 1/3 \\ 1/2 & 0 & 1/2 \\ 0 & 1/2 & 1/2 \end{bmatrix}$$

$$\left[0+\frac{N_2}{2}+0 \quad \frac{2N_1}{3}+\frac{N_3}{3} \quad \frac{N_1}{3}+\frac{N_2}{2}+\frac{N_3}{2}\right] = \left[N_1 N_2 N_3\right]$$

$$\begin{bmatrix} \frac{\overline{N}2}{2} & \frac{2\overline{N}1}{3} + \frac{\overline{N}3}{2} & \frac{\overline{N}1}{2} + \frac{\overline{N}2 + \overline{N}3}{2} \end{bmatrix} = \begin{bmatrix} \overline{N}_1 & \overline{N}_2 & \overline{N}_3 \end{bmatrix}^{-1}$$

$$\frac{\pi_2}{2} = \pi_1$$
 ;  $\frac{2\pi_1}{3} + \frac{\pi_3}{2} = \pi_2$  ;  $\frac{\pi_1}{3} + \frac{\pi_2 + \pi_3}{2} = \pi_3$ 

$$3\pi_1+\pi_3=1$$

$$\boxed{ \mathcal{T}_3 = 1 - 3 \mathcal{T}_1 } \Rightarrow (4)$$

$$\left| T_1 = 3/17 \right| \rightarrow (5)$$

Substitute (5) in (1)

27,+372-373=0

Substitute (5) in (4)

$$m_3 = 1-3(3/m)$$

[:1] = [1] 97 [ ...]

$$\therefore [n_1 \ n_2 \ n_3] = [3/17 \ 6/17 \ 8/17]$$

∴ 
$$\pi_1 + \pi_2 + \pi_3 = 1$$

$$\frac{3}{17} + \frac{6}{17} + \frac{8}{17} = 1$$

$$\frac{3+6+8}{17} = 1$$

$$\frac{17}{17} = 1$$
Hence, verified y

14.

The 2 state markov chain is given by TPM, where P = [0.3 0.7] Find steady state distribution.

Soln!

(i) 
$$\pi p = \pi$$

$$\left( \begin{array}{ccc} \pi_1 & \pi_2 \end{array} \right) \left[ \begin{array}{ccc} 0.3 & 0.7 \\ 0.4 & 0.6 \end{array} \right] = \left( \begin{array}{ccc} \pi_1 & \pi_2 \end{array} \right) \quad ,$$

[0.3×1+0.4×2 0.7×1+0.6×2] = (×1 ×2)

$$0.7\pi_1 + 0.6\pi_2 - \pi_2 = 0$$

$$0.7(\frac{4}{11}) - 0.470 = 0$$

$$\pi_1 + \frac{0.7}{0.4} \pi_1 = 1$$

$$\begin{bmatrix} 2.0 & 8.6 & 8.6 \\ 9.7 & 8.5 & 1.0 \\ 10. & 8.6 & 1.0 \end{bmatrix} = \begin{bmatrix} 2.8 & -0.4 & 0.2 & 0.0 \\ 11 & -0.4 & 0.0 & 0.0 \end{bmatrix} = 6.7$$

 $\pi_1 + \frac{7}{4}\pi_1 = 1$ 

$$\frac{2.87}{11} = 9.4 \Pi_2$$

$$T_1 + \frac{7}{4}T_1 = 1$$

$$\frac{\overline{D_2} = \frac{1}{2}}{\overline{D_2}} = \frac{1}{2}$$

$$\frac{\overline{D_2} = \frac{1}{2}$$

$$\frac{\overline{D_2} = \frac{1}{2}}{\overline{D_2}} = \frac{1}{2}$$

$$\frac{\overline{D_2} = \frac{1}$$

$$\therefore \left[ \left[ \begin{array}{c} \tau_1 \\ \tau_2 \end{array} \right] = \left[ \begin{array}{c} \frac{1}{11} \\ \frac{1}{11} \end{array} \right]$$

$$\boxed{\underline{n} = \frac{4}{11}} = 0 + 0 \qquad \text{(i.i. } \boxed{\underline{n}} = \frac{4}{11}, \frac{7}{11} = 0 + 0 = 0 = 0$$

PP = [0.61 0.42 0.49]

[1 0][1 2 ] . Ma

15.

Probability distribution based on the initial distribution, if the initial state probability distribution of a markov chain is  $p^{(0)} = \begin{bmatrix} 5, \frac{1}{6} \end{bmatrix}$  e the TPM of the chain is  $p = \begin{bmatrix} 0 & 1 \\ 3 & 3 \end{bmatrix}$ . Find the probability distribution of the chain after a steps.

