Estimating Functionals of the Joint Distribution of Potential Outcomes with Optimal Transport

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The fundamental problem of causal inference

It is impossible to observe the [treated outcome] and [untreated outcome] on the same unit and, therefore, it is impossible to observe the effect...

(Holland, 1986)

- Parameters of the joint distribution of potential outcomes are not point identified.
- ► This paper
 - shows optimal transport characterizes sharp bounds,
 - accomodates noncompliance through a standard IV model, and
 - provides simple, computationally convenient estimators.

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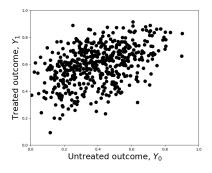
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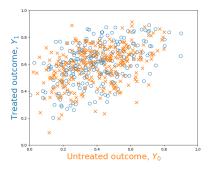
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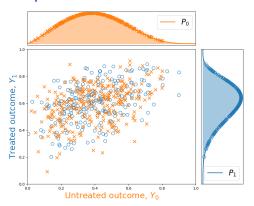


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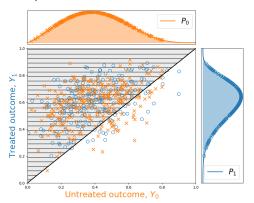
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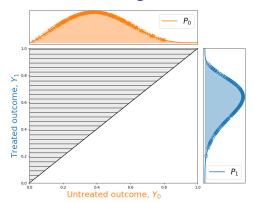


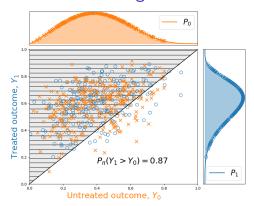
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- ▶ For example, what share of units benefit from treatment?

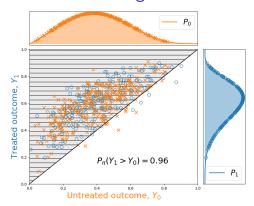
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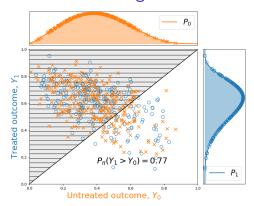
Many joint distributions π share marginal distributions P_1 , P_0 :

$$\Pi(P_1, P_0) = \{\pi : \pi_1 = P_1, \ \pi_0 = P_0\}$$



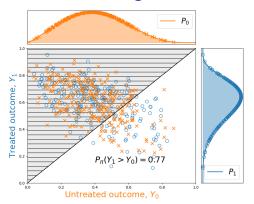
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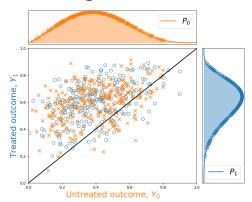


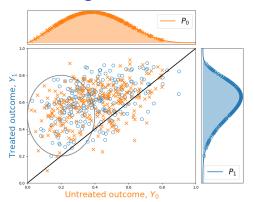
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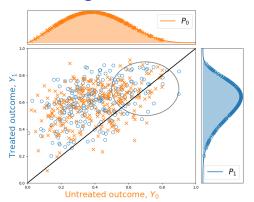
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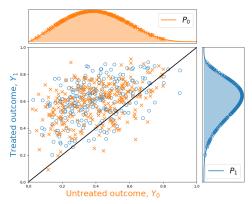
▶ Optimizing $P(Y_1 > Y_0)$ over $\Pi(P_1, P_0)$ implies bounds:

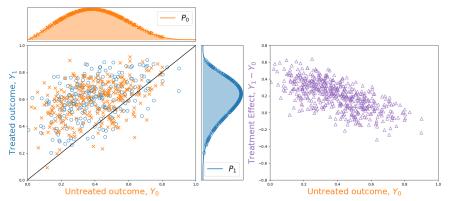
$$\min_{\pi \in \Pi(P_1, P_0)} P_\pi(Y_1 > Y_0) \qquad \qquad \max_{\pi \in \Pi(P_1, P_0)} P_\pi(Y_1 > Y_0)$$

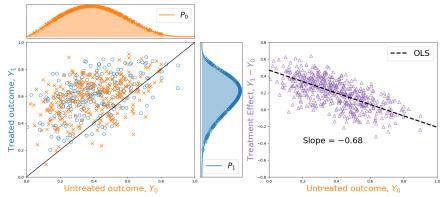


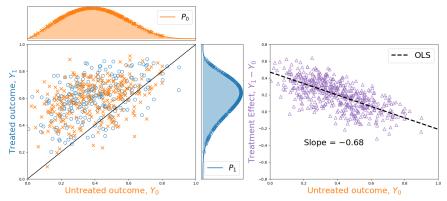




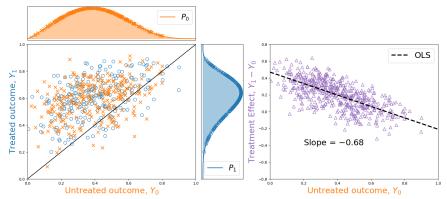




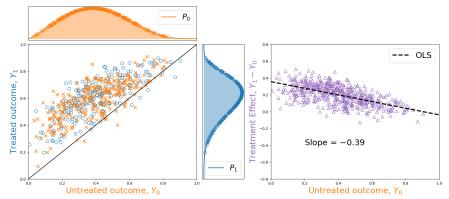




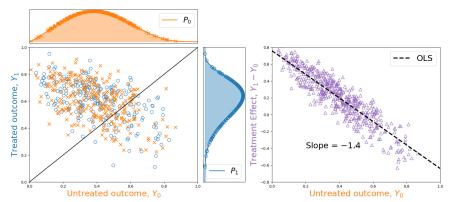
OLS slope
$$=\frac{\mathsf{Cov}(Y_1-Y_0,Y_0)}{\mathsf{Var}(Y_0)}$$



