Assignment 3

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https://github.com/diya-goyal-29/AI1103/blob/main/Assignment%203/Assignment%203.tex

1 CSIR-UGC-NET-Exam-JUNE-2015 Question 60 :

Let $X_1, X_2, ..., X_n$ be independent and identically distributed random variables having an exponential distribution with mean $\frac{1}{\lambda}$.Let $S_n = X_1 + X_2 + \cdots + X_n$ and $N = inf\{n \ge 1 : S_n > 1\}$. Then Var(N) is

- 1) 1
- λ
- 3) λ^2
- 4) ∞

2 Solution:

$$Var(N) = E[N^2] - E[N]^2$$
 (2.0.1)

Replacing n with n+1

Definition 2.1 (infimum). N is the greatest lower bound of the subset S of T (where T is the set of all natural numbers and S is the set of all natural numbers such that $S_n \ge 1$).

$$(N = inf\{n \ge 1 : S_n > 1\})$$

The event $\{N \le n\}$ implies $\{S_{n+1} > 1\}$ (or equivalently, $\{N > n\}$ implies $\{S_{n+1} \le 1\}$).

Lemma 2.1. [Tail Sum formula]: A random variable X which takes values only in N, then $E[X] = \sum_{k=1}^{\infty} Pr(X)$

Proof.

$$E[X] = \sum_{k=1}^{\infty} kPr(X=k)$$
 (2.0.2)

$$= \sum_{k=1}^{\infty} \sum_{n=1}^{k} Pr(X=k)$$
 (2.0.3)

$$= \sum_{n=1}^{\infty} \sum_{k=n}^{\infty} Pr(X=k)$$
 (2.0.4)

$$=\sum_{n=1}^{\infty} Pr(X \ge n) \tag{2.0.5}$$

Corollary 2.1.1. Summation of independent and identical exponential random variables follows Erlang Distribution.

P.D.F [Erlang distribution]:

$$f(x) = \frac{\lambda^k x^{k-1} e^{-\lambda x}}{(k-1)!}$$
 (2.0.6)

where, k is the shape λ is the rate

 $S_n = \text{Erlang}(n, \lambda)$

P.D.F of
$$S_n = f(x)_{S_n} = \frac{\lambda^{n+1} x^n e^{-\lambda x}}{(n)!}$$
 (2.0.7)

$$Pr(S_n \le 1) = \int_0^1 \frac{\lambda^{n+1} x^n e^{-\lambda x}}{(n)!} dx \qquad (2.0.8)$$

By the tail-sum formula (lemma 2.1), we can thus get the expectation of N as

$$E[N] = \sum_{n=0}^{\infty} Pr(N > n) = \sum_{n=0}^{\infty} Pr(S_n \le 1) \quad (2.0.9)$$

 $Var(N) = E[N^2] - E[N]^2$ (2.0.23)

 $= \lambda^2 + \lambda - (\lambda)^2 \tag{2.0.24}$

 $= \lambda \tag{2.0.25}$

Substituting the value

Thus Option 2 is correct.

$$E[N] = \sum_{n=0}^{\infty} \int_{0}^{1} \frac{\lambda^{n+1} x^{n} e^{-\lambda x}}{(n)!} dx$$
 (2.0.10)

$$= \int_0^1 \sum_{n=0}^\infty \frac{\lambda^{n+1} x^n e^{-\lambda x}}{(n)!} dx$$
 (2.0.11)

$$e^{x\lambda} = \sum_{n=0}^{\infty} \frac{(x\lambda)^n}{(n)!}$$
 (2.0.12)

$$E[N] = \lambda \int_0^1 e^{\lambda x} e^{-\lambda x} dx \qquad (2.0.13)$$

$$= \lambda \int_0^1 dx \tag{2.0.14}$$

$$= \lambda \tag{2.0.15}$$

By tail-sum formula (lemma 2.1) the expectation of N^2 is

$$E[N^2] = \sum_{n=0}^{\infty} Pr(N^2 \ge n)$$
 (2.0.16)

$$= \sum_{n=0}^{\infty} [(n+1)^2 - n^2] Pr(N \ge n) \qquad (2.0.17)$$

$$= \sum_{n=0}^{\infty} [2n+1] Pr(S_{n+1} \le 1)$$
 (2.0.18)

$$= \sum_{n=0}^{\infty} \int_{0}^{1} [2n+1] \frac{\lambda^{n+1} x^{n} e^{-\lambda x}}{(n)!} dx \qquad (2.0.19)$$

$$= \int_0^1 \sum_{n=0}^{\infty} [2n+1] \frac{\lambda^{n+1} x^n e^{-\lambda x}}{(n)!} dx \qquad (2.0.20)$$

$$= \int_0^1 2\lambda^2 x e^{\lambda x} e^{-\lambda x} + \lambda e^{\lambda x} e^{-\lambda x} dx \quad (2.0.21)$$

$$= \lambda^2 + \lambda \tag{2.0.22}$$

Substituting the values of E[N] and $E[N^2]$ in eq.(2.0.1)