

STAT 510

SPRING 2017

EXAM 1 SOLUTIONS

POINTS PER PROBLEM

1 a) 9

b) 9

c) 6

d) 10

e) 13

2 a) 9

b) 9

c) 5

d) 10

e) 10

3. 10

1. a) 95% Confidence Interval for $\mu_{W,4} - \mu_{B,4}$:

$$\underline{C}' \hat{\beta} \pm 1.99 \sqrt{\hat{\sigma}^2 \underline{C}' (\underline{x}' \underline{x})^{-1} \underline{C}}$$

$$\underline{x}' \underline{x} = 10 \quad I_{8 \times 8} \quad \underline{C}' = [0, 0, 0, -1, 0, 0, 0, 1]$$

$$(\underline{x}' \underline{x})^{-1} = \frac{1}{10} \quad I \quad \underline{\beta}' = [\mu_{B,1}, \mu_{B,2}, \dots, \mu_{W,4}]$$

$$\hat{\mu}_{W,4} - \hat{\mu}_{B,4} \pm 1.99 \sqrt{18.3 \left(\frac{1}{10} + \frac{1}{10} \right)}$$

$$24.5 - 22.1 \pm 1.99 \sqrt{3.66}$$

$$1 b) t = \frac{\underline{C}' \hat{\beta}}{\sqrt{\hat{\sigma}^2 \underline{C}' (\underline{x}' \underline{x})^{-1} \underline{C}}}$$

$$\underline{C}' = (1, -1, 0, 0, -1, 1, 0, 0)$$

$$t = \frac{23.0 - 27.2 - 27.8 + 23.2}{\sqrt{18.3 \cdot 4/10}}$$

$$1c) t = \frac{\hat{\beta} - \beta}{\sqrt{\hat{\sigma}^2 \hat{C}'(x'x)^{-1} C}} = \frac{\hat{\beta} - \beta + \hat{\beta} - \beta}{\sqrt{\frac{\hat{\sigma}^2}{\sigma^2} \sigma^2 \hat{C}'(x'x)^{-1} C}}$$

$$= \frac{\hat{\beta} - \beta}{\sqrt{\sigma^2 \hat{C}'(x'x)^{-1} C}} + \frac{\hat{\beta} - \beta}{\sqrt{\sigma^2 \hat{C}'(x'x)^{-1} C}}$$

$\sqrt{\hat{\sigma}^2 / \sigma^2}$

$$= \frac{Z + S}{\sqrt{W/(n-r)}} \quad \text{WHERE}$$

$$Z \equiv \frac{\hat{\beta} - \beta}{\sqrt{\sigma^2 \hat{C}'(x'x)^{-1} C}} \sim N(0, 1) \quad \text{INDEPENDENT}$$

$$\text{OR } W = \frac{(n-r)\hat{\sigma}^2}{\sigma^2} \sim \chi_{n-r}^2 \quad \text{AND}$$

$$\delta = \frac{c' \beta}{\sqrt{\sigma^2 c'(x'x)^{-1} c}}.$$

Thus, $NCP = \frac{c' \beta}{\sqrt{\sigma^2 c'(x'x)^{-1} c}}$.

For A t DISTRIBUTION, IT IS
 EASY TO DIRECTLY WRITE DOWN THE
 NCP BY REPLACING ESTIMATORS IN
 THE STATISTIC WITH THE PARAMETERS
 THEY ESTIMATE. IN THIS PROBLEM, WE
 HAVE

$$NCP = \frac{M_{B,1} - M_{B,2} - M_{W,1} + M_{W,2}}{\sqrt{2\sigma^2/5}}$$

I WAS EXPECTING STUDENTS TO WRITE DOWN THIS
 ANSWER DIRECTLY WITHOUT ALL THE WORK ABOVE.

