# A Bayesian Tutorial for Data Assimilation: Christopher K. Wikle & L. Mark Berliner

Damon McDougall

University of Warwick Mathematics Institute

March 16, 2011

- Introduction
- Bayesian Inference

- Introduction
- Bayesian Inference
- Sequential Approaches

- Introduction
- Bayesian Inference
- Sequential Approaches
- Monte Carlo Sampling and Data Assimilation

- Introduction
- Bayesian Inference
- Sequential Approaches
- Monte Carlo Sampling and Data Assimilation
- Summary

• What is DA?

- What is DA?
  - Various definitions

- What is DA?
  - Various definitions
  - Their definition: "DA is an approach for fusing data (observations)
    with prior knowledge to obtain an estimate of the distribution of the
    true state of the process"

- What is DA?
  - Various definitions
  - Their definition: "DA is an approach for fusing data (observations)
    with prior knowledge to obtain an estimate of the distribution of the
    true state of the process"
- What do we need?

- What is DA?
  - Various definitions
  - Their definition: "DA is an approach for fusing data (observations)
    with prior knowledge to obtain an estimate of the distribution of the
    true state of the process"
- What do we need?
  - Data model (measurement model)

- What is DA?
  - Various definitions
  - Their definition: "DA is an approach for fusing data (observations)
    with prior knowledge to obtain an estimate of the distribution of the
    true state of the process"
- What do we need?
  - Data model (measurement model)
  - Process model

- What is DA?
  - Various definitions
  - Their definition: "DA is an approach for fusing data (observations)
    with prior knowledge to obtain an estimate of the distribution of the
    true state of the process"
- What do we need?
  - Data model (measurement model)
  - Process model
- Why do we need these models?

- What is DA?
  - Various definitions
  - Their definition: "DA is an approach for fusing data (observations)
    with prior knowledge to obtain an estimate of the distribution of the
    true state of the process"
- What do we need?
  - Data model (measurement model)
  - Process model
- Why do we need these models?
  - We need them for Bayesian Inference

• What is Bayesian Inference and how do we do it?

- What is Bayesian Inference and how do we do it?
- There are 3 steps

- What is Bayesian Inference and how do we do it?
- There are 3 steps
  - Formulate 'full' probability model p(x, y)

- What is Bayesian Inference and how do we do it?
- There are 3 steps
  - Formulate 'full' probability model p(x, y)
  - 2 Find p(x|y)

- What is Bayesian Inference and how do we do it?
- There are 3 steps
  - 1 Formulate 'full' probability model p(x, y)
  - 2 Find p(x|y)
  - Evaluate fit and validity of model

- What is Bayesian Inference and how do we do it?
- There are 3 steps
  - Formulate 'full' probability model p(x, y)
  - 2 Find p(x|y)
  - Evaluate fit and validity of model
- Conditional probability gives:

$$p(x,y) = p(x|y)p(y) = p(y|x)p(x)$$

- What is Bayesian Inference and how do we do it?
- There are 3 steps
  - Formulate 'full' probability model p(x, y)
  - 2 Find p(x|y)
  - Evaluate fit and validity of model
- Conditional probability gives:

$$p(x,y) = p(x|y)p(y) = p(y|x)p(x)$$

Using this we obtain Bayes' Rule:

$$p(x|y) = \frac{p(y|x)p(x)}{p(y)}$$
 for  $0 < p(y) < \infty$ 



• p(x) is our prior belief or *prior distribution* 

- p(x) is our prior belief or prior distribution
- p(y|x) is the data (or measurement) model. These are our observations

- p(x) is our prior belief or prior distribution
- p(y|x) is the data (or measurement) model. These are our observations
- Using these and Bayes' Rule (on the previous slide) we can find p(x|y), the state of the system *given* the observations

• Let's say our prior knowledge is:  $X \sim \mathcal{N}(\mu, \tau^2)$ 

- Let's say our prior knowledge is:  $X \sim \mathcal{N}(\mu, \tau^2)$
- And we have *n* independent observations,  $Y = (Y_1, ..., Y_n)^T$ , where

$$Y_i|X \sim \mathcal{N}(x,\sigma^2)$$

- ullet Let's say our prior knowledge is:  $X \sim \mathcal{N}(\mu, au^2)$
- And we have *n* independent observations,  $Y = (Y_1, \dots, Y_n)^T$ , where

$$Y_i|X \sim \mathcal{N}(x,\sigma^2)$$

Equation for the prior is:

$$p(x) = \frac{1}{\sqrt{2\pi\tau^2}} \exp\left(-\frac{1}{2\tau^2}(x-\mu)^2\right)$$



As a result of independent observations, we have

• As a result of independent observations, we have

$$p(y|x) = \prod_{i=1}^{n} \left(\frac{1}{\sqrt{2\pi\sigma^2}}\right) \exp\left(-\frac{1}{2\sigma^2}(y_i - x)^2\right)$$
$$\propto \exp\left(-\frac{1}{2\sigma^2}\sum_{i=1}^{n}(y_i - x)^2\right)$$

We can work out the posterior distribution:

As a result of independent observations, we have

$$p(y|x) = \prod_{i=1}^{n} \left(\frac{1}{\sqrt{2\pi\sigma^2}}\right) \exp\left(-\frac{1}{2\sigma^2}(y_i - x)^2\right)$$
$$\propto \exp\left(-\frac{1}{2\sigma^2}\sum_{i=1}^{n}(y_i - x)^2\right)$$

