# Text S1: Derivatives of the likelihood

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As mentioned in Appendix S1, maximization is performed on the log-likelihood. In this section we give the derivations of the derivatives of the log-likelihood, which were used to aid optimization. In the main paper we used the symbol y to denote the distance from transect to object. In practice, this distance is a perpendicular distance in the case of line transects and a radial distance for point transects. Here, to avoid confusion, perpendicular distances in the line transect case are denoted by x and radial distances in the point transect case are denoted by r.

#### Line transects

Starting from the log-likelihood:

$$l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{x}, \mathbf{Z}) = \sum_{i=1}^{n} \left( \log \sum_{j=1}^{J} \phi_{j} g_{j}(x_{i}, \mathbf{Z}; \boldsymbol{\theta}_{j}) - \log \sum_{j=1}^{J} \phi_{j} \mu_{ij} \right)$$
(1)

we derive the derivatives with respect to the optimisation parameters.

#### With respect to $\beta_{0i*}$

For the intercept terms (also considering in the non-covariate case, these are just the parameters), the parameters have no effect outside of their mixture (ie.  $\beta_{0j*}$  only has an influence on mixture component j\*), so we can write:

$$\frac{\partial l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{x}, \mathbf{Z})}{\partial \beta_{0j*}} = \sum_{i=1}^{n} \frac{1}{g(x_i, \mathbf{Z}; \boldsymbol{\theta}, \boldsymbol{\phi})} \phi_{j*} \frac{\partial}{\partial \beta_{0j*}} g_{j*}(x_i, \mathbf{Z}; \boldsymbol{\theta}_{j*}) - \frac{\phi_{j*}}{\mu_i} \frac{\partial}{\partial \beta_{0j*}} \mu_{ij*}.$$

Now, to first find  $\frac{\partial}{\partial \beta_{0j_*}} g_{j_*}(x_i, \mathbf{Z}; \boldsymbol{\theta}_{j_*})$ :

$$\frac{\partial g_{j*}(x_i, \mathbf{Z}; \boldsymbol{\theta}_{j*})}{\partial \beta_{0j*}} = \frac{\partial}{\partial \beta_{0j*}} \exp\left(-\frac{x_i^2}{2\sigma_{j*}^2}\right),$$

applying the chain rule and remembering that  $\sigma_{j*}$  is a (trivial) function of the  $\beta_{0j}$ s:

$$\frac{\partial g_{j*}(x_i, \mathbf{Z}; \boldsymbol{\theta}_{j*})}{\partial \beta_{0j*}} = \left(\frac{x_i}{\sigma_{j*}}\right)^2 \exp\left(-\frac{x_i^2}{2\sigma_{j*}^2}\right)$$

Expressing  $\mu_{ij*}$  in terms of the error function, Erf:

$$\frac{\partial \mu_{ij*}}{\partial \beta_{0j*}} = \frac{\partial}{\partial \beta_{0j*}} \left( \sqrt{\frac{\pi}{2}} \sigma_{j*} \operatorname{Erf} \left( \frac{w}{\sqrt{2\sigma_{j*}^2}} \right) \right) 
= \operatorname{Erf} \left( \frac{w}{\sqrt{2\sigma_{j*}^2}} \right) \frac{\partial}{\partial \beta_{0j*}} \left( \sqrt{\frac{\pi}{2}} \sigma_{j*} \right) + \sqrt{\frac{\pi}{2}} \sigma_{j*} \frac{\partial}{\partial \beta_{0j*}} \left( \operatorname{Erf} \left( \frac{w}{\sqrt{2\sigma_{j*}^2}} \right) \right)$$
(2)

To find  $\frac{\partial}{\partial \beta_{0j*}} \operatorname{Erf}\left(\frac{w}{\sqrt{2\sigma_{i*}^2}}\right)$ , note that we can write and then apply the chain rule:

$$\frac{\partial}{\partial \beta_{0j*}} \operatorname{Erf}\left(\frac{w}{\sqrt{2\sigma_{j*}^2}}\right) = \frac{\partial}{\partial \beta_{0j*}} S(u(\sigma_{j*}))$$
$$= \frac{\partial S(u)}{\partial u} \frac{\partial u(\sigma_{j*})}{\partial \sigma_{j*}} \frac{\partial \sigma_{j*}}{\partial \beta_{0j*}}$$

where

$$S(u) = \int_0^u \exp(-t^2) dt \quad \text{and} \quad u(\sigma_{j*}) = \frac{w}{\sqrt{2\sigma_{j*}^2}}.$$

