

# Joint, conditional, and marginal distributions

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## Intended learning outcomes

- ▶ Apply definitions and theorems regarding joint, conditional, and marginal distributions.
- ▶ Recognize and explain Simpson's paradox

# Random vector

## Definition

An  $n$ -dimensional **random vector** is a function from a sample space  $\mathcal{S}$  to  $n$ -dimensional Euclidean space  $\mathcal{R}^N$  .

# Joint probability mass function

## Definition

Given a discrete bivariate random vector  $(X, Y)$ , the joint probability mass function (pmf) is defined by

$$f_{X,Y}(x, y) \triangleq P_{X,Y}(X = x, Y = y).$$

## Properties

A pmf  $f_{X,Y}(x, y)$  satisfies

$$f_{X,Y}(x, y) \geq 0 \quad \forall (x, y) \in \mathcal{R}^2 \quad \text{and} \quad \sum_{(x,y) \in \mathcal{R}^2} f_{X,Y}(x, y) = 1.$$

## Support

$f_{X,Y} : \mathcal{R} \times \mathcal{R} \rightarrow [0, 1]$  but we only defined  $P_{X,Y}$  for  $x \in \mathcal{X} \subseteq \mathcal{R}$  and  $y \in \mathcal{Y}$ .

### Definition

The **support** of a distribution  $f_X(x)$  is

$$\mathcal{X} = \{x : f_X(x) > 0\}.$$

Therefore,  $f_{X,Y} = 0$  for  $x \notin \mathcal{X}$  or  $y \notin \mathcal{Y}$ .

# Marginal probability mass function

## Definition

The marginal pmfs of random vector  $(X, Y)$  are defined by

$$f_X(x) \triangleq P_X(X = x) \quad f_Y(y) \triangleq P_Y(Y = y)$$

## Theorem

Given a discrete random vector  $(X, Y)$  with joint pmf  $f_{X,Y}(x, y)$ , the marginal pmfs of  $X \in \mathcal{X}$  and  $Y \in \mathcal{Y}$  are given by

$$f_X(x) = \sum_{y \in \mathcal{Y}} f_{X,Y}(x, y) \quad f_Y(y) = \sum_{x \in \mathcal{X}} f_{X,Y}(x, y)$$

This theorem follows from the law of total probability.

# Total law of probability

Given sample space  $\mathcal{S}$ ,  $\mathcal{A} \subseteq \mathcal{S}$ ,

$$P(\mathcal{A}) = \sum_i P(\mathcal{A} \cap \mathcal{B}_i)$$

where  $\mathcal{B}_1, \mathcal{B}_2, \dots \subseteq \mathcal{S}$  is a partition of  $\mathcal{S}$ , which is defined by

$$\mathcal{B}_i \cap \mathcal{B}_j = \emptyset \quad \forall i \neq j \quad \text{and} \quad \bigcup_i^{\infty} \mathcal{B}_i = \mathcal{S}$$

## Proof

It follows from set theory and Kolmogorov's probability axioms.

$$\mathcal{A} = \mathcal{A} \cap \mathcal{S} = \mathcal{A} \cap \left( \bigcup_i \mathcal{B}_i \right) = \bigcup_i \mathcal{A} \cap \mathcal{B}_i$$

$$P(\mathcal{A}) = P\left(\bigcup_i \mathcal{A} \cap \mathcal{B}_i\right) = \sum_i P(\mathcal{A} \cap \mathcal{B}_i) \quad (\text{additivity axiom}) \quad \blacksquare$$

## Proof: Marginal pmf

Define  $\mathcal{B}_y = \{s \in \mathcal{S} : Y(s) = y\}$ .

