

Estimating Functionals of the Joint Distribution of Potential Outcomes with Optimal Transport

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Introduction

The fundamental problem of causal inference

It is impossible to observe the [treated outcome] and [untreated outcome] on the same unit and, therefore, it is impossible to observe the effect...

(Holland, 1986)

- ▶ Parameters of the joint distribution of potential outcomes are not point identified.
- ▶ **This paper**
 - shows optimal transport characterizes sharp bounds,
 - accomodates noncompliance through a standard IV model, and
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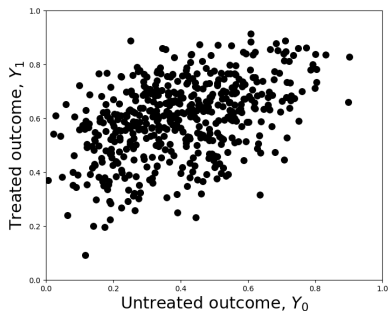
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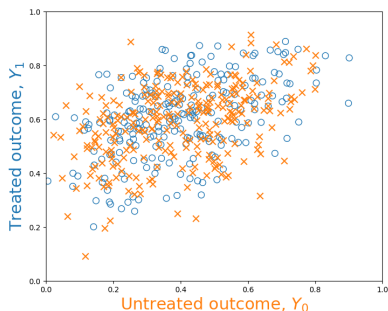
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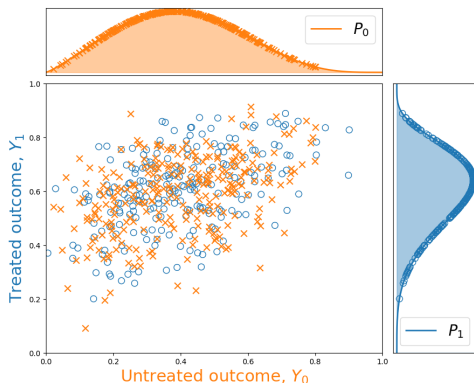
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$$\text{Observed outcome } Y = D Y_1 + (1 - D) Y_0$$

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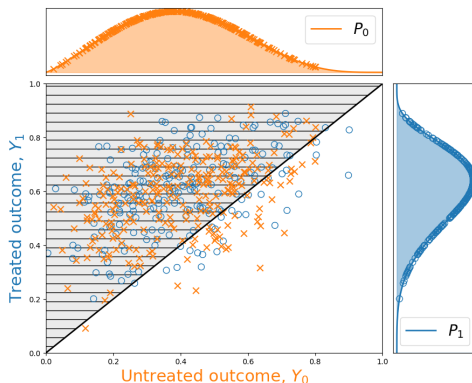


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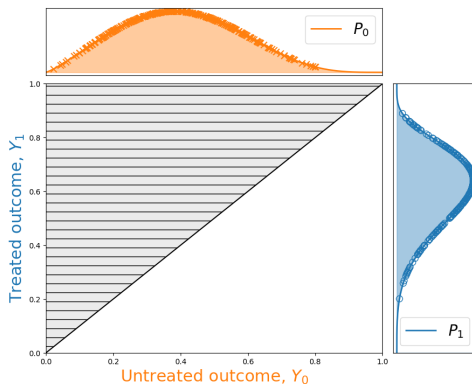


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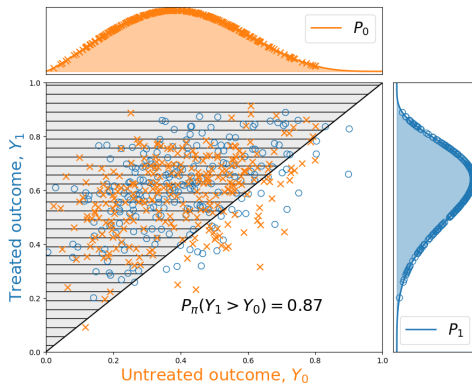
$$\text{Observed outcome } Y = DY_1 + (1 - D)Y_0$$

- ▶ The marginal distributions P_1 and P_0 are identified - but have less information.
- ▶ For example, what share of units benefit from treatment?

