

# Hidden Markov Models: lecture 6

Model selection and checking

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# HMM definition

- ▶ A Hidden Markov Model (HMM) is a Markov chain in which the sequence of states  $C_1, \dots, C_T$  is not observed but hidden
- ▶ Instead of observing the sequence of states, we observe the emissions  $X_1, \dots, X_T$
- ▶ A HMM is defined by two quantities:
  - ▶ The transition matrix  $\Gamma$  of elements  $\gamma_{ij}$  where  $i$  and  $j$  are states:

$$\gamma_{ij} = p(C_t = j | C_{t-1} = i)$$

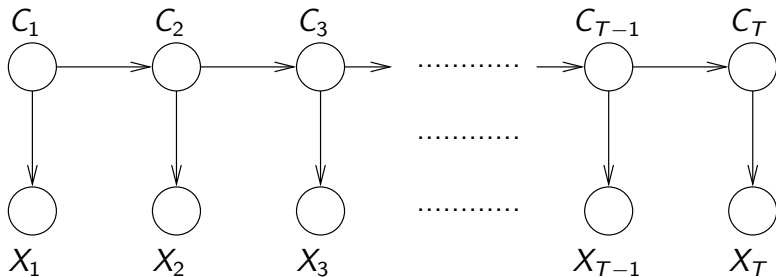
- ▶ The emission probabilities  $p_i(x)$  where  $i$  is a state and  $x$  is an emission:

$$p_i(x) = p(X_t = x | C_t = i)$$

- ▶ The unconditional distribution at  $t$  is denoted  $\mathbf{u}(t)$  and the initial distribution is  $\mathbf{u}(1)$

$$\mathbf{u}(t) = (p(C_t = 1), p(C_t = 2), \dots, p(C_t = m))$$

## Dependency graph of a hidden Markov model



$$p(\mathbf{X}^{(T)}, \mathbf{C}^{(T)}) = p(C_1) \prod_{k=2}^T p(C_k | C_{k-1}) \prod_{k=1}^T p(X_k | C_k)$$

$$p(\mathbf{x}^{(T)}, \mathbf{c}^{(T)}) = u_{c_1}(1) \prod_{k=2}^T \gamma_{c_{k-1}c_k} \prod_{k=1}^T p_{c_k}(x_k)$$

# Parameter estimation

- ▶ In the previous lectures, we discussed how to calculate the likelihood, how to estimate the hidden states, and how to estimate the parameters of a HMM
- ▶ This assumed that we know the (unparametrised) structure of the HMM to use, and in particular the number of hidden states
- ▶ In this lecture we are concerned with two closely related questions
- ▶ How can we **select** between two models? For example, a model with 2 states and a model with 3 states?
- ▶ How can we **check** that a model fits well with the data, without comparison with other models
- ▶ These are very common problems when modelling, not just for HMM, and even though we focus on HMM some of the solutions we will describe apply to more general models

## Model selection

- ▶ The likelihood  $L$  is an indication of how well a model fits
- ▶ A HMM with more states  $m$  will always have a better likelihood  $L$
- ▶ Two criteria are in common use to account for the complexity of models when comparing them
- ▶ The Akaike information criterion:

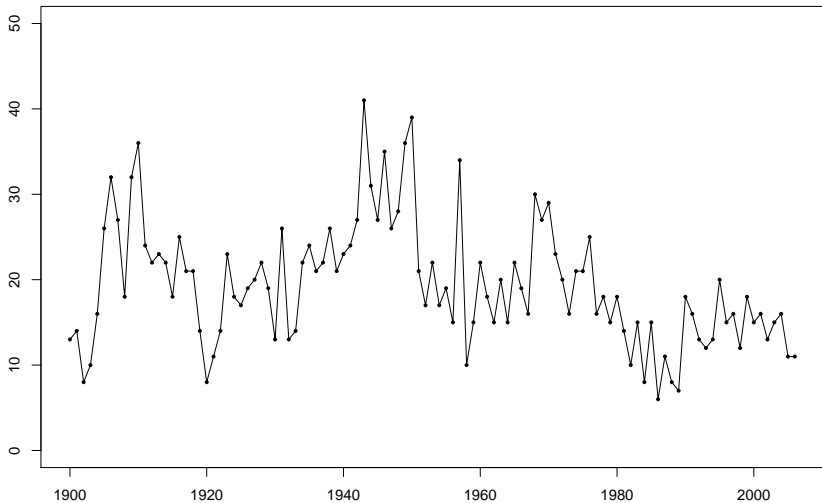
$$\text{AIC} = -2\log(L) + 2p$$

- ▶ The Bayesian information criterion:

$$\text{BIC} = -2\log(L) + p\log(T)$$

- ▶  $p$  is the number of parameters of the model, and  $T$  the number of observations
- ▶ The model with the smallest AIC or BIC is selected
- ▶ In the case of a HMM with  $m$  states and Poisson emissions, we have  $m^2 - m$  parameters in the transition matrix  $\mathbf{\Gamma}$  and  $m$  parameters for each of the Poisson rates and so  $p = m^2$