1 d) Let $X_1 = \underline{1}$ and X_2 be the model matrix that allows a distinct mean for each group 1, 2, 3, and 4. $\text{RANK}(X_2) - \text{RANK}(X_1) = 3$.

$$\begin{aligned}
 \mathbf{y}'(\mathbf{P}_2 - \mathbf{P}_1)\mathbf{y} &= \| \mathbf{P}_2 \mathbf{y} - \mathbf{P}_1 \mathbf{y} \|^2 \\
 &= 20 (\bar{y}_{\cdot 1} - \bar{y}_{\cdot \cdot})^2 + 20 (\bar{y}_{\cdot 2} - \bar{y}_{\cdot \cdot})^2 + 20 (\bar{y}_{\cdot 3} - \bar{y}_{\cdot \cdot})^2 \\
 &\quad + 20 (\bar{y}_{\cdot 4} - \bar{y}_{\cdot \cdot})^2 \\
 &= 20 [(25.4 - 24.75)^2 + (25.2 - 24.75)^2 + (25.1 - 24.75)^2 + (23.3 - 24.75)^2] \\
 &= 57
 \end{aligned}$$

BECAUSE DATA ARE BALANCED WE KNOW ANOVA F-STAT FOR GROUP MAIN EFFECTS:

$$F = \frac{\mathbf{y}'(\mathbf{P}_2 - \mathbf{P}_1)\mathbf{y} / 3}{MSE} = \frac{57 / 3}{18.3}$$

WHAT FOLLOWS IS AN ALTERNATIVE APPROACH TO FINDING F BASED ON TEST OF $H_0: C\beta = 0$.

$$H_0: \bar{\mu}_{.1} = \bar{\mu}_{.2} = \bar{\mu}_{.3} = \bar{\mu}_{.4}$$

$\Leftrightarrow \bar{\mu}_{.1} = \bar{\mu}_{.2}$ AND $\bar{\mu}_{.3} = \bar{\mu}_{.4}$ AND

$$\frac{\bar{\mu}_{.1} + \bar{\mu}_{.2}}{2} = \frac{\bar{\mu}_{.3} + \bar{\mu}_{.4}}{2}$$

$\Leftrightarrow C\beta = 0$ WHERE

$$C = \begin{bmatrix} 1 & -1 & 0 & 0 & 1 & -1 & 0 & 0 \\ 0 & 0 & 1 & -1 & 0 & 0 & 1 & -1 \\ 1 & 1 & -1 & -1 & 1 & 1 & -1 & -1 \end{bmatrix}$$

$$\beta' = [\mu_{B,1}, \mu_{B,2}, \mu_{B,3}, \mu_{B,4}, \mu_{w,1}, \mu_{w,2}, \mu_{w,3}, \mu_{w,4}]$$

$$C(x'x)^{-1}C' = \frac{1}{10} CC' = \frac{1}{10} \begin{bmatrix} 4 & 0 & 0 \\ 0 & 4 & 0 \\ 0 & 0 & 8 \end{bmatrix}$$

$$(C(x'x)^{-1}C')^{-1} = \frac{5}{4} \begin{bmatrix} 2 & 0 & 0 \\ 0 & 2 & 0 \\ 0 & 0 & 1 \end{bmatrix}$$

$$\hat{C\beta} = \begin{bmatrix} 50.8 - 50.4 \\ 50.2 - 46.6 \\ 50.8 + 50.4 - 50.2 - 46.6 \end{bmatrix} = \begin{bmatrix} 0.4 \\ 3.6 \\ 4.4 \end{bmatrix}$$

$$(\hat{C\beta})' [C(x'x)^{-1} C']^{-1} \hat{C\beta}$$

$$= 5/4 (2*4^2 + 2*3.6^2 + 4.4^2)$$

$$F = \frac{5/4 (2*4^2 + 2*3.6^2 + 4.4^2) / 3}{18.3} = \frac{57/3}{18.3}$$

<u>Source</u>	<u>DF</u>	<u>SS</u>
RACE	1	
GROUP	3	57 ← From PART d)
RACE × GROUP	3	
ERROR	72	18.3 * 72 ← Given MSE
C. TOTAL	79	

