a) L(p; x) = TT, F(x,) = TT, Px(1-p)'-x;

NAME David Levy f(x)=(px(1-p)1-x, x=0,1

$$\begin{aligned} \log L(p; X) &= \sum_{i=1}^{p} \log \left[p^{x_{i}} (1-p)^{1-x_{i}} \right] \\ &= \sum_{i=1}^{p} \left[x_{i} \log p + (1-x_{i}) \log (1-p) \right] \\ &= \sum_{i=1}^{p} \left[\frac{x_{i}}{p} - \frac{(1-x_{i})}{1-p} \right] \\ &= \sum_{i=1}^{p} \left[\frac{x_{i}}{p} - \frac{(1-x_{i})}{1-p} \right] \\ &= \left(\sum_{i=1}^{p} \left[\frac{x_{i}}{p(1-p)} - \frac{p(1-x_{i})}{p(1-p)} \right] = \sum_{i=1}^{p} \left[\frac{x_{i}-p}{p(1-p)} \right] \\ &= \left(\frac{1}{p(1-p)} \right) \sum_{i=1}^{p} X_{i} - \frac{n}{1-p} \end{aligned}$$

set equal to 0

$$\hat{p}_{mLE} = \frac{1}{n} \sum_{i} X_{i} = \overline{X}_{n}$$

The MLE is the sample mean, Xn.

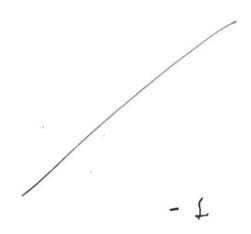
A sum of Bernoullis has a bin (n,p) distribution.

mean up, variance up(1-p).

$$X_n \xrightarrow{d} N(np, p(1-p))$$
 -1

convergence statements industries b) We want to show $\gamma_n \xrightarrow{P} P$ for consistency, so we need $\lim_{n\to\infty} P(|\gamma_n-P|\geq E) = 0$ = 1- P(Yn-P

February 12, 2018



c) For asymptotic efficiency we would need to show that as $n\to\infty$, the estimator attains the Rao-Cramer lower bound for variance. Given that Yn is a statistic, we can find the lower bound by

Var $(T_n) = \frac{1}{n T(p)}$ Where we find T(p) by $-E\left[\frac{\partial^2 \log f(p;x)}{\partial p^2}\right]$ or set it's the hozolar cay to 80

5 -

$$X_{c} \sim N(\mu_{r}, \sigma^{2})$$
 $S_{n}^{2} = \frac{1}{n} \sum_{c} (X_{c} - \overline{X}_{n})^{2}$

So we should be able to state that
$$\frac{1}{X_n} \xrightarrow{P} \frac{1}{\mu}$$

We also know that
$$S_n^2$$
 is a biased estimator of population variance, but that asymptotically the bias disappears. So

$$S_n^2 \xrightarrow{P} \sigma^2$$
, and therefore $S_n^2 \xrightarrow{d} \sigma^2$

we use delta method for
$$\sqrt{arc((x:-u)^2)} = \lambda$$

$$\sqrt{nr(s^2 - \sigma^2)} \xrightarrow{d} N(0,1) \times$$

$$\int_{\infty}^{\infty} \sqrt{x_n} \int_{\infty}^{\infty} dx \int_{\infty}^{\infty} \sqrt{x_n} \left(s - s \right) \xrightarrow{d} N \left(0; \frac{1}{4} \right)$$

$$S_{n}^{2} = \frac{1}{n} \sum_{i} (X_{i} - \overline{X}_{n})^{2} = \frac{1}{n} \left[\sum_{i} X_{i}^{2} - \sum_{i} 2X_{i} \overline{X}_{n} + n \overline{X}_{n}^{2} \right]$$

$$= \frac{1}{n} \sum_{i} X_{i}^{2} - 2 \overline{X}^{2} + \overline{X}_{n}^{2}$$

$$= \frac{1}{n} \sum_{i} X_{i}^{2} - \overline{X}_{n}^{2}$$

Sh

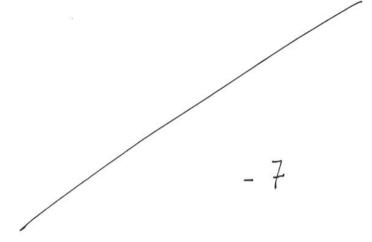
a) E(YL)

Y ~ Bernoulli (p.)

so joint distribution is

fr (y) = TT P2 (1-Pi)-y

b) The score function is what we get when we take the derivative of the log of the density function. I suppose we're meant to use what we derived in problem 1?



We want to show that charging $Y_n \xrightarrow{P} C + C + V_n \xrightarrow{P} Y$ does not strengthen Slutsky's Theorem

If instead of letting Yn converge to a constant, C, we let it converge to a normally distributed T.V., then we have