Example 1: the share benefiting from treatment



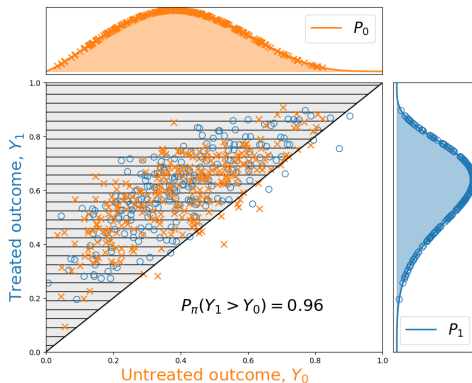
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- Many joint distributions π share marginal distributions P_1 , P_0 :

$$\Pi(P_1, P_0) = \{\pi : \pi_1 = P_1, \pi_0 = P_0\}$$

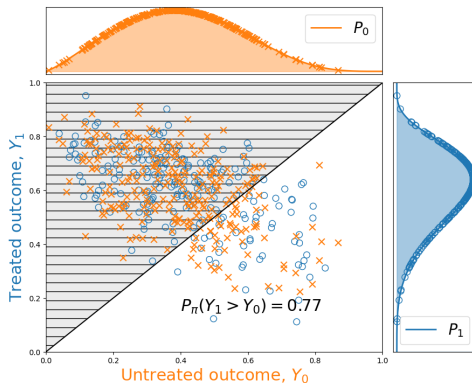
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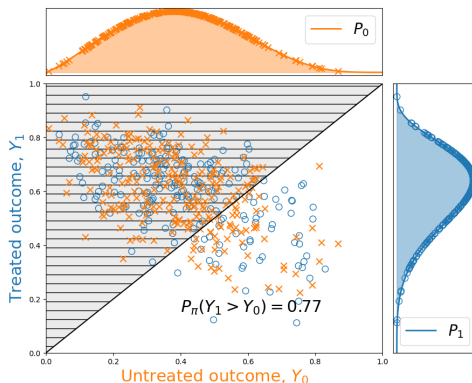
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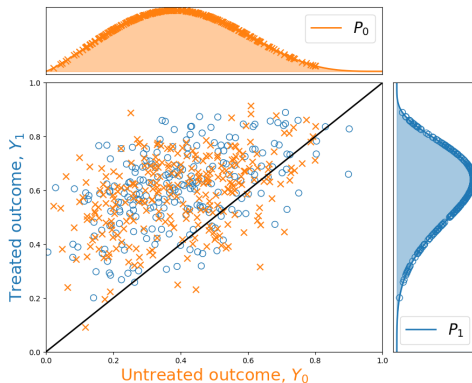
$$\Pi(P_1, P_0) = \{\pi : \pi_1 = P_1, \pi_0 = P_0\}$$

- ▶ Optimizing $P(Y_1 > Y_0)$ over $\Pi(P_1, P_0)$ implies bounds:

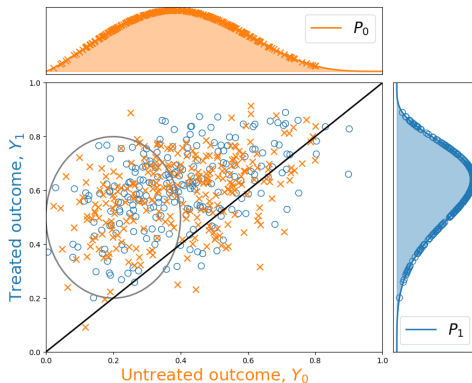
$$\min_{\pi \in \Pi(P_1, P_0)} P_\pi(Y_1 > Y_0)$$

