
Mobile Communications

Performance Evaluation

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Techniques Available for Performance Evaluation

- ♦ Analytical Stochastic Models
- ♦ Deterministic analysis
- ♦ Bottleneck analysis
- ♦ Simulation

Techniques Available for Performance Evaluation

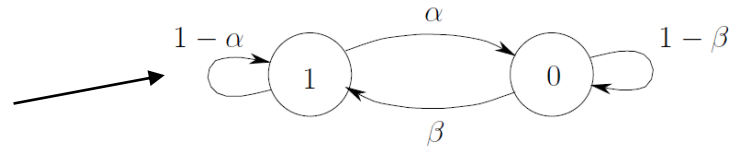
- ♦ **Analytical Stochastic Models**
- ♦ Deterministic analysis
- ♦ Bottleneck analysis
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Analytical Models

- ◆ Stochastic process $X(\omega, t)$, Poisson process $P_n(t) = \frac{(\lambda t)^n}{n!} e^{-\lambda t} \equiv P[N(t) = n]$

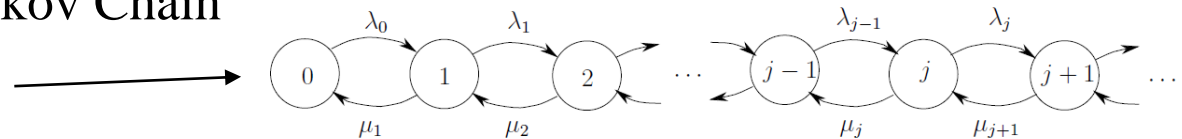
- ◆ Markov Chains

- » Discrete-Time Markov Chain



- » Continuous-Time Markov Chain

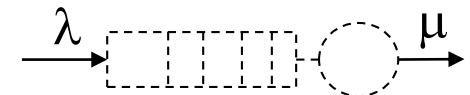
- » Birth-Death Chains



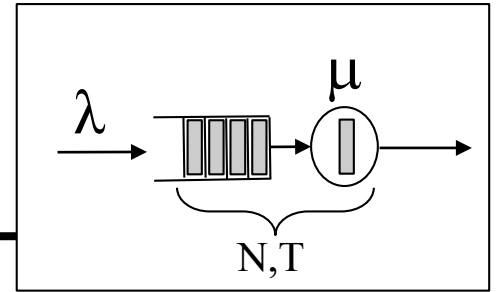
- ◆ Queueing Theory

- » M/G/1, M/M/1, M/M/1/K, M/D/1, M/M/s, M/G/1/PS

- » Open Queueing, Closed Queueing Networks



M/M/1 Queue



- ♦ Average Queue size N

$$N = \sum_{n=0}^{\infty} nP(n) = \sum_{n=0}^{\infty} n\rho^n(1-\rho) = \frac{\rho}{1-\rho} \quad N = \sum_{n=0}^{\infty} nP(n) = \frac{\rho}{1-\rho} = \frac{\lambda/\mu}{1-\lambda/\mu} = \frac{\lambda}{\mu-\lambda}$$

- ♦ Average amount of time the client spends in the system, T

» Little's formula, $T=N/\lambda \quad \Rightarrow \quad T = \frac{1}{\mu-\lambda}$

- ♦ Average waiting time $T_w \Rightarrow T_w = T - T_s = \frac{1}{\mu-\lambda} - \frac{1}{\mu} = \frac{\rho}{\mu(1-\rho)}$

- ♦ Average number of clients waiting in the queue, N_w

$$N_w = T_w \lambda = \frac{\lambda}{\mu-\lambda} - \frac{\lambda}{\mu} = N - \rho$$

M/M/1/B Queue

- ♦ M/M/1 queue has limited capacity (B buffers)
 - » Packets can be lost
 - » Probability of packet being lost = $P(B)$ → Queue is full
- ♦ Analysis similar to M/M/1

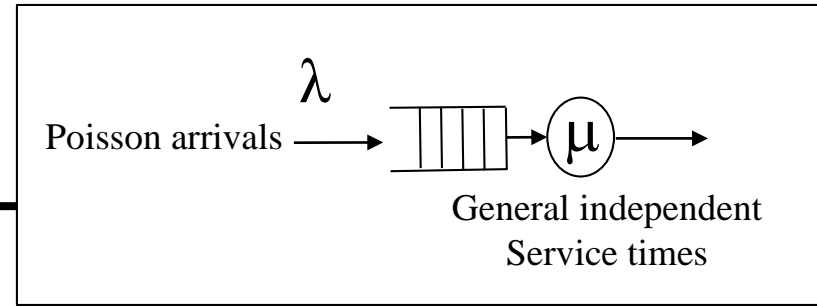
$$\sum_{i=0}^B P(i) = 1 \qquad P(n) = \rho^n P(0)$$

$$P(0) = \frac{1 - \rho}{1 - \rho^{B+1}} \qquad P(B) = \frac{(1 - \rho)\rho^B}{1 - \rho^{B+1}}$$

- ♦ Particular cases

$$\rho = 1, \quad P(B) = \frac{1}{B+1} \qquad \rho \gg 1, \quad P(B) \approx \frac{\rho - 1}{\rho} = \frac{\lambda - \mu}{\lambda}$$

M/G/1 Queue



- ♦ Poisson arrivals at rate λ
- ♦ Service time X has arbitrary distribution with given $E[X]$ and $E[X^2]$
 - » Service times Independent and Identically Distributed (IID)
 - » Independent of arrival times
 - » $E[\text{service time}] = E[X] = 1/\mu$
 - » Single Server queue

