

# CSCI567 Fall16 Homework 2

Dinh Nguyen  
dinhnguy@usc.edu

Oct 2nd 2016

## 1. Logistic Regression

- (a) Consider a binary logistic regression model, given  $n$  training examples  $(\mathbf{x}_1, y_1), (\mathbf{x}_2, y_2), \dots$ , write down the negative log likelihood (as loss function):

For cleaner equation, I append 1 to  $\mathbf{x}_i$  and  $b$  to  $\mathbf{w}$  and assume possible values for  $y_i$  to be  $\{0, 1\}$ ;

$$\begin{aligned}\mathcal{L}(\mathbf{w}) &= -\log \left( \prod_{i=1}^n P(Y = y_i | \mathbf{X} = \mathbf{x}_i) \right) \\ &= -\log \prod_{i=1}^n \left( \frac{1}{1 + \exp(-\mathbf{w}^\top \mathbf{x}_i)} \right)^{y_i} \left( 1 - \frac{1}{1 + \exp(-\mathbf{w}^\top \mathbf{x}_i)} \right)^{1-y_i} \\ &= \sum_{i=1}^n y_i \log(1 + \exp(-\mathbf{w}^\top \mathbf{x}_i)) - (1 - y_i) \log \left( \frac{\exp(-\mathbf{w}^\top \mathbf{x}_i)}{1 + \exp(-\mathbf{w}^\top \mathbf{x}_i)} \right) \\ &= \sum_{i=1}^n y_i \log(1 + \exp(-\mathbf{w}^\top \mathbf{x}_i)) - (1 - y_i) \log(-\exp(\mathbf{w}^\top \mathbf{x}_i)) \\ &\quad + (1 - y_i) \log(1 + \exp(-\mathbf{w}^\top \mathbf{x}_i)) \\ &= \sum_{i=1}^n \log(1 + \exp(-\mathbf{w}^\top \mathbf{x}_i)) + (1 - y_i) \mathbf{w}^\top \mathbf{x}_i\end{aligned}$$

- (b) Use Gradient Descent Method to find the update rule for  $\mathbf{w}$ .

$$\begin{aligned}\frac{\partial \mathcal{L}}{\partial \mathbf{w}} &= \sum_{i=1}^n \frac{-\exp(-\mathbf{w}^\top \mathbf{x}_i) \mathbf{x}_i}{1 + \exp(-\mathbf{w}^\top \mathbf{x}_i)} + (1 - y_i) \mathbf{x}_i \\ &= \sum_{i=1}^n \left( -\frac{\exp(-\mathbf{w}^\top \mathbf{x}_i)}{1 + \exp(-\mathbf{w}^\top \mathbf{x}_i)} + 1 - y_i \right) \mathbf{x}_i \\ &= \sum_{i=1}^n \left( \frac{1}{1 + \exp(-\mathbf{w}^\top \mathbf{x}_i)} - y_i \right) \mathbf{x}_i\end{aligned}$$

Update rule:

$$\mathbf{w}_{new} = \mathbf{w}_{old} - \eta \sum_{i=1}^n \left( \frac{1}{1 + \exp(-\mathbf{w}_{old}^\top \mathbf{x}_i)} - y_i \right) \mathbf{x}_i$$

for some appropriate  $\eta$ .

This solution will not converge to a global minimum since it will only go toward the closest local minimum and stop there since the gradient at that point would be  $\mathbf{0}$ .

(c) The negative log likelihood  $\mathcal{L}(\mathbf{w}_1, \dots, \mathbf{w}_K)$

$$\begin{aligned}\mathcal{L}(\mathbf{w}_1, \dots, \mathbf{w}_K) &= -\log \left( \prod_{i=1}^n P(Y = y_i | \mathbf{X} = \mathbf{x}_i) \right) \\ &= -\sum_{i=1}^n \log \left\{ \prod_{k=1}^{K-1} \left( \frac{\exp(\mathbf{w}_k^\top \mathbf{x}_i)}{1 + \sum_{t=1}^{K-1} \exp(\mathbf{w}_t^\top \mathbf{x}_i)} \right)^{I_{(y_i=k)}} \right. \\ &\quad \left. \times \left( \frac{1}{1 + \sum_{t=1}^{K-1} \exp(\mathbf{w}_t^\top \mathbf{x}_i)} \right)^{I_{(y_i=K)}} \right\}\end{aligned}$$