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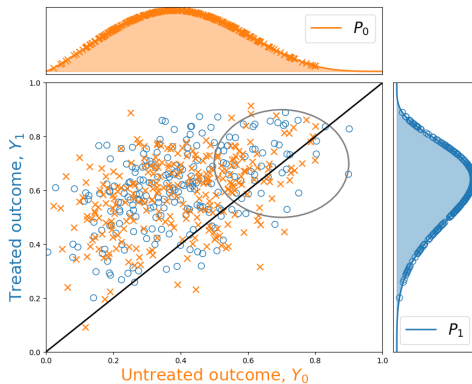
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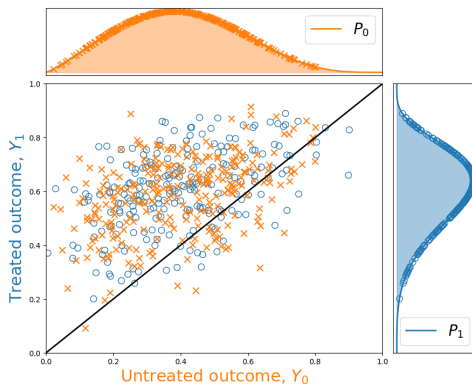
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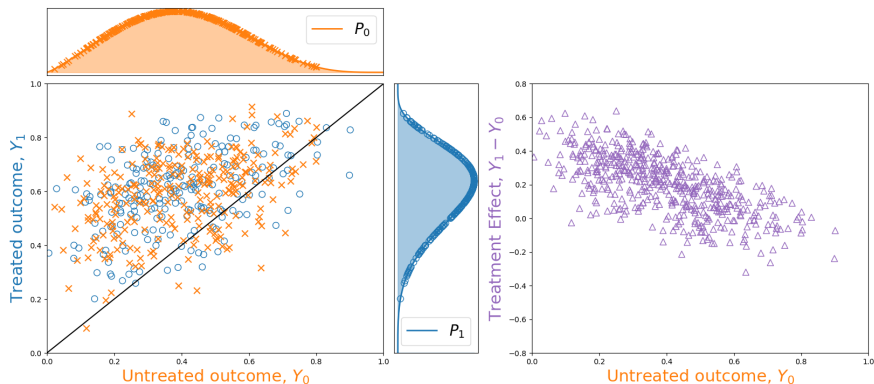