Their derivatives being

$$\frac{\partial S(u)}{\partial u} = \frac{2}{\sqrt{\pi}} \exp(-u^2), \quad \frac{\partial u(\sigma_{j*})}{\partial \sigma_{j*}} = -\frac{w}{\sqrt{2}} \sigma_{j*}^{-2}.$$

Given these terms, it is just a case of multiplying them:

$$\frac{\partial S(u)}{\partial u} \frac{\partial u(\sigma_{j*})}{\partial \sigma_{j*}} \frac{\partial \sigma_{j*}}{\partial \beta_{0j*}} = -\sqrt{\frac{2}{\pi}} \frac{w}{\sigma_{j*}} \exp\left(-\frac{w^2}{2\sigma_{j*}^2}\right)$$

Substituting into (2):

$$\frac{\partial \mu_{ij*}}{\partial \beta_{0j*}} = \mu_{ij*} - w \exp\left(-\frac{w^2}{2\sigma_{j*}^2}\right)$$

Finally, the derivative is:

$$\frac{\partial l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{x}, \mathbf{Z})}{\partial \beta_{0j*}} = \sum_{i=1}^n \left(\frac{x_i}{\sigma_{j*}}\right)^2 \phi_{j*} \frac{g_{j*}(x_i, \mathbf{Z}; \boldsymbol{\theta}_{j*})}{g(x_i, \mathbf{Z}; \boldsymbol{\theta}, \boldsymbol{\phi})} - \frac{\phi_{j*}}{\mu_i} (\mu_{ij*} - wg_{j*}(w, \mathbf{Z}; \boldsymbol{\theta}_{j*})).$$

## With respect to $\beta_{k*}$

Derivatives with respect to the common covariate parameters are found in a similar way to above. The expressions are slightly more complicated since the  $\beta_k$ s effect all of the mixture components.

$$\frac{\partial l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{x}, \mathbf{Z})}{\partial \beta_{k*}} = \sum_{i=1}^{n} \left( \frac{1}{g(x_i, \mathbf{Z}; \boldsymbol{\theta}, \boldsymbol{\phi})} \sum_{j=1}^{J} \phi_j \frac{\partial}{\partial \beta_{k*}} g_j(x_i, \mathbf{Z}; \boldsymbol{\theta}_j) - \frac{1}{\mu_i} \sum_{j=1}^{J} \phi_j \frac{\partial}{\partial \beta_{k*}} \mu_{ij} \right)$$

Every  $\sigma_j$  is a function of the  $\beta_k$ s, so:

$$\frac{\partial \sigma_j}{\partial \beta_{k*}} = \frac{\partial}{\partial \beta_{k*}} \exp\left(\beta_{0j} + \sum_{k=1}^K z_{ik} \beta_k\right),$$
$$= z_{ik*} \sigma_j.$$

Hence:

$$\frac{\partial}{\partial \beta_{k*}} \exp\left(-\frac{x_i^2}{2\sigma_j^2}\right) = z_{k*} \left(\frac{x_i}{\sigma_j}\right)^2 \exp\left(-\frac{x_i^2}{2\sigma_j^2}\right) = z_{k*} \left(\frac{x_i}{\sigma_j}\right)^2 g_j(x_i, \mathbf{Z}; \boldsymbol{\theta}_j). \tag{3}$$

And so for the  $\mu_{ij}$ s:

$$\frac{\partial \mu_{ij}}{\partial \beta_{k*}} = z_{ik*} \left( \mu_{ij} - w \exp\left(-\frac{w^2}{2\sigma_i^2}\right) \right)$$

The derivative is then:

$$\frac{\partial l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{x}, \mathbf{Z})}{\partial \beta_{k*}} = \sum_{i=1}^{n} \left( \frac{1}{g(x_i, \mathbf{Z}; \boldsymbol{\theta}, \boldsymbol{\phi})} \sum_{j=1}^{J} \phi_j z_{k*} \left( \frac{x_i}{\sigma_j} \right)^2 g_j(x_i, \mathbf{Z}; \boldsymbol{\theta}_j) - \frac{1}{\mu_i} \sum_{j=1}^{J} \phi_j z_{ik*} (\mu_{ij} - w g_j(x_i, \mathbf{Z}; \boldsymbol{\theta}_j)) \right)$$