By letting  $\mathbf{w}_K = \mathbf{0}$ , we have  $\exp(\mathbf{w}_K^\top \mathbf{x}_i) = 1$  so we can combine the  $K$  term into to product as well as the sum in the denominator:

$$\begin{aligned}\mathcal{L}(\mathbf{w}_1, \dots, \mathbf{w}_K) &= -\sum_{i=1}^n \log \left\{ \prod_{k=1}^K \left( \frac{\exp(\mathbf{w}_k^\top \mathbf{x}_i)}{\sum_{t=1}^K \exp(\mathbf{w}_t^\top \mathbf{x}_i)} \right)^{I_{(y_i=k)}} \right\} \\ &= -\sum_{i=1}^n \sum_{k=1}^K I_{(y_i=k)} \log \left( \frac{\exp(\mathbf{w}_k^\top \mathbf{x}_i)}{\sum_{t=1}^K \exp(\mathbf{w}_t^\top \mathbf{x}_i)} \right) \\ &= -\sum_{i=1}^n \sum_{k=1}^K I_{(y_i=k)} \left( \log(\exp(\mathbf{w}_k^\top \mathbf{x}_i)) - \log\left(\sum_{t=1}^K \exp(\mathbf{w}_t^\top \mathbf{x}_i)\right) \right) \\ &= -\sum_{i=1}^n \sum_{k=1}^K I_{(y_i=k)} \left( \mathbf{w}_k^\top \mathbf{x}_i - \log\left(\sum_{t=1}^K \exp(\mathbf{w}_t^\top \mathbf{x}_i)\right) \right)\end{aligned}$$

(d) Compute the gradient of the negative log likelihood:

$$\begin{aligned}\frac{\partial \mathcal{L}(\mathbf{w}_1, \dots, \mathbf{w}_K)}{\partial \mathbf{w}_k} &= -\sum_{i=1}^n I_{(y_i=k)} \left( \mathbf{x}_i - \frac{\exp(\mathbf{w}_k^\top \mathbf{x}_i) \mathbf{x}_i}{\sum_{t=1}^K \exp(\mathbf{w}_t^\top \mathbf{x}_i)} \right) \\ &= -\sum_{i=1}^n I_{(y_i=k)} \left( 1 - \frac{\exp(\mathbf{w}_k^\top \mathbf{x}_i)}{\sum_{t=1}^K \exp(\mathbf{w}_t^\top \mathbf{x}_i)} \right) \mathbf{x}_i\end{aligned}$$

Update rule:

$$\mathbf{w}_k^{new} = \mathbf{w}_k^{old} + \eta \sum_{i=1}^n I_{(y_i=k)} \left( 1 - \frac{\exp(\mathbf{w}_k^{(old)\top} \mathbf{x}_i)}{\sum_{t=1}^K \exp(\mathbf{w}_t^{(old)\top} \mathbf{x}_i)} \right) \mathbf{x}_i$$

## 2. Linear/ Gaussian Discriminant

(a) Write the log likelihood function  $\mathcal{L}(\mathcal{D})$

Since  $y_n \in \{1, 2\}$ ,

$$\begin{aligned}\mathcal{L}(\mathcal{D}) = \sum_{n=1}^N I(y_n = 1) & \left( \log(p_1) - (1/2) \log(2\pi\sigma_1^2) - \frac{(x_n - \mu_1)^2}{2\sigma_1^2} \right) \\ & + I(y_n = 2) \left( \log(p_2) - (1/2) \log(2\pi\sigma_2^2) - \frac{(x_n - \mu_2)^2}{2\sigma_2^2} \right)\end{aligned}$$

Use MLE to find  $(p_1^*, p_2^*, \mu_1^*, \mu_2^*, \sigma_1^*, \sigma_2^*)$

Since  $p_1 + p_2 = 1$ , we can rewrite  $p_2$  as  $1 - p_1$  in the likelihood function before taking the derivative with respect to  $p_1$ :

$$\frac{\partial L}{\partial p_1} = \sum_{n=1}^N I(y_n = 1)/p_1 - I(y_n = 2)/(1 - p_1) = 0$$

Solve for  $p_1$ : we have

$$p_1^* = \frac{\sum_{n=1}^N I(y_n = 1)}{N}$$

since  $\sum_{n=1}^N I(y_n = 1) + I(y_n = 2) = N$ . And  $p_2^* = 1 - p_1^* = 1 - \frac{\sum_{n=1}^N I(y_n=1)}{N} = \frac{\sum_{n=1}^N I(y_n=2)}{N}$

For  $\mu_i^*$ , we have:

$$\frac{\partial L}{\partial \mu_i} = \sum_{n=1}^N I(y_n = i) \frac{x_n - \mu_i}{\sigma_i^2} = 0$$

Solve for  $\mu_i^*$  we have

$$\mu_i^* = \frac{\sum_{n=1}^N I(y_n = i) x_n}{\sum_{n=1}^N I(y_n = i)}$$

i.e we take the average of all the  $x_n$  where  $y_n = i$ .

