

HW1

Saturday, February 9, 2019 9:55 AM

Problems : Grinstead Snell (7, 10, 16, 18, 25, 27, 28, 38)

7). Probability of size showing = $\frac{1}{6}$.

It follows geometric distribution.

$$a) p(T) = \left(\frac{5}{6}\right)^{T-1} \left(\frac{1}{6}\right) \quad \left\{ \begin{array}{l} T-1 \text{ failures} \\ \text{followed by 1 success} \end{array} \right\}.$$

$$\begin{aligned} b) p(T > 3) &= 1 - \sum_{i=1}^3 p(T = i) \\ &= 1 - \left(\frac{1}{6}\right) - \left(\frac{5}{6}\right)\left(\frac{1}{6}\right) - \left(\frac{25}{36}\right)\left(\frac{1}{6}\right) \\ &= 1 - \frac{1}{6} \left(1 + \frac{5}{6} + \frac{25}{36}\right) = 1 - \frac{91}{216} = \frac{125}{216} \end{aligned}$$

c) $p(T > 6 | T > 3) \Rightarrow$ We are sure of first 3 outcomes.

so this is similar to :

$$1 - \underbrace{p(T=1)}_4 + \underbrace{p(T=2)}_5 + \underbrace{p(T=3)}_6 = \left(\frac{125}{216}\right) \quad \text{from previous.}$$

10) $N \leftarrow$ Total population

$n_1 \leftarrow$ people counted 1st time.

$n_2 \leftarrow$ people counted 2nd time.

$n_{12} \leftarrow$ people counted both time.

	<u>Census 1</u>	<u>Census 2</u>	total counted
N	n_1	n_2	repeated.
	0	n_{12}	Not counted
	$N - n_1$	$N - n_1 - n_2 + n_{12}$	

a) Probability of $n_{12} = k$.

$$= \left(\text{Probability of selecting } k \text{ out of } n_1 \right) \times \left(\text{Probability of selecting } n_2 - k \text{ out of } N - n_1 \right)$$

(Total probability of selecting n_2 out of N)

$$= \frac{\binom{n_1}{k} \binom{N - n_1}{n_2 - k}}{\binom{N}{n_2}}$$

— Answer.

b) Let $x = n_{12}$

N which maximizes $x = n_{12}$

Probability of hypergeometric is highest at mean.

so. $n_{12} = \left(\frac{n_1}{N} \right)$ gives $N = \frac{n_1}{n_{12}}$
= Ans

16) $\lambda = \text{mean} = (0.01) \times (60 \times 5) = 3.$

$$P(\text{miss at most 1}) = P(0) + P(1).$$

Approximating to poison.

$$P(x=c) = \frac{\lambda^c}{c!} e^{-\lambda}$$

$$P(0) = \frac{3^0}{0!} e^{-3} = \frac{1}{e^3} = 0.0497 = 4.97\%$$

$$P(1) = \frac{3^1}{1!} e^{-3} = \frac{3}{e^3} = 0.1493 = 14.93\%$$

$$P(\text{at most 1 call}) = \underline{19.9\%} \quad \text{Ans}$$

18) 600 raisins 400 chocolate chips
500 cookies.

$$\lambda_{\text{Raisin}} = \frac{600}{500} = 1.2$$

$$\lambda_{\text{Choc.}} = \frac{400}{500} = 0.8$$

$$\text{a) } P_{\text{no raisin}} = \frac{(1.2)^0}{0!} e^{-1.2} = \frac{1}{e^{1.2}} = 0.3011$$

$$= \underline{30.11\%} \quad \text{Ans}$$

$$\text{b) } P_2 \text{ choc.} = \frac{(0.8)^2}{2!} e^{-0.8} = \frac{0.64}{2} e^{-0.8} = 0.1437$$

$$= \underline{14.37\%} \quad \text{Ans}$$

$$c) P(\text{at least 2 bits}) = 1 - P(0 \text{ bit} \vee 1 \text{ bit})$$

$$P(0 \text{ bit}) = P_{\text{no choc}} \times P_{\text{no raisin}}.$$

$$= \frac{1}{e^{0.8}} \times \frac{1}{e^{1.2}} = 0.1352 \\ = 13.52\%$$

$$P(1 \text{ bit only}) = P_{1 \text{ choc}} \times P_{\text{raisin}} + P_{0 \text{ choc}} \times P_{1 \text{ raisin}}$$

$$= \frac{(0.8)^1}{e^{0.8}} \times \frac{1}{e^{1.2}} + \frac{1.2}{e^{1.2}} \times \frac{1}{e^{0.8}} = 0.108 + 0.1624$$