We can work out the posterior distribution:

$$p(x|y) \propto \exp\left(-rac{1}{2}\left[\sum_{i=1}^{n}rac{(y_i-x)^2}{\sigma^2}+rac{(x-\mu)^2}{ au^2}
ight]
ight) \ \propto \exp\left(-rac{1}{2}\left[\sum_{i=1}^{n}rac{-2xy_i+x^2}{\sigma^2}+rac{x^2-2x\mu}{ au^2}
ight]
ight)$$



$$= \exp\left(-\frac{1}{2}\left(\frac{n}{\sigma^2} + \frac{1}{\tau^2}\right)x^2 - 2\left(\sum_{i=1}^n \frac{y_i}{\sigma^2} + \frac{\mu}{\tau^2}\right)x\right)$$
$$= \exp\left(-\frac{1}{2}(ax^2 - 2bx)\right)$$
$$\propto \exp\left(-\frac{a}{2}\left(x - \frac{b}{a}\right)^2\right)$$

$$= \exp\left(-\frac{1}{2}\left(\frac{n}{\sigma^2} + \frac{1}{\tau^2}\right)x^2 - 2\left(\sum_{i=1}^n \frac{y_i}{\sigma^2} + \frac{\mu}{\tau^2}\right)x\right)$$
$$= \exp\left(-\frac{1}{2}(ax^2 - 2bx)\right)$$
$$\propto \exp\left(-\frac{a}{2}\left(x - \frac{b}{a}\right)^2\right)$$

So we have:

$$p(x|y) = \mathcal{N}\left(\left(\frac{n}{\sigma^2} + \frac{1}{\tau^2}\right)^{-1}\left(\sum_{i=1}^n \frac{y_i}{\sigma^2} + \frac{\mu}{\tau^2}\right), \left(\frac{n}{\sigma^2} + \frac{1}{\tau^2}\right)^{-1}\right)$$



Write

#### Write

$$\bar{y} = \sum_{i=1}^{n} y_i / n$$

$$\omega_y = \frac{n\tau^2}{n\tau^2 + \sigma^2}$$

$$\omega_\mu = \frac{\sigma^2}{n\tau^2 + \sigma^2}$$

Write

$$\bar{y} = \sum_{i=1}^{n} y_i / n$$

$$\omega_y = \frac{n\tau^2}{n\tau^2 + \sigma^2}$$

$$\omega_\mu = \frac{\sigma^2}{n\tau^2 + \sigma^2}$$

Then we have

$$\mathbb{E}(X|y) = \omega_y \bar{y} + \omega_\mu \mu$$



ullet  $au^2$  or  $n o\infty\Rightarrow\omega_y o 1$  and  $\omega_\mu o 0$  (so data dominates prior)

# Example: Univariate Normal-Normal Case

- $\circ$   $au^2$  or  $n o\infty\Rightarrow\omega_y o 1$  and  $\omega_\mu o 0$  (so data dominates prior)
- For small  $au^2$ , the prior is critical for small  $n \; (\omega_y o 0)$

#### Example: Univariate Normal-Normal Case

- ullet  $au^2$  or  $n o\infty\Rightarrow\omega_y o 1$  and  $\omega_\mu o 0$  (so data dominates prior)
- For small  $\tau^2$ , the prior is critical for small n ( $\omega_v \to 0$ )
- Also, can write

$$\mathbb{E}(X|y) = \mu + \omega_y(\bar{y} - \mu)$$
$$= \mu + K(\bar{y} - \mu)$$

Gain K adjusts prior mean towards sample mean.

### Example: Univariate Normal-Normal Case

- $au^2$  or  $n o \infty \Rightarrow \omega_{\scriptscriptstyle V} o 1$  and  $\omega_{\scriptscriptstyle \mu} o 0$  (so data dominates prior)
- For small  $\tau^2$ , the prior is critical for small n ( $\omega_v \to 0$ )
- Also, can write

$$\mathbb{E}(X|y) = \mu + \omega_y(\bar{y} - \mu)$$
$$= \mu + K(\bar{y} - \mu)$$

Gain K adjusts prior mean towards sample mean.

Similar for variance

$$\mathsf{Var}(X|y) = (1 - K)\tau^2$$



Idea: update posterior sequentially

- Idea: update posterior sequentially
- Notation

$$Y_{1:t} = \{Y_1, \dots, Y_t\}, \quad X_{0:t} = \{X_0, \dots, X_t\}$$

- Idea: update posterior sequentially
- Notation

$$Y_{1:t} = \{Y_1, \dots, Y_t\}, \quad X_{0:t} = \{X_0, \dots, X_t\}$$

Assumptions:

$$p(x_{0:T}) = p(x_0) \prod_{t=1}^{T} p(x_t|x_{t-1})$$
 $p(y_{1:T}|x_{0:T}) = \prod_{t=1}^{T} p(y_t|x_t)$ 

- Idea: update posterior sequentially
- Notation

$$Y_{1:t} = \{Y_1, \dots, Y_t\}, \quad X_{0:t} = \{X_0, \dots, X_t\}$$

Assumptions:

$$p(x_{0:T}) = p(x_0) \prod_{t=1}^{T} p(x_t|x_{t-1})$$
 $p(y_{1:T}|x_{0:T}) = \prod_{t=1}^{T} p(y_t|x_t)$ 

