

PDC Project Theory

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(a) (i)

$$\begin{aligned} Y &= e^{j\Theta}x + Z \\ &= e^{j\Theta}(x + Z'), Z' = e^{-j\Theta}Z \end{aligned}$$

From the first exercise set of the course, we know that $\arg(Z) \sim U[0, 2\pi[$. And with :

$$\left. \begin{array}{l} |Z'| = |Z| \\ \arg(Z') = \arg(Z) - \Theta \mod (2\pi) \sim U[0, 2\pi[\end{array} \right\} \implies Z' \text{ does not depend on } \Theta \implies Z' \perp \Theta$$

We want to show that $|Y| \perp \arg(Y)$. We have :

$$\begin{aligned} |Y| &= |x + Z'| \\ \arg(Y) &= \Theta + \arg(x + Z') \mod (2\pi) \sim U[0, 2\pi[\end{aligned}$$

We know that $\Theta \sim U[0, 2\pi[$. Therefore, independently of the distribution of $\arg(x + Z')$, $\arg(Y) \mod (2\pi)$ will also be uniformly distributed over $[0, 2\pi[$. We observe that $|Y|$ does not depend on $\arg(Y)$ nor Θ , thus $|Y| \perp \arg(Y)$

- (ii) We also proved on the previous point that $\arg(Y) \sim U[0, 2\pi[$
(iii)

$$Y = \begin{cases} Y_1 = |x|\cos(\Theta + \arg(x)) + \Re(Z) \\ Y_2 = |x|\sin(\Theta + \arg(x)) + \Im(Z) \end{cases}$$

We can create a dummy random variable :

$$Z' = \begin{cases} Z'_1 = \Re(Z) - |x|\cos(\Theta + \arg(x)) \\ Z'_2 = \Im(Z) - |x|\sin(\Theta + \arg(x)) \end{cases}$$

$$Z' \sim \mathcal{N}((-|x|\cos(\Theta + \arg(x)), -|x|\sin(\Theta + \arg(x)))^T, \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix})$$

$|Y| < \rho \leftrightarrow Z$ brings $e^{j\Theta}x$ inside the circle centred in $(0,0)$ of radius ρ .

We observe that it is equivalent to $|Z'| < \rho$

$$Pr(U < \rho) = \int_{Z'_1} \int_{Z'_2} f_{Z'} dz'_1 dz'_2$$

We compute the gaussian bivariate :

$$\begin{aligned} P(U < \rho) &= \int_{Z'_1} \int_{Z'_2} f_{Z'} dz'_1 dz'_2 \\ &= \int_{Z'_1} \int_{Z'_2} \frac{1}{2\pi} \exp\left(-\frac{1}{2}(\mathbf{z} - \boldsymbol{\mu})^T(\mathbf{z} - \boldsymbol{\mu})\right) dz'_1 dz'_2 \\ (\mathbf{z} - \boldsymbol{\mu})^T(\mathbf{z} - \boldsymbol{\mu}) &= (-|x|\cos(\arg(x) + \Theta) + |z|\cos(\arg(z)))^2 + (-|x|\sin(\arg(x) + \Theta) + |z|\sin(\arg(z)))^2 \\ &= |x|^2 + |z|^2 - 2|x||z|\cos(\arg(x) + \Theta)\cos(\arg(z)) - 2|x||z|\sin(\arg(x) + \Theta)\sin(\arg(z)) \\ &= |x|^2 + |z|^2 - 2|x||z|\cos(\arg(x) + \Theta - \arg(z)) \end{aligned}$$

Because we will integrate over $[0, 2\pi[$ for $\arg(z)$ and \cos is even, 2π periodic, $\arg(x) + \Theta$ can be omitted and $-\arg(z)$ replaced by $\arg(z)$.

