## **MarkovProcess**

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ROBEM I Assume  $(\mathscr{F}_t:t\geq 0,t\in\mathbb{R})$  is a filtration. For  $t\geq 0$  we let  $\mathscr{F}_{t+}:=\bigcap_{s>t}\mathscr{F}_s$ . Prove that  $\mathscr{F}_t\subset\mathscr{F}_{t+}$  and  $(\mathscr{F}_{t+}:t\geq 0)$  is a filtration.

SOLION. To prove  $\mathscr{F}_t \subset \mathscr{F}_{t+} = \bigcap_{s>t} \mathscr{F}_s$ , we only need to prove  $\forall s>t, \mathscr{F}_t \subset \mathscr{F}_s$ . By the definition of filtration it's obvious. Now we will prove  $(\mathscr{F}_{t+}:t\geq 0)$  is a filtration. Only need to prove  $\forall t,s\in\mathbb{R} \land t\leq s, \mathscr{F}_{t+}\subset \mathscr{F}_{s+}$ . By the definition of  $\mathscr{F}_{c+}$  we know that  $\mathscr{F}_{t+}=\bigcap_{x>t} \mathscr{F}_x=\bigcap_{x>s} \mathscr{F}_x\cap$ 

$$\bigcap_{x:t < x \le s} \mathscr{F}_x \subset \bigcap_{x > s} \mathscr{F}_x = \mathscr{F}_{s+}. \text{ So } (\mathscr{F}_{t+} : t \ge 0) \text{ is a filtration.}$$

ROBEM II Assume  $(X_t : t \ge 0, t \in \mathbb{R})$  is a stochastic process on probability space  $(\Omega, \mathscr{F}, \mathbb{P})$ . Prove that  $\forall s, t \ge 0, \varepsilon > 0, \{\rho(X_s, X_t) \ge \varepsilon\} \in \mathscr{F}$ .

SOLTON. Easily  $\{\rho(X_s, X_t) \geq \varepsilon\} = \bigcup_{k=1}^{\infty} \{\rho(X_s, X_t) > \varepsilon - \frac{1}{k}\varepsilon\}$ . So we only need to prove  $\forall k \in \mathbb{N}^+, \{\rho(X_s, X_t) > \varepsilon(1 - \frac{1}{k})\} \in \mathscr{F}$ . Take  $\delta = \varepsilon(1 - \frac{1}{k})$ , only need to prove  $\forall \delta > 0, \{\rho(X_s, X_t) > \delta\} \in \mathscr{F}$ .

 $\forall t \geq 0, X_t : \Omega \to E \text{ is measurable, where } E \subset \mathbb{R}^d. \text{ So we can find a countable dense set}$  in  $\mathbb{R}^d$ , write D. We will prove that  $\{\rho(X_s, X_t) > \delta\} = \bigcup_{q \in D} \{\rho(X_s, q) - \rho(X_t, q) > \delta\}$ . On one hand, easily  $\rho(X_s, q) - \rho(X_t, q) > \delta \implies \rho(X_s, X_t) > \delta$  from triangle inequality. So we easily get  $\{\rho(X_s, X_t) > \delta\} \supset \bigcup_{q \in D} \{\rho(X_s, q) - \rho(X_t, q) > \delta\}$ . On the other hand, assume for certain  $\omega \in \Omega$  we have  $\rho(X_s(\omega), X_t(\omega)) > \delta$ , we will prove  $\exists q \in D, \rho(X_s(\omega), q) - \rho(X_t(\omega), q) > \delta$ . For convenience, we omit  $(\omega)$  from now on to the end of this paragraph. Since  $\rho(X_s, X_t) > \delta$ , we know  $\gamma := \frac{\rho(X_s, X_t) - \delta}{2} > 0$ . Since D is dense, we obtain  $\exists q \in D, \rho(X_t, q) < \gamma$ . So from triangle inequality we get  $\rho(X_s, q) \geq \rho(X_s, X_t) - \rho(X_t, q) > 2\gamma + \delta - \gamma = \gamma + \delta$ . So we get  $\rho(X_s, q) - \rho(X_t, q) > \gamma + \delta - \gamma = \delta$ . Finally, we get  $\{\rho(X_s, X_t) > \delta\} = \bigcup_{q \in D} \{\rho(X_s, q) - \rho(X_t, q) > \delta\}$ .