Using the above we have:

$$p(x_{0:T}|y_{1:T}) \propto p(x_0) \prod_{t=1}^{T} p(y_t|x_t) p(x_t|x_{t-1})$$



• Filtering: use only past data to update posterior

- Filtering: use only past data to update posterior
- Given  $p(x_{t-1}|y_{1:t-1})$ , find forecast  $p(x_t|y_{1:t-1})$  and analysis  $p(x_t|y_{1:t})$

- Filtering: use only past data to update posterior
- Given  $p(x_{t-1}|y_{1:t-1})$ , find forecast  $p(x_t|y_{1:t-1})$  and analysis  $p(x_t|y_{1:t})$
- Markovian assumption gives:

$$p(x_t|y_{1:t-1}) = \int p(x_t|x_{t-1})p(x_{t-1}|y_{1:t-1}) dx_{t-1}$$

- Filtering: use only past data to update posterior
- Given  $p(x_{t-1}|y_{1:t-1})$ , find forecast  $p(x_t|y_{1:t-1})$  and analysis  $p(x_t|y_{1:t})$
- Markovian assumption gives:

$$p(x_t|y_{1:t-1}) = \int p(x_t|x_{t-1})p(x_{t-1}|y_{1:t-1}) dx_{t-1}$$

Bayes' Rule gives:

$$p(x_t|y_{1:t}) = p(x_t|y_t, y_{1:t-1})$$

$$\propto p(y_t|x_t, y_{1:t-1})p(x_t|y_{1:t-1})$$

$$= p(y_t|x_t)p(x_t|y_{1:t-1})$$

- Filtering: use only past data to update posterior
- Given  $p(x_{t-1}|y_{1:t-1})$ , find forecast  $p(x_t|y_{1:t-1})$  and analysis  $p(x_t|y_{1:t})$
- Markovian assumption gives:

$$p(x_t|y_{1:t-1}) = \int p(x_t|x_{t-1})p(x_{t-1}|y_{1:t-1}) dx_{t-1}$$

Bayes' Rule gives:

$$p(x_t|y_{1:t}) = p(x_t|y_t, y_{1:t-1})$$

$$\propto p(y_t|x_t, y_{1:t-1})p(x_t|y_{1:t-1})$$

$$= p(y_t|x_t)p(x_t|y_{1:t-1})$$

 So we can iterate between forecast and analysis distributions to update the posterior distribution



• Smoothing: use all the data to update posterior

- Smoothing: use all the data to update posterior
- Want  $p(x_{0:T}|y_{1:T})$  in sequential fashion

- Smoothing: use all the data to update posterior
- Want  $p(x_{0:T}|y_{1:T})$  in sequential fashion
- Smoothing distribution is given by:

$$p(x_t|y_{1:T}) = \int p(x_t|x_{t+1}, y_{1:T}) p(x_{t+1}|y_{1:T}) dx_{t+1}$$

- Smoothing: use all the data to update posterior
- Want  $p(x_{0:T}|y_{1:T})$  in sequential fashion
- Smoothing distribution is given by:

$$p(x_t|y_{1:T}) = \int p(x_t|x_{t+1}, y_{1:T}) p(x_{t+1}|y_{1:T}) dx_{t+1}$$

• Assume:  $p(x_t|x_{t+1}, y_{1:T}) = p(x_t|x_{t+1}, y_{1:t})$ 



- Smoothing: use all the data to update posterior
- Want  $p(x_{0:T}|y_{1:T})$  in sequential fashion
- Smoothing distribution is given by:

$$p(x_t|y_{1:T}) = \int p(x_t|x_{t+1}, y_{1:T}) p(x_{t+1}|y_{1:T}) dx_{t+1}$$

- Assume:  $p(x_t|x_{t+1}, y_{1:T}) = p(x_t|x_{t+1}, y_{1:t})$
- Bayes' Rule gives:

$$p(x_t|x_{t+1}, y_{1:t}) \propto p(x_{t+1}|x_t, y_{1:t})p(x_t|y_{1:t})$$
  
=  $p(x_{t+1}|x_t)p(x_t|y_{1:t})$ 



- Forward Filtering-Backward Smoothing Algorithm
- For t = T 1 to 1

- Forward Filtering-Backward Smoothing Algorithm
- For t = T 1 to 1
  - multiply filtered analysis distribution  $p(x_t|y_{1:t})$  with  $p(x_{t+1}|x_t)$  to obtain  $p(x_t|x_{t+1},y_{1:t})$

- Forward Filtering-Backward Smoothing Algorithm
- For t = T 1 to 1
  - multiply filtered analysis distribution  $p(x_t|y_{1:t})$  with  $p(x_{t+1}|x_t)$  to obtain  $p(x_t|x_{t+1},y_{1:t})$
  - use smoothing distribution definition to obtain the smoothing distribution  $p(x_t|y_{1:T})$  making use of smoothing distribution at time t+1,  $p(x_{t+1}|y_{1:T})$ , obtained at previous iteration

- Forward Filtering-Backward Smoothing Algorithm
- For t = T 1 to 1
  - multiply filtered analysis distribution  $p(x_t|y_{1:t})$  with  $p(x_{t+1}|x_t)$  to obtain  $p(x_t|x_{t+1},y_{1:t})$
  - use smoothing distribution definition to obtain the smoothing distribution  $p(x_t|y_{1:T})$  making use of smoothing distribution at time t+1,  $p(x_{t+1}|y_{1:T})$ , obtained at previous iteration
- Aside: For a linear model, linear observation operator and Gaussian prior, the posterior is Gaussian. Its mean and covariance can be updated iteratively in a similar manner. This is known as the Kalman Filter and Kalman Smoother.

