Probabilistic Graphical Models

IKHAN CHOI

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1. Introduction

1.1. Statistical model.

Definition 1.1. A *statistical model* is an approximation scheme for unknown joint probability distribution.

The general purpose of many statistical models is to estimate the joint probability distribution of several random variables. The joint probability distribution contains data about correlations among the random variables. For example, suppose that we have to obtain the most possible value of Y when given X = x, and we have already estimated the joint distribution function $f_{X,Y}$. Then, since the function $y \mapsto f_{X,Y}(x,y)$ describes the distribution of the random variable Y|X=x, what we want to find can be defined reasonably as

$$\hat{y} = \arg\max_{y} f_{X,Y}(x,y).$$

Example 1.1. A random field, which we have not defined yet, is a way to represent several random variables together with their dependecies. Therefore, a paramterized random field gives a statistical model. In this case, training means an approximating process to find the best parameter, usually written as β .

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1.2. Random fields.

Definition 1.2 (Random field). A random field is a set of random variables parametrized by a topological space or a (directed or undirected) graph.

Definition 1.3. In this note, a term *random field* or *network* will be used to refer to random fields on a graph.

Example 1.2 (Markov chain). Define a graph G = (V, E) and S such that:

$$V = \mathbb{Z}_{t \geq 0},$$

$$E = \{(t, t+1)\}_{t \in V},$$

$$S = \text{a finite set.}$$

An element $t \in V$ denotes the time t. Then, the set of S-valued random variables $\{X_t\}_{t\in V}$ indexed by V defines a random field.

The Markov property is given by

$$X_t \perp \!\!\! \perp X_s \mid X_{t-1}$$

for s < t. Since S is finite, alternatively we may write it by

$$\Pr(X_t = x_t \mid X_{t-1} = x_{t-1}) = \Pr(X_t = x_t \mid X_{t-1} = x_{t-1}, \dots, X_0 = x_0)$$

for all $x_i \in S$ with $i = 0, \dots, t$. With abusing notation, it is often written as

$$p(x_t|x_{t-1}) = p(x_t|x_{t-1}, \cdots, x_0).$$

Example 1.3 (Boltzman distribution). Let X be a S-valued random field on a graph G = (V, E) such that:

$$V = \{n\}_{n=1}^{N},$$

$$E = \emptyset,$$

$$S = \text{a discretized space}$$

for a large natural number N. Also, suppose we have energy function $E: S \to \mathbb{R}$ and parametrized family of functions $\phi: S \times (0, \infty) \to \mathbb{R}$ defined by

$$\phi(s_n; \beta) = e^{-\beta E(s_n)}.$$

The functions ϕ are called *Boltzmann factors* and the paramter β is sometimes called *coldness*. As a typical example, we may choose

$$S = \mathbb{Z}_x^3 \times \mathbb{Z}_p^3$$
, $E(x,p) = \frac{\|p\|^2}{2m}$ (*m*: mass of particle)

for the Maxwell-Boltzmann statistics of ideal gas. In this case, the set V is the set of ideal gas particles, and the space of states S embodies the discretized phase space.

At each particle $n \in V$ is attached an S-valued random variable X_n . Our ultimate goal is to describe the joint probability distribution of X_n 's; equivalently, the distribution of a S^N -valued random variable $X = (X_1, \dots, X_N)$. Elements of S^N are called microstates.

First, we give the probability distribution of an individual random variable X_n . The assumption for Boltzmann factors states that if given β then the probability for a particle n to be in a state $s_n \in S$ is proportional to the Boltzmann factor:

$$\Pr(X_n = s_n) \propto_{\beta} \phi(s_n; \beta) = e^{-\beta E(s_n)}$$

for each state $s_n \in S$ and a fixed particle n. Thus, we can write

$$\Pr(X_n = s_n) = \frac{\phi(s_n; \beta)}{\sum_{s'_n \in S} \phi(s'_n; \beta)} =: \frac{\phi(s_n; \beta)}{Z_n(\beta)}.$$

Next, consider the entire random field $X:\Omega\to S^N$. If we assume the independence of X_n 's, then we get the disjoint probability distribution

$$\Pr(X = s) = \frac{\phi(s; \beta)}{\sum_{s' \in S^N} \phi(s'; \beta)} =: \frac{\phi(s; \beta)}{Z(\beta)},$$

where $Z(\beta) = \prod_{n=1}^{N} Z_n(\beta)$ and the definitions of ϕ and E are extended as follows:

$$\phi(s;\beta) = \prod_{n=1}^{N} \phi(s_n;\beta) = e^{-\beta \sum_{n=1}^{N} E(s_n)} = e^{-\beta E(s)}.$$

The denominator $Z: \mathbb{R}_{\beta>0} \to \mathbb{R}^+$ is called the partition function.