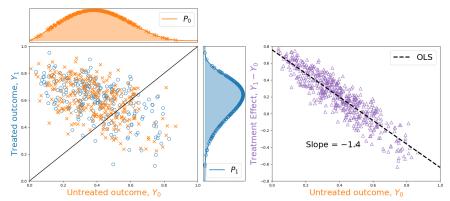
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▶ Do those with smaller Y_0 see larger $Y_1 - Y_0$?

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• Optimizing $E[(Y_1 - Y_0)Y_0]$ over $\Pi(P_1, P_0)$ implies bounds on OLS slope:

$$\min_{\pi \in \Pi(P_1, P_0)} E_{\pi}[(Y_1 - Y_0)Y_0]$$

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Parameter of interest:

$$\gamma = g(\theta, \eta) \in \mathbb{R},$$

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- Propose and study sample analogue estimators of the bounds.
- ▶ Empirical application: who sees larger benefits from the NSW job training?

Related literature

- Joint distribution of potential outcomes
 - CDF or quantiles of $Y_1 Y_0$: Manski (1997), Heckman et al. (1997), Firpo (2007), Fan and Park (2010), Fan and Park (2012), Firpo and Ridder (2019), Callaway (2021), Frandsen and Lefgren (2021).
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- 2 Identification
- Estimators
- 4 Simulations
- 6 Application

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$$Y\in\mathcal{Y}\subseteq\mathbb{R}, \hspace{1cm} D\in\{0,1\}, \hspace{1cm} X\in\mathcal{X}=\{x_1,\dots,x_M\}$$

generated from a distribution satisfying

- (i) Potential outcomes: $Y = DY_1 + (1 D)Y_0$
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- In the paper, binary IV satisfying monotonicity condition (Imbens and Angrist, 1994).

Setting w/IV

Parameter of interest:

$$\gamma=g(\theta,\eta)\in\mathbb{R}$$
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Assumption 2 (Cost function) Either

- (i) $c(y_1, y_0)$ is Lipschitz continuous and $\mathcal Y$ is compact, or
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<u>Remark:</u> If $c(y_1, y_0) = \mathbb{1}\{y_1 - y_0 \le \delta\}$ but $F_{d|x}(\cdot)$ are not continuous, inference remains valid for an outer identified set.

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Assumption 3 (Function of moments, simplified)

- (i) $\eta_1(Y)$ and $\eta_0(Y)$ have finite second moments,
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Remark: Assumption 3 (iii) is relaxed in the paper.



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- ▶ Quantiles of $Y_1 Y_0$
 - Median is more representative than mean when distribution is skewed.

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$$OT_c(P_1, P_0) = \inf_{\pi \in \Pi(P_1, P_0)} E_{\pi}[c(Y_1, Y_0)]$$

- Choose a joint distribution with given marginals to minimize costs.
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Daniel Ober-Revnolds (UCLA)

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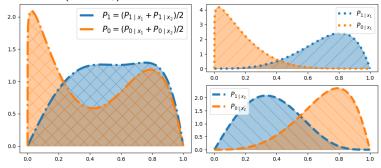
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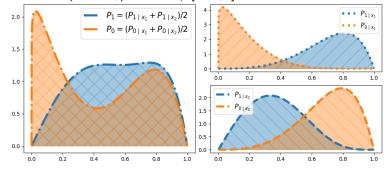
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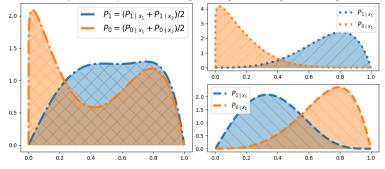
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- **Dounds on** $P(Y_1 > Y_0)$: not sharp [0.25, 1], sharp: [0.44, 0.68].



Theorem: identification

For continuous c,

$$\begin{array}{ll} \text{Bounds on } \theta_{x}: & \theta_{x}^{L} = OT_{c}(P_{1|x}, P_{0|x}), & \theta_{x}^{H} = -OT_{-c}(P_{1|x}, P_{0|x}) \\ \text{Bounds on } \theta: & \theta^{L} = E[\theta_{X}^{L}] & \theta^{H} = E[\theta_{X}^{H}] \\ \text{Bounds on } \gamma: & \gamma^{L} = \min_{t \in [\theta^{L}, \theta^{H}]} g(t, \eta), & \gamma^{H} = \max_{t \in [\theta^{L}, \theta^{H}]} g(t, \eta) \\ \end{array}$$

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Theorem (identification)

Suppose assumptions 1, 2, and 3 are satisfied. Then the sharp identified set for $\gamma = g(\theta, \eta)$ is $[\gamma^L, \gamma^H]$.