• What if the model (or observation operator) is nonlinear?

- What if the model (or observation operator) is nonlinear?
  - Linearise locally (Extended Kalman Filter)

- What if the model (or observation operator) is nonlinear?
  - Linearise locally (Extended Kalman Filter)
  - Use Monte Carlo sampling

- What if the model (or observation operator) is nonlinear?
  - Linearise locally (Extended Kalman Filter)
  - Use Monte Carlo sampling
- Basic MC is to estimate integrals. Say we want to calculate:

$$\mathbb{E}(f(X_{0:t})|y_{1:t}) = \int f(x_{0:t})p(x_{0:t}|y_{1:t}) dx_{0:t}$$

$$= \frac{\int f(x_{0:t})p(y_{1:t}|x_{0:t})p(x_{0:t}) dx_{0:t}}{\int p(y_{1:t}|x_{0:t})p(x_{0:t}) dx_{0:t}}$$

- What if the model (or observation operator) is nonlinear?
  - Linearise locally (Extended Kalman Filter)
  - Use Monte Carlo sampling
- Basic MC is to estimate integrals. Say we want to calculate:

$$\mathbb{E}(f(X_{0:t})|y_{1:t}) = \int f(x_{0:t})p(x_{0:t}|y_{1:t}) dx_{0:t}$$

$$= \frac{\int f(x_{0:t})p(y_{1:t}|x_{0:t})p(x_{0:t}) dx_{0:t}}{\int p(y_{1:t}|x_{0:t})p(x_{0:t}) dx_{0:t}}$$

We can use m pseudo-random realisations,  $x_{0:t}^i$ , from  $p(x_{0:t}|y_{1:t})$  and compute:

$$\widehat{\mathbb{E}}(f(X_{0:t})|y_{1:t}) = \frac{1}{M} \sum_{i=1}^{M} f(x_{0:t}^{i})$$



- What if the model (or observation operator) is nonlinear?
  - Linearise locally (Extended Kalman Filter)
  - Use Monte Carlo sampling
- Basic MC is to estimate integrals. Say we want to calculate:

$$\mathbb{E}(f(X_{0:t})|y_{1:t}) = \int f(x_{0:t})p(x_{0:t}|y_{1:t}) dx_{0:t}$$

$$= \frac{\int f(x_{0:t})p(y_{1:t}|x_{0:t})p(x_{0:t}) dx_{0:t}}{\int p(y_{1:t}|x_{0:t})p(x_{0:t}) dx_{0:t}}$$

We can use m pseudo-random realisations,  $x_{0:t}^i$ , from  $p(x_{0:t}|y_{1:t})$  and compute:

$$\widehat{\mathbb{E}}(f(X_{0:t})|y_{1:t}) = \frac{1}{M} \sum_{i=1}^{M} f(x_{0:t}^{i})$$

ullet Can approximate  $p(x_{0:t}|y_{1:t})$  by  $p^m(x_{0:t}|y_{1:t}) = \sum_{i=1}^M \delta_{x_{0:t}^i}$ 



• Use ISMC when drawing from  $p(x_{0:t}|y_{1:t})$  is hard

- Use ISMC when drawing from  $p(x_{0:t}|y_{1:t})$  is hard
- Sample from  $g(x_{0:t}|y_{1:t}) \gg p(x_{0:t}|y_{1:t})$ , but weight each sample and calculate:

$$\hat{\mathbb{E}}(f(X_{0:t})|y_{1:t}) = \sum_{i=1}^{M} w^{i} f(x_{0:t}^{i})$$

- Use ISMC when drawing from  $p(x_{0:t}|y_{1:t})$  is hard
- Sample from  $g(x_{0:t}|y_{1:t}) \gg p(x_{0:t}|y_{1:t})$ , but weight each sample and calculate:

$$\hat{\mathbb{E}}(f(X_{0:t})|y_{1:t}) = \sum_{i=1}^{M} w^{i} f(x_{0:t}^{i})$$

where

$$w^{i} = \frac{p(x_{0:t}^{i}|y_{1:t})/g(x_{0:t}^{i}|y_{1:t})}{\sum_{j=1}^{M} p(x_{0:t}^{j}|y_{1:t})/g(x_{0:t}^{j}|y_{1:t})}$$

- Use ISMC when drawing from  $p(x_{0:t}|y_{1:t})$  is hard
- Sample from  $g(x_{0:t}|y_{1:t}) \gg p(x_{0:t}|y_{1:t})$ , but weight each sample and calculate:

$$\hat{\mathbb{E}}(f(X_{0:t})|y_{1:t}) = \sum_{i=1}^{M} w^{i} f(x_{0:t}^{i})$$

where

$$w^{i} = \frac{p(x_{0:t}^{i}|y_{1:t})/g(x_{0:t}^{i}|y_{1:t})}{\sum_{j=1}^{M} p(x_{0:t}^{j}|y_{1:t})/g(x_{0:t}^{j}|y_{1:t})}$$

ullet Can approximate  $p(x_{0:t}|y_{1:t})$  by  $p^m(x_{0:t}|y_{1:t}) = \sum_{i=1}^M w^i \delta_{x_{0:t}^i}$ 



• Usually choose  $g(x_{0:t}|y_{1:t}) = p(x_{0:t})$ 

- Usually choose  $g(x_{0:t}|y_{1:t}) = p(x_{0:t})$
- Given a sample from  $p(x_{0:t})$ , simulate the forward model to obtain Monte Carlo trajectories