Lastly, for  $\sigma_i^*$ :

$$\frac{\partial L}{\partial \sigma_i} = \sum_{n=1}^N I(y_n = i) \frac{(x_n - \mu_i)^2 - \sigma_i^2}{\sigma_i^2} = 0$$

Solve for  $\sigma_i$ :

$$\sigma_i^* = \sqrt{\frac{\sum_{n=1}^N I(y_n = i) (x_n - \mu_i^*)^2}{\sum_{n=1}^N I(y_n = i)}}$$

(b) Since the samples are iid, we know the covariance matrix for both classes is diagonal with  $\sigma^2$  entries. For the mean vector,  $\mu_1 = \mathbf{0}$  and  $\mu_2 = (0, \dots, 0, \delta, \dots, \delta)$ .

Plug those in the formula of  $P(y = c_1|x, \mu, \Sigma)$  we have from last hw (problem 2) and set it equal to 1/2. We have

$$P(y = c_1|\mathbf{x}, \mu, \Sigma) = \frac{1}{1 + p_2/p_1 \exp((\sum_{n=D}^{2D} \delta \sigma^2 x_n) - D\delta^2 \sigma^2/2)} = \frac{1}{2}$$

So

$$p_1 = p_2 \exp((\sum_{n=D}^{2D} \delta \sigma^2 x_n) - D\delta^2 \sigma^2/2)$$

Take log of both sides and we have the linear discriminant:

$$\sum_{n=D}^{2D} \delta \sigma^2 x_n = \log(p_1) - \log(p_2) + D\delta^2 \sigma^2/2$$

which clearly depends on  $\delta$  so it changes when  $\delta$  changes.

(c) Let  $P(y = c_i) = p_i$ :

$$\begin{aligned} P(y = 1|\mathbf{x}) &= \frac{p_1 P(\mathbf{x}|y = 1)}{p_1 P(\mathbf{x}|y = 1) + p_2 P(\mathbf{x}|y = 2)} \\ &= \frac{1}{1 + \frac{p_2 P(\mathbf{x}|y=2)}{p_1 P(\mathbf{x}|y=1)}} \end{aligned}$$

Since  $p_i P(\mathbf{x}|y = c_i)$  follows a multivariate Gaussian distribution with the same covariance matrix, we can cancel out the  $(2\pi)^{D/2} |\Sigma|^{-1/2}$ . So we have:

$$\begin{aligned} \frac{P(\mathbf{x}|y = 2)}{P(\mathbf{x}|y = 1)} &= \frac{\exp(-1/2(\mathbf{x} - \mu_2)^\top \Sigma^{-1}(\mathbf{x} - \mu_2))}{\exp(-1/2(\mathbf{x} - \mu_1)^\top \Sigma^{-1}(\mathbf{x} - \mu_1))} \\ &= \exp(-1/2(\mu_2^\top \Sigma^{-1} \mu_2 - \mu_1^\top \Sigma^{-1} \mu_1) - (\Sigma^{-1}(\mu_1 - \mu_2))^\top \mathbf{x}) \end{aligned}$$

So by setting  $\theta = (\log(p_1) - \log(p_2) + 1/2(\mu_2^\top \Sigma^{-1} \mu_2 - \mu_1^\top \Sigma^{-1} \mu_1), \Sigma^{-1}(\mu_1 - \mu_2))$  and append 1 to  $\mathbf{x}$  we have the form of a logistic function.

### 3. Perceptron and Online Learning

By setting  $\mathbf{w}_{i+1} = \mathbf{w}_i + (y_i - \mathbf{w}_i^\top \mathbf{x}_i) \frac{\mathbf{x}_i}{|\mathbf{x}_i|^2}$ . Since we used  $\mathbf{w}_i^\top \mathbf{x}_i$  to evaluate  $y_i$  and it didn't work (otherwise we would not update), by adding the difference to  $w_i$  with a normalized  $\mathbf{x}_i$ , we guarantee to have get  $y_i$  when using  $w_{i+1}$ :

$$\mathbf{w}_{i+1}^\top \mathbf{x}_i = \mathbf{w}_i^\top \mathbf{x}_i + y_i - \mathbf{w}_i^\top \mathbf{x}_i \frac{\mathbf{x}_i^\top \mathbf{x}_i}{|\mathbf{x}_i|^2} = y_i$$

since  $\mathbf{x}_i^\top \mathbf{x}_i = |\mathbf{x}_i|^2$ . Since we add the exact distance to  $\mathbf{w}_i$ ,  $\mathbf{w}_{i+1}$  is the closest vector to  $\mathbf{w}_i$  that classifies correctly.

### 4. Programming All the answers are in the output of the scripts. Thanks.