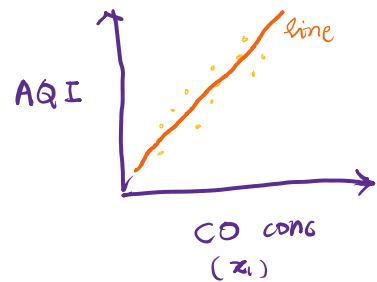
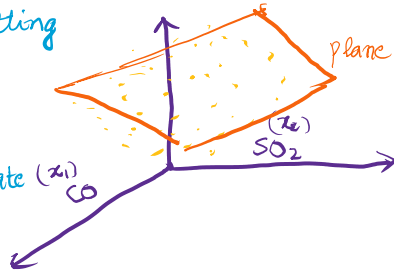


Lets begin with an example: AQI (Air Quality Index)

$$AQI = \max \{f_1(x), \dots, f_n(x_n)\} \quad \text{where } f_i(x_i) \text{ is a function associated with } i^{\text{th}} \text{ pollutant.}$$

The device for measuring AQI, may not be measuring all pollutants, but some of them and use the data to estimate AQI-

This is done by fitting the perfect AQI data with the limited observation and estimate AQI.



### Linear Regression

- "Simple but powerful"
- Works on transformation of raw data
- Interpretable

? How will we best fit the given data?

Measure the goodness of the fit using an error fn.

$$\text{Error/Loss/cost/Energy} : E(f, D) \quad \text{where } D = \{(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)\}$$

$f$  is the estimated function

Some possible E-functions:

- ①  $\sum_{i=1}^n (f(x_i) - y_i)$  (signed, not a good candidate for E-fn)
- ②  $\sum_{i=1}^n |f(x_i) - y_i|$  (unsigned, a good candidate)
- ③  $\sum_{i=1}^n (f(x_i) - y_i)^2$  (Squared error - a good candidate)
- ④  $\sum_{i=1}^n (f(x_i) - y_i)^3$  (signed, not a good candidate)

### Squared error function

$$\sum_{\{i: x_i, y_i \in D\}} (f(x_i) - y_i)^2$$

This function is useful as it is:

- Continuous, differentiable
- Geometrically, it is similar to reducing euclidean distance b/w estimated  $f^*$  and data set. Hence visualisable in Euclidean space
- Mathematical analysis becomes easier

Let  $x_i \in \mathbb{R}^d$ ,  $y \in \mathbb{R}$  :-  $x_i = \begin{bmatrix} x_{i1} \\ \vdots \\ x_{id} \end{bmatrix}$

$$X = \begin{bmatrix} x_{11} & x_{12} & \dots & x_{1d} \\ \vdots & \vdots & & \vdots \\ x_{n1} & x_{n2} & \dots & x_{nd} \end{bmatrix} = \begin{bmatrix} x_1^T \\ \vdots \\ x_n^T \end{bmatrix}$$

$$y = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}$$

### General regression problem

Find a  $f^n$   $f^*$  such that  $f^*(x)$  is the best predictor of  $y$  w.r.t  $D$

$$f^* \in \arg\min_{f \in \mathcal{F}} E(f, D)$$

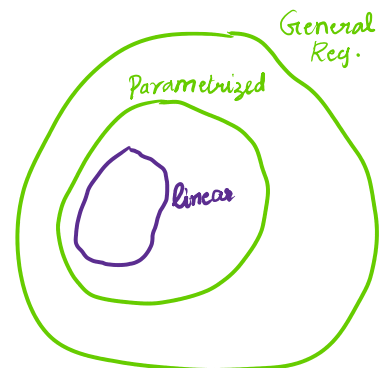
### Parametrized Regression

Here we will find a  $f^n$   $f \equiv f(x, w)$

Eg: i.  $f(x, (\alpha, \lambda)) = \alpha e^{-\lambda^T x}$

ii.  $f(x, w) = w_0 + w_1 x + \dots + w_n x^n$

Here,  $f^* \in \arg\min_w E(f(x, w), D)$



### Linear Regression

Here we've  $f(x, w) = w^T x + w_0 = \bar{w}^T \bar{x}$ , where  $w = \begin{bmatrix} w_1 \\ \vdots \\ w_d \end{bmatrix} \in \mathbb{R}^d$  and  $\bar{x} \in \mathbb{R}^d$

NB: Let  $f(x, w) = w_0 + w_1 x + w_2 x^2 + \dots + w_n x^n$ . by suitably replacing 'x' we can solve this problem as linear regression

### Least square Optimization for Linear Regression

$$n \text{ } \leftarrow \begin{matrix} n \\ \leftarrow \end{matrix} \leftarrow \begin{matrix} d \\ \leftarrow \end{matrix} \leftarrow \begin{matrix} 2 \\ \leftarrow \end{matrix} \dots \dots n = \{(x, y_1), \dots, (x, y_n)\}$$

$$\omega^* \in \underset{\omega}{\operatorname{argmin}} \sum_{i=1}^n \left( \sum_{j=1}^d \omega_j x_{ij} - y_i \right)^2 \quad \text{where } D = \{(x_1, y_1), \dots, (x_n, y_n)\}$$

Prediction:  $\hat{y}_i = \sum_{j=1}^d \omega_j^* x_{ij}$

d=1 LR:

$$E(\omega, D) = \sum_{i=1}^n (y_i - \omega_0 - \omega_1 x_i)^2$$

We need to find  $\omega = (\omega_0, \omega_1)$  such that  $\frac{\partial E}{\partial \omega_0} = \frac{\partial E}{\partial \omega_1} = 0$

$$\frac{\partial E}{\partial \omega_0} = -2 \sum_{i=1}^n (y_i - \omega_0 - \omega_1 x_i) = 0$$

$$\Rightarrow \omega_0 = \frac{\sum y_i - \omega_1 \sum x_i}{n} = \bar{y} - \omega_1 \bar{x}$$

where  $\bar{y} = \sum y_i / n$  and  $\bar{x} = \sum x_i / n$

Similarly,  $\frac{\partial E}{\partial \omega_1} = 0 \Rightarrow \omega_1 = \frac{\sum x_i y_i - \omega_0 \sum x_i}{\sum x_i^2}$

Let  $\alpha = \frac{\sum x_i y_i}{\sum x_i^2}$ ,  $\beta = \frac{\sum x_i^2}{n}$ .

We've:  $\omega_1 = \frac{\alpha \beta - \bar{x} \bar{y}}{\beta - \bar{x}^2} = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})^2}$

## d-D Data LR

$$x_i = \begin{bmatrix} x_{i1} \\ \vdots \\ x_{id} \end{bmatrix} \quad y = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix} \quad X = \begin{bmatrix} x_1^T \\ \vdots \\ x_n^T \end{bmatrix}^T$$

$$w^* \in \underset{w}{\operatorname{argmin}} \sum_{i=1}^n (y_i - w^T x_i)^2 = \sum_{i=1}^n z_i^2$$

where

$$z = \begin{bmatrix} z_1 \\ \vdots \\ z_n \end{bmatrix} = \begin{bmatrix} y_1 - x_1^T w \\ \vdots \\ y_n - x_n^T w \end{bmatrix} = y - Xw$$

Hence

$$\underset{w}{\operatorname{argmin}} (Xw - y)^T (Xw - y)$$

$$E(w, D) = w^T X^T X w - 2y^T X w + y^T y$$

$$\nabla_w E = 0 \Rightarrow 2X^T X w - 2X^T y = 0$$

$$\Rightarrow w^* = (X^T X)^{-1} (X^T y)$$

Also,  $\hat{y} = w^{*T} x$