

*Lemma 1.* Assume  $(N_t : t \geq 0)$  is a random process, and  $\mathbb{P}(\forall t \in [0, \infty), \lim_{s \rightarrow t+} N_s = N_t, \forall t \in (0, \infty), \lim_{s \rightarrow t-} N_s \in \mathbb{R}) = 1$ . Assume  $\alpha > 0$ , and  $\forall t, s \geq 0$ , we have  $N_{t+s} - N_s \sim \text{Poisson}(\alpha t)$ . Then  $(N_t : t \geq 0)$  is a Poisson process  $\iff \forall 0 \leq s \leq t, N_t - N_s \perp \mathcal{F}_s$ , where  $\mathcal{F}_s := \sigma(N_x : x \leq s)$ .

*证明.* “ $\implies$ ”: To prove  $N_t - N_s \perp \mathcal{F}_s$ , only need to prove for  $t_1 < t_2 < \dots < t_{n-1} = s < t = t_n$ , we have  $N_t - N_s \perp \sigma(N_{t_k} : k = 1, \dots, n-1)$ . Easily  $N_t - N_s \perp \sigma(N_{t_{k+1}} - N_{t_k}, N_{t_1} : k = 1, \dots, n-2) = \sigma(N_{t_k} : k = 1, \dots, n-1)$ , so  $N_t - N_s \perp \mathcal{F}_s$ .

“ $\impliedby$ ”: For  $t_1 < \dots < t_n$ , we need to prove  $N_{t_1}, N_{t_{k+1}} - N_{t_k} : 1 \leq k \leq n-1$  are independent. Use MI to  $n$ . When  $n = 1$  it's obvious. Assume we have proved it for certain  $n \geq 1$ , now consider  $n+1$ . Since  $N_{t_{k+1}} - N_{t_k} \in \mathcal{F}_{t_n}, k = 1, \dots, n-1$ , we have  $N_{t_{n+1}} - N_{t_n} \perp \sigma(N_{t_1}, N_{t_{k+1}} - N_{t_k} : k = 1, \dots, n-1)$ . So  $\mathbb{P}(N_{t_1} \in A_1, N_{t_{k+1}} - N_{t_k} \in A_{k+1}, k = 1, \dots, n) = \mathbb{P}(N_{t_1} \in A_1, N_{t_{k+1}} - N_{t_k} \in A_{k+1}, k = 1, \dots, n-1) \mathbb{P}(N_{t_{n+1}} - N_{t_n} \in A_{n+1})$ . By Induction assumption we get  $\mathbb{P}(N_{t_1} \in A_1, N_{t_{k+1}} - N_{t_k} \in A_{k+1}, k = 1, \dots, n) = \mathbb{P}(N_{t_1} \in A_1) \prod_{k=1}^n \mathbb{P}(N_{t_{k+1}} - N_{t_k} \in A_{k+1})$ . So finally we get  $N_{t_1}, N_{t_{k+1}} - N_{t_k} : 1 \leq k \leq n$  are independent.  $\square$

**PROBLEM I** Assume  $(N_t : t \geq 0)$  is a Poisson process with parameter  $\alpha$ . Let  $P(t) := \mathbb{P}(2 \nmid N_t), Q(t) := \mathbb{P}(2 \mid N_t)$ . Prove that  $P(t) = e^{-\alpha t} \sinh(\alpha t), Q(t) = e^{-\alpha t} \cosh(\alpha t)$ .

*SOLUTION.* Easily to get

$$P(t) = \sum_{k=0}^{\infty} \mathbb{P}(N_t = 2k+1) = \sum_{k=0}^{\infty} \frac{(\alpha t)^{2k+1}}{(2k+1)!} e^{-\alpha t}$$

Noting that  $\sinh(\alpha t) = \frac{1}{2} \left( \sum_{k=0}^{\infty} \frac{(\alpha t)^k}{k!} - \sum_{k=0}^{\infty} \frac{(-\alpha t)^k}{k!} \right) = \sum_{k=0}^{\infty} \frac{(\alpha t)^{2k+1}}{(2k+1)!}$ , we easily obtain  $P(t) = e^{-\alpha t} \sinh(\alpha t)$ . Noting  $P(t) + Q(t) = 1$ , we easily get  $Q(t) = 1 - P(t) = e^{-\alpha t} \cosh(\alpha t)$ .  $\square$

**PROBLEM II** Assume  $(N_t : t \geq 0)$  is a Poisson process with parameter  $\alpha$ . Prove that  $\lim_{t \rightarrow \infty} \frac{N_t}{t} = \alpha, a.s..$

*Lemma 2.* Assume  $(N_t : t \geq 0)$  is a Poisson process with parameter  $\alpha$ . Then  $\mathbb{P}(\forall 0 \leq s \leq t, N_s \leq N_t) = 1$ .