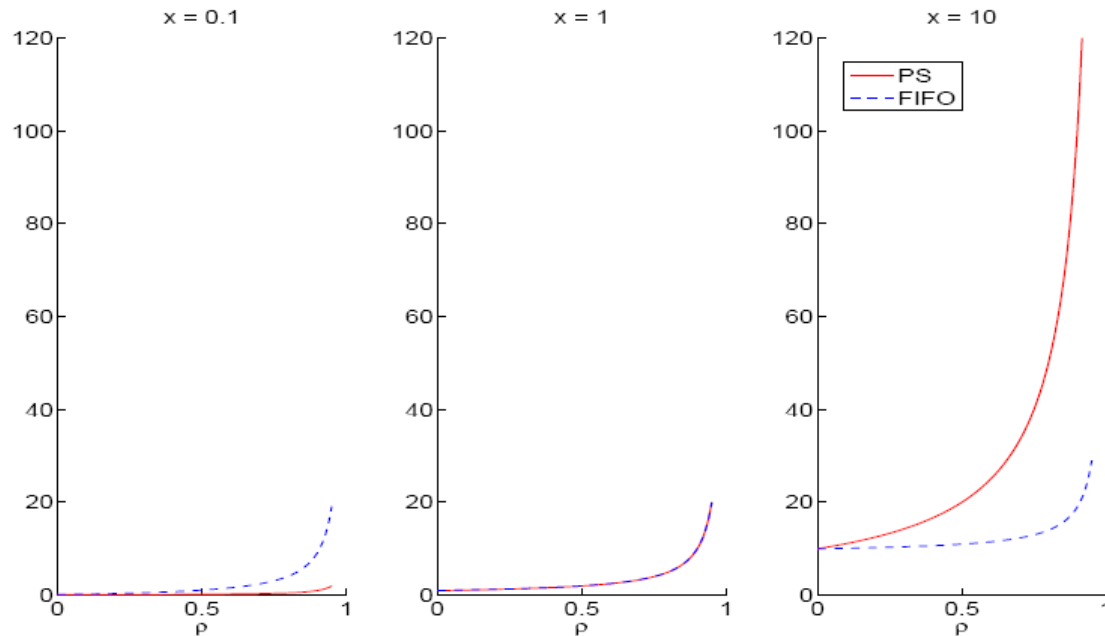
$$T_w = \frac{\lambda E[X^2]}{2(1 - \rho)}$$

- ♦ where $\rho = \lambda/\mu = \lambda E[X]$ = line utilization
- ♦ From Little's Theorem
 - » $N_w = \lambda T_w$
 - » $T = T_w + E[X] = T_w + 1/\mu$
 - » **$N = \lambda T = \lambda(T_w + 1/\mu) = N_w + \rho$**

The Processor Sharing Queue M/G/1/PS

- ♦ Non-FIFO queue
- ♦ Server is equally shared between all customers present

$$E(\text{response time} \mid \text{service time} = x) = \frac{x}{1 - \rho}$$



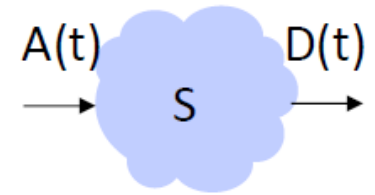
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Deterministic Analysis - Cumulative Functions

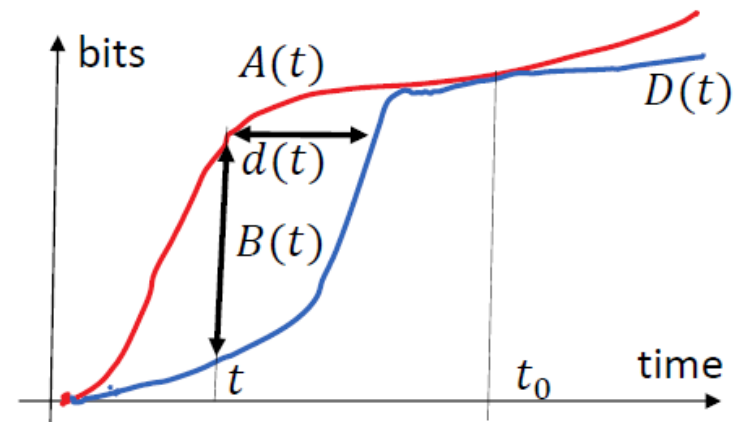
◆ Cumulative functions

- » $A(t)$: amount of bits input to system S in $[0, t]$
- » $D(t)$: amount of bits output from system S in $[0, t]$



◆ Assuming no loss

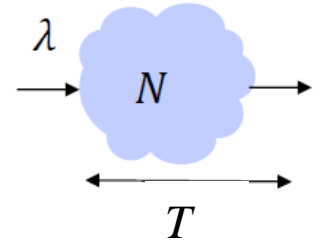
- » $B(t)$ = buffer content at time t
- » $d(t)$ = virtual delay



Techniques Available for Performance Evaluation

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Little's Law



- ♦ Consider a discrete system
 - » λ = customer arrival rate [customer/s]
 - » T = average time spent in system by an arbitrary customer [s]
 - » N = average number of customers in system [customers]

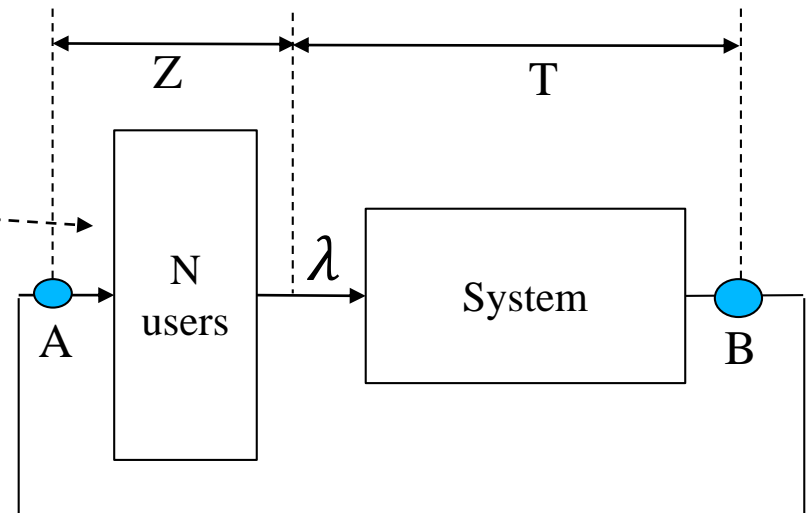
- ♦ Little's law: $\lambda T = N$
 - » Applies to average values

- ♦ Note: customer = packet

Interactive User Model - General

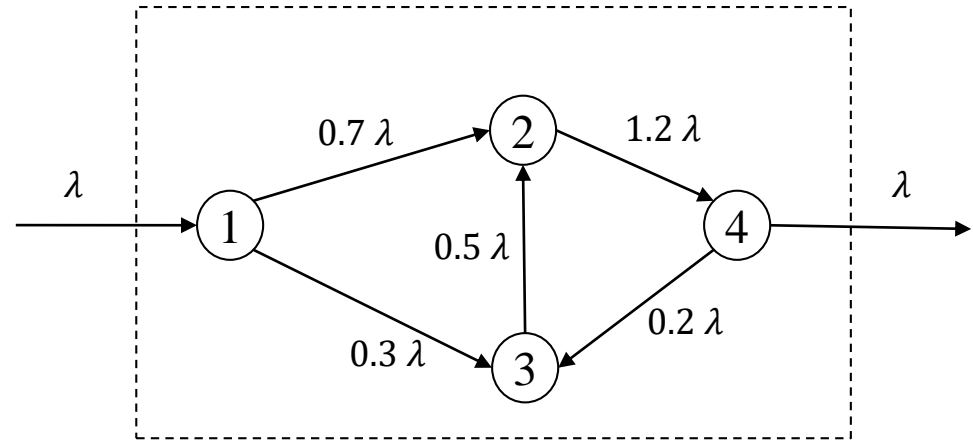
- ◆ N users alternate between think time and visit to the system

