# Multiple covariate distance sampling (MCDS)

• Aim: Model the effect of additional covariates on detection probability, in addition to distance, while assuming probability of detection at zero distance is 1

#### • References:

- Marques (F) and Buckland (2004) Covariate models for the detection function. Chapter 3 in Buckland *et al.* (eds). Advanced Distance Sampling.
- Marques (T) et al. (2007) Improving estimates of bird density using multiple covariate distance sampling.
  The Auk 127: 1229-1243.
- Section 5.3 of Buckland et al. (2015) Distance Sampling: Methods and Applications





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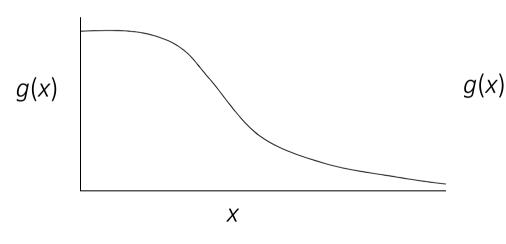


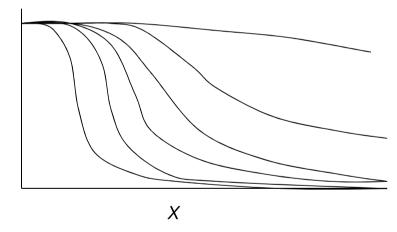


### Why additional covariates?

In conventional distance sampling (CDS) analysis all factors affecting detectability, except distance, are ignored

In reality, many factors may affect detectability





Sources of heterogeneity:

Object : species, sex, cluster size

Effort: observer, habitat, weather





#### Examples of heterogeneity 1

Effect of time of day on Rufous Fantail birds in Micronesia (point transects). Ramsey et. al. 1987. Biometrics 43:1-11

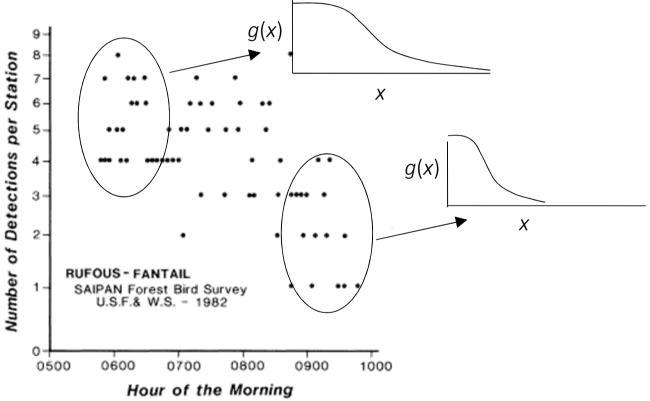




Figure 1. Station counts of Rufous Fantails on Saipan appear higher in the early morning hours than in the late morning (n = 64, r = -.60).



#### Examples of heterogeneity 2

Effect of sea state (and other covariates) on sea turtles in the Eastern Tropical Pacific (shipboard line transects). Beavers and Ramsey, 1998, J. Wildl. Manage. 63: 948-957

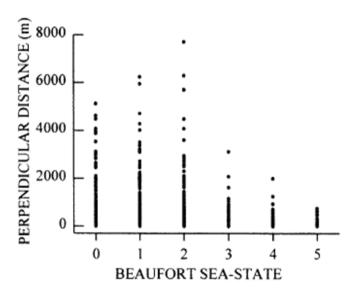


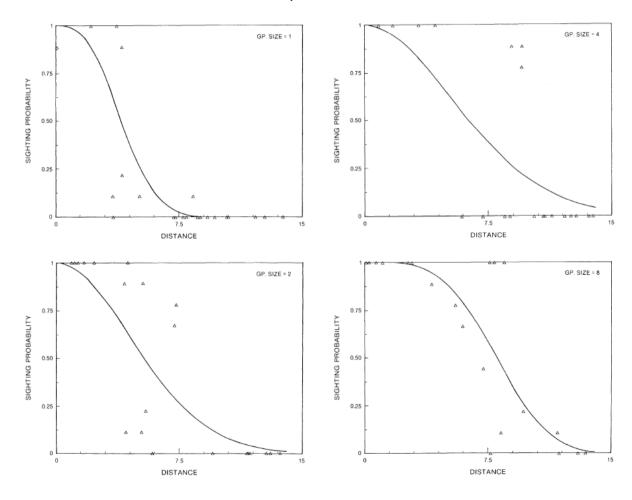
Fig. 2. Covariates of air temperature, sea surface temperature, and Beaufort sea-state plotted against unadjusted, ungrouped perpendicular sighting distances (m) of sea turtles in the eastern tropical Pacific, 1989–90.