# Earthquake example

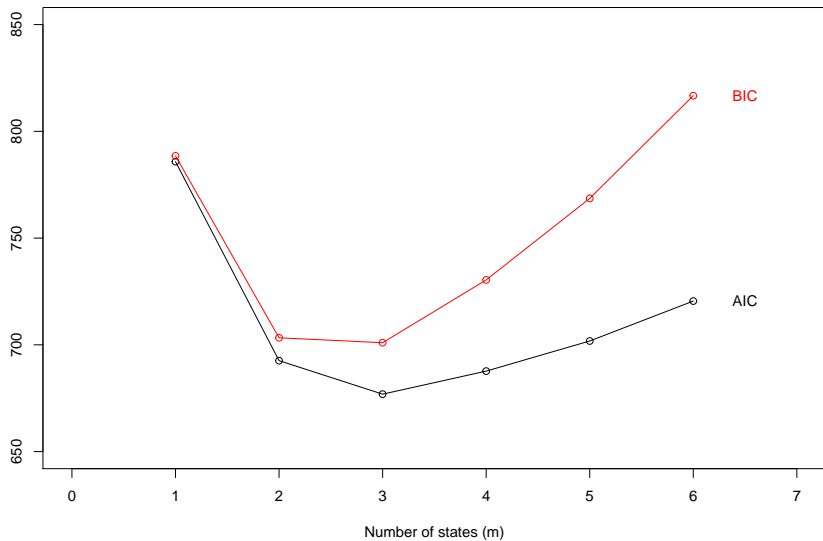


## Application to earthquakes data

- ▶ HMM = hidden Markov model ( $p = m^2$ )
- ▶ IM = independent mixture ( $p = 2m - 1$ )

Model	m	p	logL	AIC	BIC
HMM	1	1	-391.9189	785.8	788.5
HMM	2	4	-342.3183	692.6	703.3
HMM	3	9	-329.4603	676.9	701.0
HMM	4	16	-327.8316	687.7	730.4
HMM	5	25	-325.9000	701.8	768.6
HMM	6	36	-324.2270	720.5	816.7
IM	1	1	-391.9189	785.8	788.5
IM	2	3	-360.3690	726.7	734.8
IM	3	5	-356.8489	723.7	737.1
IM	4	7	-356.7337	727.5	746.2

## Application to earthquakes data





## Application to earthquakes data

- ▶ Both AIC and BIC select the HMM with  $m = 3$  states
- ▶ More generally, BIC and AIC do not always agree
- ▶ When  $T > e^2$  (as is usually the case) the BIC penalizes larger models more than the AIC
- ▶ Independent mixture models are not as good as HMM for this dataset, despite the higher number of parameters in HMM

# Model checking

- ▶ Even after selection of the best model, there remains the question of how good the model is in absolute (not relative) terms
- ▶ Need to assess the goodness of fit of the model
- ▶ This is something commonly done for simpler model and that we need to adapt for HMM
- ▶ For example, in a simple linear regression model we have:

$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$$

with  $\epsilon_i \sim \text{Norm}(0, \sigma^2)$

- ▶ The residuals  $y_i - \beta_0 - \beta_1 x_i$  are therefore expected to be independently and identically distributed as  $\text{Norm}(0, \sigma^2)$  and this can be used to check the model
- ▶ For a HMM, what is the residual for each observation  $X_t$ ?

## HMM pseudo-residuals

- ▶ We consider that  $X_t$  is continuous (similar results can be derived for the discrete case)
- ▶ If  $X_t$  is from a distribution with cumulative density  $F_{X_t}$  then we can consider the pseudo-residual:

$$z_t = \Phi^{-1}(F_{X_t}(x_t))$$

where  $\Phi$  is the cumulative density of a Normal(0,1)

- ▶ Since  $X_t$  has cumulative density  $F_{X_t}$  we have that  $F_{X_t}(x_t)$  should be distributed as Unif(0,1) and therefore  $z_t$  should be distributed as Normal(0,1)
- ▶ What is the distribution of  $X_t$ ?