- Usually choose  $g(x_{0:t}|y_{1:t}) = p(x_{0:t})$
- Given a sample from  $p(x_{0:t})$ , simulate the forward model to obtain Monte Carlo trajectories
- With this g, we get:

$$w^{i} \propto rac{p(x_{0:t}^{i}|y_{1:t})}{p(x_{0:t}^{i})} \ \propto p(y_{1:t}|x_{0:t}^{i})$$

- Usually choose  $g(x_{0:t}|y_{1:t}) = p(x_{0:t})$
- Given a sample from  $p(x_{0:t})$ , simulate the forward model to obtain Monte Carlo trajectories
- With this g, we get:

$$w^{i} \propto \frac{p(x_{0:t}^{i}|y_{1:t})}{p(x_{0:t}^{i})} \\ \propto p(y_{1:t}|x_{0:t}^{i})$$

This is called an ensemble smoother



Sequential MC algorithm follows from Filtering algorithm

- Sequential MC algorithm follows from Filtering algorithm
- Have seen:  $g = prior \Rightarrow w^i \propto likelihood$

- Sequential MC algorithm follows from Filtering algorithm
- Have seen:  $g = prior \Rightarrow w^i \propto likelihood$
- Let  $w_t^i$  be weight on ensemble member i at time t

- Sequential MC algorithm follows from Filtering algorithm
- Have seen:  $g = prior \Rightarrow w^i \propto likelihood$
- Let  $w_t^i$  be weight on ensemble member i at time t
- Independent data gives:

$$w_t^i \propto p(y_{1:t}|x_{0:t}^i) \propto p(y_t|x_t^i)w_{t-1}^i$$

- Sequential MC algorithm follows from Filtering algorithm
- Have seen:  $g = prior \Rightarrow w^i \propto likelihood$
- Let  $w_t^i$  be weight on ensemble member i at time t
- Independent data gives:

$$w_t^i \propto p(y_{1:t}|x_{0:t}^i) \propto p(y_t|x_t^i)w_{t-1}^i$$

\* Approximate analysis distribution at time t-1 by

$$p^{m}(x_{t-1}|y_{1:t-1}) = \sum_{i=1}^{M} w_{t-1}^{i} \delta_{x_{t-1|t-1}^{i}}$$



- Sequential MC algorithm follows from Filtering algorithm
- Have seen:  $g = prior \Rightarrow w^i \propto likelihood$
- Let  $w_t^i$  be weight on ensemble member i at time t
- Independent data gives:

$$w_t^i \propto p(y_{1:t}|x_{0:t}^i) \propto p(y_t|x_t^i)w_{t-1}^i$$

ullet Approximate analysis distribution at time t-1 by

$$p^{m}(x_{t-1}|y_{1:t-1}) = \sum_{i=1}^{M} w_{t-1}^{i} \delta_{x_{t-1|t-1}^{i}}$$

where  $x_{t-1|t-1}^i$  are random draws from  $p(x_{t-1}|y_{1:t-1})$ 



 In practice one considers a kernel-density approximation of this and so we can estimate the forecast:

 In practice one considers a kernel-density approximation of this and so we can estimate the forecast:

$$p^{m}(x_{t}|y_{1:t-1}) \propto \sum_{i=1}^{M} p(x_{t}|x_{t-1:t-1}^{i})w_{t-1}^{i}$$

• In practice one considers a kernel-density approximation of this and so we can estimate the forecast:

$$p^{m}(x_{t}|y_{1:t-1}) \propto \sum_{i=1}^{M} p(x_{t}|x_{t-1:t-1}^{i})w_{t-1}^{i}$$

Using the weight-update above we have:

$$p^{m}(x_{t}|y_{1:1}) \propto p(y_{t}|x_{t}) \sum_{i=1}^{M} p(x_{t}|x_{t-1:t-1}^{i}) w_{t-1}^{i}$$
$$= \sum_{i=1}^{M} p(x_{t}|x_{t-1:t-1}^{i}) w_{t}^{i}$$

• In practice one considers a kernel-density approximation of this and so we can estimate the forecast:

$$p^{m}(x_{t}|y_{1:t-1}) \propto \sum_{i=1}^{M} p(x_{t}|x_{t-1:t-1}^{i})w_{t-1}^{i}$$

Using the weight-update above we have:

$$p^{m}(x_{t}|y_{1:1}) \propto p(y_{t}|x_{t}) \sum_{i=1}^{M} p(x_{t}|x_{t-1:t-1}^{i}) w_{t-1}^{i}$$

$$= \sum_{i=1}^{M} p(x_{t}|x_{t-1:t-1}^{i}) w_{t}^{i}$$

Various methods to resolve problems from dimensionality...