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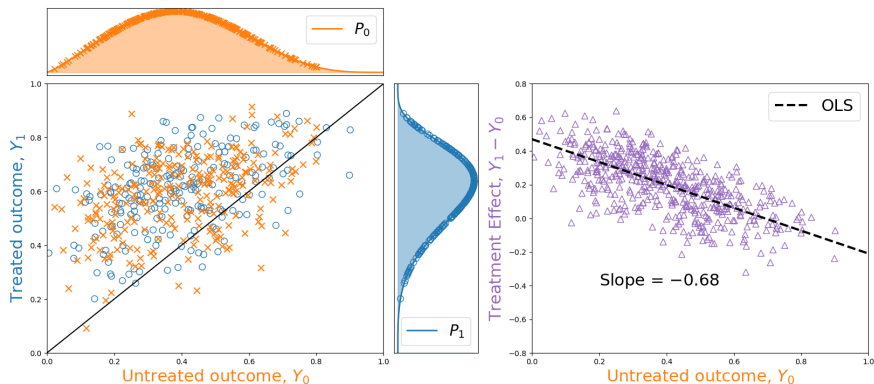
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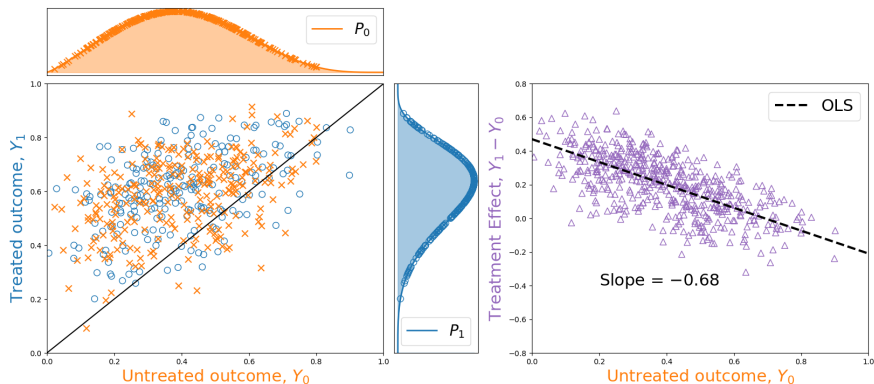
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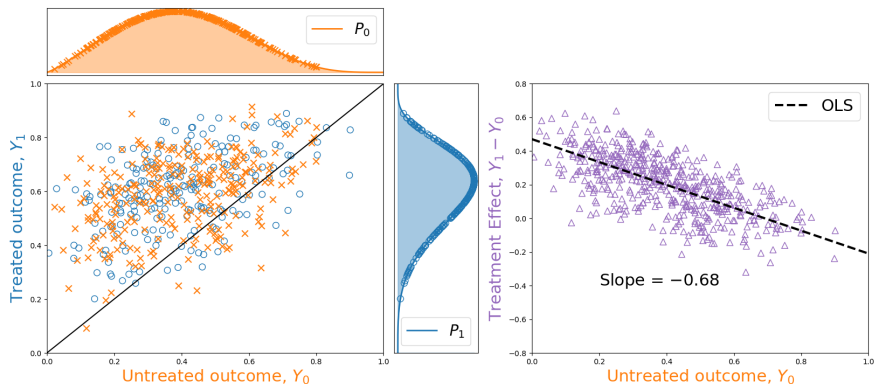
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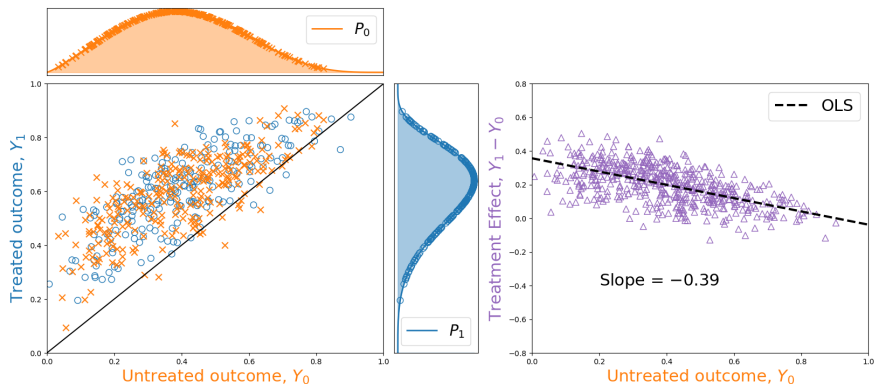
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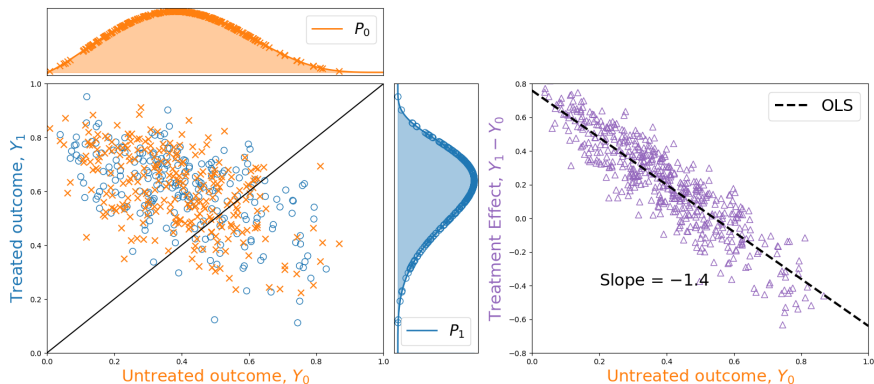
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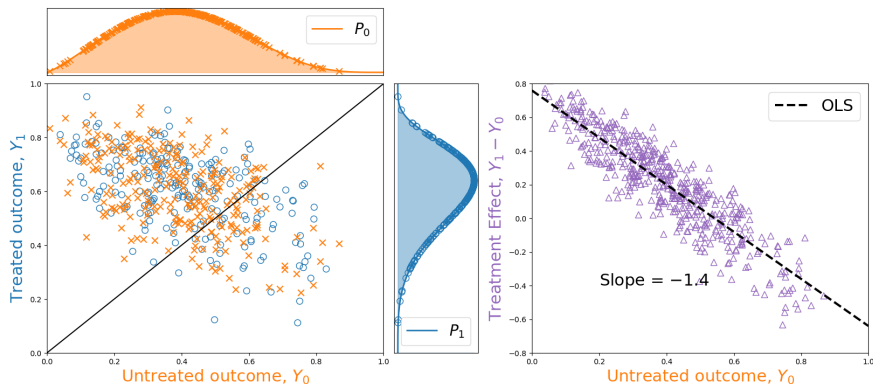
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- Optimizing $E[(Y_1 - Y_0)Y_0]$ over $\Pi(P_1, P_0)$ implies bounds on OLS slope:

$$\min_{\pi \in \Pi(P_1, P_0)} E_{\pi}[(Y_1 - Y_0)Y_0]$$

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$$OT_c(P_1, P_0) = \min_{\pi \in \Pi(P_1, P_0)} E_{\pi}[c(Y_1, Y_0)]$$

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- ▶ Propose and study sample analogue estimators of the bounds.
- ▶ Empirical application: who sees larger benefits from the NSW job training?

Related literature

► Joint distribution of potential outcomes

- CDF or quantiles of $Y_1 - Y_0$: Manski (1997), Heckman et al. (1997), Firpo (2007), Fan and Park (2010), Fan and Park (2012), Firpo and Ridder (2019), Callaway (2021), Frandsen and Lefgren (2021).
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Overview

- 1 Setting and parameter class
- 2 Identification
- 3 Estimators
- 4 Simulations
- 5 Application

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Assumption 1 (Setting, simplified) $\{Y_i, D_i, X_i\}_{i=1}^n$ is an i.i.d. sample with

$$Y \in \mathcal{Y} \subseteq \mathbb{R}, \quad D \in \{0, 1\}, \quad X \in \mathcal{X} = \{x_1, \dots, x_M\}$$

generated from a distribution satisfying

- (i) Potential outcomes: $Y = DY_1 + (1 - D)Y_0$
- (ii) Unconfoundedness: $(Y_1, Y_0) \perp D \mid X$
- (iii) $P(D = d, X = x) > 0$ for each (d, x)

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- In the paper, **binary IV satisfying monotonicity condition** (Imbens and Angrist, 1994).

Setting w/IV

Parameter class

- ▶ Parameter of interest:

$$\gamma = g(\theta, \eta) \in \mathbb{R}$$

where $\theta = E[c(Y_1, Y_0)] \in \mathbb{R}$ and $\eta = (E[\eta_1(Y_1)], E[\eta_0(Y_0)]) \in \mathbb{R}^{K_1+K_0}$

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Assumption 2 (Cost function) Either

- (i) $c(y_1, y_0)$ is Lipschitz continuous and \mathcal{Y} is compact, or
- (ii) $c(y_1, y_0) = \mathbb{1}\{y_1 - y_0 \leq \delta\}$ and the CDFs $F_{d|x}(y) = P(Y_d \leq y \mid X = x)$ are continuous.

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Remark: If $c(y_1, y_0) = \mathbb{1}\{y_1 - y_0 \leq \delta\}$ but $F_{d|x}(\cdot)$ are not continuous, inference remains valid for an outer identified set.

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- (i) $\eta_1(Y)$ and $\eta_0(Y)$ have finite second moments,
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Remark: Assumption 3 (iii) is relaxed in the paper.

Full assumption 3

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 - Friebe et al. (2023): employee referral programs increase grocery store profit.
- ▶ Who benefits more from treatment? $\text{Cov}(Y_1 - Y_0, Y_0)/\text{Var}(Y_0)$
 - **Application**: NSW job experience increases post-training annual income.