Quantile details

Overview

- Setting and parameter class
- 2 Identification
- Stimators
- 4 Simulations
- 6 Application

$$OT_c(P_1, P_0) = \underbrace{\inf_{\pi \in \Pi(P_1, P_0)} E_{\pi}[c(Y_1, Y_0)]}_{\text{Primal Problem}}$$

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- ▶ The primal problem is used in identification.
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Estimators: recall identification

Distributions of Y_d | X = x ~ P_{d|x}:

$$E_{P_{d|x}}[f(Y_d)] = \frac{E[f(Y)1\{D=d,X=x\}]}{P(D=d,X=x)}$$

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▶ The identified set for γ is $[\gamma^L, \gamma^H]$, where for c continuous,

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Estimators: sample analogues

• Estimate $P_{d|x}$ with sample analogues $\hat{P}_{d|x}$:

$$E_{\hat{P}_{d|x}}[f(Y_d)] = \frac{\frac{1}{n} \sum_{i=1}^{n} f(Y_i) \mathbb{1} \{D_i = d, X_i = x\}]}{\frac{1}{n} \sum_{i=1}^{n} \mathbb{1} \{D_i = d, X_i = x\}}$$

Using strong duality,

$$OT_c(\hat{P}_{1|x}, \hat{P}_{0|x}) = \max_{(\varphi, \psi) \in \Phi_c} E_{\hat{P}_{1|x}}[\varphi(Y_1)] + E_{\hat{P}_{0|x}}[\psi(Y_0)].$$

Estimate the endpoints of $[\gamma^L, \gamma^H]$ with plug-in estimators. For c continuous,

$$\begin{split} \hat{\theta}_{x}^{L} &= OT_{c}(\hat{P}_{1|x}, \hat{P}_{0|x}), & \hat{\theta}_{x}^{H} &= -OT_{-c}(\hat{P}_{1|x}, \hat{P}_{0|x}) \\ \hat{\theta}^{L} &= \frac{1}{n} \sum_{i=1}^{n} \hat{\theta}_{X_{i}}^{L}, & \hat{\theta}^{H} &= \frac{1}{n} \sum_{i=1}^{n} \hat{\theta}_{X_{i}}^{H} \\ \hat{\gamma}^{L} &= \min_{t \in [\hat{\theta}^{L}, \hat{\theta}^{H}]} g(t, \hat{\eta}), & \hat{\gamma}^{H} &= \max_{t \in [\hat{\theta}^{L}, \hat{\theta}^{H}]} g(t, \hat{\eta}) \end{split}$$

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▶ To evaluate $E_{\hat{P}_{d|x}}[f(Y_d)]$ for any function f,

$$E_{\hat{P}_{d|x}}[f(Y_d)] = \sum_{i=1}^n \omega_{d,x,i} \times f(Y_i), \qquad \omega_{d,x,i} = \frac{\mathbb{1}\{D_i = d, X_i = x\}/n}{\frac{1}{n} \sum_{j=1}^n \mathbb{1}\{D_j = d, X_j = x\}}.$$

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► Computating $OT_c(\hat{P}_{1|x}, \hat{P}_{0|x})$ is straightforward linear programming:

$$\begin{split} OT_c(\hat{P}_{1|x},\hat{P}_{0|x}) &= \max_{\{\varphi_i,\psi_i\}_{i=1}^n} \sum_{i=1}^n \omega_{1,x,i} \times \varphi_i + \sum_{i=1}^n \omega_{0,x,i} \times \psi_i \\ \text{s.t. } \varphi_i + \psi_j &\leq c(Y_i,Y_j) \text{ for all } 1 \leq i,j \leq n, \end{split}$$

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▶ Dimension is reduced by ignoring φ_i , ψ_i , and constraints where $\omega_{d,x,i} = 0$.

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Convergence in distribution: theorem

Let P be the distribution of an observation, and \mathbb{P}_n the empirical distribution.

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$$(\hat{\gamma}^L, \hat{\gamma}^H) = T(\mathbb{P}_n), \qquad (\gamma^L, \gamma^H) = T(P)$$

 $T(\cdot)$ details

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Theorem (Weak convergence)

Suppose assumptions 1, 2, and 3 hold. Then

$$\sqrt{n}((\hat{\gamma}^L, \hat{\gamma}^H) - (\gamma^L, \gamma^H)) \stackrel{L}{\to} T_P'(\mathbb{G})$$

where $\sqrt{n}(\mathbb{P}_n - P) \stackrel{L}{\to} \mathbb{G}$ and $T'_P(\cdot)$ is the Hadamard directional derivative of $T(\cdot)$ at P.

 $T(\cdot)$ details

Proof sketch

Estimating the asymptotic distribution is necessary for inference.

