ROBEM I A radio is powered by one battery, the lifetime of the battery obey the distribution of exponential distribution with parameter  $\lambda = \frac{1}{30}$ . In long term, in which frequence should we change the battery?

SOLTON. Easy to get that  $\lim_{t\to\infty}\frac{m(t)}{t}=\frac{1}{\mathbb{E}(\xi_1)}=\frac{1}{30}$ . So we change battery every 30 hours in average.

ROBEM II Consider a primitive renewing process with average renewing internal time  $\mu$ . Assume every renewing time is recorded by probability p, and each record and each renew are independence. Let  $N_r(t)$  be the times of renewing by recorded until time t.  $\{N_r(t): t \geq 0\}$  is a renewing process or not? And calculate  $\lim_{t\to\infty} \frac{N_r(t)}{t}$ .

SOUTHON. Assume  $X_n: n \in \mathbb{N}$  are i.i.d r.v and  $X_0 \sim Geo(p)$ , and  $(X_n: n \in \mathbb{N}) \perp (N(t): t \geq 0)$ . Let  $Y_n:=\sum_{k=1}^n X_n$ , and  $Y_0=0$ . Let  $\xi_r(n):=\sum_{k=Y_{n-1}+1}^{Y_n} \xi_k$ . Then  $\xi_r(n): n \in \mathbb{N}^+$  is update time of  $N_r$ . Since  $(X_n: n \in \mathbb{N}) \perp (N(t): t \geq 0)$ , we get that  $(\xi_r(n): n \in \mathbb{N}^+)$  are i.i.d. And  $\mathbb{E}(\xi_r(1))=\mathbb{E}(X_1)\mathbb{E}(\xi_1)=\frac{\mu}{p}$ . So  $\lim_{t\to\infty}\frac{N_r(t)}{t}=\frac{\mu}{p}$ .

ROBEM III Assume  $(U_n: n \in \mathbb{N}^+)$  are i.i.d r.v. and  $U_1 \sim U(0,1)$ . Assume  $X_{n,m}: n, m \in \mathbb{N}^+$  are r.v. and  $X_{n,m} \mid U_n \sim B(U_n)$ . And  $(X_{n,m} \mid U_n : m \in \mathbb{N}^+)$  are i.i.d. Let  $\xi_n := \inf\{m \in \mathbb{N}^+ : X_{n,m} = 1\}$  be the *n*-th update time of N(t). Find  $\lim_{t\to\infty} \frac{N(t)}{t}$ .

SOUTON. Easy to find that  $\mathbb{E}(\xi_1) = \int_0^1 \mathbb{E}(\xi_1 \mid U_1 = x) \, \mathrm{d}x = \int_0^1 \frac{\mathrm{d}x}{x} = \infty$ . So easy to find that  $\lim_{t \to \infty} \frac{N(t)}{t} = \infty$ .

ROBEM IV Assume  $(\xi_n : n \in \mathbb{N}^+)$  is i.i.d r.v. ranging in  $\mathbb{N}$  is update time of N(t). Let  $A_n$  be the event that at time n there is an update. Assume  $a = \lim_{n \to \infty} \mathbb{P}(A_n)$  exists. Prove that  $a = \frac{1}{\mathbb{E}(\xi_1)}$ .

SOUTON. Since  $N(n) = \sum_{k=1}^n \mathbbm{1}(A_k)$ , we know that  $\mathbb{E}(N(n)) = \sum_{k=1}^n \mathbb{P}(A_k)$ . Noting that  $\lim_{n \to \infty} \frac{N(n)}{n} = \frac{1}{\mathbb{E}(\xi_1)}$ , we obtain that  $\lim_{n \to \infty} \mathbb{E}(\frac{N(n)}{n}) = \frac{1}{\mathbb{E}(\xi_1)}$ . So  $\lim_{n \to \infty} \frac{\sum_{k=1}^n \mathbb{P}(A_k)}{n} = \frac{1}{\mathbb{E}(\xi_1)}$ . By stolz, we can get that  $\lim_{n \to \infty} \frac{\sum_{k=1}^n \mathbb{P}(A_k)}{n} = a$ . So  $a = \frac{1}{\mathbb{E}(\xi_1)}$ .

ROBEM V Assume  $N_1(t), N_2(t)$  are two independent updating process with update time distribution E(1), U(0, 2). Find an estimate of  $\mathbb{P}(N_1(100) + N_2(100) \ge 190)$ .

SOLION. Easy to know the expectation and varience of the update time are  $\mu_1 = 1, \sigma_1^2 = 1, \mu_2 = 1, \sigma_2^2 = \frac{1}{3}$ . So by the central limit theorem of updating process we know that

$$\frac{N_1(100) - 100}{\sqrt{100}}, \frac{N_2(100) - 100}{\sqrt{\frac{100}{3}}} \sim N(0, 1)$$

So  $\frac{N_1(100)+N_2(100)-200}{\sqrt{\frac{400}{2}}} \sim N(0,1)$ . So  $\mathbb{P}(N_1(100)+N_2(100)\geq 190) \approx \mathbb{P}(N(0,1)\geq -\frac{\sqrt{3}}{2})$ .