• Throw away samples with small weight

- Throw away samples with small weight
- Multiply particles with high weight

- Throw away samples with small weight
- Multiply particles with high weight
- Use Bootstrap Algorithm:

- Throw away samples with small weight
- Multiply particles with high weight
- Use Bootstrap Algorithm:
  - 1 Initialisation, t = 0

- Throw away samples with small weight
- Multiply particles with high weight
- Use Bootstrap Algorithm:
  - 1 Initialisation, t = 0
    - $\bullet$  for  $i=0,\ldots,m$  sample  $x_{0|0}^i\sim p(x_0)$  and set t=1

- Throw away samples with small weight
- Multiply particles with high weight
- Use Bootstrap Algorithm:
  - Initialisation, t = 0
    - $\bullet$  for  $i=0,\ldots,m$  sample  $x_{0|0}^i\sim p(x_0)$  and set t=1
  - Importance sampling step

- Throw away samples with small weight
- Multiply particles with high weight
- Use Bootstrap Algorithm:
  - Initialisation, t = 0
    - $\bullet$  for  $i=0,\ldots,m$  sample  $x_{0|0}^i\sim p(x_0)$  and set t=1
  - Importance sampling step
    - for  $i=0,\ldots,m$  sample  $ilde{x}_t^i \sim p(x_t|x_{t-1}^i)$  and set  $ilde{x}_t^i = \{x_{t-1}^i, ilde{x}_t^i\}$

- Throw away samples with small weight
- Multiply particles with high weight
- Use Bootstrap Algorithm:
  - 1 Initialisation, t = 0
    - $\bullet$  for  $i=0,\ldots,m$  sample  $x_{0|0}^i\sim p(x_0)$  and set t=1
  - Importance sampling step
    - for  $i=0,\ldots,m$  sample  $ilde{x}_t^i \sim p(x_t|x_{t-1}^i)$  and set  $ilde{x}_t^i = \{x_{t-1}^i, ilde{x}_t^i\}$
    - for  $i=0,\ldots,m$  evaluate importance weights  $\tilde{w}_t^i=p(y_t|\tilde{x}_t^i)$

- Throw away samples with small weight
- Multiply particles with high weight
- Use Bootstrap Algorithm:
  - 1 Initialisation, t = 0
    - $\circ$  for  $i=0,\ldots,m$  sample  $x_{0|0}^i\sim p(x_0)$  and set t=1
  - Importance sampling step
    - for  $i=0,\ldots,m$  sample  $ilde{x}_t^i \sim p(x_t|x_{t-1}^i)$  and set  $ilde{x}_t^i = \{x_{t-1}^i, ilde{x}_t^i\}$
    - for  $i=0,\ldots,m$  evaluate importance weights  $ilde{w}_t^i=p(y_t| ilde{x}_t^i)$
    - normalise weights

- Throw away samples with small weight
- Multiply particles with high weight
- Use Bootstrap Algorithm:
  - 1 Initialisation, t = 0
    - $\bullet$  for  $i=0,\ldots,m$  sample  $x_{0|0}^i\sim p(x_0)$  and set t=1
  - Importance sampling step
    - for  $i=0,\ldots,m$  sample  $ilde{x}_t^i \sim p(x_t|x_{t-1}^i)$  and set  $ilde{x}_t^i = \{x_{t-1}^i, ilde{x}_t^i\}$
    - for  $i=0,\ldots,m$  evaluate importance weights  $\tilde{w}_t^i=p(y_t|\tilde{x}_t^i)$
    - normalise weights
  - Selection step

- Throw away samples with small weight
- Multiply particles with high weight
- Use Bootstrap Algorithm:
  - 1 Initialisation, t = 0
    - $\bullet$  for  $i=0,\ldots,m$  sample  $x_{0|0}^i\sim p(x_0)$  and set t=1
  - 2 Importance sampling step
    - for  $i=0,\ldots,m$  sample  $ilde{x}_t^i \sim p(x_t|x_{t-1}^i)$  and set  $ilde{x}_t^i = \{x_{t-1}^i, ilde{x}_t^i\}$
    - for  $i=0,\ldots,m$  evaluate importance weights  $ilde{w}_t^i=p(y_t| ilde{x}_t^i)$
    - normalise weights
  - Selection step
    - resample with replacement m particles  $\{x_i^i:i=1,\ldots,m\}$  from the set  $\{\tilde{x}_i^i:i=1,\ldots,m\}$  according to importance weights



- Throw away samples with small weight
- Multiply particles with high weight
- Use Bootstrap Algorithm:
  - 1 Initialisation, t = 0
    - $\bullet$  for  $i=0,\ldots,m$  sample  $x_{0|0}^i\sim p(x_0)$  and set t=1
  - Importance sampling step
    - for  $i=0,\ldots,m$  sample  $ilde{x}_t^i \sim p(x_t|x_{t-1}^i)$  and set  $ilde{x}_t^i = \{x_{t-1}^i, ilde{x}_t^i\}$
    - for i = 0, ..., m evaluate importance weights  $\tilde{w}_t^i = p(y_t | \tilde{x}_t^i)$
    - normalise weights
  - Selection step
    - resample with replacement m particles  $\{x_i^i:i=1,\ldots,m\}$  from the set  $\{\tilde{x}_i^i:i=1,\ldots,m\}$  according to importance weights
    - set t = t + 1 and go to step 2



• Can use the Ensmeble Kalman Filter in high dimensional problems

- Can use the Ensmeble Kalman Filter in high dimensional problems
- Approach:

- Can use the Ensmeble Kalman Filter in high dimensional problems
- Approach:
  - Use MC samples to approximate forecast mean and covariance while still using nonlinear forward model

- Can use the Ensmeble Kalman Filter in high dimensional problems
- Approach:
  - Use MC samples to approximate forecast mean and covariance while still using nonlinear forward model
  - Use the these in linear Kalman Filter update formulas to obtain analysis distribution

- Can use the Ensmeble Kalman Filter in high dimensional problems
- Approach:
  - Use MC samples to approximate forecast mean and covariance while still using nonlinear forward model
  - Use the these in linear Kalman Filter update formulas to obtain analysis distribution
- \* Assume we have m independent samples from analysis distribution at time t-1,  $x_{t-1|t-1}^{j}$ .