- ◆ Little's law applied to $A \rightarrow B$
 - » $\lambda(\bar{Z} + \bar{T}) = N$



- » Suppose we have N packets to continuously transmit through the system
- » *Think time* forces a packet to wait for Z sec before reintroduced in the system
- » Z is a random delay and $\bar{Z} = E[Z]$
- » N is kept constant

Network Laws



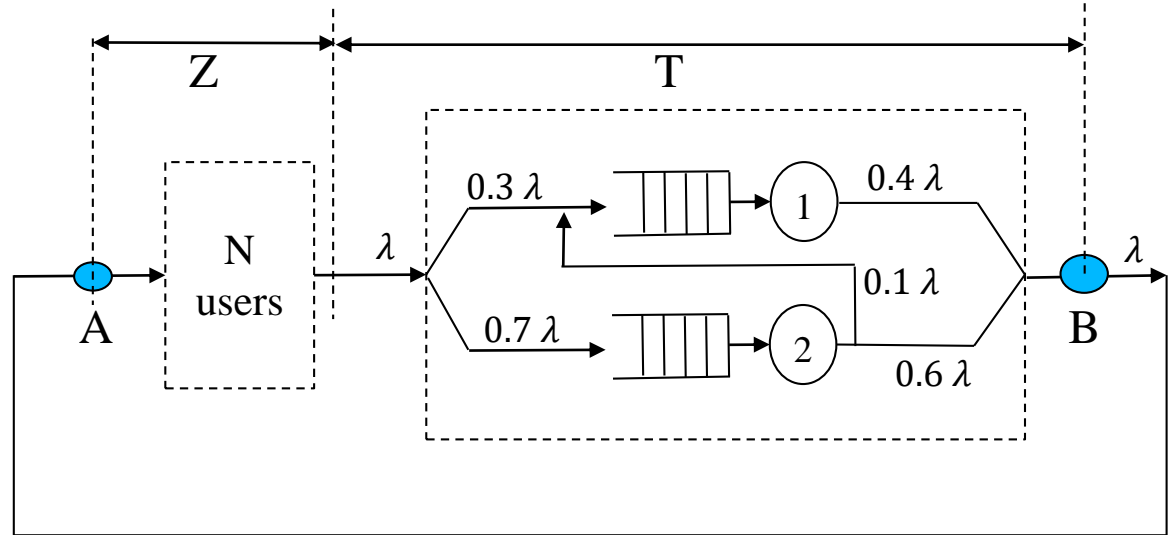
- ♦ $\lambda_k = \lambda V_k$
 - » λ_k : expected number of customer arriving per second to node k
 - » V_k : expected number of visits to node k by an arbitrary customer
 - » $\lambda_1 = \lambda \rightarrow V_1 = 1 \quad | \quad \lambda_2 = 1.2 \lambda \rightarrow V_2 = 1.2$
 - » $\lambda_3 = 0.5 \lambda \rightarrow V_3 = 0.5 \quad | \quad \lambda_4 = 1.2 \lambda \rightarrow V_4 = 1.2$

- ♦ $\bar{T} = \sum \bar{T}_k V_k$
 - » \bar{T} : expected total response time seen by a customer
 - » \bar{T}_k : expected response time seen by a customer which visits node k
 - » $\bar{T} = \bar{T}_1 V_1 + \bar{T}_2 V_2 + \bar{T}_3 V_3 + \bar{T}_4 V_4 = \bar{T}_1 + 1.2 \bar{T}_2 + 0.5 \bar{T}_3 + 1.2 \bar{T}_4$

Bottleneck Analysis

- ◆ Powerful approach to analyze complex queuing systems
- ◆ Based on Little's Formula
- ◆ Considers two (extreme) scenarios:
 - » **Very low utilization:** no waiting time,
 $T_w=0 \rightarrow T = T_w + T_s = T_s$
 - » **Very high utilization:** $\rho = \frac{\lambda}{\mu} = \lambda \cdot T_s \sim 1$ (for a single server)

Bottleneck Analysis



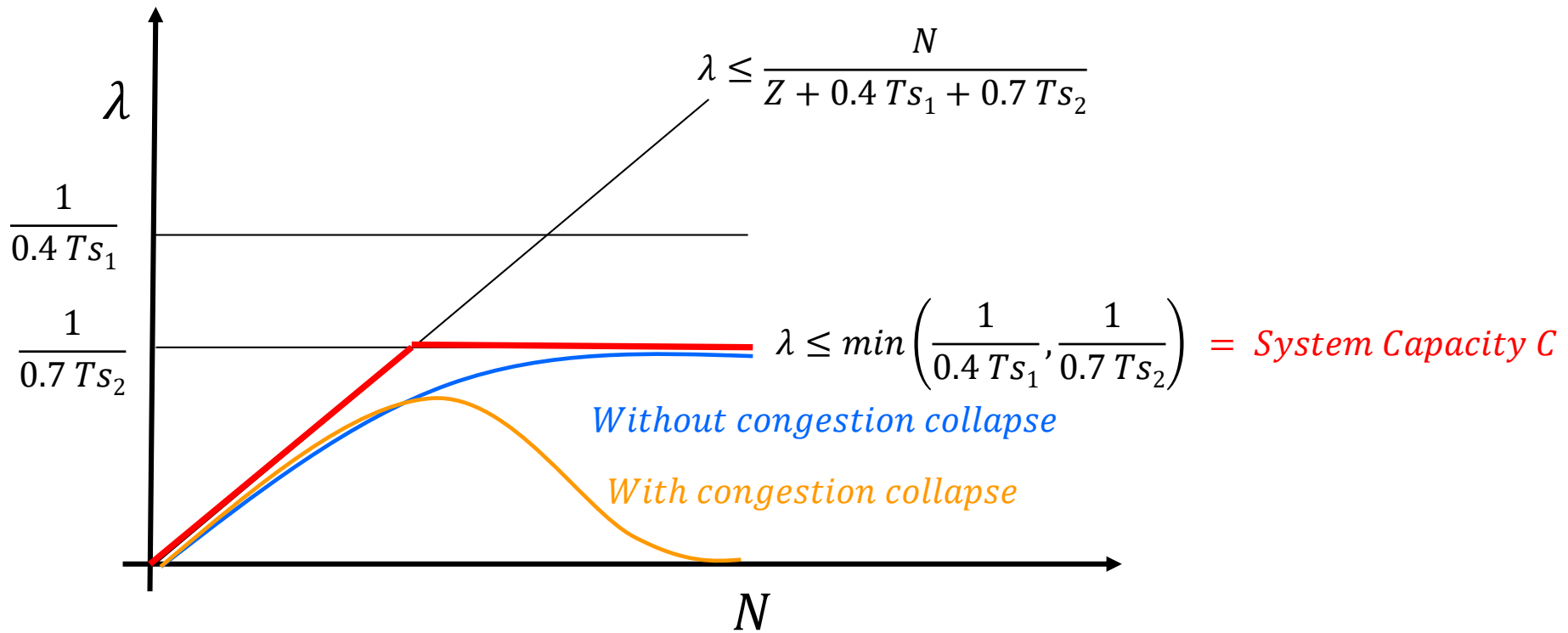
♦ $\lambda(Z + T) = \lambda(Z + 0.4 T_1 + 0.7 T_2) = N$