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 - **Application**: NSW job experience increases post-training annual income.
- ▶ Expected percent change: $E\left[\frac{Y_1 - Y_0}{Y_0}\right]$
 - This parameter is often approximated with $E[\log(Y_1) - \log(Y_0)]$.

Parameter class: motivating examples

- ▶ Share benefiting: $P(Y_1 > Y_0)$
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- ▶ Quantiles of $Y_1 - Y_0$
 - Median is more representative than mean when distribution is skewed.

Overview

1 Setting and parameter class

2 Identification

3 Estimators

4 Simulations

5 Application

Optimal transport

$$OT_c(P_1, P_0) = \inf_{\pi \in \Pi(P_1, P_0)} E_{\pi}[c(Y_1, Y_0)]$$

- ▶ Choose a joint distribution with given marginals to minimize costs.
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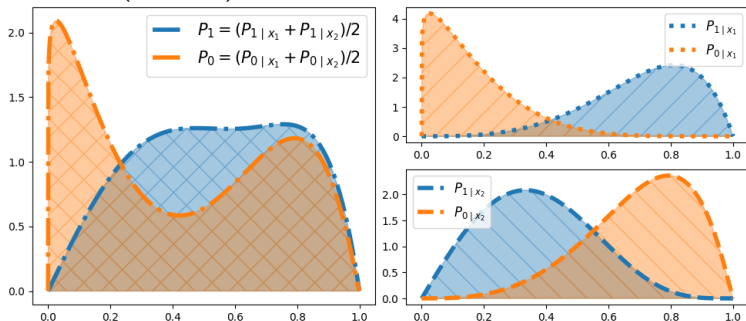
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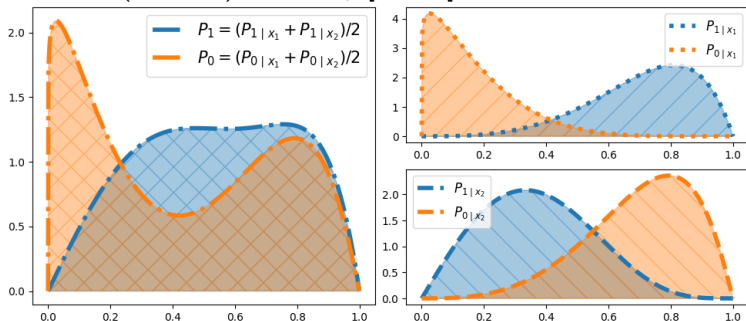
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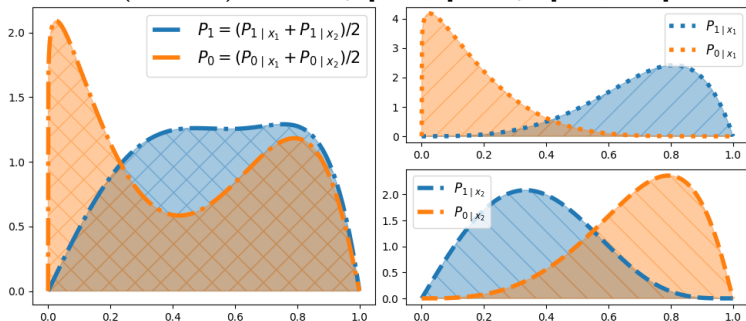
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Theorem: identification

► For continuous c ,

$$\text{Bounds on } \theta_x : \quad \theta_x^L = OT_c(P_{1|x}, P_{0|x}), \quad \theta_x^H = -OT_{-c}(P_{1|x}, P_{0|x})$$

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Theorem (identification)

Suppose assumptions 1, 2, and 3 are satisfied. Then the sharp identified set for $\gamma = g(\theta, \eta)$ is $[\gamma^L, \gamma^H]$.

CDF?

IV Aside

Quantile details

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- ▶ The **primal problem** is used in identification.
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- ▶ **Strong duality** holds under the cost function assumptions. Each problem is attained, too.