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- ► The bootstrap provides an attractive procedure.
 - Bootstrap draw: $\{Y_i^*, D_i^*, X_i^*\}_{i=1}^n$
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- ► The bootstrap provides an attractive procedure.
 - Bootstrap draw: $\{Y_i^*, D_i^*, X_i^*\}_{i=1}^n$
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- ► Compute $T(\mathbb{P}_n^*)$ the same way as $T(\mathbb{P}_n)$: let $\omega_{d,x,i}^* = \frac{\mathbb{1}\{D_i^* = d, X_i^* = x\}/n}{\frac{1}{n}\sum_{j=1}^n \mathbb{1}\{D_j^* = d, X_j^* = x\}}$,

$$\begin{split} OT_c(\hat{P}_{1|x}^*, \hat{P}_{0|x}^*) &= \max_{\{\varphi_i, \psi_i\}_{i=1}^n} \sum_{i=1}^n \omega_{1,x,i}^* \varphi_i + \sum_{i=1}^n \omega_{0,x,i}^* \psi_i \\ \text{s.t. } \varphi_i + \psi_j &\leq c(Y_i, Y_j) \text{ for all } 1 \leq i, j \leq n \end{split}$$

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Precise assumption 4

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Theorem (Bootstrap consistency)

Suppose assumptions 1, 2, 3, and 4 hold. Then $T_P'(\mathbb{G})$ is bivariate normal, and conditional on $\{Y_i, D_i, X_i\}_{i=1}^n$,

$$\sqrt{n}(T(\mathbb{P}_n^*) - T(\mathbb{P}_n)) \stackrel{L}{\to} T'_P(\mathbb{G})$$

in outer probability.

Precise assumption 4

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Lemma (Unique solutions) Suppose that

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Assumption 4 may hold without this lemma's conditions.

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- ► The paper shows a consistent alternative.
 - Follows Fang and Santos (2019): estimating the derivative $T_P'(\cdot)$.
 - Implementation is more involved, but still computationally tractable.

Overview

- Setting and parameter class
- 2 Identification
- Estimators
- 4 Simulations
- 6 Application

Simulations: parameter and DGP

Parameter $\gamma = \theta = P(Y_1 - Y_0 \le \delta)$ has simple bounds:

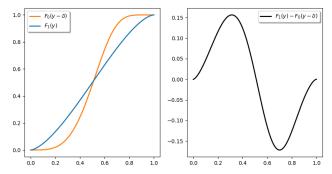
$$\gamma^{L} = \sup_{y} \left\{ F_{1}(y) - F_{0}(y - \delta) \right\}, \qquad \gamma^{H} = 1 + \inf_{y} \left\{ F_{1}(y) - F_{0}(y - \delta) \right\}$$

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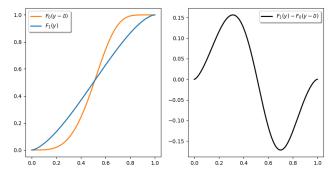


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▶ Unique solutions ⇒ bootstrap is valid.

▶ Asymptotic $1 - \alpha$ confidence set for $[\gamma^L, \gamma^H]$:

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 - (i) Using $\{Y_i, D_i, X_i\}_{i=1}^n$, compute estimators:

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(ii) For each $b=1,\ldots,B$, draw $\{Y_{i,b}^*,D_{i,b}^*,X_{i,b}^*\}_{i=1}^n$ to define $\mathbb{P}_{n,b}^*$ and compute:

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(iii) Let $\hat{c}_{1-\alpha}$ be the $1-\alpha$ quantile of $\{\max\{\sqrt{n}(\hat{\gamma}_b^{L*}-\hat{\gamma}),-\sqrt{n}(\hat{\gamma}_b^{H*}-\hat{\gamma}^H)\}\}_{b=1}^B$, and

$$CI = [\hat{\gamma}^L - \hat{c}_{1-\alpha}/\sqrt{n}, \hat{\gamma}^H + \hat{c}_{1-\alpha}/\sqrt{n}]$$

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- Bootstrap bias correction (Efron and Tibshirani, 1994; Horowitz, 2001):

$$(\widehat{bias}^{L}, \widehat{bias}^{H}) = \frac{1}{B} \sum_{b=1}^{B} (\hat{\gamma}^{L*}, \hat{\gamma}^{H*}) - (\hat{\gamma}^{L}, \hat{\gamma}^{H}),$$
$$\hat{\gamma}_{BC}^{L} = \hat{\gamma}^{L} - \widehat{bias}^{L}, \qquad \qquad \hat{\gamma}_{BC}^{H} = \hat{\gamma}^{H} - \widehat{bias}^{H}$$

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$$\begin{split} &(\widehat{\textit{bias}}^L, \widehat{\textit{bias}}^H) = \frac{1}{B} \sum_{b=1}^B (\hat{\gamma}^{L*}, \hat{\gamma}^{H*}) - (\hat{\gamma}^L, \hat{\gamma}^H), \\ &\hat{\gamma}^L_{BC} = \hat{\gamma}^L - \widehat{\textit{bias}}^L, & \hat{\gamma}^H_{BC} = \hat{\gamma}^H - \widehat{\textit{bias}}^H \end{split}$$