#### Examples of heterogeneity 3

Effect of cluster size on beer can detectability. Otto and Pollock, 1990, Biometrics 46: 239-245







## Why worry about heterogeneity?

In CDS, we use models that are pooling robust, so why worry about heterogeneity?

- Pooling robustness works for all but extreme levels of heterogeneity
- Potential bias if density is estimated at a 'lower level' than detection function (e.g. density by geographic region, detection function global)
- Could potentially increase precision of detection function estimate
- Interest in sources of heterogeneity in their own right (e.g. group size)





## Dealing with heterogeneity

#### Stratification

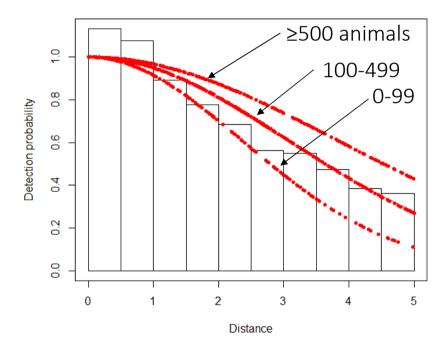
Requires estimating separate detection function parameters for each stratum,

often not possible due to lack of data

Model as covariates in detection function

Allows a more parsimonious approach:

- can model effect of numerical covariates
- can 'share information' about detection function shape between covariate levels







### Multiple covariate models Recap of CDS models

g(x) = Pr[animal at distance x is detected]

$$= k(x) \left[ 1 + \sum_{j=1}^{m} a_j p_j(x_s) \right] / c$$

Key function

*j*<sup>th</sup> series adjustment term

Scaling constant to ensure g(0) = 1





## CDS models continued

Key functions

Shape parameter

Hazard rate 
$$k(x) = 1 - \exp\left[-\left(\frac{x}{\sigma}\right)^{-b}\right]$$

Half-normal 
$$k(x) = \exp\left(\frac{-x^2}{2\sigma_x^2}\right)$$

Uniform

$$k(x) = 1$$

Scale parameter

Series adjustments

Cosine  $\cos(j\pi x_s)$ 

Polynomial  $x_s^j$ 

Hermite poly.  $H_i(x_s)$ 

 $x_s$  are scaled distances





## Modelling with covariates

 $g(x, \mathbf{z}) = \text{Pr}[\text{animal at distance } x \text{ and covariates } \mathbf{z} \text{ is detected}]$ 

Assume the covariates affect the *scale* of the key function, not its *shape*. So choose key functions with a scale parameter

Let 
$$\sigma(z) = \exp\left(\beta_0 + \sum_{j=1}^{J} \beta_j z_j\right)$$

e.g. Hazard rate 
$$k(x, z) = 1 - \exp \left[ -\left(\frac{x}{\sigma(z)}\right)^{-b} \right]$$