### With respect to $\alpha_{j*}$

First note that we can write the likelihood (1) as:

$$l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{x}, \mathbf{Z}) = \sum_{i=1}^{n} \left( \log \left( \sum_{j=1}^{J-1} \phi_j g_j(x_i, \mathbf{Z}; \boldsymbol{\theta}_j) + (1 - \sum_{j=1}^{J-1} \phi_j) g_J(x_i, \mathbf{Z}; \boldsymbol{\theta}_J) \right) - \log \left( \sum_{j=1}^{J-1} \phi_j \mu_{ij} + (1 - \sum_{j=1}^{J-1} \phi_j) \mu_{ij} \right) \right)$$

The derivatives with respect to the  $\alpha_{j*}$  of this expression are then:

$$\frac{\partial l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{x}, \mathbf{Z})}{\partial \alpha_{j*}} = \left(\sum_{i=1}^{n} \frac{1}{g(x_{i}, \mathbf{Z}; \boldsymbol{\theta}, \boldsymbol{\phi})} \left(\sum_{j=1}^{J-1} g_{j}(x_{i}, \mathbf{Z}; \boldsymbol{\theta}_{j}) \frac{\partial \phi_{j}}{\partial \alpha_{j*}} - g_{J}(x_{i}, \mathbf{Z}; \boldsymbol{\theta}_{J}) \sum_{j=1}^{J-1} \frac{\partial \phi_{j}}{\partial \alpha_{j*}}\right) - \frac{1}{\mu_{i}} \left(\sum_{j=1}^{J-1} \mu_{ij} \frac{\partial \phi_{j}}{\partial \alpha_{j*}} - \mu_{iJ} \sum_{j=1}^{J-1} \frac{\partial \phi_{j}}{\partial \alpha_{j*}}\right)\right) \tag{4}$$

Finding the derivatives is then simply a matter of finding the derivatives of  $\phi_j$  with respect to  $\alpha_{j*}$  and substituting them back into (4).

$$\frac{\partial \phi_j}{\partial \alpha_{j*}} = \frac{\partial}{\partial \alpha_{j*}} F(\sum_{p=1}^j e^{\alpha_p}) - \frac{\partial}{\partial \alpha_{j*}} F(\sum_{p=1}^{j-1} e^{\alpha_p}).$$

Looking at each of the terms:

$$\frac{\partial}{\partial \alpha_{j*}} F(\sum_{p=1}^{j} e^{\alpha_p}) = A_j = \begin{cases} e^{\alpha_{j*}} f(\sum_{p=1}^{j} e^{\alpha_p}) & \text{for } j \ge j*, \\ 0 & \text{for } j < j*. \end{cases}$$

and

$$\frac{\partial}{\partial \alpha_{j*}} F(\sum_{p=1}^{j-1} e^{\alpha_p}) = A_{(j-1)} = \begin{cases} e^{\alpha_{j*}} f(\sum_{p=1}^{j-1} e^{\alpha_p}) & \text{for } j-1 \ge j*, \\ 0 & \text{for } j-1 < j*. \end{cases}$$

So

$$\frac{\partial \phi_j}{\partial \alpha_{j*}} = A_j - A_{j-1}.$$

Substituting these back into (4) and re-arranging gives:

$$\frac{\partial l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{x}, \mathbf{Z})}{\partial \alpha_{j*}} = \sum_{i=1}^{n} \left( \frac{1}{g(x_i, \mathbf{Z}; \boldsymbol{\theta}, \boldsymbol{\phi})} \sum_{j=1}^{J-1} (A_j - A_{j-1}) (g_j(x, \mathbf{Z}; \boldsymbol{\theta}_j) - g_J(x, \mathbf{Z}; \boldsymbol{\theta}_J)) - \frac{1}{\mu_i} \sum_{j=1}^{J-1} (A_j - A_{j-1}) (\mu_{ij} - \mu_{iJ}) \right)$$

#### Point transects

We now provide the corresponding quantities for point transects, starting from the log-likelihood:

$$l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{r}, \mathbf{Z}) = n \log 2\pi + \sum_{i=1}^{n} \left( \log r_i + \log \sum_{j=1}^{J} \phi_j g_j(r_i, \mathbf{Z}; \boldsymbol{\theta}_j) - \log \sum_{j=1}^{J} \phi_j \nu_{ij} \right).$$
 (5)