♦ **Very low utilization**, $T w_i = 0 \rightarrow \lambda \leq \frac{N}{Z + 0.4 T s_1 + 0.7 T s_2}$

♦ **Very high utilization**, $\lambda T_s \leq 1$,

$$0.4 \lambda T s_1 \leq 1 \text{ and } 0.7 \lambda T s_2 \leq 1 \rightarrow \lambda \leq \min \left(\frac{1}{0.4 \cdot T s_1}, \frac{1}{0.7 \cdot T s_2} \right)$$

Bottleneck Analysis



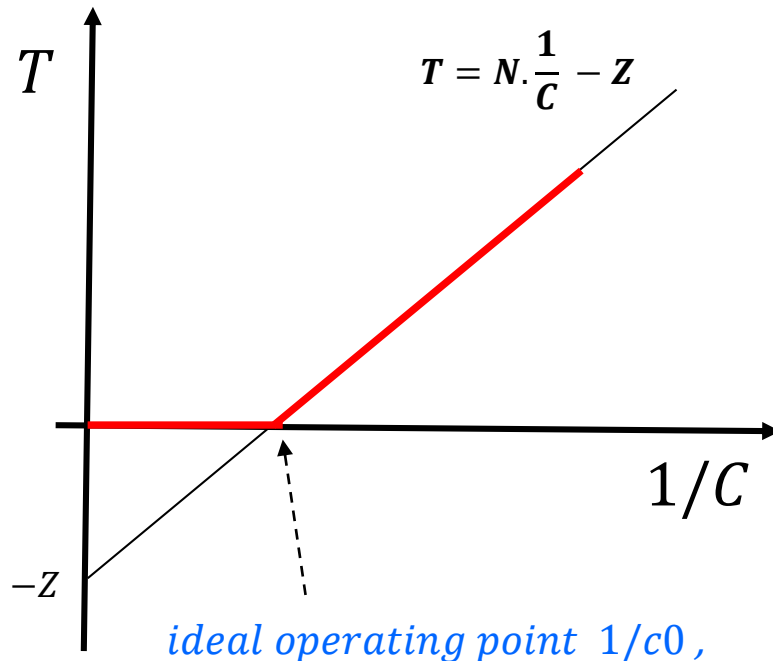
- ◆ Real curve is either blue or orange

Delay vs Capacity

- ◆ Let us assume the system is operating near the system capacity

- » $\lambda = C$

- » Then, $\lambda(Z + T) = N \rightarrow C(Z + T) = N \rightarrow T = N \cdot \frac{1}{C} - Z$



*ideal operating point $1/c_0$,
the minimum capacity at which we experiment no relevant delay*

Techniques Available for Performance Evaluation

- ♦ Analytical Stochastic Models
- ♦ Deterministic analysis
- ♦ Bottleneck analysis
- ♦ **Simulation**

Discrete Event Simulation

Introduction

- ◆ Precise analytical solutions for complex systems: difficult to obtain
- ◆ **Simulation**
 - » process which evaluates a system model numerically
 - » estimates quantities of interest
- ◆ Discrete Event Simulation
 - » Change of system state is caused by events
 - » An execution path of the system is obtained

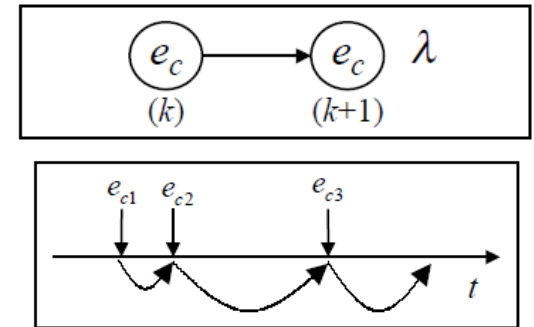
Simulation Process

- ◆ Initialization
 - » Time, initial system state and other counters are set to initial values
 - » Future events and their time of occurrence are calculated
- ◆ Time advance
 - » Time advances for the time of the next event (the one with smallest time)
 - » Next event is processed and the system state is updated
 - » Counters and other variables are updated
 - » Future events are calculated and stored in Events List
- ◆ Events List
 - » List of events $E(\text{OccurrenceTime}, \text{EventType}, \dots)$
 - » Events ordered by their increasing occurrence time
- ◆ End of simulation
 - » Counters and other variables are kept for analysis
 - » Simulations may be repeated with different sequence of random numbers
 - » New independent samples increase the confidence in the results obtained

Simulation of the Poisson Process

♦ Poisson Process

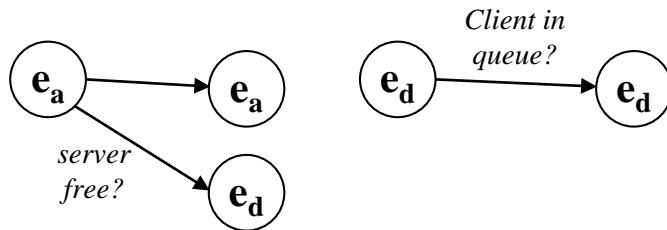
- » λ : client arrival rate (customer/s)
- » Event e_c : client arrives
- » Processing event $k \rightarrow$
determines occurrence of event $k+1$