# HMM pseudo-residuals

- ▶ What is the distribution of  $X_t$ ?
- ▶ One approach is to use the conditional distribution given all other data:

$$f_{X_t}(x) = p(X_t = x | \mathbf{X}^{(-t)} = \mathbf{x}^{(-t)})$$

- ▶ Using our previous calculation of the likelihood, we find:

$$f_{X_t}(x) = \frac{\mathbf{u}(1)\mathbf{P}(x_1)\mathbf{\Gamma} \dots \mathbf{P}(x_{t-1})\mathbf{\Gamma} \mathbf{P}(x)\mathbf{\Gamma} \mathbf{P}(x_{t+1}) \dots \mathbf{\Gamma} \mathbf{P}(x_T)\mathbf{1}'}{\mathbf{u}(1)\mathbf{P}(x_1)\mathbf{\Gamma} \dots \mathbf{P}(x_{t-1})\mathbf{\Gamma} \mathbf{\Gamma} \mathbf{P}(x_{t+1}) \dots \mathbf{\Gamma} \mathbf{P}(x_T)\mathbf{1}'}$$

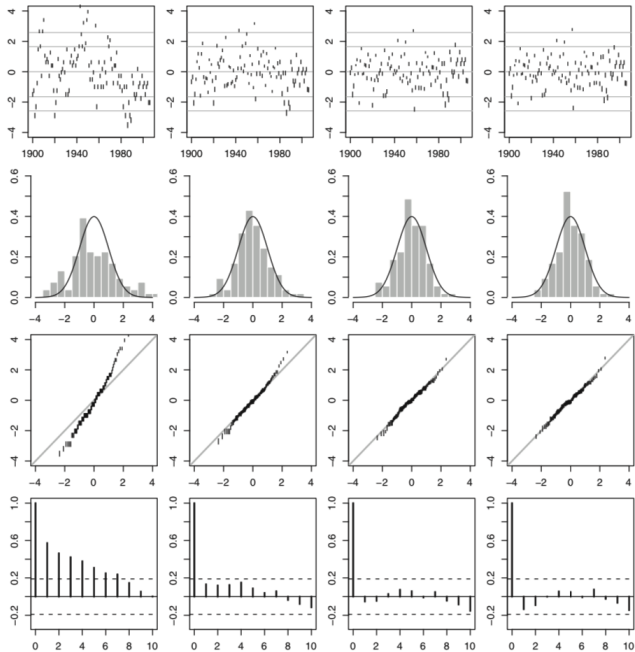
- ▶ Using the definitions of the forward and backward vectors:

$$f_{X_t}(x) \propto \alpha_{t-1} \mathbf{\Gamma} \mathbf{P}(x) \beta'_t$$

# HMM pseudo-residuals

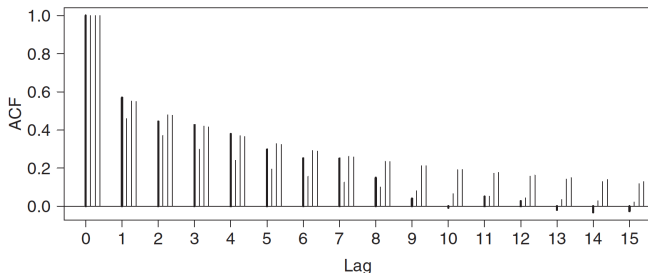
- ▶ Use forward-backward algorithm to calculate vectors  $\alpha_t$  and  $\beta_t$
- ▶ Compute distribution of  $X_t$  given all other observations
- ▶ Compute pseudo-residual  $z_t$  for each  $x_t$
- ▶ This can be used to detect outliers
- ▶ Can also be used to check validity of a model
- ▶ Plot distribution of pseudo-residuals vs Normal distribution
- ▶ Q-Q plot of observed (y-axis) vs expected (x-axis)
- ▶ ACF of pseudo-residuals

# Model checking for earthquakes dataset



# Observed vs expected ACF

- ▶ The observed ACF in the data can be compared to the ACF expected under the HMM
- ▶  $\text{Corr}(X_t, X_{t+k})$  can be computed analytically or simulated
- ▶ In the earthquake example, ACF of real data (bold) vs HMM with  $m = 1, 2, 3$  states:



# Conclusions

- ▶ Model selection is used to choose the best model amongst a set of candidates
- ▶ Comparing likelihoods directly is unfair since models with more parameters (eg HMM states) will always have better likelihoods
- ▶ AIC and BIC criteria penalize more complex models to make comparison more fair
- ▶ If all candidates are bad, the least bad is selected but this might still be very bad
- ▶ Model checking assesses the goodness of fit of a model (without comparison to other models)
- ▶ We can compute pseudo-residuals for a HMM and use them to do model checking
- ▶ So far we have considered HMM only within a classical statistics framework
- ▶ In our next lecture we will see how to use HMM in a Bayesian framework