*证明.* For  $s, t \in \mathbb{Q} \wedge 0 \leq s \leq t$ , we have  $\mathbb{P}(N_s > N_t) = 0$  since  $N_t - N_s \sim \text{Poisson}(\alpha(t-s))$ . So we get  $\mathbb{P}(\exists s, t \in \mathbb{Q}, 0 \leq s \leq t, N_s > N_t) = 0$ .

Now we will prove  $\exists s, t \in \mathbb{R}, 0 \leq s \leq t, N_s > N_t \implies \exists a, b \in \mathbb{Q}, 0 \leq a \leq b, N_a > N_b$ . Let  $a_n = \frac{[ns]}{n}, b_n = \frac{[nt]}{n}$ . Then  $\lim a_n = s, \lim b_n = t$ . Easily  $a_n \geq s, b_n \geq t$ . So since  $N$  is continuous we get  $\lim N_{a_n} = N_s, \lim N_{b_n} = N_t$ . Since  $N_s > N_t$ , we get  $\exists n, N_{a_n} > N_{b_n}$ . Let  $a = a_n, b = b_n$  will work.

So  $\mathbb{P}(\forall 0 \leq s \leq t, N_s \leq N_t) = 1 - \mathbb{P}(\exists 0 \leq s \leq t, N_s > N_t) = 1 - \mathbb{P}(\exists s, t \in \mathbb{Q}, 0 \leq s \leq t, N_s > N_t) = 1 - 0 = 1$ .  $\square$

*SOLUTION.* Consider  $N_n : n \in \mathbb{N}$ . Let  $X_n := N_n - N_{n-1}, n \geq 1$ . Then easily  $(X_n : n \in \mathbb{N}^+)$  is i.i.d sequence and  $X_1 \sim \text{Poisson}(\alpha)$ . So from the strong law of large numbers we get  $\lim_{n \rightarrow \infty} \frac{N_n}{n} = \alpha$ . From Lemma 2 we get  $\frac{[t]}{t} \frac{N_{[t]}}{[t]} \leq \frac{N_t}{t} \leq \frac{N_{[t]}}{[t]} \frac{[t]}{t}, \forall t \in \mathbb{R}$ , let  $t \rightarrow \infty$  we get  $[t], [t] \rightarrow \infty$ , and  $[t] \sim t \sim [t]$ . So finally we get  $\lim_{t \rightarrow \infty} \frac{N_t}{t} = \alpha$ .  $\square$

**PROBLEM III** Assume  $(N_t : t \geq 0)$  is a Poisson process with parameter  $\alpha > 0$ . Prove that  $\frac{N_t - \alpha t}{\sqrt{\alpha t}} \xrightarrow{d} N(0, 1)$ .

**SOLUTION**. Consider  $N_n : n \in \mathbb{N}$ . Let  $X_n := N_n - N_{n-1}, n \geq 1$ . Then easily  $(X_n : n \in \mathbb{N}^+)$  is i.i.d sequence and  $X_1 \sim \text{Poisson}(\alpha)$ . Easily  $\mathbb{V}(X_n) = \alpha < \infty, \mathbb{E}(X_n) = \alpha$ . So from the central limit theorem we get  $\frac{N_n - \alpha n}{\sqrt{\alpha n}} \xrightarrow{d} N(0, 1)$ . Noting  $\frac{N_t - \alpha t}{\sqrt{\alpha t}} = \frac{N_{[t]} - \alpha[t]}{\sqrt{\alpha[t]}} \frac{\sqrt{[t]}}{\sqrt{t}} + \frac{N_t - N_{[t]} - \alpha(t - [t])}{\sqrt{\alpha t}}$ . Let  $t \rightarrow \infty$  we get  $[t] \rightarrow \infty$ , and  $[t] \sim t$ . Noting  $N_t - N_{[t]} \stackrel{d}{=} N_{t - [t]}$ , and  $t - [t] \leq 1$ , we easily get  $\frac{N_t - N_{[t]}}{\sqrt{\alpha t}} \xrightarrow{d} 0$ . Easily  $\frac{\alpha(t - [t])}{\alpha t} \rightarrow 0$ , so finally we get that  $\frac{N_t - \alpha t}{\sqrt{\alpha t}} \xrightarrow{d} N(0, 1)$   $\square$

**PROBLEM IV** Assume  $(X_t : t \geq 0), (Y_t : t \geq 0)$  are two independent Poisson processes with parameter  $\alpha, \beta$  respectively. Prove that  $(X_t + Y_t : t \geq 0)$  is Poisson process with parameter  $\alpha + \beta$ .