Bootstrap bias corrected confidence interval:

$$CI_{BC} = [\hat{\gamma}_{BC}^L - \hat{c}_{1-\alpha}/\sqrt{n}, \hat{\gamma}_{BC}^H + \hat{c}_{1-\alpha}/\sqrt{n}]$$

Simulations: results

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n	Bias		St.	Dev.	Emp. Coverage
	$\hat{\gamma}^L$	$\hat{\gamma}^H$	$\hat{\gamma}^L$	$\hat{\gamma}^H$	CI
100	0.047	-0.051	0.065	0.066	0.900
200	0.031	-0.031	0.049	0.049	0.917
300	0.030	-0.021	0.040	0.040	0.893

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n	Bias		St.	Dev.	Emp. Coverage		
	$\hat{\gamma}_{BC}^{L}$	$\hat{\gamma}^{H}_{BC}$	$\hat{\gamma}^{L}_{BC}$	$\hat{\gamma}^{H}_{BC}$	CI _{BC}		
100	0.021	-0.026	0.071	0.071	0.927		
200	0.013	-0.015	0.052	0.051	0.953		
300	0.015	-0.007	0.042	0.042	0.957		

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► The National Supported Work Demonstration Program (NSW)

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Table: Balance table

	base inc.	age	yrs. educ.	HS dropout	black	hispanic	married
control	3672.49 (6521.53)	24.45	10.19	0.81	0.80	0.11	0.16
	(6521.53)	(6.59)	(1.62)	(0.39)	(0.40)	(0.32)	(0.36)
treated	3571.00 (5773.13)	24.63 (6.69)	10.38	0.73	0.80	0.09	0.17
	(5773.13)	(6.69)	(1.82)	(0.44)	(0.40)	(0.29)	(0.37)

Note: Standard deviations in parentheses.

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- ▶ Parameter: The OLS slope coefficient $Y_1 Y_0 = \alpha + \gamma Y_0 + \varepsilon$

$$\gamma = \frac{\text{Cov}(Y_1 - Y_0, Y_0)}{\text{Var}(Y_0)} = \underbrace{\frac{E[(Y_1 - Y_0)Y_0]}{E[Y_0]} - (E[Y_1] - E[Y_0])E[Y_0]}_{E[Y_0^2] - (E[Y_0])^2}$$

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$$\gamma = \frac{\mathsf{Cov}(Y_1 - Y_0, Y_0)}{\mathsf{Var}(Y_0)} = \frac{\overbrace{E[(Y_1 - Y_0)Y_0]}^{\theta} - (E[Y_1] - E[Y_0])E[Y_0]}{E[Y_0^2] - (E[Y_0])^2}$$

Interpretation: $\gamma < 0$ implies workers with below average Y_0 tend to see above average $Y_1 - Y_0$

- ▶ Discretized age and baseline income are informative covariates.
 - age bins: $[16, 23], (23, \infty)$
 - baseline income bins: [0,0], (0,4000], $(4000,\infty)$

Table: Estimates of bounds for γ , the OLS Slope

	Lower Bound	Upper Bound	95% <i>CI</i>
No Covariates			
Disc. Age and Inc.			
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NSW results: conditional on covariate values

Table: Estimates conditional on covariate values

age	base inc.	$\hat{\gamma}^{\it L}_{\it BC}$	$\hat{\gamma}^{H}_{BC}$	95% CI _{BC}	n
	0	-1.97	0.28	[-2.26, 0.56]	140
(16, 23]	(0, 4000]	-1.74	-0.15	[-1.9, 0.01]	141
	$(4000, \infty)$	-1.45	-0.44	[-1.63, -0.27]	90
	0	-2.13	0.81	[-2.65, 1.33]	187
$(23, \infty)$	(0, 4000]	-1.39	-0.16	[-1.93, 0.38]	56
	(4000, ∞)	-1.66	0.03	[-2.08, 0.45]	108

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- lacktriangle Among young men with + base income, low Y_0 is associated with high Y_1-Y_0 .
- ► This subset's vulnerable individuals see larger benefits from treatment.

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 - Parameters of the joint distribution of potential outcomes are not point identified.
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- Ongoing and future work:
 - Accomodate plausible support restrictions, such as $Y_1 \ge Y_0$.
 - Support function approach to consider parameters depending on more than one joint moment.

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Appendix: full setting

Assumption 1 (Setting). $\{Y_i, D_i, Z_i, X_i\}_{i=1}^n$ is an i.i.d. sample, with

$$Y \in \mathcal{Y} \subseteq \mathbb{R}, \qquad D \in \{0,1\}, \qquad Z \in \{0,1\}, \qquad X \in \mathcal{X} = \{x_1, \dots, x_M\}$$

generated from a distribution satisfying

- (i) Potential outcomes: $Y = DY_1 + (1 D)Y_0$,
- (ii) Potential treatment statuses: $D=ZD_1+(1-Z)D_0$, with $D_z\in\{0,1\}$,
- (iii) Instrument exogeneity: $(Y_1, Y_0, D_1, D_0) \perp Z \mid X$,
- (iv) Monotonicity: $D_1 \ge D_0$ almost surely,
- (v) Existence of compliers: $P(D_1 > D_0, X = x) > 0$ for each x, and
- (vi) P(X = x, Z = z) > 0 for each (x, z)
- ► Terminology: always-taker, complier, defier, never-taker.