Noting  $\bigcup_{q\in D} \{\rho(X_s,q) - \rho(X_t,q) > \delta\} = \bigcup_{q\in D} \bigcup_{p\in \mathbb{Q}^+} \{\rho(X_s,q) > \delta + p, \rho(X_t,q) < p\}$ , and  $D, \mathbb{Q}^+$  are countable, so we only need to check  $\{\rho(X_s,q) > \delta + p, \rho(X_t,q) < p\} \in \mathscr{F}, \forall q \in D, p \in \mathbb{Q}^+$ . Noting  $\{\rho(X_s,q) > \delta + p, \rho(X_t,q) < p\} = \{\rho(X_s,q) > \delta + p\} \cap \{\rho(X_t,q) < p\}$ , and  $X_s, X_t$  are measurable from  $\Omega$  to E, we obtain  $\{\rho(X_s,q) > \delta + p\}, \{\rho(X_t,q) < p\} \in \mathscr{F}$ . So we proved  $\{\rho(X_s,X_t) > \delta\} \in \mathscr{F}, \forall s,t \geq 0, \forall \delta > 0$ .

Finally, we obtain  $\{\rho(X_s, X_t) \ge \varepsilon\} \in \mathscr{F}, \forall s, t \ge 0, \varepsilon > 0.$ 

ROBEM III Let  $\mathscr{D}_X := \{\mu_J^X : J \in S(I)\}$  be the family of finite-dimentional distributions of a stochastic process  $(X_t : t \geq 0, t \in \mathbb{R})$ .  $\forall (s_1, s_2) \in S(I)$  and  $J = (t_1, \dots, t_n) \in S(I)$ , write  $K_1 := (s_1, s_2, t_1, \dots, t_n) \in S(I)$ ,  $K_2 := (s_2, s_1, t_1, \dots, t_n) \in S(I)$ . Take  $A_1, A_2 \in \mathscr{E}, B \in \mathscr{E}^n$ , prove that

$$\mu_{K_1}^X(A_1 \times A_2 \times B) = \mu_{K_2}^X(A_2 \times A_1 \times B)$$

and

$$\mu_{K_1}^X(E \times E \times B) = \mu_{K_2}^X(E \times E \times B) = \mu_J^X(B)$$

SOUTHON. By the definition of finite-dimentional distributions we get

$$\mu_{K_1}^X(A_1 \times A_2 \times B) = \mathbb{P}((X_{s_1}, X_{s_2}, X_{t_1}, \cdots, X_{t_n}) \in A_1 \times A_2 \times B)$$

$$= \mathbb{P}(X_{s_1} \in A_1) \mathbb{P}(X_{s_2} \in A_2) \mathbb{P}((X_{t_1}, X_{t_2}, \cdots, X_{t_n}) \in B)$$

For the same reason, we obtain

$$\mu_{K_2}^X(A_2 \times A_1 \times B)$$

$$= \mathbb{P}((X_{s_2}, X_{s_1}, X_{t_1}, \cdots, X_{t_n}) \in A_2 \times A_1 \times B)$$

$$= \mathbb{P}(X_{s_2} \in A_2) \mathbb{P}(X_{s_1} \in A_1) \mathbb{P}((X_{t_1}, X_{t_2}, \cdots, X_{t_n}) \in B)$$

So we get

$$\mu_{K_1}^X(A_1 \times A_2 \times B) = \mu_{K_2}^X(A_2 \times A_1 \times B)$$

Also, let  $A_1 = A_2 = E$ , we get

$$\mu_{K_1}^X(E \times E \times B) = \mu_{K_2}^X(E \times E \times B)$$

$$= \mathbb{P}(X_{s_1} \in E)\mathbb{P}(X_{s_2} \in E)\mathbb{P}((X_{t_1}, \dots, X_{t_n}) \in B)$$

$$= \mathbb{P}((X_{t_1}, \dots, X_{t_n}) \in B)$$

By the definition of finite-dimentional distributions we get

$$\mu_J^X(B) = \mathbb{P}((X_{t_1}, \cdots, X_{t_n}) \in B)$$

So finally we get

$$\mu_{K_1}^X(E \times E \times B) = \mu_{K_2}^X(E \times E \times B) = \mu_J^X(B)$$

ROBEM IV Assume  $(\tau_k : k \in \mathbb{N}^+)$  is an i.i.d sequence of r.v. with exponential distribution with parameter  $\alpha > 0$ . Let  $S_n := \sum_{k=1}^n \tau_k$ . For  $t \geq 0, t \in \mathbb{R}$ , let:

$$N_t := \sum_{n=1}^{\infty} \mathbb{1}_{\{S_n \le t\}}, X_t := \sum_{n=1}^{\infty} \mathbb{1}_{\{S_n < t\}}$$

Prove that N and X are modifications of each other, but they are not indistinguishable.

SOUTON. First we prove N and X are modifications of each other. Fix  $t \in [0, \infty)$ , we need to prove  $\mathbb{P}(N_t = X_t) = 1$ . Noting  $N_t - X_t = \sum_{n=1}^{\infty} \mathbb{1}_{S_n \le t} - \mathbb{1}_{S_n < 1} = \sum_{n=1}^{\infty} \mathbb{1}_{S_n = t}$ , we get  $\mathbb{P}(N_t = X_t) = 1$ .

$$X_t$$
) =  $\mathbb{P}(\sum_{n=1}^{\infty} \mathbb{1}_{S_n=t}) = \mathbb{P}(\forall n \in \mathbb{N}^+, S_n \neq t)$ . So we only need to prove  $\mathbb{P}(S_n = t) = 0, \forall n \in \mathbb{N}^+$ .