We use the following change of variables :

$$\begin{aligned} |Z| &= r \\ \theta &= \arg(z) \end{aligned}$$

Thus we have $|\mathbb{J}| = r$. And therefore :

$$Pr(U < \rho) = (2\pi)^{-1} \int_0^{2\pi} \int_0^\rho \exp\left(-\frac{1}{2}[r^2 + |x|^2 - 2r|x|\cos(\theta)]\right) r dr d\theta$$

Finally we can conclude that for any $u \geq 0$:

$$\begin{aligned}
f_{U|X}(u|x) &= \frac{d}{du} \Pr(U < u | X = x) \\
&= \frac{1}{2\pi} \int_0^{2\pi} \frac{d}{du} \left[\int_0^u \exp\left(-\frac{1}{2}[r^2 + |x|^2 - 2r|x|\cos(\Theta)]\right) r dr \right] d\Theta \\
&= \frac{1}{2\pi} \int_0^{2\pi} \exp\left(-\frac{1}{2}[u^2 + |x|^2 - 2u|x|\cos(\Theta)]\right) u d\Theta \\
&= \frac{1}{2\pi} \int_0^{2\pi} \exp\left(-\frac{u^2 + |x|^2}{2}\right) \exp(-2u|x|\cos(\Theta)) u d\Theta \\
&= u \exp\left(-\frac{u^2 + |x|^2}{2}\right) \frac{1}{2\pi} \int_0^{2\pi} \exp(-2u|x|\cos(\Theta)) d\Theta \\
f_{U|X}(u|x) &= u \exp\left(-\frac{u^2 + |x|^2}{2}\right) I_0(u|x|)
\end{aligned}$$

- (b) To show that $U = T(Y) = |Y|$ is a sufficient statistic, we will show that

$$H \rightarrow T(Y) \rightarrow Y \iff P(Y|T(Y), H) = P(Y|T(Y))$$

or in other word, we want to show that the probability of Y knowing $T(Y) = |Y|$ is independent of H .

Using the formula $Y = |Y|e^{j\arg(Y)}$, we see than we can show the above equality by showing that H is independent from $\arg(Y)$.

Indeed, if we know $|Y|$, the only part we don't know about Y is $e^{j\arg Y}$.

Thus, to show that $P(Y|T(Y), H) = P(Y|T(Y))$ we simply need to show that H is independent from $e^{j\arg(Y)}$ or equivalently, to show that H is independent from $\arg(Y)$.

Finally, this independence is a direct result of part ii) of exercise a) where we showed that, knowing x (equivalent of knowing H), the argument of Y ($\arg(Y)$) is uniformly distributed on $[0, 2\pi[$.

Thus, $U = |Y|$ is a sufficient statistic.

- (c) The output of the channel Y will have the same statistic for $|c_j| = |c_l|$, for $j \neq l$ because the multiplication by $\exp(j\Theta)$ shifts the phase by Θ . Θ is uniformly distributed between 0 and 2π . Hence, we lose any information about $\arg(X)$, only the magnitude matters. It is impossible to distinguish 2 hypotheses that have the same magnitude.
- (d) To show that replacing c_j by $|c_j|$ has no impact on the performance of an ML decoder, we will show that the distribution of Y is not affected by this modification. Let also note that replacing c_j by $|c_j|$ is equivalent to setting $\arg c_j$ to 0.

Let $W = \theta + \arg c_j \mod 2\pi$. We can then rewrite Y as follows

$$Y = |c_j|e^{j(\theta + \arg c_j)} + Z = |c_j|e^{jW} + Z$$

We can clearly see that the only modified random variable in this new system is W as only $\arg c_j$ is modified. Thus, in order to show that the distribution of Y does not change in this new system, we can simply show that the distribution of W does not change.

Since $\theta \sim \mathcal{U}(0, 2\pi)$, we also have that $W \sim \mathcal{U}(0, 2\pi)$. This result hold because W is uniform on the unit circle in the complex plane. Thus, we can use the 2π modulation as shown above without modifying the distribution. Finally, in the case of $\arg c_j = 0$, we have $W = \theta \bmod 2\pi$ which is again uniformly distributed between 0 and 2π .

To conclude, we have shown that replacing c_j by $|c_j|$ does not affect the distribution of Y , thus an ML decoder on this new system will achieve the same performance as on the old system.