	$D_0=1$	$D_0=0$
$D_1 = 1$	Always-takers	Compliers
$D_1 = 0$	Defiers	Never-takers

Monotonicity rules out defiers. Focus on distribution of compliers.





Appendix: identification of $P(Y_1 - Y_0 < \delta)$

 $ightharpoonup OT_c(P_1, P_0)$ is well behaved (attained, strong duality holds, etc) when $c(y_1, y_0)$ is bounded and lower semicontinuous

▶ If
$$c(y_1, y_0) = \mathbb{1}\{y_1 - y_0 \le \delta\}$$
, let
$$c_L(y_1, y_0) = \mathbb{1}\{y_1 - y_0 < \delta\}, \qquad c_H(y_1, y_0) = \mathbb{1}\{y_1 - y_0 > \delta\}$$
$$\theta_x^L = OT_{c_L}(P_{1|x}, P_{0|x}), \qquad \theta_x^H = 1 - OT_{c_H}(P_{1|x}, P_{0|x})$$

The form of the bounds remains the same:

$$\begin{aligned} \theta^L &= E[\theta_X^L], & \theta^H &= E[\theta_X^H] \\ \gamma^L &= \min_{t \in [\theta^L, \theta^H]} g(t, \eta), & \gamma^H &= \max_{t \in [\theta^L, \theta^H]} g(t, \eta) \end{aligned}$$

Identified sets are still sharp when CDFs are continuous:

$$F_{d|x}(y) = P(Y_d \le y \mid X = x)$$









Appendix: aside, CDF results are conservative when continuity fails

$$OT_c(P_1, P_0) = \inf_{\pi \in \Pi(P_1, P_0)} E_{\pi}[c(Y_1, Y_0)]$$

▶ Bounds on $\theta = P(Y_1 - Y_0 \le \delta)$ are found with

$$c_{L}(y_{1}, y_{0}) = \mathbb{1}\{y_{1} - y_{0} < \delta\}, \qquad c_{H}(y_{1}, y_{0}) = \mathbb{1}\{y_{1} - y_{0} > \delta\},\$$

$$\theta^{L} = OT_{c_{L}}(P_{1}, P_{0}), \qquad \theta^{H} = 1 - OT_{c_{H}}(P_{1}, P_{0})$$

Using OT results, show that if marginal CDFs F_d are continuous then $\Theta_{ID} = [\theta^L, \theta^H]$.







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▶ As a byproduct, recover the famed Makarov bounds studied by Fan and Park (2010)

$$\theta^{L} = \sup_{y} \{F_{1}(y) - F_{0}(y - \delta)\}, \qquad \qquad \theta^{H} = 1 + \inf_{y} \{F_{1}(y) - F_{0}(y - \delta)\}$$









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▶ Furthermore, $\mathbb{1}\{y_1 - y_0 < \delta\} \le \mathbb{1}\{y_1 - y_0 \le \delta\}$ implies the bounds are conservative: $\Theta_{ID} \subseteq [\theta^L, \theta^H]$ whether or not F_d are continuous.



Identification Thm

Estimators

Appendix: full assumption 3

Parameter of interest:

$$\gamma = \mathbf{g}(\theta, \eta) \in \mathbb{R}$$

where $\theta = E[c(Y_1, Y_0)] \in \mathbb{R}$ and $\eta = (E[\eta_1(Y_1)], E[\eta_0(Y_0)]) \in \mathbb{R}^{K_1 + K_0}$.

Assumption 3 (Function of moments)

- (i) $E[\|\eta_d(Y)\|^2] < \infty$ for d = 1, 0,
- (ii) $g(\cdot, \eta)$ is continuous, and
- (iii) the functions

$$g^{L}(t^{L}, t^{H}, e) = \min_{t \in [t^{L}, t^{H}]} g(t, e),$$
 $g^{H}(t^{L}, t^{H}, e) = \max_{t \in [t^{L}, t^{H}]} g(t, e)$

are continuously differentiable at $(t^L, t^H, e) = (\theta^L, \theta^H, \eta)$.

Remark: A3 (ii), (iii) implied by g continuously differentiable and $g(\cdot, \eta)$ monotonic



Appendix: quantiles

Suppose the parameter of interest is q_{τ} solving

$$P(Y_1 - Y_0 \le q_\tau) = \tau$$

▶ View CDF bounds as a function: $\theta(\delta) = P(Y_1 - Y_0 \le \delta)$

$$\begin{split} c_{L,\delta}(y_1,y_0) &= \mathbb{I}\{y_1 - y_0 < \delta\}, \\ \theta_x^L(\delta) &= OT_{c_L}(P_{1|x}, P_{0|x}), \\ \theta^L(\delta) &= E[\theta_X^L(\delta)] \end{split} \qquad \begin{aligned} c_{H,\delta}(y_1,y_0) &= \mathbb{I}\{y_1 - y_0 > \delta\}, \\ \theta_x^H(\delta) &= 1 - OT_{c_H}(P_{1|x}, P_{0|x}), \\ \theta^H(\delta) &= E[\theta_X^H(\delta)] \end{aligned}$$

and let $Q_{I, au}$ be the sharp identified set for $q_ au.$

Lemma (Identification of q_{τ}). Suppose assumptions 1 and 2(ii) hold. Then $q \in Q_{l,\tau}$ if and only if $\theta^{L}(q) < \tau < \theta^{H}(q)$.