Since  $\tau_k, k \in \mathbb{N}^+$  are continuous-distributed, we know  $S_n = \sum_{k=1}^n \tau_k$  is continuous-distributed, so  $\mathbb{P}(S_n = t) = 0$ . So we proved N and X are modifications of each other.

Next we will prove they are not indistinguishable. Only need to prove  $\mathbb{P}(\forall t \in [0, \infty), X_t = N_t) = 0 \neq 1$ . Since  $N_t - X_t = \sum_{n=1}^{\infty} \mathbb{1}_{S_n \leq t} - \mathbb{1}_{S_n \leq t} = \sum_{n=1}^{\infty} \mathbb{1}_{S_n = t}$ , we know  $\forall t, N_t = X_t \iff$ 

 $\forall t, \forall n \in \mathbb{N}^+, S_n \neq t$ . But  $S_n$  is ranged in  $[0, \infty)$ , so R.H.S is an impossible event. So we finally get  $\mathbb{P}(\forall t \in [0, \infty), X_t = N_t) = 0$  and thus X and N are not indistinguishable.

ROBEM V Assume T is non-negetive r.v. with distribution function F continuous on  $\mathbb{R}$ . Let  $X_t = \mathbb{1}_{\{T \leq t\}}$ . Prove that X is stochastically continuous.

SOUTION. Only need to check  $\forall t \geq 0, X_s \stackrel{\mathbb{P}}{\to} X_t, s \to t$ . Take  $\varepsilon > 0$ , we need to prove  $\lim_{s \to t} \mathbb{P}(\rho(X_s, X_t) > \varepsilon) = 0$ . For  $u > v \geq 0$ , we have  $X_u - X_v = \mathbb{1}_{v < T \leq u}$ . So  $\mathbb{P}(\rho(X_u - X_v) > \varepsilon) \leq \mathbb{P}(X_u \neq X_v) = \mathbb{P}(v < T \leq u) \leq \mathbb{P}(T \in [v, u])$ . So we easily get  $\lim_{s \to t+} \mathbb{P}(\rho(X_s, X_t) > \varepsilon) \leq \lim_{s \to t+} \mathbb{P}(T \in [t, s]) = 0$  and  $\lim_{s \to t-} \mathbb{P}(\rho(X_s, X_t) > \varepsilon) \leq \lim_{s \to t-} \mathbb{P}(T \in [s, t]) = 0$ . So  $\lim_{s \to t} \mathbb{P}(\rho(X_s, X_t) > \varepsilon) = 0$ .

ROBEM VI Assume  $I = \mathbb{Z}^+$ , then the stochastic process  $X = (X_0, X_1, \cdots)$  is a r.v. from  $\Omega$  to  $E^{\infty}$ . Define the distribution of X,  $\mu_X$ , as follows:

$$\mu_X(A) = \mathbb{P}(X \in A), A \in \mathscr{E}^{\infty}$$

Then stochastic process X, Y are equivalent  $\iff \mu_X = \mu_Y$ .

SOUTON. "  $\Longrightarrow$  ":Assume X,Y are equivalent, now we will prove  $\mu_X = \mu_Y$ . Let  $\mathscr{A} := \{A \in \mathscr{P}(E^{\infty}) : \exists n \in \mathbb{N}^+, A = A_1 \times A_2 \times \cdots \times A_n \times E \times E \times \cdots \}$ . Then we can get  $\mu_X(A) = \mu_{(1,2,\cdots,n)}^X(A_1 \times \cdots \times A_n)$ . So for  $A \in \mathscr{A}$  we know  $\mu_X(A) = \mu_Y(A)$ . By the definition of  $\mathscr{E}^{\infty} = \sigma(\mathscr{A})$ , and noting  $\mathscr{A}$  is a Semiset algebra, by the Measure extension theorem we get  $\mu_X = \mu_Y$ .

"  $\Leftarrow$  ": Assume  $\mu_X = \mu_Y$ , then easily  $\mu_{(s_1, \dots, s_n)}^X(A_{s_1} \times \dots \times A_{s_n}) = \mu_X(\prod_{k \in \mathbb{N}^+} B_k)$ , where  $B_k =$ 

 $A_{s_t}$  for  $k=s_t$  and  $B_k=E$  for  $k\neq s_t, \forall t=1,\cdots,n$ . So easily  $\mu_J^X=\mu_J^Y, \forall J\subset I \land |J|<\infty$ .