1.3. **Notations.** We review the definition of probability distribution. In this subsection, let $(\Omega, \mathcal{F}, Pr)$ be a probability space and (S, \mathcal{S}) a measurable space.

Definition 1.4. Let $X : \Omega \to S$ be a random variable. The *probability distribution* of X is the pushforward measure $X_* \Pr$ on S defined as $X_* \Pr(A) = \Pr(X^{-1}(A))$ for measurable $A \in S$. We often also write $\Pr(X \in A)$ for $X_* \Pr(A)$.

Definition 1.5. Let S be a Lebesgue measurable subset of a Euclidean space. Let $X: \Omega \to S$ be a random variable. If the probability distribution of X is absolutely continuous with respect to the Lebesgue measure, then we have the Radon-Nikodym derivative. We call the derivative *probability distribution function* or shortly pdf of X.

Definition 1.6. Furthermore, let S be finite or countable with discrete measurable structure. Let $X: \Omega \to S$ be a random variable. We call a function $p(s) = \Pr(X = s)$ the *probability mass function* or shortly pmf of X.

We will not consider a random variable that has neither pdf nor pmf. If we do not restrict the regularity of distributions, we must resolve intractable technical issues. Also note that pdf is more general notion than pmf.

We mainly deal with several random variables and their joint distribution. The notations then become dirty and equations get longer due to the number of random variables, so we introduce alternative notations. Suppose $X = \{X_n\}_{n=1}^N$ is a set of random variables and let (S_n, S_n) be the codomain of X_n . The pdf and pmf of an individual random variable X_i will be denoted by $f(x_i)$ and $p(x_i)$ using lower case alphabets for the random variables. Moreover, we will admit the abuse of notations more generally without precise definitions, but for examples, just note that

- $f(x_1)$ is a pdf on S_1 ,
- $f(x_1, x_2) = f(x_{\{1,2\}})$ is a pdf on $S_1 \times S_2$,

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- f(x) is a pdf on $S = \prod_{n=1}^{N} S_n$ (joint distribution), $f(x_2|x_1)$ is a pdf on S_2 parametrized by S_1 , or more generally S_1 ,
- $f(x_A|x_B)$ is a pdf on $\prod_{n\in A} S_n$ parametrized by $\prod_{n\in B} S_n$.

Every notation above in fact can be recognized as a real-valued function on S via projections. For instance, the function $f(x_1): S_1 \to \mathbb{R}_{>0}$ can have extended domain like $f(x_1): S \to S_1 \to \mathbb{R}_{>0}$ that is constant for all variables except X_1 .

According to this notation, our goal is always to estimate the joint distribution f(x)of a random field.

2. Bayesian networks

2.1. Factorization of probability. Recall the definition of descendants/ancestors and children/parents of nodes.

Definition 2.1 (Bayesian network). Let X be a random field on a directed acyclic graph G. We say X is a Bayesian network or satisfies the local Markov property if

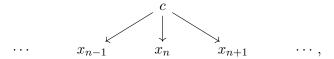
$$X_v \perp \!\!\! \perp X_{V \setminus de(v)} \mid X_{pa(v)}.$$

Definition 2.2 (Factorization). Let X be a random field on a directed acyclic graph G. We say X factorizes over G if

$$f(x) = \prod_{v \in V} f(x_v | x_{pa(v)}).$$

Theorem 2.1. A random field X on a directed acyclic graph G is a bayesian network if and only if it factorizes over G.

Example 2.2 (NBC, Naive Bayesian Classifier). Consider a random field on the following directed acyclic graph(DAG):



where $n = 1, \dots, N$. The joint distribution is

$$p(c, x) = p(c) \prod_{n=1}^{N} p(x_n|c).$$

It is represented by factors p(c) and $p(x_n|c)$'s; the conditional probability p(c|x)p(c,x)/p(x) can be computed if we assume that the input feature x has been given and the probability p(c), $p(x_n|c)$ had been estimated.

If we view $p(c) \in [0,1]^{S_c}$ and $p(x_n|c) \in [0,1]^{S_x \times S_c}$ as parameters which parametrize the joint distribution $p(c,x) \in [0,1]^{S_c \times S_x^N}$, then we can say the NBC defines a family of distribution functions for approximating p(x,c) that is parametrized by $(|S_c|-1+$ $N(|S_x|-1)|S_c|$ -dimensional parameter, which is much smaller than $|S_c||S_x|^N-1$. The NBC provides with a statistical model via "conditional parametrization".