Colu;

Given,  $p^{(0)} = \left[\frac{5}{6}, \frac{1}{6}\right]$ ;  $p = \begin{bmatrix} 0 & 1 \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix}$ 

$$p^{(i)} = p^{(0)}, p$$

$$p^{(i)} = \left(\frac{5}{6} \frac{1}{6}\right) \begin{bmatrix} 0 & 1\\ \frac{1}{2} & \frac{1}{2} \end{bmatrix}$$

$$p^{(i)} = \left(0 + \frac{1}{12} \frac{5}{6} + \frac{1}{12}\right)$$

$$p^{(i)} = \left(\frac{1}{12} \frac{11}{12}\right)$$

$$p^{(2)} = p^{(1)}p$$

$$p^{(2)} = \begin{bmatrix} \frac{1}{12} & \frac{11}{12} \end{bmatrix} \begin{bmatrix} 0 & 1 \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix} \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix}$$

$$p^{(2)} = \begin{bmatrix} 0 + \frac{11}{24} & \frac{1}{2} + \frac{11}{24} \\ \frac{1}{24} & \frac{1}{24} \end{bmatrix}$$

$$p^{(2)} = \begin{bmatrix} \frac{11}{24} & \frac{13}{24} \\ \frac{12}{24} & \frac{13}{24} \end{bmatrix}$$

16. If the initial state probability distribution of a markov chain is 
$$p^{(0)} = [1 \ 0] e$$
 the TPM of  $P = \begin{bmatrix} 0.6 \ 0.4 \\ 0.2 \ 0.8 \end{bmatrix}$ . Find  $p^{(2)}$ 

$$p^{(1)} = p^{(0)}.p$$

$$p^{(1)} = (10) \begin{cases} 0.6 & 0.4 \\ 0.2 & 0.8 \end{cases}$$

$$p^{(1)} = (0.640 & 0.440)$$

$$p^{(1)} = [0.60.4]$$

$$p^{(0)} = \begin{bmatrix} 1 & 0 \end{bmatrix}; \quad p = \begin{bmatrix} 0.6 & 0.4 \\ 0.2 & 0.8 \end{bmatrix}$$

$$p^{(1)} = p^{(0)}.p$$

$$p^{(1)} = \begin{bmatrix} 1 & 0 \end{bmatrix} \begin{bmatrix} 0.6 & 0.4 \\ 0.2 & 0.8 \end{bmatrix} \qquad p^{(2)} = \begin{bmatrix} 0.6 & 0.4 \end{bmatrix} \begin{bmatrix} 0.6 & 0.4 \\ 0.2 & 0.8 \end{bmatrix}$$

$$p^{(1)} = \begin{bmatrix} 0.6+0 & 0.4+0 \end{bmatrix} \qquad p^{(2)} = \begin{bmatrix} 0.36+0.08 & 0.24+0.32 \end{bmatrix}$$

$$p^{(1)} = \begin{bmatrix} 0.6 & 0.4 \end{bmatrix} \qquad p^{(2)} = \begin{bmatrix} 0.44 & 0.56 \end{bmatrix}$$

The initial process of the markov TPM is given by, 
$$P = \begin{bmatrix} 0.2 & 0.3 & 0.5 \\ 0.1 & 0.3 & 0.7 \end{bmatrix}$$
 with initial probabilities  $p_1^{(0)} = 0.4$ ,  $p_2^{(0)} = 0.3$ ,  $0.6 & 0.3 & 0.1 \end{bmatrix} p_3^{(0)} = 0.8$ . Find  $p_1^{(1)}$ ,  $p_2^{(1)} \in p_3^{(1)}$ .

Soln:

ver),
$$p^{(0)} = \begin{bmatrix} 0.4 & 0.3 & 0.8 \end{bmatrix}, \quad p = \begin{bmatrix} 0.1 & 0.3 & 0.5 \\ 0.1 & 0.2 & 0.7 \\ 0.6 & 0.3 & 0.1 \end{bmatrix}$$

$$p^{(1)} = [0.4 \ 0.3 \ 0.8] \begin{bmatrix} 0.2 \ 0.3 \ 0.5 \\ 0.1 \ 0.2 \ 0.7 \\ 0.6 \ 0.3 \ 0.1 \end{bmatrix}$$

$$p^{(1)} = [0.08 + 0.03 + 0.48] \begin{bmatrix} 0.12 + 0.06 + 0.24 \\ 0.12 + 0.06 + 0.24 \end{bmatrix}$$

$$p^{(1)} = [0.59 \ 0.42 \ 0.49]$$

$$P_{3}^{(1)} = 0.59$$

$$P_{3}^{(1)} = 0.42$$

$$P_{3}^{(1)} = 0.49$$