Half normal 
$$k(x, z) = \exp\left(\frac{-x^2}{2\sigma(z)^2}\right)$$





## Modelling with covariates

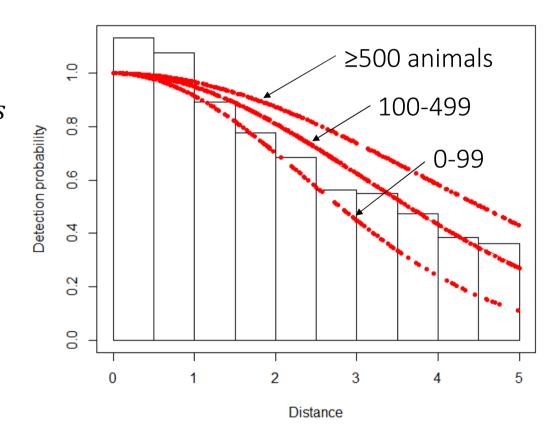
Example: Dolphin tuna vessel data

Model: half-normal, with no adjustments

Covariate: cluster size as factor (3 levels), s

$$g(x,s) = exp\left(\frac{-x^2}{2\sigma(s)^2}\right)$$

$$\sigma(s) = exp(\beta_0 + \beta_1 s)$$







## Estimating abundance without covariates using Horvitz-Thompson estimator

$$\widehat{N} = \sum_{i=1}^{n} \frac{1}{Pr[animal\ included]} = \sum_{i=1}^{n} \frac{1}{\left[\frac{2wL\widehat{P}_{a}}{A}\right]} = \frac{nA}{2wL\widehat{P}_{a}}$$

Recall that f(x) = pdf of observed x's =  $\frac{g(x)}{\int g(x)dx} = \frac{g(x)}{\mu} = \frac{g(x)}{wP_a}$ 

Remember: x's are the distances and  $P_a = {}^{\mu}/_{w}$ 

Because g(0)=1 by assumption, then f(0) =  $g(0)/\mu$  =  $1/\mu$  =  $1/wP_a$ 

So 
$$\widehat{N} = \frac{nA}{2wL\widehat{P}_a} = \frac{nA}{2L}.\widehat{f}(0)$$





## Estimating abundance with covariates

$$\widehat{N} = \sum_{i=1}^{n} \frac{1}{Pr[animal\ included]} = \sum_{i=1}^{n} \frac{1}{\left[\frac{2wL\widehat{P}_{a}(z_{i})}{A}\right]} = \frac{A}{2wL} \sum_{i=1}^{n} \frac{1}{\widehat{P}_{a}(z_{i})}$$

Now 
$$f(x|\mathbf{z}) = \frac{g(x,\mathbf{z})}{\int g(x,\mathbf{z})dx} = \frac{g(x,\mathbf{z})}{\mu(\mathbf{z})} = \frac{g(x,\mathbf{z})}{wP_a(\mathbf{z})}$$

Because  $g(0,\mathbf{z})=1$  by assumption, then  $f(0|\mathbf{z})=\frac{g(0,\mathbf{z})}{\mu(\mathbf{z})}=\frac{1}{\mu(\mathbf{z})}=\frac{1}{wP_a(\mathbf{z})}$ 

So

$$\widehat{N} = \frac{A}{2wL} \sum_{i=1}^{n} \frac{1}{\widehat{P}_a(0|\mathbf{z}_i)} = \frac{A}{2L} \sum_{i=1}^{n} \widehat{f}(0|\mathbf{z}_i)$$

Note similarity to CDS estimator





#### MCDS in Distance

In ds command, specify covariates in formula argument

ds(data, key, formula)

#### Covariate type:

- Factor covariates classify the data into distinct classes or levels. Can be numerical or text. One parameter per factor level.
- Non-factor (i.e., continuous) covariates must be numerical (integer or decimal). One parameter per covariate + 1 for the intercept.





#### Complications 1. Clustered populations

When cluster size is a covariate:

• Distance recognizes cluster size because column is called `size` (i.e. reserved word)

E.g. ds(data=Dolphin, key="hn", formula=~size)

$$\widehat{N}_{group} = \sum_{i=1}^{n} \frac{1}{Pr[group \ i \ included]} \qquad \widehat{N} = \sum_{i=1}^{n} \frac{size \ of \ group \ i}{Pr[group \ i \ included]}$$

Estimate of group size is given by 
$$\widehat{E}[s] = \frac{\widehat{N}}{\widehat{N}_{group}}$$





#### MCDS analysis guidelines

#### Choose covariates that are:

- independent of distance
- not strongly correlated with each other

#### Specifying the model:

- factor covariates generally harder to fit
- check convergence and monotonicity
- add only one covariate at a time
- where necessary, use starting values and bounds for parameters
- consider reducing the truncation distance, w, if more than 5% of the  $P_a(z_i)$  are <0.2, or if any are less than 0.1