- Can use the Ensmeble Kalman Filter in high dimensional problems
- Approach:
  - Use MC samples to approximate forecast mean and covariance while still using nonlinear forward model
  - Use the these in linear Kalman Filter update formulas to obtain analysis distribution
- Assume we have m independent samples from analysis distribution at time t-1,  $x_{t-1|t-1}^{j}$ .
- Also assume  $w_{t-1}^i = 1/m$

- Can use the Ensmeble Kalman Filter in high dimensional problems
- Approach:
  - Use MC samples to approximate forecast mean and covariance while still using nonlinear forward model
  - Use the these in linear Kalman Filter update formulas to obtain analysis distribution
- Assume we have m independent samples from analysis distribution at time t-1,  $x_{t-1|t-1}^{i}$ .
- Also assume  $w_{t-1}^i = 1/m$
- Use these samples and basic MC to approximate forecast:

$$p^{m}(x_{t}|y_{1:t-1}) = (1/m) \sum_{i=1}^{M} p(x_{t}|x_{t-1:t-1}^{i})$$



Now assume forecast distribution can be characterised by first two moments (or Guassian) with mean  $x_{t|t-1}^i$  and estimated covariance matrix  $\hat{P}_{t|t-1}$ . Then analysis (update) distribution is given by:

Now assume forecast distribution can be characterised by first two moments (or Guassian) with mean  $x_{t|t-1}^i$  and estimated covariance matrix  $\hat{P}_{t|t-1}$ . Then analysis (update) distribution is given by:

$$p^{m}(x_{t}|y_{1:t}) \propto (1/m)p(y_{t}|x_{t}) \sum_{i=1}^{M} \mathcal{N}(x_{t|t-1}^{i}, \hat{P}_{t|t-1})$$

Now assume forecast distribution can be characterised by first two moments (or Guassian) with mean  $x_{t|t-1}^i$  and estimated covariance matrix  $\hat{P}_{t|t-1}$ . Then analysis (update) distribution is given by:

$$p^{m}(x_{t}|y_{1:t}) \propto (1/m)p(y_{t}|x_{t}) \sum_{i=1}^{M} \mathcal{N}(x_{t|t-1}^{i}, \hat{P}_{t|t-1})$$

Assume a linear (or linearised) observation model:

$$y_t = H_t x_t + e_t$$

Now assume forecast distribution can be characterised by first two moments (or Guassian) with mean  $x_{t|t-1}^i$  and estimated covariance matrix  $\hat{P}_{t|t-1}$ . Then analysis (update) distribution is given by:

$$p^m(x_t|y_{1:t}) \propto (1/m)p(y_t|x_t) \sum_{i=1}^M \mathcal{N}(x_{t|t-1}^i, \hat{P}_{t|t-1})$$

Assume a linear (or linearised) observation model:

$$y_t = H_t x_t + e_t$$

where  $Cov(e_t) = R_t$ 



Now assume forecast distribution can be characterised by first two moments (or Guassian) with mean  $x_{t|t-1}^i$  and estimated covariance matrix  $\hat{P}_{t|t-1}$ . Then analysis (update) distribution is given by:

$$p^{m}(x_{t}|y_{1:t}) \propto (1/m)p(y_{t}|x_{t}) \sum_{i=1}^{M} \mathcal{N}(x_{t|t-1}^{i}, \hat{P}_{t|t-1})$$

Assume a linear (or linearised) observation model:

$$y_t = H_t x_t + e_t$$

where  $Cov(e_t) = R_t$ 

• Let  $x_{t|t-1}$  and  $P_{t|t-1}$  be mean and covariance of  $p(x_t|y_{1:t-1})$ 



Now assume forecast distribution can be characterised by first two moments (or Guassian) with mean  $x_{t|t-1}^i$  and estimated covariance matrix  $\hat{P}_{t|t-1}$ . Then analysis (update) distribution is given by:

$$p^m(x_t|y_{1:t}) \propto (1/m) p(y_t|x_t) \sum_{i=1}^M \mathcal{N}(x_{t|t-1}^i, \hat{P}_{t|t-1})$$

Assume a linear (or linearised) observation model:

$$y_t = H_t x_t + e_t$$

where  $Cov(e_t) = R_t$ 

- Let  $x_{t|t-1}$  and  $P_{t|t-1}$  be mean and covariance of  $p(x_t|y_{1:t-1})$
- Want a viable ensemble from  $p(x_t|y_{1:t})$ , or of  $x_{t|t}$  and  $P_{t|t}$



Algorithm:

### Algorithm:

• Evolve  $x_{t-1|t-1}^i$  forward using:

$$x_{t|t-1}^i = \mathcal{M}(x_{t-1|t-1}^i) + \eta_t^i, \quad \eta_t^i \sim \mathcal{N}(0, Q)$$

### Algorithm:

• Evolve  $x_{t-1|t-1}^i$  forward using:

$$x_{t|t-1}^i = \mathcal{M}(x_{t-1|t-1}^i) + \eta_t^i, \quad \eta_t^i \sim \mathcal{N}(0, Q)$$

