

Lecture - 18.

BLUE:- $\hat{\theta} = \frac{S^T C^{-1} x}{S^T C^{-1} S}$

$$\text{var}(\hat{\theta}) = \frac{1}{S^T C^{-1} S}$$

$x[n] = A + w[n]$

$w[n]$: zero mean uncorrelated noise with variance σ_w^2

$$\hat{A} = \frac{\sum_n \frac{x[n]}{\sigma_w^2[n]}}{\sum_n \frac{1}{\sigma_w^2[n]}}$$

$$\text{var}(\hat{A}) = \frac{1}{\sum_{n=1}^N \frac{1}{\sigma_w^2}}$$

Vector Parameter:-

$$\hat{\theta} = \vec{A} x$$

$$E[\hat{\theta}] = A E[x] = A \theta$$

$$E[x] = H \theta$$

$$AH = I$$

$$H = \begin{bmatrix} s[0] \\ s[1] \\ \vdots \\ s[N-1] \end{bmatrix}$$

$$\vec{C}_{\theta} = A^T \vec{C} A \quad \text{s.t.} \quad AH = I$$

\vec{C} : Covariance matrix of \vec{x}

$$\hat{\theta} = A x = (H^T C^{-1} H)^{-1} H^T C^{-1} x$$

$$C_{\theta} = (H^T C^{-1} H)^{-1}$$

$$x = H \theta + w$$

$$\hat{\theta} =$$

Gauss - Markov Theorem:-

If the data are of general linear model

$$\text{form } \vec{x} = \vec{H} \vec{\theta} + \vec{w} \quad \left[\begin{array}{l} \vec{H}: N \times P \text{ observation matrix} \\ \vec{\theta}: P \times 1 \text{ vector} \\ \vec{w}: N \times 1 \text{ noise with} \end{array} \right.$$

then the BLUE of $\hat{\theta}$ is $\hat{\theta} = (H^T C^{-1} H)^{-1} H^T C^{-1} x$ zero mean and covariance \bar{C}

$$\text{var}(\hat{\theta}_i) = [(H^T C^{-1} H)^{-1}]_{ii}$$

$$\text{cov}(\hat{\theta}) = (H^T C^{-1} H)^{-1}$$

Maximum - Likelihood Estimation (MLE):-

$$* \quad x[n] = A + w[n] \quad [w[n] \sim N(0, A)]$$

① MVUE using CRAB:-

$$P(x; A) = \frac{1}{(2\pi A)^{N/2}} \exp \left[-\frac{1}{2A} \sum_{n=0}^{N-1} (x[n] - A)^2 \right]$$

$$\frac{\partial \ln P(x; A)}{\partial A} = -\frac{N}{2A} + \frac{1}{A} \sum_n (x[n] - A) + \frac{1}{2A^2} \sum_n (x[n] - A)^2$$

$$\stackrel{?}{=} I(\theta)(g(x) - \theta) \stackrel{?}{=} I(A)(\hat{A} - A)$$

$$\text{var}(\hat{A}) \geq B \quad [B = ?]$$

(i) Using sufficient statistic

$$\frac{1}{A} \sum_n (x[n] - A)^2 = \frac{1}{A} \sum_n x^2[n] - 2N\bar{x} + NA$$

$$p(x; A) = \frac{1}{(2\pi A)^{N/2}} \exp\left[-\frac{1}{2} \left(\frac{1}{A} \sum_n x^2[n] + NA\right)\right]$$

$\underbrace{\quad}_{g(\sum_n x^2[n], A)} \quad \underbrace{\exp[N\bar{x}]}_{h(x)}$

$$T(x) = \sum_n x^2[n]$$

Assuming $T(x)$ is complete

$$(i) \quad E[g(T(x))] = A$$

$$E\left[g\left(\sum_n x^2[n]\right)\right] = A$$

$$\begin{aligned} E\left[\sum_n x^2[n]\right] &= N E[x^2[n]] \\ &= N \left[\text{var}(x[n]) + (E[x[n]])^2 \right] \\ &= N [A + A^2] \neq A \end{aligned}$$

$$(ii) \quad E(\hat{A} | T(x)) \quad [\hat{A} = \text{unbiased estimator}]$$

$$E \hat{A} = x[0]$$

$$E(x[0] | \sum_n x^2[n])$$

For MLR:-

If $N \rightarrow \infty : E(\hat{A}) \rightarrow A : \text{Var}(\hat{A}) \rightarrow 0$

$\hat{\theta} = \underset{\theta}{\text{arg max}} P(x; \theta)$ for x fixed.

$$\# \quad P(x; A) = \frac{1}{(2\pi A)^{N/2}} \exp \left[-\frac{1}{2A} \sum_n (x[n] - A)^2 \right]$$

$$\frac{\partial \ln P(x; A)}{\partial A} = -\frac{N}{2A} + \frac{1}{A} \sum_n (x[n] - A) + \frac{1}{2A^2} \sum_n (x[n] - A)^2 \rightarrow 0$$

$$\frac{ax^2 + bx + c = 0}{x = \frac{-b \pm \sqrt{b^2 - 4ac}}{2a}}$$

$$\hat{A}^2 + \hat{A} - \frac{1}{N} \sum_n x^2[n] = 0$$

$$\hat{A} = -\frac{1}{2} \pm \sqrt{\frac{1}{N} \sum_n x^2[n] + \frac{1}{4}}$$

$$\hat{A} = -\frac{1}{2} + \sqrt{\frac{1}{N} \sum_n x^2[n] + \frac{1}{4}} \quad \text{to make } A > 0$$

$$\# \quad x[n] = A + w[n] \quad [w[n] \sim \mathcal{N}(0, \sigma^2)]$$

$$P(x; A) = \frac{1}{(2\pi\sigma^2)^{N/2}} \exp \left[-\frac{1}{2\sigma^2} \sum_{n=0}^{N-1} (x[n] - A)^2 \right]$$

$$\frac{\partial \ln P(x; A)}{\partial A} = \frac{1}{\sigma^2} \sum_n (x[n] - A) \rightarrow 0$$

$$\hat{A} = \frac{1}{N} \sum_n x[n]$$

* If efficient exists, the MLE will produce it.

Asymptotic Property of the MLE:

If the PDF $p(x; \theta)$ satisfies some regularity conditions, then the MLE of the unknown parameter θ is "asymptotically distributed according to"

$$\hat{\theta} \sim N(\theta, I^{-1}(\theta))$$

$I(\theta)$: Fisher information evaluated at θ .

$$\frac{\partial \ln p(x; \theta)}{\partial \theta} = I(\theta)(g(x) - \theta)$$

$$\hat{\theta}_{MLE} = g(x)$$

$$I(\theta)(g(x) - \theta) = 0$$

$$\hat{\theta}_{MLE} = g(x) = \hat{\theta}_{MLE}$$