# With respect to $\beta_{0j}$

From (5), one can see that we obtain:

$$\frac{\partial l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{r}, \mathbf{Z})}{\partial \beta_{0j*}} = \sum_{i=1}^{n} \left( \frac{\partial}{\partial \beta_{0j*}} \log \sum_{j=1}^{J} \phi_{j} g_{j}(r_{i}, \mathbf{Z}; \boldsymbol{\theta}_{j}) - \frac{\partial}{\partial \beta_{0j*}} \log \sum_{j=1}^{J} \phi_{j} \nu_{ij} \right)$$

$$= \sum_{i=1}^{n} \left( \frac{\phi_{j*} \frac{\partial}{\partial \beta_{0j*}} g_{j*}(r_{i}, \mathbf{Z}; \boldsymbol{\theta}_{j})}{g(r_{i}, \mathbf{Z}; \boldsymbol{\theta}, \boldsymbol{\phi})} - \frac{\phi_{j*} \frac{\partial}{\partial \beta_{0j*}} \nu_{ij*}}{\sum_{j=1}^{J} \phi_{j} \nu_{ij}} \right)$$

the first part of which (the derivatives of the detection function) are as in the line transect case. The derivatives of  $\nu_{ij}$  are simpler in the point transect case, since there is an easy analytic expression for  $\nu_{ij}$  when  $g_j$  is half-normal:

$$\nu_{ij} = 2\pi\sigma_{ij}^2(1 - \exp(-w^2/2\sigma_{ij}^2))$$

then simply applying the product rule yields:

$$\frac{\partial \nu_{ij}}{\partial \beta_{0j*}} = 2(\nu_{ij*} + \pi w^2 g_{j*}(w)).$$

Substituting this into the above expression:

$$\frac{\partial l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{r}, \mathbf{Z})}{\partial \beta_{0j*}} = \sum_{i=1}^n \Big( \frac{\phi_{j*}(r_i/\sigma_{j*})^2 g_{j*}(r_i, \mathbf{Z}; \boldsymbol{\theta}_{j*})}{g(r_i, \mathbf{Z}; \boldsymbol{\theta}, \boldsymbol{\phi})} - \frac{\phi_{j*} 2(\nu_{j*} + \pi w g_{j*}(w))}{\sum_{j=1}^J \phi_j \nu_{ij}} \Big)$$

#### With respect to $\beta_{k*}$

Again working from (5), we obtain:

$$\begin{split} \frac{\partial l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{r}, \mathbf{Z})}{\partial \beta_{k*}} &= \sum_{i=1}^{n} \left( \frac{\partial}{\partial \beta_{k*}} \log \sum_{j=1}^{J} \phi_{j} g_{j}(r_{i}, \mathbf{Z}; \boldsymbol{\theta}_{j}) - \frac{\partial}{\partial \beta_{k*}} \log \sum_{j=1}^{J} \phi_{j} \nu_{ij} \right) \\ &= \sum_{i=1}^{n} \left( \frac{\sum_{j=1}^{J} \phi_{j} \frac{\partial}{\partial \beta_{k*}} g_{j}(r_{i}, \mathbf{Z}; \boldsymbol{\theta}_{j})}{g(r_{i}, \mathbf{Z}; \boldsymbol{\theta}, \boldsymbol{\phi})} - \frac{\sum_{j=1}^{J} \phi_{j} \frac{\partial}{\partial \beta_{k*}} \nu_{ij}}{\sum_{j=1}^{J} \phi_{j} \nu_{ij}} \right) \end{split}$$

The derivatives of  $g_i$  are as in (3). For  $\nu_{ij}$ :

$$\frac{\partial \nu_{ij}}{\partial \beta_{k*}} = 2z_{ik*}(\nu_{ij} - \pi w^2 g_j(w))$$

Putting that together:

$$\frac{\partial l(\boldsymbol{\theta}, \boldsymbol{\phi}; \mathbf{r}, \mathbf{Z})}{\partial \beta_{k*}} = \sum_{i=1}^{n} \Big( \frac{\sum_{j=1}^{J} \phi_{j} z_{k*} \left(\frac{x_{i}}{\sigma_{j}}\right)^{2} g_{j}(x_{i}, \mathbf{Z}; \boldsymbol{\theta}_{j})}{g(r_{i}, \mathbf{Z}; \boldsymbol{\theta}, \boldsymbol{\phi})} - \frac{\sum_{j=1}^{J} \phi_{j} 2 z_{ik*} (\nu_{ij} - \pi w^{2} g_{j}(w))}{\sum_{j=1}^{J} \phi_{j} \nu_{ij}} \Big).$$