Use evolved samples to calculate a sample forecast covariance matrix,  $\hat{P}_{t|t-1}$ :

$$\hat{P}_{t|t-1} = \frac{1}{m-1} \sum_{i=1}^{m} (x_{t|t-1}^{i} - \hat{x}_{t|t-1}^{i}) (x_{t|t-1}^{i} - \hat{x}_{t|t-1}^{i})^{T}$$

### Algorithm:

• Evolve  $x_{t-1|t-1}^i$  forward using:

$$x_{t|t-1}^i = \mathcal{M}(x_{t-1|t-1}^i) + \eta_t^i, \quad \eta_t^i \sim \mathcal{N}(0, Q)$$

Use evolved samples to calculate a sample forecast covariance matrix,  $\hat{P}_{t|t-1}$ :

$$\hat{P}_{t|t-1} = \frac{1}{m-1} \sum_{i=1}^{m} (x_{t|t-1}^{i} - \hat{x}_{t|t-1}^{i}) (x_{t|t-1}^{i} - \hat{x}_{t|t-1}^{i})^{T}$$

where 
$$\hat{x}_{t|t-1}^{i} = (1/m) \sum_{i=1}^{m} x_{t|t-1}^{i}$$

### Algorithm:

• Evolve  $x_{t-1|t-1}^i$  forward using:

$$x_{t|t-1}^i = \mathcal{M}(x_{t-1|t-1}^i) + \eta_t^i, \quad \eta_t^i \sim \mathcal{N}(0, Q)$$

Use evolved samples to calculate a sample forecast covariance matrix,  $\hat{P}_{t|t-1}$ :

$$\hat{P}_{t|t-1} = \frac{1}{m-1} \sum_{i=1}^{m} (x_{t|t-1}^{i} - \hat{x}_{t|t-1}^{i}) (x_{t|t-1}^{i} - \hat{x}_{t|t-1}^{i})^{T}$$

where 
$$\hat{x}_{t|t-1}^i = (1/m) \sum_{i=1}^m x_{t|t-1}^i$$

 Use Kalman update equations to update each forecast sample given the sampled observations:

$$x_{t|t}^{i} = x_{t|t-1}^{i} + K_{t}(y_{t} + e_{t} - H_{t}x_{t|t-1}^{i})$$



### Algorithm:

• Evolve  $x_{t-1|t-1}^i$  forward using:

$$x_{t|t-1}^i = \mathcal{M}(x_{t-1|t-1}^i) + \eta_t^i, \quad \eta_t^i \sim \mathcal{N}(0, Q)$$

Use evolved samples to calculate a sample forecast covariance matrix,  $\hat{P}_{t|t-1}$ :

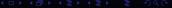
$$\hat{P}_{t|t-1} = \frac{1}{m-1} \sum_{i=1}^{m} (x_{t|t-1}^{i} - \hat{x}_{t|t-1}^{i}) (x_{t|t-1}^{i} - \hat{x}_{t|t-1}^{i})^{T}$$

where 
$$\hat{x}_{t|t-1}^i = (1/m) \sum_{i=1}^m x_{t|t-1}^i$$

 Use Kalman update equations to update each forecast sample given the sampled observations:

$$x_{t|t}^{i} = x_{t|t-1}^{i} + K_{t}(y_{t} + e_{t} - H_{t}x_{t|t-1}^{i})$$

where  $K_t = \hat{P}_{t|t-1}H_t^T(H_t\hat{P}_{t|t-1}H_t^T + R)^{-1}$  and  $e_t^i \sim \mathcal{N}(0,R)$ 



### We have seen:

How to use Bayes' Rule to infer on posterior

- How to use Bayes' Rule to infer on posterior
- How to update posterior sequentially (and analytically):

- How to use Bayes' Rule to infer on posterior
- How to update posterior sequentially (and analytically):
  - Filtering

- How to use Bayes' Rule to infer on posterior
- How to update posterior sequentially (and analytically):
  - Filtering
  - Smoothing

- How to use Bayes' Rule to infer on posterior
- How to update posterior sequentially (and analytically):
  - Filtering
  - Smoothing
- Forward-Filtering Backward-Smoothing

- How to use Bayes' Rule to infer on posterior
- How to update posterior sequentially (and analytically):
  - Filtering
  - Smoothing
- Forward-Filtering Backward-Smoothing
- Basic and Importance Monte Carlo sampling:

- How to use Bayes' Rule to infer on posterior
- How to update posterior sequentially (and analytically):
  - Filtering
  - Smoothing
- Forward-Filtering Backward-Smoothing
- Basic and Importance Monte Carlo sampling:
  - Sequential Monte Carlo

- How to use Bayes' Rule to infer on posterior
- How to update posterior sequentially (and analytically):
  - Filtering
  - Smoothing
- Forward-Filtering Backward-Smoothing
- Basic and Importance Monte Carlo sampling:
  - Sequential Monte Carlo
  - Bootstrap Algorithm

- How to use Bayes' Rule to infer on posterior
- How to update posterior sequentially (and analytically):
  - Filtering
  - Smoothing
- Forward-Filtering Backward-Smoothing
- Basic and Importance Monte Carlo sampling:
  - Sequential Monte Carlo
  - Bootstrap Algorithm
  - EnKF

- How to use Bayes' Rule to infer on posterior
- How to update posterior sequentially (and analytically):
  - Filtering
  - Smoothing
- Forward-Filtering Backward-Smoothing
- Basic and Importance Monte Carlo sampling:
  - Sequential Monte Carlo
  - Bootstrap Algorithm
  - EnKF
- Problems with these methods