- » Inter client arrival interval c has an exponential distribution $F(c)$
 - $F(c) = 1 - e^{-\lambda c}$ $E[c] = 1/\lambda$
- » If u is randomly generated in $[0,1]$ (uniform distribution, random function) then a sample of c may be obtained as

$$u = F(c) \Leftrightarrow c = F^{-1}(u)$$
$$c = -\frac{1}{\lambda} \ln(1-u) \Leftrightarrow c = -\frac{1}{\lambda} \ln(u)$$

A Simple Queue Model



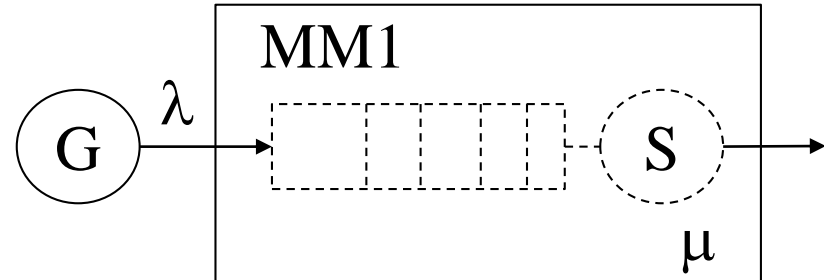
```

main()
{
    < create sch, g, m >
    sch.schedule(0.4, g.send() );
    sch.schedule( 10, sch.stop() );
    sch.loop();
    <clean up>
}
  
```

Scheduler

```

schedule()
loop()
getFirst()
stop()
  
```



G

send()

MM1

```

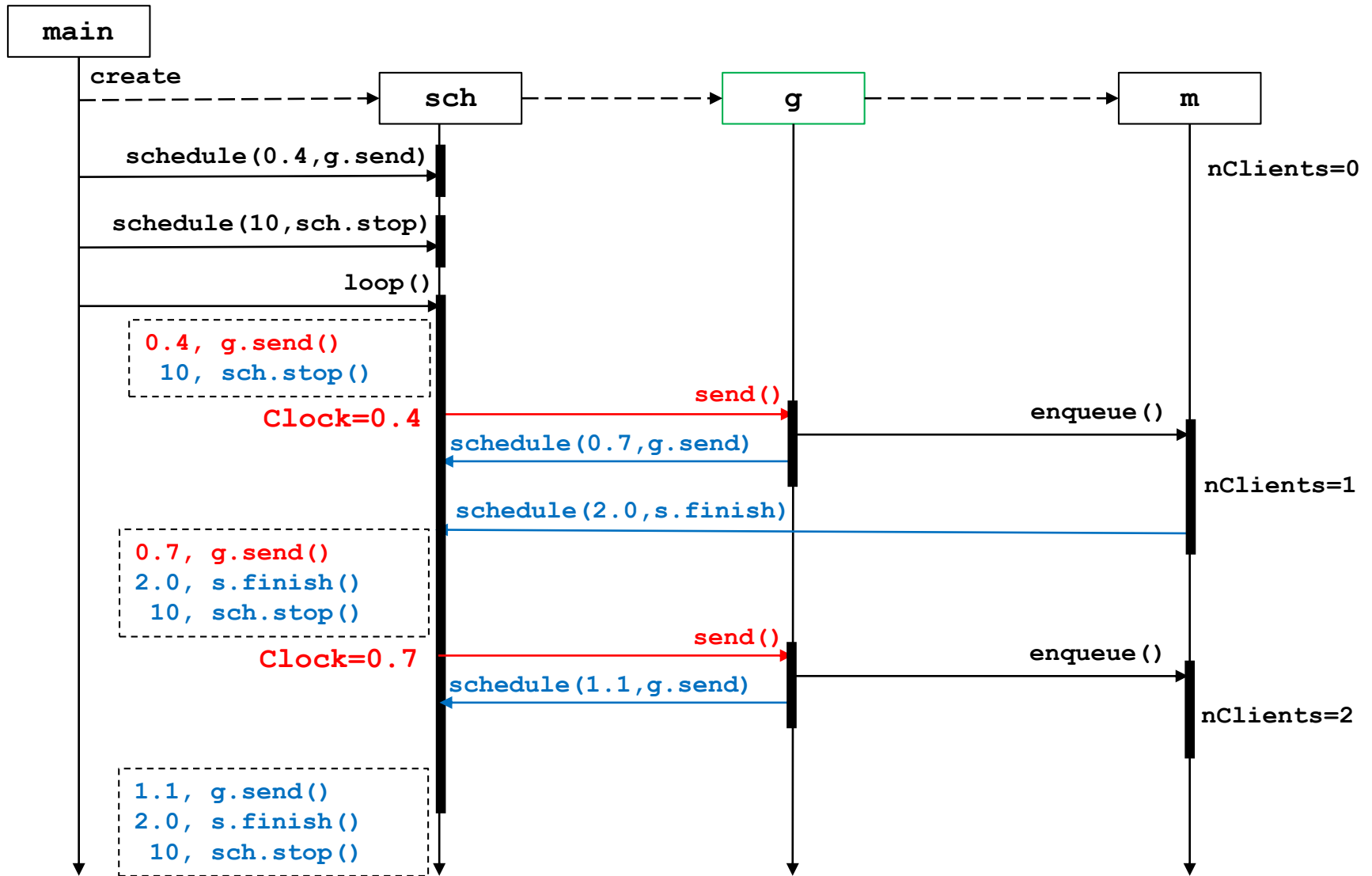
nClients
enqueue()
finish()
  
```

Scheduler::loop()