**SOLUTION**. Write  $Z_t = X_t + Y_t$ . First we prove  $Z_{t+s} - Z_s \sim \text{Poisson}((\alpha + \beta)t)$ . Since  $X_{t+s} - X_s \sim \text{Poisson}(\alpha t), Y_{t+s} - Y_s \sim \text{Poisson}(\beta t)$ , and  $X_{t+s} - X_s \perp Y_{t+s} - Y_s$ , easily to get  $Z_{t+s} - Z_s = X_{t+s} - X_s + Y_{t+s} - Y_s \sim \text{Poisson}((\alpha + \beta)t)$ .

Second we prove  $\forall 0 \leq s \leq t, Z_t - Z_s \perp \mathcal{H}_s$ , where  $\mathcal{H}_s = \sigma(Z_x : 0 \leq x \leq s)$ . Easily  $Z_t - Z_s \in \sigma(X_t - X_s, Y_t - Y_s)$ . Easily  $X_t - X_s \perp \mathcal{F}_s := \sigma(X_x : 0 \leq x \leq s)$  since  $(X_x : x \geq 0)$  is Poisson process. Since  $(X_x : x \geq 0) \perp (Y_x : x \geq 0)$ , we get  $X_t - X_s \perp \mathcal{G}_s := \sigma(Y_x : 0 \leq x \leq s)$ . For the same reason we get  $Y_t - Y_s \perp \mathcal{F}_s, \mathcal{G}_s$ . So we get  $Z_t - Z_s \perp \sigma(\mathcal{F}_s, \mathcal{G}_s) \supset \mathcal{H}_s$ .

Finally, we prove that  $\mathbb{P}(\forall t \in [0, \infty), \lim_{s \rightarrow t+} Z_s = Z_t, \forall t \in (0, \infty), \lim_{s \rightarrow t-} Z_s \in \mathbb{R}) = 1$ . Since  $Z_t = X_t + Y_t$ , and  $\mathbb{P}(\forall t \in [0, \infty), \lim_{s \rightarrow t+} Y_s = Y_t, \forall t \in (0, \infty), \lim_{s \rightarrow t-} Y_s \in \mathbb{R}) = 1$ ,  $\mathbb{P}(\forall t \in [0, \infty), \lim_{s \rightarrow t+} X_s = X_t, \forall t \in (0, \infty), \lim_{s \rightarrow t-} X_s \in \mathbb{R}) = 1$ , obvious we get the requirement.

All in all,  $(X_t + Y_t : t \geq 0)$  is a Poisson process with parameter  $\alpha + \beta$ .  $\square$

**PROBLEM V** Assume  $(\xi_n : n \in \mathbb{N}^+)$  is a sequence of i.i.d. random variable ranging in  $\mathbb{Z}^d$ . Let  $X_n = X_0 + \sum_{k=1}^n \xi_k$ , and  $X_0 \perp (\xi_n : n \in \mathbb{N}^+)$  ranging in  $\mathbb{Z}^d$ , too. Assume  $(N_t : t \geq 0)$  is a Poisson process with parameter  $\alpha > 0$ . Discuss  $\frac{X_{N_t}}{t}$  when  $t \rightarrow \infty$ .

**SOLUTION**. First we prove that  $\lim_{t \rightarrow \infty} N_t = \infty, a.s.$ . We have  $\mathbb{P}(\sup_t N_t \geq n) \geq \mathbb{P}(N_t \geq n), \forall t, \forall n \in \mathbb{N}$ . Easily  $\lim_{t \rightarrow \infty} \mathbb{P}(N_t \geq n) = 1$ , so  $\mathbb{P}(\sup_t N_t \geq n) = 1, \forall n \in \mathbb{N}$ . So  $\mathbb{P}(\sup_t N_t = \infty) = 1$ . Noting Lemma 2 we easily get  $\mathbb{P}(\lim_{t \rightarrow \infty} N_t = \infty) = 1$ .

Now we can decompose  $\frac{X_{N_t}}{t}$  into  $\frac{X_{N_t}}{N_t} \frac{N_t}{t}$ . We have proved that  $\frac{N_t}{t} \rightarrow \alpha, a.s.$  in Problem II, so we only need to find  $\frac{X_{N_t}}{N_t}$ . Since  $N_t \rightarrow \infty, a.s.$ , we only need to find  $\frac{X_n}{n}$  when  $n \rightarrow \infty$ .

If  $\mathbb{E}(\xi_1)$  exists, then easily  $\frac{X_n}{n} \rightarrow \mathbb{E}(\xi_1), a.s.$ . Then we easily get  $\frac{X_{N_t}}{t} \rightarrow \alpha \mathbb{E}(\xi_1), a.s.$   $\square$