Since  $Y$  is a map from  $\mathcal{S}$  to  $\mathcal{Y}$ , there exists some  $y$  for every  $s \in \mathcal{S}$ . Then,  $\bigcup_{y \in \mathcal{Y}} \mathcal{B}_y = \mathcal{S}$ . Therefore,  $\mathcal{B}_1, \mathcal{B}_2, \dots$  is a partition of  $\mathcal{S}$ .

$$\begin{aligned} f_X(x) &= P_X(X = x) \\ &= P(\{s \in \mathcal{S} : X(s) = x\}) \\ &= \sum_{y \in \mathcal{Y}} P(\{s \in \mathcal{S} : X(s) = x\} \cap \mathcal{B}_y) \quad (\text{Law of total prob}) \\ &= \sum_{y \in \mathcal{Y}} P_{X,Y}(X = x, Y = y) = \sum_{y \in \mathcal{Y}} f_{X,Y}(x, y) \end{aligned}$$

The proof for  $f_Y(y)$  follows similarly as above. ■



# Marginalization as sweeping

# Joint probability density function

## Definition

Given a *continuous* random vector  $(X, Y)$ , a joint probability density function (pdf) is a function  $f_{X,Y}(x, y)$  such that, for every subset  $\mathcal{A} \subset \mathcal{R}^2$ ,

$$P((X, Y) \in \mathcal{A}) = \int \int_{\mathcal{A}} f_{X,Y}(x, y) dx dy.$$

The notation  $\int \int_{\mathcal{A}}$  means that the limits of integration are set so that the function is integrated over all  $(x, y) \in \mathcal{A}$ .

## Properties

A pdf  $f_{X,Y}(x, y)$  defined above also satisfies

$$f(x, y) \geq 0 \quad \forall (x, y) \in \mathcal{R}^2 \quad \text{and} \quad \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1.$$

# Marginal probability density function

## Theorem

Given a *continuous* random vector  $(X, Y)$  with joint pdf  $f_{X,Y}(x, y)$ , the marginal pdf of  $X$  and  $Y$  are given by

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy \quad x \in \mathcal{X},$$
$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx \quad y \in \mathcal{Y}.$$

# Conditional probability

## Definition

Given events  $A$  and  $B$ , if  $P(B) > 0$ , then

$$P(A | B) \triangleq \frac{P(A \cap B)}{P(B)}.$$

# Conditional probability distributions

## Definition

Given a discrete (or continuous) random vector  $(X, Y)$  with joint pmf (or pdf)  $f_{X,Y}(x, y)$  and marginal pmfs (or pdfs)  $f_X(x)$  and  $f_Y(y)$ , for any  $x$  such that  $f_X(x) > 0$ , the conditional pmf (or pdf) of  $Y$  given that  $X = x$  is defined by

$$f_{Y|X}(y | x) = \frac{f_{X,Y}(x, y)}{f_X(x)}.$$

Similarly, for any  $y$  such that  $f_Y(y) > 0$ ,

$$f_{X|Y}(x | y) = \frac{f_{X,Y}(x, y)}{f_Y(y)}.$$

## Conditioning as slicing

## Chain rule of probability

By re-arranging definition of conditional probability, we have

$$P(A_1 \cap A_2) = P(A_1 | A_2) P(A_2)$$

Of course, we can apply the definition of  $P(A_2 | A_1)$  as well:

$$P(A_1 \cap A_2) = P(A_2 | A_1) P(A_1).$$

## Chain rule of probability

If we have a sequence of  $J$  events,  $A_1, A_2, \dots, A_J$  in some *arbitrary order*, we can keep applying the definition of conditional probability.

$$\begin{aligned} P(A_1 \cap A_2 \cap \dots A_J) \\ &= P(A_1 \mid A_2, A_3, \dots, A_J) P(A_2 \cap A_3 \cap \dots A_J) \\ &= P(A_1 \mid A_2, A_3, \dots, A_J) P(A_2 \mid A_3 \cap \dots A_J) P(A_3 \cap A_4 \cap \dots A_J) \\ &= \dots \end{aligned}$$

This then gives the **chain rule**:

$$P\left(\bigcap_{j=1}^J A_j\right) = \prod_{j=1}^J P\left(A_j \mid \bigcap_{k=1}^{j-1} A_k\right).$$

This applies to probability distributions as well.