Identification extends easily to IV.

Ident. Thm.

- Identification extends easily to IV.
- Consider the binary IV potential outcomes framework of Abadie (2003):

$$D = ZD_1 + (1 - Z)D_0$$
 $(Y_1, Y_0, D_1, D_0) \perp Z \mid X,$ $D_1 \geq D_0$

units with $D_1 > D_0$ are known as *compliers*.

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Same identification applies to parameters conditional on compliance. E.g.,

$$P(Y_1 > Y_0 \mid D_1 > D_0)$$



Appendix: definition of T

▶ Proof defines a set of universally bounded functions

$$\mathcal{F} \subseteq \{f: \mathcal{Y} \times \{0,1\} \times \mathcal{X} \to \mathbb{R}\}$$

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▶ View \mathbb{P}_n , P as bounded functions on \mathcal{F} :

$$\ell^\infty(\mathcal{F}) = \left\{ g: \mathcal{F} o \mathbb{R} \; ; \; \|g\|_\infty = \sup_{f \in \mathcal{F}} |g(f)| < \infty
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ight\}$$

▶ The map $T:\ell^\infty(\mathcal{F}) \to \mathbb{R}^2$ is described by $P \mapsto (P_{1|x}, P_{0|x}, \eta)$ and

$$\begin{aligned} \theta_x^L &= OT_c(P_{1|x}, P_{0|x}), & \theta_x^H &= -OT_{-c}(P_{1|x}, P_{0|x}) \\ \theta^L &= E[\theta_X^L], & \theta^H &= E[\theta_X^H] \\ \gamma^L &= \min_{t \in [\theta^L, \theta^H]} g(t, \eta), & \gamma^H &= \max_{t \in [\theta^L, \theta^H]} g(t, \eta) \end{aligned}$$

Appendix: proof sketch (1/3)

1. Will view P, \mathbb{P} as maps in $\ell^{\infty}(\mathcal{F})$ for Donsker set \mathcal{F} (defined later), and $T:\ell^{\infty}(\mathcal{F})\to\mathbb{R}^2$.

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- 2. To show $T(\cdot)$ is (Hadamard) directionally differentiable, suffices to show OT_c is directionally differentiable.
- 3. By strong duality,

$$\begin{split} OT_c(P_{1|x}, P_{0|x}) &= \sup_{(\varphi, \psi) \in \Phi_c} E_{P_{1|x}}[\varphi(Y_1)] + E_{P_{0|x}}[\psi(Y_0)] \\ \Phi_c &= \{(\varphi, \psi) : \varphi(y_1) + \psi(y_0) \le c(y_1, y_0)\} \end{split}$$

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$$\sup_{(\varphi,\psi)\in\Phi_c} E_{P_1}[\varphi(Y_1)] + E_{P_0}[\psi(Y_0)] = \sup_{(\varphi,\psi)\in\Phi_c\cap(\mathcal{F}_c\times\mathcal{F}_c^c)} E_{P_1}[\varphi(Y_1)] + E_{P_0}[\psi(Y_0)] \quad (1)$$

- (i) if $c(y_1, y_0)$ is L-Lip. and \mathcal{Y} is compact, \mathcal{F}_c and \mathcal{F}_c^c are L-Lip. and universally bounded.
- (ii) if $c(y_1, y_0) = \mathbb{1}\{y_1 y_0 \le \delta\}$, \mathcal{F}_c is the set of intervals, \mathcal{F}_c^c the complements of intervals.

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 - \implies OT_c , and therefore $T(\cdot)$, are Hadamard directionally differentiable.

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Appendix: proof sketch (3/3)

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- 8. \mathcal{F} is Donsker $\implies \sqrt{n}(\mathbb{P}_n P) \stackrel{L}{\to} \mathbb{G}$ in $\ell^{\infty}(\mathcal{F})$.
- 9. Functional delta method implies the result,

$$\sqrt{n}((\hat{\gamma}^L, \hat{\gamma}^H) - (\gamma^L, \gamma^H)) = \sqrt{n}(T(\mathbb{P}_n) - T(P)) \stackrel{L}{\to} T'_P(\mathbb{G}).$$

$$OT_c(P_1, P_0) = \sup_{(\varphi, \psi) \in \Phi_c} \underbrace{E_{P_1}[\varphi(Y_1)] + E_{P_0}[\psi(Y_0)]}_{J(\varphi, \psi)},$$

▶ Define the *c*-transforms:

$$\varphi^{c}(y_{0}) = \inf_{y_{1}} \{c(y_{1}, y_{0}) - \varphi(y_{1})\}, \qquad \qquad \psi^{c}(y_{1}) = \inf_{y_{0}} \{c(y_{1}, y_{0}) - \psi(y_{0})\}$$

call φ^c (and ψ^c) c-concave functions.