$$= 0.2706 = \underline{27.06\%} \text{ Ans}$$

$$25) \text{ Money) parking meter} = 0.1 \times 100 = 10\$\text{}$$

$$\text{p) get caught} = 0.05$$

$$\text{mean) caught every 100 times} = 0.05 \times 100 = 5.$$

$$P(0) = \frac{5^0}{0!} e^{-5} = \frac{1}{e^5} = 0.0067 = 0.67\%.$$

$$P(1) = \frac{5}{1} e^{-5} = \frac{5}{e^5} = 0.0336 \approx 3.36\%.$$

$$P(2) = \frac{25}{2} e^{-5} = 12.5/e^5 = 0.084 = 8.4\%.$$

$$P(>2) = 1 - \left(\frac{18.5}{e^5} \right) = 87.6\%$$

(Expected) money when caught \Rightarrow

$$= 0 \times (P(0) + P(1)) + 2 \times P(2) + \sum_{x=3}^{100} [5(x-2)+2] P(x)$$

$$= 0 \times 4\% + 2 \times 8.4\% - 8(87.6\%) + 5 \sum_{x=3}^{100} x P(x)$$

$$= 0.168 - 7.008 + 5 \left(\underbrace{\sum_{x=0}^{100} x P(x)}_{\lambda N} \right) - 5 \left(\underbrace{\sum_{x=0}^2 x P(x)}_{\lambda^2} \right)$$

$$\approx 0.168 - 7.008 + 5 \times (0.05 \times 100) - 5 \left(0 \times P(0) + 1 \times P(1) + 2 \times P(2) \right)$$

$$= 18.16 - 5(0.033 + 2 \times 0.084)$$

$$= 18.16 - 5(0.201) = \underline{\underline{17.16 \$}}$$

Thus, Reese Prosser is expected to pay 17.16\\$ in the long run. It's better for him to follow the law.

27) Solving with Poisson.

$$\text{mean} = \lambda = 100 \times 0.001 = 0.01.$$

$P(\text{at least 1 accident})$

$$= 1 - [P(0 \text{ accident})].$$

$$P(0 \text{ accident}) = \frac{1}{e^{(0.1)}} = \frac{1}{1.105} = 0.904$$

$$P(\text{at least 1 such accident}) = 1 - 0.904 = 0.096 = \underline{\underline{9.6\%}}$$

(b) Solving using binomial

$$P(0 \text{ accident}) = (0.999)^{100}$$

$$P(\text{at least 1}) = 1 - (0.999)^{100}$$

$$= 1 - 0.9047 = 0.0953 = \underline{\underline{9.53\%}}$$

Ans

28) $P(\text{all fund seat}) = 1 - P(100 \text{ vgg show up}).$

$$P(100 \text{ show up}) = (1 - p \text{ not show})^{100}$$

$$= (0.96)^{100} = 0.01687 \\ = 1.68\%$$

$$\begin{aligned}
 P(\text{gg show up}) &= \binom{100}{1} (0.04)^1 (0.96)^{99} \\
 &= 100 \times 0.04 \times (0.96)^{99} \\
 &= 0.0702 \Rightarrow 7.02\%
 \end{aligned}$$

$$\begin{aligned}
 \text{So } P(\text{all fund seat}) &= 1 - 0.01687 = 0.9133 \\
 &\quad - 0.0702 \\
 &= \underline{\underline{91.33\%}} \text{ Ans}
 \end{aligned}$$

38) Total population = 20

Defective = 5

number selected = 5

$P(1 \text{ item defective}) ?$

a) Without replacement \Rightarrow .

Hypergeometric distribution can be used.

$$P(1 \text{ defect}) = \frac{\left[\begin{matrix} 15 \\ 4 \end{matrix} \right] \left[\begin{matrix} 5 \\ 1 \end{matrix} \right]}{\left[\begin{matrix} 20 \\ 5 \end{matrix} \right]}$$

4 good items
selected from
15 good items
1 bad item
selected from 5
bad item

5 items
selected from
20.

$P(1 \text{ defective w/o replacement})$

$$= \frac{15 \times 14 \times 13 \times 12}{4 \times 3 \times 2 \times 1} \times \frac{5}{\frac{20 \times 19 \times 18 \times 17 \times 16}{5 \times 4 \times 3 \times 2 \times 1}}$$

$$15C_4 = \frac{15 \times 14 \times 13 \times 12}{4 \times 3 \times 2 \times 1} = 0.4402 = \underline{\underline{44.02\%}}$$

Ans

b) With replacement.

We can use binomial distribution to model.