THE ABOVE IS STRAIGHTFORWARD. ALSO, WE KNOW
 $SS_{RACE} + SS_{GROUP} + SS_{RACE \times GROUP} = 2.95 \times 7 \times 18.3$
BECAUSE THE OVERALL F STATISTIC IS

$$2.95 = \frac{\gamma' (P_4 - P_1) \gamma / 7}{MSE}$$

GIVEN IN
PROBLEM

$$= \frac{[\gamma' (P_4 - P_3) \gamma + \gamma' (P_3 - P_2) \gamma + \gamma' (P_2 - P_1) \gamma] / 7}{MSE}$$

$$= \frac{[SS_{RACE \times Group} + SS_{RACE} + SS_{Group}] / 7}{MSE}$$

18.3

IF WE CAN GET SS_{RACE} , THE REST WILL FOLLOW.

$$C = [1, 1, 1, 1, -1, -1, -1, -1]$$

$$\begin{aligned} C\hat{\beta} &= 23 + 27.2 + 27.3 + 22.1 - 27.8 - 23.2 - 22.9 - 24.5 \\ &= 99.6 - 98.4 = 1.2 \end{aligned}$$

$$C(X'X)^{-1}C^T = 8/10 \Rightarrow [C(X'X)^{-1}C^T]^{-1} = 5/4$$

$$SS_{RACE} = (\hat{\beta})' [C(X'X)^{-1}C^T]^{-1} \hat{\beta} = \frac{5 \times 1.2^2}{4} = 1.8$$

SO WE HAVE

<u>Source</u>	<u>DF</u>	<u>SS</u>
RACE	1	1.8
GROUP	3	57
RACE × GROUP	3	$2.95 \times 7 \times 18.3 - SS_{RACE} - SS_{GROUP}$
ERROR	72	18.3×72
C. TOTAL	79	Sum of All The Above OR $18.3 \times 72 + 2.95 \times 7 \times 18.3$

$$2 a) SS_{C.TOTAL} = 9.63 + 6.37 = 16 = \gamma'(I - P_1)\gamma$$

THIS IS SSE FOR INTERCEPT ONLY MODEL.

$$\gamma'(P_2 - P_1)\gamma = 0.19$$

$$\begin{aligned} \gamma'(I - P_2)\gamma &= \gamma'(I - P_1)\gamma - \gamma'(P_2 - P_1)\gamma \\ &= 16 - 0.19 = 15.81 \end{aligned}$$

$$G^2 = \frac{\gamma'(I - P_2)\gamma}{q - 2} = \frac{15.81}{7}$$

$$2b) F = \frac{0.19/1}{15.81/7}$$

$$2c) 142.124$$

<u>Source</u>	<u>DF</u>	<u>SS</u>
TRT	2	9.63
X	1	?
<u>ERROR</u>	<u>5</u>	<u>0.27</u>
C. TOTAL	8	16.0

$$SS_x = 16 - (9.63 + 0.27) = 16 - 9.9 = 6.1$$

$$F = \frac{6.1}{0.27/5} \Rightarrow |t| = \sqrt{\frac{30.5}{0.27}}$$

INSPECTION OF DATA SHOWS THAT Y DECREASES AS X INCREASES WITHIN EACH TREATMENT GROUP. THUS, $\hat{\alpha}_2 < 0$, AND $t = -\sqrt{\frac{30.5}{0.27}}$

2e) THIS PROBLEM ASKS FOR A BRIEF REPORT "FOR THE RESEARCHERS". THUS, I WAS HOPING YOU WOULD FOCUS ON CONCLUSIONS RELEVANT FOR THE RESEARCHERS, AND EXPLAIN THOSE CONCLUSIONS IN A WAY THAT MIGHT MAKE SENSE TO A SCIENTIFICALLY LITERATE NON-STATISTICIAN.