Proof sketch

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For any $(\varphi, \psi) \in \Phi_c = \{(\varphi, \psi) ; \varphi(y_1) + \psi(y_0) \le c(y_1, y_0)\},$

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Proof sketcl

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Proof sketc

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 - (iii) $J(\varphi, \psi) \leq J(\varphi, \varphi^c)$ by monotonicity of $E_{P_d}[\cdot]$.

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- c-concave functions often inherit properties of c:

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$$\varphi^{c}(y_{0}) = \inf_{y_{1}} \{c(y_{1}, y_{0}) - \varphi(y_{1})\}, \qquad \qquad \psi^{c}(y_{1}) = \inf_{y_{0}} \{c(y_{1}, y_{0}) - \psi(y_{0})\}$$

call φ^c (and ψ^c) c-concave functions.

- For any $(\varphi, \psi) \in \Phi_c = \{(\varphi, \psi) ; \varphi(y_1) + \psi(y_0) \le c(y_1, y_0)\},$
 - (i) $(\varphi, \varphi^c), \in \Phi_c$
 - (ii) If $(\varphi, \psi) \in \Phi_c$, then $\psi(y_0) \le \varphi^c(y_0)$ for all y_0 , so
 - (iii) $J(\varphi, \psi) \leq J(\varphi, \varphi^c)$ by monotonicity of $E_{P_d}[\cdot]$.
 - ⇒ The dual problem can be restricted to *c*-concave functions.
- c-concave functions often inherit properties of c:
 - Lipschitz continuity, boundedness, etc.

$$OT_c(P_1, P_0) = \sup_{(\varphi, \psi) \in \Phi_c} \underbrace{E_{P_1}[\varphi(Y_1)] + E_{P_0}[\psi(Y_0)]}_{J(\varphi, \psi)},$$

Define the *c*-transforms:

$$\varphi^{c}(y_{0}) = \inf_{y_{1}} \{c(y_{1}, y_{0}) - \varphi(y_{1})\}, \qquad \qquad \psi^{c}(y_{1}) = \inf_{y_{0}} \{c(y_{1}, y_{0}) - \psi(y_{0})\}$$

call φ^c (and ψ^c) c-concave functions.

- For any $(\varphi, \psi) \in \Phi_c = \{(\varphi, \psi) ; \varphi(y_1) + \psi(y_0) \le c(y_1, y_0)\},$
 - (i) $(\varphi, \varphi^c), \in \Phi_c$
 - (ii) If $(\varphi, \psi) \in \Phi_c$, then $\psi(y_0) \le \varphi^c(y_0)$ for all y_0 , so
 - (iii) $J(\varphi, \psi) \leq J(\varphi, \varphi^c)$ by monotonicity of $E_{P_d}[\cdot]$.
 - ⇒ The dual problem can be restricted to *c*-concave functions.
- c-concave functions often inherit properties of c:
 - · Lipschitz continuity, boundedness, etc.
 - These properties are used to define \mathcal{F}_c and \mathcal{F}_c^c

Appendix: formal assumption 4

- ▶ Let P be the distribution of an observation: $(Y, D, Z, X) \sim P$.
- ▶ Let $\mathcal{Y}_{d,x}$ be the support of $Y \mid D = d, X = x$, and $\mathbb{1}_{\mathcal{Y}_{d,x}}(y) = \mathbb{1}\{y \in \mathcal{Y}_{d,x}\}$
- \triangleright Define c_L , c_H :
 - (i) If assumption 2 (i) holds, let $c_L = c(y_1, y_0)$ and $c_H(y_1, y_0) = -c(y_1, y_0)$.
 - (ii) If assumption 2 (ii) holds, let $c_L(y_1, y_0) = \mathbb{1}\{y_1 y_0 < \delta\}$ and $c_H(y_1, y_0) = \mathbb{1}\{y_1 y_0 > \delta\}$.

Assumption 4 (Unique solutions) For each $x \in \mathcal{X}$, each $c \in \{c_L, c_H\}$, and any

$$(\varphi_1,\psi_1),(\varphi_2,\psi_2)\in \mathop{\arg\max}_{(\varphi,\psi)\in\Phi_c\cap(\mathcal{F}_c\times\mathcal{F}_c^c)} E_{P_1|_{\mathcal{X}}}[\varphi(Y_1)]+E_{P_0|_{\mathcal{X}}}[\psi(Y_0)],$$

there exists $s \in \mathbb{R}$ such that

$$\mathbb{1}_{\mathcal{Y}_{1,x}}\times\varphi_1=\mathbb{1}_{\mathcal{Y}_{1,x}}\times(\varphi_2+s),\ P-\text{a.s.},\quad \mathbb{1}_{\mathcal{Y}_{0,x}}\times\psi_1=\mathbb{1}_{\mathcal{Y}_{0,x}}\times(\psi_2-s),\ P-\text{a.s.}$$