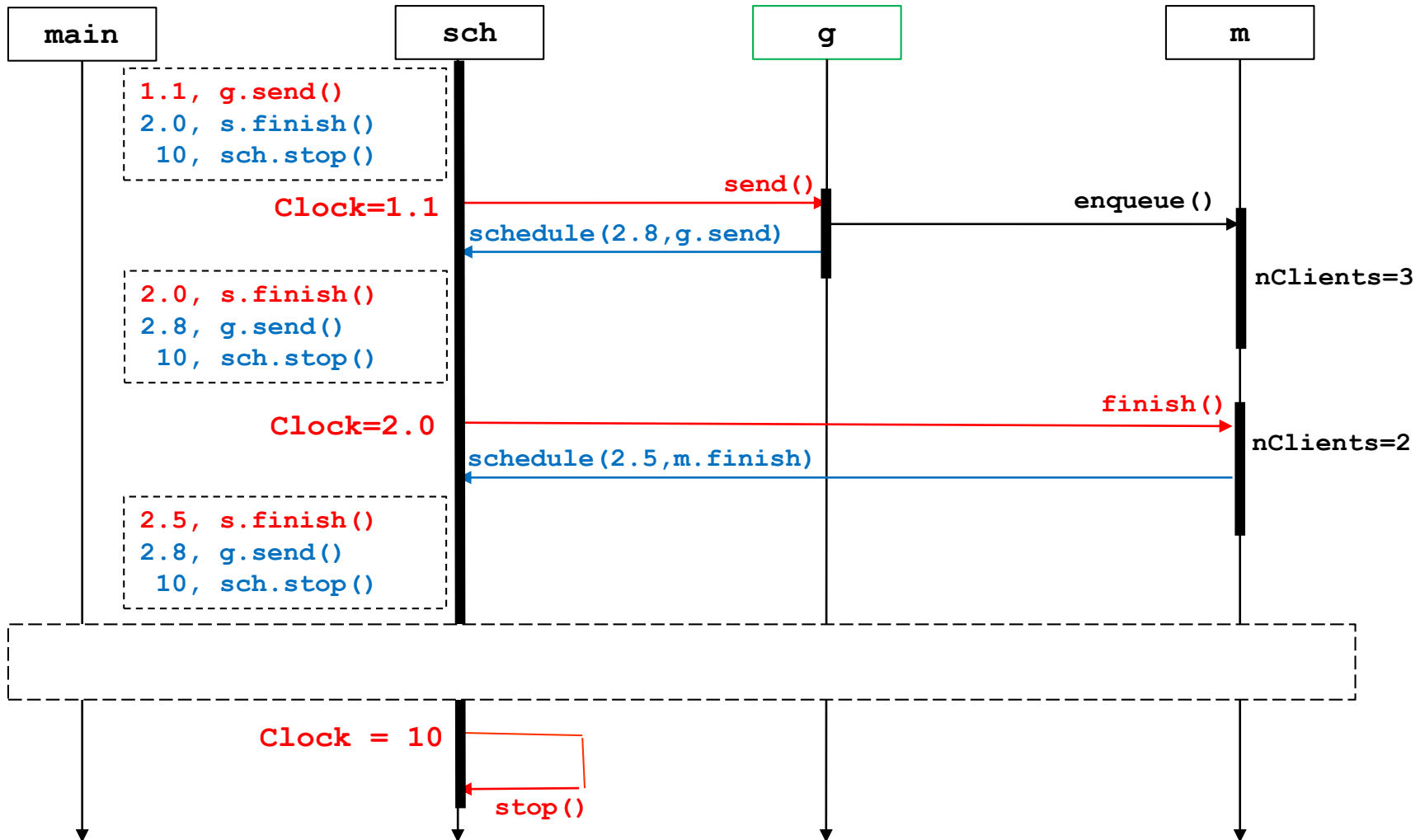
```

{
    do
    {
        event=getFirst() // fist event;
        clock = event.scheduledTime;
        event.function();
    } while (eventListNotEmpty && noStop)
}
  
```


Discrete Event Simulation in Execution (./..)



Discrete Event Simulation in Execution (../..)



What to Study?

- ◆ Formulate a question
 - » E.g.: How long does it take to serve N customers ?
- ◆ Answer
 - » This quantity is a random variable → no single answer exists
 - » It depends on
 - the various inter-arrival times for these customers
 - their corresponding service times
- ◆ Let us suppose we carry out one simulation
 - » ending after N customers are served
 - » and observe the termination time T_N
- ◆ T_N is an “estimate” of the quantity of interest

Performance Measures

- ♦ Expected Average System Time
- ♦ Server Utilization
- ♦ Mean Queue Length

Expected Average System Time

- ♦ Random variable S_N
- ♦ Time spent in the system by every single customer S_1, S_2, \dots, S_N
- ♦ The estimate is
$$\hat{S}_N = \frac{1}{N} \sum_{k=1}^N S_k$$

Server Utilization

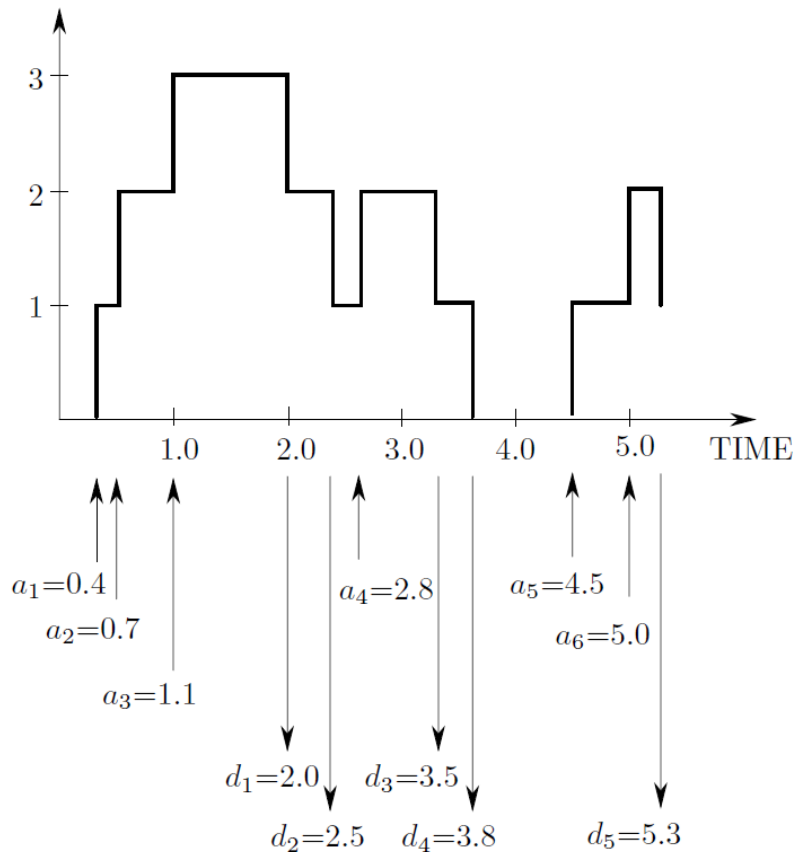
- ♦ Probability the server is busy during total time it takes to serve N customers: ρ_N
- ♦ Fraction of time during which the queue length is positive
- ♦ $T(i)$
 - » total observed time during which the queue length is i , $i = 0, 1, \dots$
- ♦ Estimate $\hat{\rho}_N$

$$\hat{\rho}_N = \frac{\sum_{i=1}^{\infty} T(i)}{T_N} = 1 - \frac{T(0)}{T_N} \qquad T_N = \sum_{i=0}^{\infty} T(i)$$