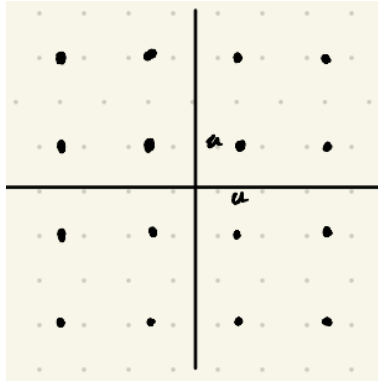
- (e) $U = |Y|$ is a sufficient statistic. Hence, the error probability when observing U is the same as when observing Y . Here we have $||x| - u|$ is small compared to x because the only difference between $|x|$ and u is due to Z and we know that $a \gg 1$.

Our system is then similar to a m-PAM with an AWGN channel. So we have that $P_e = (2 - \frac{2}{2^k}) Q(\frac{d}{2\sigma}) =$ with $d = 2a$ and $\sigma = 1$ we therefore have $(2 - \frac{2}{2^k}) Q(\frac{d}{2\sigma}) \approx 2Q(a)$ for k big enough.

- (f)

$$\begin{aligned} \mathcal{E} &= \sum_{i=0}^{m-1} P_H(i) ||c_i||^2 = \frac{1}{m} \sum_{i=0}^{m-1} ((2i+1)a)^2 \\ &= \frac{1}{m} a^2 \sum_{i=0}^{m-1} (2i+1)^2 = \frac{a^2}{m} \frac{m(4m^2-1)}{3} \\ &= \frac{a^2(4 \cdot 2^{2k} - 1)}{3} = \frac{a^2(4^{k+1} - 1)}{3} \\ \mathcal{E}_b &= \frac{\mathcal{E}}{k} = \frac{a^2(4^{k+1} - 1)}{3k} \approx \frac{a^2(4^{k+1})}{3k} \end{aligned}$$

- (g) With the knowledge of Θ , we can use the QAM constellation. The c_j 's will have the form $\pm(2p-1)a \pm i(2q-1)a$ with p, q in $1, \dots, 2^{k-1}$. This gives us 2^{2k} different c_j (enough to transmit $2k$ bits).



Now we show that the energy per bit is $\frac{a^2(4^{k+1})}{3k}$:

By symmetry we only consider the first quadrant and multiply by 4:

$$\begin{aligned}
\mathcal{E} &= 4 \sum_i P_H(i) \|c_i\|^2 = 4 \frac{1}{2^{2k}} \sum_{p=1}^{2^{k-1}} \sum_{q=1}^{2^{k-1}} ((2p+1)a)^2 + ((2q+1)a)^2 \\
&= \frac{a^2}{2^{2k-2}} \sum_{p=1}^{2^{k-1}} 2^{k-1} (2p-1)^2 + \frac{2^{k-1} (4(2^{k-1})^2 - 1)}{3} \\
&= \frac{a^2}{2^{2k-2}} 2 \frac{(2^{k-1})^2 (4 \cdot 2^{2k-2} - 1)}{3} = \frac{2a^2 \cdot 2^{2k-2} (2^{2k} - 1)}{2^{k-2} 3} = \frac{2a^2 (4^k - 1)}{3} \\
\mathcal{E}_b &= \frac{\mathcal{E}}{2k} = \frac{2a^2 (4^k - 1)}{3 \cdot 2k} = \frac{a^2 (4^k - 1)}{3k}
\end{aligned}$$

We still need to show that the error probability is less than $4Q(a)$. Recall : the real and imaginary parts of Z are i.i.d $\mathcal{N}(0, 1)$

P_c = probability of being correct.