EVEN WITHOUT A CALCULATOR, YOU SHOULD BE ABLE TO SEE THAT THE t-STAT COMPUTERED IN PART (d) IS VERY LARGE IN MAGNITUDE. THUS, BOTH PRE-TREATMENT WEIGHT AND TREATMENT (SEE PART c) ARE HIGHLY STATISTICALLY SIGNIFICANT IN A MODEL THAT INCLUDES BOTH. WE CAN ALSO SEE THAT AN ADDITIVE MODEL ~~SEEMS~~ ADEQUATE COMPARED TO A MORE COMPLEX MODEL THAT ALLOWS FOR INTERACTION BETWEEN PRE-TREATMENT WEIGHT AND TREATMENT. THIS ADDITIVE MODEL SAYS THE EXPECTED VALUE OF THE RESPONSE IS A LINEAR FUNCTION OF PRETREATMENT WEIGHT WITHIN EACH TREATMENT GROUP, THUS,

THESE ARE THREE REGRESSION LINES, ONE FOR EACH TREATMENT GROUP. THE SLOPES OF THE THREE LINES ARE THE SAME IN THE ADDITIVE MODEL, BUT THE INTERCEPTS DIFFER ACROSS TREATMENT GROUPS. WE COULD SAY SOMETHING LIKE THE FOLLOWING TO THE RESEARCHERS.

"BOTH PRE-TREATMENT WEIGHT AND TREATMENT HAVE A STATISTICALLY SIGNIFICANT ASSOCIATION WITH THE CHEMICAL LEVEL RESPONSE. WITHIN EACH TREATMENT GROUP, LARGER PRE-TREATMENT WEIGHTS ARE ASSOCIATED WITH LOWER LEVELS OF THE CHEMICAL. FOR ANY PARTICULAR PRE-TREATMENT WEIGHT HELD CONSTANT ACROSS THE TREATMENT GROUPS, THE MEAN LEVEL OF THE CHEMICAL DIFFERS IN A STATISTICALLY SIGNIFICANT WAY ACROSS TREATMENT GROUPS, AND THE DIFFERENCES AMONG TREATMENT GROUPS IN MEAN RESPONSE APPEAR SIMILAR, REGARDLESS OF WHICH PRE-TREATMENT WEIGHT WE CONSIDER."

3. A IS THE MODEX MATRIX FOR A QUADRATIC REGRESSION FUNCTION
FOR DATA WITH TWO OBSERVATIONS AT EACH OF $x=1$, $x=2$, AND $x=3$.

BECAUSE A QUADRATIC FUNCTION CAN PASS THROUGH THE THREE POINTS

$(1, M_1)$, $(2, M_2)$, $(3, M_3)$ FOR ANY M_1, M_2, M_3 , THIS IS JUST ANOTHER VERSION OF THE CELL MEANS MODEX MATRIX WITH THREE MEANS, i.e.,

$C(A) = C(X)$, WHERE

$$X = \begin{pmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \end{pmatrix} . \text{ THUS, } P_A = P_X$$

$$\begin{aligned} P_X &= X(X'X)^{-1}X' \\ &= X \frac{1}{2} I X' = \frac{1}{2} \begin{pmatrix} 1 & 1 & 0 & 0 & 0 & 0 \\ 1 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 1 & 0 & 0 \\ 0 & 0 & 1 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 1 \\ 0 & 0 & 0 & 0 & 1 & 1 \end{pmatrix} \end{aligned}$$

Now Note That $C(A) \subset C(B)$ because

$$B \begin{bmatrix} 1 & 2 & 4 \\ 0 & 1 & 4 \\ 0 & 0 & 1 \\ 0 & 0 & 0 \end{bmatrix} = A.$$

This can also be seen by noting that the first 3 columns of B are like A but with a centered X variable $(X - \bar{X})$.

Thus, $P_A P_B = P_A = P_X$.