Mean Queue Length

- ♦ Q_N : average queue length over the time interval required to serve N customers
- ♦ $p_N(i)$ = probability that queue length is i over the time interval required to serve N customers, $i = 0, 1, \dots$
- ♦ By definition $Q_N = \sum_{i=0}^{\infty} i p_N(i)$
- ♦ *Estimate* $\hat{Q}_N \quad \hat{p}_N(i) = \frac{T(i)}{T_N} \quad \hat{Q}_N = \frac{1}{T_N} \sum_{i=0}^{\infty} i T(i)$
- ♦ During simulation we should maintain data variables $T(0), T(1), \dots$ for all observed queue lengths $i = 0, 1, \dots$

Sample path for $N=5$



QUEUE LENGTH OCCUPANCY TIMES	
T(0)	1.1
T(1)	1.4
T(2)	1.9
T(3)	0.9
T(4)	0.0

$$T_N = 5.3 \text{ s}$$

Arrival and departure times are a_k and d_k

Estimates of the Performance Measures

◆ Expected Average System Time $\hat{S}_5 = \frac{1.6 + 1.8 + 2.4 + 1.0 + 0.8}{5} = 1.52$

◆ Server Utilization $\hat{\rho}_5 = 1 - \frac{1.1}{5.3} \approx 0.79$

◆ Mean Queue Length
$$\begin{aligned}\hat{Q}_5 &= \frac{1}{5.3} \sum_{i=0}^{\infty} iT(i) \\ &= \frac{(0 \cdot 1.1) + (1 \cdot 1.4) + (2 \cdot 1.9) + (3 \cdot 0.9) + 0 + \dots}{5.3} \approx 1.49\end{aligned}$$

Output Analysis

- ♦ Simulation output data is used to estimate random variables
 - » Moments, percentiles
 - » Distribution functions, if possible
- ♦ Random variable **X**
 - » Function of the state of the system
 - » **Becomes a stochastic process $\{X(t)\}$ \rightarrow $X(t)$ distribution depends on t**

Non-terminating Simulation

- ♦ Study the steady-state behavior of a system
- ♦ The longer the simulation, the better, but there are limitations
- ♦ A stopping rule
 - » Let simulation length be T and the estimate $X(T)$
 - » Extend simulation length to $2T$ and obtain $X(2T)$
 - » If $|X(2T) - X(T)| < \text{Error}$, STOP
 - » Otherwise, extend the simulation to $3T$ and repeat for $|X(3T) - X(2T)|$
 - » and so on ...

Point estimation

- ♦ X_1, X_2, \dots, X_n is a **stochastic sequence of IID random variables**
 - » There is a **single probability distribution function**
 - » Let θ be the mean of that distribution
 - » We aim to estimate θ from X_1, X_2, \dots, X_n
 - » We can obtain the point estimate $\hat{\theta}_n$ based on n samples

$$\hat{\theta}_n = \frac{1}{n} \sum_{i=1}^n X_i \quad E[\hat{\theta}_n] = \frac{1}{n} \sum_{i=1}^n E[X_i] = \frac{1}{n} \sum_{i=1}^n \theta = \theta$$

- ♦ For the stochastic process $\{X(t)\}$ $\hat{\theta}_T = \frac{1}{T} \int_0^T X(t) dt$

Point Estimation

- ◆ The sample mean $\hat{\theta}_n$ is also a random variable

$$\text{Var}[\hat{\theta}_n] = E[(\hat{\theta}_n - \theta)^2] = E \left[\left(\frac{1}{n} \sum_{i=1}^n X_i - \theta \right)^2 \right] = \frac{1}{n^2} E \left[\left(\sum_{i=1}^n (X_i - \theta) \right)^2 \right] = \frac{\sigma^2}{n}$$

- ◆ We resort to the *sample variance*

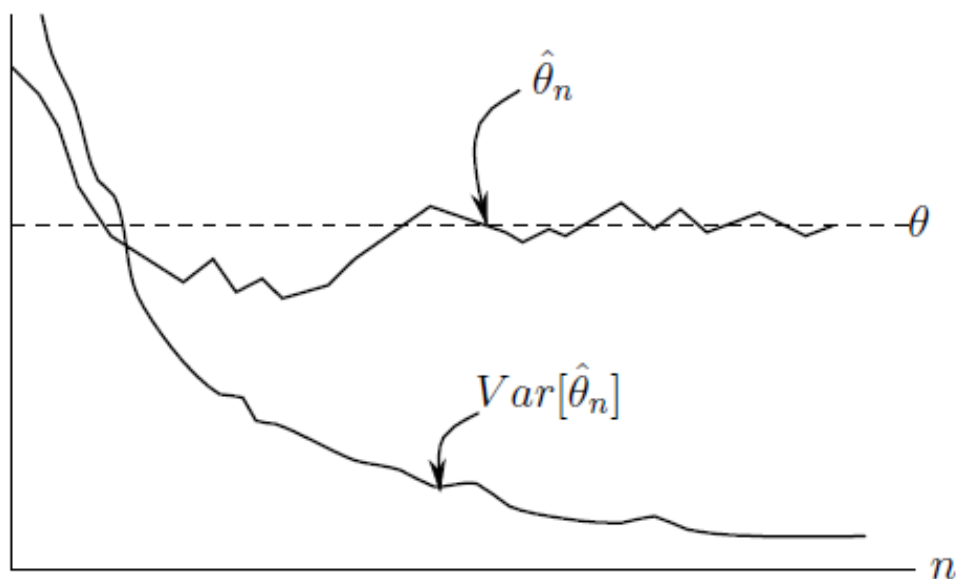
$$\hat{\sigma}^2(\hat{\theta}_n) = \frac{S_n^2}{n} = \frac{1}{n(n-1)} \sum_{i=1}^n (X_i - \hat{\theta}_n)^2 \qquad S_n^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \hat{\theta}_n)^2$$

Conclusion

Let X_1, X_2, \dots, X_n form an iid sequence of random variables with mean $\theta < \infty$. Then,

$$\hat{\theta}_n \rightarrow \theta \quad \text{with probability 1, as } n \rightarrow \infty$$

where $\hat{\theta}_n$ is the sample mean of X_1, X_2, \dots, X_n .