$P(1 \text{ defective in 5 trials})$.

$$P(1 \text{ defective}) = 5/20 = 1/4.$$

$$P(1 \text{ def. in 5}) = \binom{5}{1} \left(\frac{1}{4}\right)^1 \left(\frac{3}{4}\right)^4$$

$$= 5 \times \frac{1}{4} \times \frac{3^4}{4^4} = \frac{5 \times 3^4}{2^{10}} = \frac{5 \times 81}{1024}$$

$$= 0.3955 = \underline{\underline{39.55\%}}$$

Ans

Section 5.2

Problem 1, 7, 21, 37

1. $U \sim \text{Unif}[0, 1]$

a) $Y = U + 2$.

$$= \text{Unif}[0, 1] + 2. = \text{Unif}[2, 3]$$

$$f(Y) = 1 \quad 2 \leq Y \leq 3$$

$$F(Y) = \int_{-\infty}^Y f(Y)$$

$$= \int_2^Y f(Y) = [Y]_2^Y = \underline{\underline{Y-2}}$$

b) $Y = U^3$

$$\begin{aligned} F_Y(y) &= P(Y \leq y) \\ &= P(U^3 \leq y) = P(U \leq y^{\frac{1}{3}}) = F_U(y^{\frac{1}{3}}). \end{aligned}$$

$$f_Y(y) = \frac{d}{dy} F_Y(y) = \frac{d}{dy} F_U(y^{\frac{1}{3}})$$

$$= \frac{d}{du} F_U(y^{\frac{1}{3}}) \cdot \frac{du}{dy}.$$

$$= \cancel{P_U(y^{\frac{1}{3}})} \cdot \frac{d(y^{\frac{1}{3}})}{dy} = \frac{1}{3} y^{-\frac{2}{3}}.$$

$$f_Y(y) = \frac{1}{3} y^{-2/3}.$$

(can also be derived from $f_Y(y) = f_X(\phi^{-1}(y)) \frac{d(\phi^{-1}(y))}{dy}$).

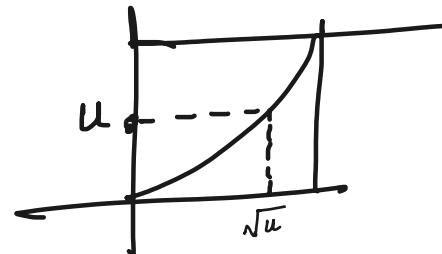
$$F_Y(y) = \int_0^y \frac{1}{3} y^{-2/3} dy = \frac{1}{3} \left(-\frac{1}{3} \right) y^{1/3} \Big|_0^y = y^{1/3}$$

— Answer

$$7. F_x \begin{cases} 0 & x < 0 \\ x^2 & 0 \leq x \leq 1 \\ 1 & x > 1 \end{cases}$$

$$x = F_x^{-1}(\text{rnd})$$

$$u = F_x(x) = x^2$$



$$\sqrt{u} = x,$$

To simulate this density:

1) Draw from uniform density between 0,1

2) Take the root of drawn rnd variable.

— Answer

17.

$$F_x = \begin{cases} 0 & x < 0 \\ \sin^2\left(\frac{\pi x}{2}\right) & 0 \leq x \leq 1 \\ 1 & x > 1 \end{cases}$$

a) $P_{x.} = \frac{d}{dx}(F_x)$

$$\begin{cases} 0 & x < 0 \\ 2 \sin\left(\frac{\pi x}{2}\right) \cos\left(\frac{\pi x}{2}\right) \left(\frac{\pi}{2}\right) & 0 \leq x \leq 1 \\ 0 & x > 1 \end{cases}$$

$$= \frac{\pi}{2} \sin(\pi x)$$

b) $P(x \leq 1/4)$

$$F_x(1/4) = \sin^2\left(\frac{\pi}{2}\left(\frac{1}{4}\right)\right) = \sin^2\left[\frac{\pi}{8}\right]$$

$$= 0.1464 \quad \sim 14.6\%$$

Answer

21

To prove :

$Y = F(X)$ where X is cumulative density,
is uniformly distributed in $0,1$.

Range of $Y = [0,1]$: from range of cumm.
density.

To prove. $P(Y) = y -$

$$P(F(X) \leq y)$$

since F is monotonically
increasing.

$$= P(X \leq F^{-1}(y))$$

[from theorem 5.1].

$$= F(F^{-1}(y))$$

[From def. of cumm.
density].

$$= y. \quad \underline{\text{Answer}}$$

37. Find log normal density of variable Y : $y = e^x$

where x is normal density.

$$f_Y(y) = f_X(\phi^{-1}(y)) \cdot \frac{d(\phi^{-1}(y))}{dy}$$

$$y = e^x$$

$$x = \log(y)$$

$$\phi^{-1}(y) = \log(y)$$

$$f_y(y) = f_x(\log(y)) \cdot \frac{d(\log(y))}{dy}.$$

$$= f_x(\log(y)) \times \frac{1}{y}.$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\left(\frac{\log(y)^2}{2}\right)} \frac{1}{y}$$

$$= \frac{1}{y\sqrt{2\pi}} e^{-\frac{(\log y)^2}{2}}$$

Answer.