For the four corner parts, we have $P_c(\text{corner}) = \Pr\{(Z_R \leq a)(Z_C \geq -a)\}$
 $= \Pr\{(Z_R \leq a)\} \Pr\{(Z_C \geq -a)\} = (1 - Q(a))Q(-a) = (1 - Q(a))^2$

For the points on the edges, we find:

$P_c(\text{edge}) = \Pr\{-a \leq Z_R \leq a\} \Pr\{(Z_C \geq -a)\} = (1 - 2Q(a))(1 - Q(-a))$

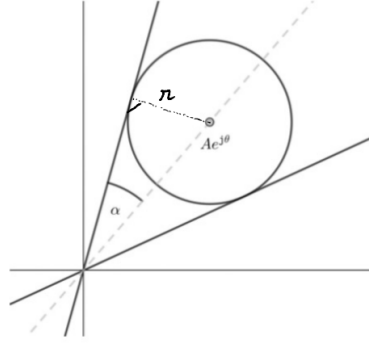
For all the points in the middle :

$P_c(\text{middle}) = \Pr\{-a \leq Z_R \leq a\} \Pr\{-a \leq Z_C \leq a\} = (1 - 2Q(a))^2$

Overall,

$$\begin{aligned}
P_C &= \sum_{i=1}^{2^{2k}} P_H(i) P_C(i) \\
&= \frac{1}{2^{2k}} [4(1 - Q(a))^2 + 4(2^k - 2)(1 - Q(a))(1 - 2Q(a)) + (2^k - 2)^2(1 - 2Q(a))^2] \\
&= \frac{1}{2^{2k}} [4(Q(a)^2 - 2Q(a) + 1) + 4(2^k - 2)(1 - 3Q(a) + 2Q(a)^2) \\
&\quad + (2^{2k} - 2^{k+2} + 4)^2(1 - 4Q(a) + 4Q(a)^2)] \\
&= 1 - \frac{8 + 12(2^k - 2) + 4(2^{2k} - 2^{k+2} + 4)}{2^{2k}} Q(a) + \frac{4 + 8(2^k - 2) + 4(2^{2k} - 2^{k+2} + 4)}{2^{2k}} Q(a)^2 \\
\implies P_e &= \frac{4 \cdot 2^{2k} - 4 \cdot 2^k}{2^{2k}} Q(a) - \frac{4 \cdot 2^{2k} - 8 \cdot 2^k + 4}{2^{2k}} Q(a)^2 \\
&= \left(4 - \frac{1}{2^{k-2}}\right) Q(a) - \left(4 - \frac{1}{2^{k-3}} + \frac{4}{2^{2k}}\right) Q(a)^2 \\
&\leq 4Q(a)
\end{aligned}$$

- (h) The method to estimate Θ is simply to take $\arg(Y_1)$ as an estimate : $\hat{\Theta} = \arg(Y_1)$.
The probability that Y_1 falls outside of a circle of radius r centered around $Ae^{j\Theta}$ is $\exp(-r^2/2)$ as Z_1 is a 0 mean complex Gaussian (the real and imaginary parts are i.d.d $\mathcal{N}(0, 1)$). If we want that the circle to inside the sector that contains all values that have an argument Θ' s.t. $|\Theta' - \Theta| < \alpha$, then $r = A \sin(\alpha)$.
So we have that
 $\Pr(|\hat{\Theta} - \Theta| > \alpha)$ (i.e. the probability that Y_1 is outside of our sector) $\leq \exp(-A^2 \sin^2(\alpha)/2)$
(i.e. the probability that Y_1 is outside of our circle)



- (i) We compute $Y = (Y_1 + Y_2 + \dots + Y_{n_0})/\sqrt{n_0}$ and we estimate Θ with the phase of Y : $\hat{\Theta} = \arg(Y)$. It is the same situation as in (h) except this time our circle is centered around $\frac{n_0 A}{\sqrt{n_0}} e^{j\Theta} = \sqrt{n_0} A \cdot e^{j\Theta}$.
So we have that
 $\Pr(|\hat{\Theta} - \Theta| > \alpha)$ (i.e. the probability that Y_1 is outside of our sector) $\leq \exp(-n_0 A^2 \sin^2(\alpha)/2)$
(i.e. the probability that Y_1 is outside of our circle)