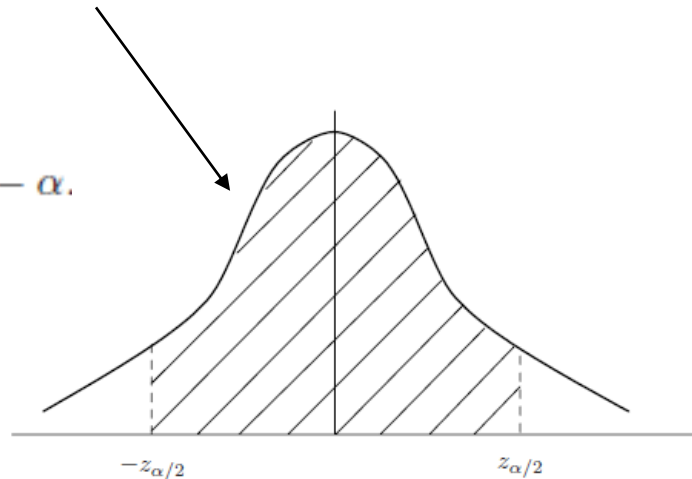
Interval estimation

- ♦ X_1, X_2, \dots, X_n , form an IID sequence $\hat{\theta}_n \xrightarrow{P} \theta$ as $n \rightarrow \infty$

$$Z_n = \frac{\hat{\theta}_n - E[\hat{\theta}_n]}{\sqrt{\text{Var}[\hat{\theta}_n]}} = \frac{\hat{\theta}_n - \theta}{\sqrt{\sigma^2/n}} \quad n \rightarrow \infty,$$

Z_n comes arbitrarily close to a normal distribution with mean zero and variance equal to 1.

$$P[-z_{\alpha/2} \leq Z \leq z_{\alpha/2}] = 1 - \alpha.$$



Interval Estimation

- ◆ $(1 - \alpha)$ confidence interval for θ is given by

$$P \left[\hat{\theta}_n - z_{\alpha/2} \sqrt{S_n^2/n} \leq \theta \leq \hat{\theta}_n + z_{\alpha/2} \sqrt{S_n^2/n} \right] \approx 1 - \alpha$$

- ◆ Confidence interval is specified by its lower and upper endpoints

$$L_{n,\alpha} = \hat{\theta}_n - z_{\alpha/2} \sqrt{S_n^2/n} \quad \text{and} \quad U_{n,\alpha} = \hat{\theta}_n + z_{\alpha/2} \sqrt{S_n^2/n}$$

- ◆ If n is not *very large*, **we should use Student's distribution**

$$P \left[\hat{\theta}_n - t_{n-1,\alpha/2} \sqrt{S_n^2/n} \leq \theta \leq \hat{\theta}_n + t_{n-1,\alpha/2} \sqrt{S_n^2/n} \right] \approx 1 - \alpha$$

- ◆ Where

- » n is the number of samples
- » $1 - \alpha$ is the confidence interval (ex. 90%)

Estimated standard deviation



For Non-IID Sequences ...

- ♦ IF X_1, X_2, \dots, X_n do not form IID sequence
 - » E.g. X_k is the time customer k spends in system and it depends on X_{k-1}
- ♦ We should define some performance measure of interest
 - » $L(X_1, X_2, \dots, X_M)$
 - » For instance
$$L(X_1, X_2, \dots, X_M) = \frac{1}{M} \sum_{i=1}^M X_i$$
- ♦ By repeating simulation n times, using different random numbers
 - » We can assume L_1, L_2, \dots, L_n is an IID sequence
 - » estimate
$$\hat{\theta}_n = \frac{1}{n} \sum_{j=1}^n L_j$$
 - » and proceed as before (confidence interval, Student's distribution)

Confidence Interval - Example

♦ Example: 10 samples:

- » $X_1=1.20; X_2=1.50; X_3=1.68; X_4=1.89; X_5=0.95;$
- » $X_6=1.49; X_7=1.58; X_8=1.55; X_9=0.50; X_{10}=1.05$

» $\hat{\theta}_{10} = \frac{1}{10} \sum_{i=1}^{10} X_i = 1.34$

» $\hat{\sigma}(\hat{\theta}_{10}) = \sqrt{\frac{1}{10*(10-1)} \sum_{i=1}^{10} [X_i - 1.34]^2} = 0.13$

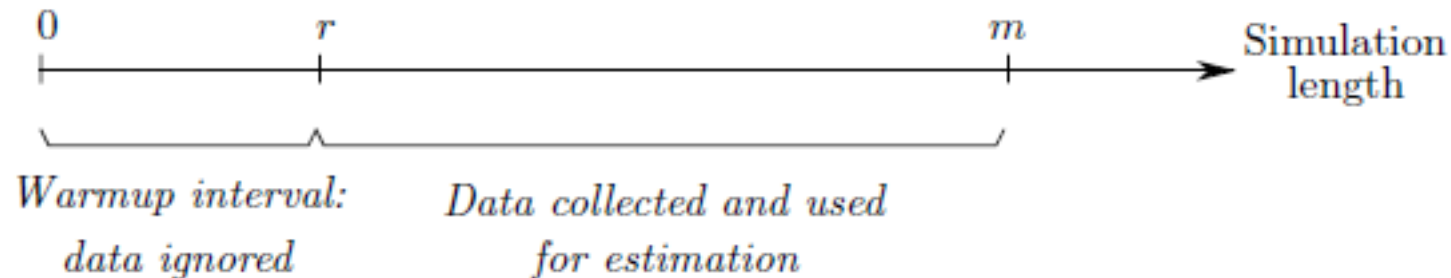
» $CI(1 - \alpha) = \hat{\theta}_{10} \pm z_{\alpha/2} \hat{\sigma}(\hat{\theta}_{10}) = CI(90\%) = 1.34 \pm 1.65 \times 0.13 = 1.34 \pm 0.21$

$(1 - \alpha)$	$z_{\alpha/2}$
80%	1.28
90%	1.65
95%	1.96
98%	2.33
99%	2.58

- ←
- Assuming a Normal distribution
 - For a small number of samples we should use *t-student's* distribution (see Wikipedia)

Non-terminating simulations – Replication with Deletions

- ◆ We are interested in the steady state
- ◆ Avoid transient state!



$$\hat{\theta}_{m,r} = \frac{1}{m-r} \sum_{i=r+1}^m X_i$$

References

- ♦ Christos G. Cassandras, Stéphane Lafortune, Introduction to Discrete Event Systems, Springer, 2008
- ♦ Jean-Yves Le Boudec, Performance Evaluation of Computer and Communication Systems, EPFL Press, 2010