Example 2.3 (HMM, Hidden Markov Model). Consider a random field on the following DAG:

The joint distribution is

$$p(y,x) = p(x_0|y_0)p(y_0)\prod_{t=1}^{T-1} p(x_t|y_t)p(y_t|y_{t-1}).$$

The terms $p(y_0)$, $p(y_t|y_{t-1})$, and $p(x_t|y_t)$ are called *start probability*, transition probability, and *emmission probability* respectively, and they play a role of parameters.

Example 2.4 (MEMM, Maximaum entropy Markov Model). Consider a random field on the following DAG:

$$y_0 \longrightarrow y_1 \longrightarrow \cdots \longrightarrow y_{T-2} \longrightarrow y_{T-1}$$

$$\uparrow \qquad \qquad \uparrow \qquad \qquad \uparrow$$

$$x_0 \qquad x_1 \qquad \qquad x_{T-2} \qquad x_{T-1}.$$

The joint distribution is

$$p(y,x) = p(x_0)p(y_0|x_0) \prod_{t=1}^{T-1} p(x_t)p(y_t|y_{t-1},x_t).$$

The MEMM introduces exponential modeling

$$p(y_i|y_{i-1},x_i) \propto_{y_{i-1},x,\beta} e^{-\beta \cdot H(y_i,y_{i-1},x_i)}$$

with a appropriately designed energy function $H: S_y^2 \times S_x \to \mathbb{R}^d$. The normalizing constant Z depends only on y_{i-1} and x_i , hence

$$p(y_i|y_{i-1},x_i) = \frac{e^{-\beta \cdot H(y_i,y_{i-1},x_i)}}{Z(y_{i-1},x_i;\beta)}.$$

Estimate of the joint distribution is done by adjusting the d-dimensional parameter β .

3. Markov networks

In this section, we discuss random fields on undirected graphs. As in Bayesian networks, we expect the whole joint probability to be factorized into a computable form with smaller dimensional parameters. In this factorization, a modified version of Markov property on the graph that is different from the classical Markov process is required.

It would seem that the initial and general definitions are too abstract that they are hardly applied in practical problems..... But...

Definition 3.1 (Markov properties on a graph). There are three Markov properties:

(1) Pairwise Markov property

Definition 3.2 (Markov network). Let X be a random field on a undirected graph G. Markov networks are sometimes called MRF, Markov random field.

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Example 3.1 (Ising model).

Example 3.2 (CRF, Conditional Random Field). Consider a network with a graph G such that vertices are divided into two classes.

$$p(y|x) \propto_{x,\beta} e^{-\beta \cdot H(y,x)}$$
.

For the Viterbi algorithm, the energy function H is frequently taken to have the form

$$H(y,x) = \sum_{i} E(y_i, y_{i-1}, x).$$

4. Neural networks

Probabilistic graphical models provide effective explanations of the neural networks, but neural networks are not confined only to graphical models.

Definition 4.1 (Neural network). *Neural network* cannot be defined mathematically. It indicates statistical models that can solve problems with a collection of artificial neurons by adjusting connection strength among them.

Example 4.1 (MLP, Multi-layer Perceptron).

Example 4.2 (RNN, Recurrent Neural Network).

4.1. Maximum likelihood estimate.

Definition 4.2. Let f be a distribution function on a measure space X. Let $\{f_{\theta}\}_{\theta}$ be a parametrized family of distribution functions on X. The *likelihood* $L_n(\theta): \Omega^n \to \mathbb{R}_{\geq 0}$ for a fixed parameter θ is a random variable defined by

$$L_n(\theta) := \prod_{i=1}^n f_{\theta}(x_i)$$

where $\{x_i\}_i$ is a family of i.i.d. X-valued random variables with a distribution f.

The objective of the likelihood function is to find θ such that f_{θ} approximates the unknown distribution f. Write

$$\frac{1}{n}\log L_n(\theta) = \frac{1}{n}\sum_{i=1}^n \log f_{\theta}(x_i).$$

By the law of large numbers, $\frac{1}{n} \log L_n(\theta)$ converges to a constant function

$$\mathbb{E}(\log f_{\theta}(x)) = \int_{X} f \log f_{\theta}$$

in measure as $n \to \infty$. This constant function is exactly what we call *cross entropy*.

The Kullback-Leibler divergence is a kind of asymmetric distance function defined from the difference with cross entropy

$$D_{KL}(f||f_{\theta}) := \int_{X} f \log f - \int_{X} f \log f_{\theta}.$$

It is proved to be always nonnegative by the Jensen inequality:

$$\int_{X} f \log f_{\theta} - \int_{X} f \log f = \int_{X} f \log \frac{f_{\theta}}{f} \le \log \left(\int_{X} f \frac{f_{\theta}}{f} \right) = 0.$$

Here, we exclude the region f=0 from the integration region. Then, we can say, bigger $L_n(\theta)$ is, closer f_{θ} and f are.