## Conditioning direction

Typically, we try to define the conditional probabilities based on our assumptions about the causal relationships. If we believe  $A$  causes  $B_1, B_2, \dots, B_J$ , then in our model, it would be easier to define  $P(B_1 | A), \dots, P(B_J | A)$ .

We can also choose to define  $P(A | B_1, B_2, \dots, B_J)$  instead, but this can make the derivations more difficult if it is inconsistent with the underlying causal relationship.

# Independence

## Definition

Given random vector  $(X, Y)$  with joint pmf (or pdf)  $f_{X,Y}(x, y)$  and marginal pmfs (or pdfs)  $f_X(x)$  and  $f_Y(y)$ ,  $X$  and  $Y$  are **independent** if and only if (iff), for every  $x \in \mathcal{X}$  and  $y \in \mathcal{Y}$ ,

$$f_{X,Y}(x, y) = f_X(x)f_Y(y).$$

If  $X$  and  $Y$  are independent, we write  $X \perp Y$ .

$X \perp Y \Rightarrow f_{Y|X}(y | x) = f_Y(y)$  for every  $x \in \mathcal{X}$  and  $y \in \mathcal{Y}$ .

## Lemma

Given random vector  $(X, Y)$ ,  $X$  and  $Y$  are independent if and only if (iff) there exists functions  $g(x)$  and  $h(y)$  such that, for every  $x \in \mathcal{X}$  and  $y \in \mathcal{Y}$ ,

$$f_{X,Y}(x, y) = g(x)h(y).$$

## Simpson's paradox

$$P(Y = 1 \mid X = x_1) > P(Y = 1 \mid X = x_2).$$

However, conditioned on  $Z = z$  for some  $z$ ,

$$P(Y = 1 \mid X = x_1, Z = z) < P(Y = 1 \mid X = x_2, Z = z).$$

## Example: Gender bias?

$Y$  represents undergraduate admission.  $X$  represents gender.  
 $Z$  represents department.

$$P(Y = 1 \mid X = x_1) > P(Y = 1 \mid X = x_2).$$

Table 1: Admission rates by gender

All	$x_1$	$x_2$
41%	<b>44%</b>	35%

$$P(Y = 1 \mid x_1, z) < P(Y = 1 \mid x_2, z), \quad z \in \{1, 2, 4, 6\}.$$

$$P(Y = 1 \mid x_1, z) > P(Y = 1 \mid x_2, z), \quad z \in \{3, 5\}.$$

Table 2: Admission rates by gender and department

Department	All	$x_1$	$x_2$
1	64%	62%	<b>82%</b>
2	63%	63%	<b>68%</b>
3	35%	<b>37%</b>	34%
4	34%	33%	<b>35%</b>
5	25%	<b>28%</b>	24%
6	6%	6%	<b>7%</b>
...			
All	41%	<b>44%</b>	35%

$$P(Y = 1 \mid x_1, z) < P(Y = 1 \mid x_2, z), \quad z \in \{1, 2, 4, 6\}.$$

$$P(Y = 1 \mid x_1, z) > P(Y = 1 \mid x_2, z), \quad z \in \{3, 5\}.$$

Table 3: Number of applicants by gender and department

Department	All	$x_1$	$x_2$
1	933	<b>825</b>	108
2	585	<b>560</b>	25
3	918	325	<b>593</b>
4	792	<b>417</b>	375
5	584	191	<b>393</b>
6	714	<b>373</b>	341
...			
All	12763	8442	4321

## Causal effects?

$$P(Y = 1 \mid X = x_1) > P(Y = 1 \mid X = x_2) \Rightarrow X \not\perp Y.$$

Suppose that we know that  $Y$  does *not* affect  $X$ .

Does this mean that  $X$  affects  $Y$ ?

# Summary

Symbolic logic  $\rightarrow$  Set theory  $\rightarrow$  Probability theory  
 $\rightarrow$  Measure theory  $\rightarrow$  Statistics  $\rightarrow$  Data modelling

Casella & Berger 2002, sections 4.1-4.2.

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