Estimators: recall identification

- Distributions of $Y_d \mid X = x \sim P_{d|x}$:

$$E_{P_{d|x}}[f(Y_d)] = \frac{E[f(Y)\mathbb{1}\{D = d, X = x\}]}{P(D = d, X = x)}$$

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- The identified set for γ is $[\gamma^L, \gamma^H]$, where for c continuous,

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Estimators: sample analogues

- Estimate $P_{d|x}$ with **sample analogues** $\hat{P}_{d|x}$:

$$E_{\hat{P}_{d|x}}[f(Y_d)] = \frac{\frac{1}{n} \sum_{i=1}^n f(Y_i) \mathbb{1}\{D_i = d, X_i = x\}}{\frac{1}{n} \sum_{i=1}^n \mathbb{1}\{D_i = d, X_i = x\}}$$

- Using strong duality,

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- Estimate the endpoints of $[\gamma^L, \gamma^H]$ with plug-in estimators. For c continuous,

$$\hat{\theta}_x^L = OT_c(\hat{P}_{1|x}, \hat{P}_{0|x}), \quad \hat{\theta}_x^H = -OT_{-c}(\hat{P}_{1|x}, \hat{P}_{0|x})$$

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Estimators: computing $OT_c(\hat{P}_{1|x}, \hat{P}_{0|x})$

$$OT_c(\hat{P}_{1|x}, \hat{P}_{0|x}) = \max_{(\varphi, \psi) \in \Phi_c} E_{\hat{P}_{1|x}}[\varphi(Y_1)] + E_{\hat{P}_{0|x}}[\psi(Y_0)].$$

- To evaluate $E_{\hat{P}_{d|x}}[f(Y_d)]$ for any function f ,

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- Dimension is reduced by ignoring φ_i , ψ_i , and constraints where $\omega_{d,x,i} = 0$.

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Theorem (Weak convergence)

Suppose assumptions 1, 2, and 3 hold. Then

$$\sqrt{n}((\hat{\gamma}^L, \hat{\gamma}^H) - (\gamma^L, \gamma^H)) \xrightarrow{L} T'_P(\mathbb{G})$$

where $\sqrt{n}(\mathbb{P}_n - P) \xrightarrow{L} \mathbb{G}$ and $T'_P(\cdot)$ is the Hadamard directional derivative of $T(\cdot)$ at P .

[T\(·\) details](#)

[Proof sketch](#)

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 - Bootstrap draw: $\{Y_i^*, D_i^*, X_i^*\}_{i=1}^n$
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- ▶ Compute $T(\mathbb{P}_n^*)$ **the same way** as $T(\mathbb{P}_n)$: let $\omega_{d,x,i}^* = \frac{\mathbb{1}\{D_i^*=d, X_i^*=x\}/n}{\frac{1}{n} \sum_{j=1}^n \mathbb{1}\{D_j^*=d, X_j^*=x\}}$,

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Theorem (Bootstrap consistency)

Suppose assumptions 1, 2, 3, and 4 hold. Then $T'_P(\mathbb{G})$ is bivariate normal, and conditional on $\{Y_i, D_i, X_i\}_{i=1}^n$,

$$\sqrt{n}(T(\mathbb{P}_n^*) - T(\mathbb{P}_n)) \xrightarrow{L} T'_P(\mathbb{G})$$

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- ▶ Assumption 4 may hold without this lemma's conditions.

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- Follows Fang and Santos (2019): estimating the derivative $T'_P(\cdot)$.
- Implementation is more involved, but still computationally tractable.

Overview

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- 2 Identification
- 3 Estimators
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Simulations: parameter and DGP

- ▶ Parameter $\gamma = \theta = P(Y_1 - Y_0 \leq \delta)$ has simple bounds:

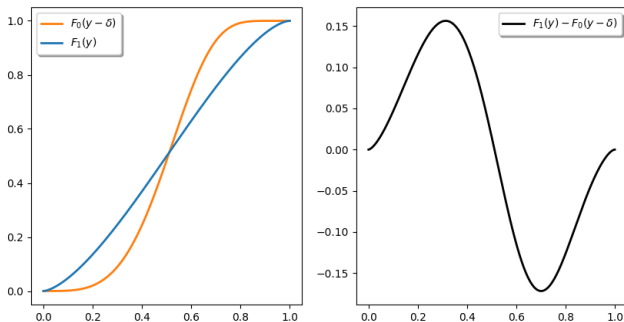
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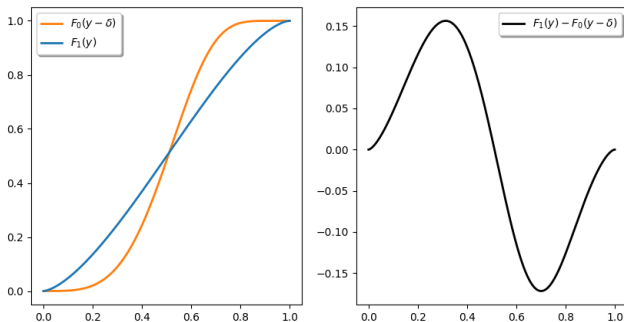


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- ▶ Unique solutions \implies bootstrap is valid.

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(iii) Let $\hat{c}_{1-\alpha}$ be the $1 - \alpha$ quantile of $\{\max\{\sqrt{n}(\hat{\gamma}_b^{L*} - \hat{\gamma}), -\sqrt{n}(\hat{\gamma}_b^{H*} - \hat{\gamma}^H)\}\}_{b=1}^B$, and

$$CI = [\hat{\gamma}^L - \hat{c}_{1-\alpha}/\sqrt{n}, \hat{\gamma}^H + \hat{c}_{1-\alpha}/\sqrt{n}]$$

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	$\hat{\gamma}^L$	$\hat{\gamma}^H$	$\hat{\gamma}^L$	$\hat{\gamma}^H$	
100	0.047	-0.051	0.065	0.066	0.900
200	0.031	-0.031	0.049	0.049	0.917
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100	0.021	-0.026	0.071	0.071	0.927
200	0.013	-0.015	0.052	0.051	0.953
300	0.015	-0.007	0.042	0.042	0.957

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A randomized job training experiment

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Table: Balance table

	base inc.	age	yrs. educ.	HS dropout	black	hispanic	married
control	3672.49 (6521.53)	24.45 (6.59)	10.19 (1.62)	0.81 (0.39)	0.80 (0.40)	0.11 (0.32)	0.16 (0.36)
treated	3571.00 (5773.13)	24.63 (6.69)	10.38 (1.82)	0.73 (0.44)	0.80 (0.40)	0.09 (0.29)	0.17 (0.37)

Note: Standard deviations in parentheses.

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► *Interpretation:* $\gamma < 0$ implies workers with below average Y_0 tend to see above average $Y_1 - Y_0$

NSW results

► Discretized age and baseline income are informative covariates.

- age bins: $[16, 23]$, $(23, \infty)$
- baseline income bins: $[0, 0]$, $(0, 4000]$, $(4000, \infty)$

Table: Estimates of bounds for γ , the OLS Slope

	Lower Bound	Upper Bound	95% <i>CI</i>
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NSW results: conditional on covariate values

Table: Estimates conditional on covariate values

age	base inc.	$\hat{\gamma}_{BC}^L$	$\hat{\gamma}_{BC}^H$	95% CI_{BC}	n
(16, 23]	0	-1.97	0.28	[-2.26, 0.56]	140
	(0, 4000]	-1.74	-0.15	[-1.9, 0.01]	141
	(4000, ∞)	-1.45	-0.44	[-1.63, -0.27]	90
(23, ∞)	0	-2.13	0.81	[-2.65, 1.33]	187
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- ▶ This subset's vulnerable individuals see larger benefits from treatment.

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- Sample analogue estimators are **computationally and analytically attractive**.

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Appendix: full setting

Assumption 1 (Setting). $\{Y_i, D_i, Z_i, X_i\}_{i=1}^n$ is an i.i.d. sample, with

$$Y \in \mathcal{Y} \subseteq \mathbb{R}, \quad D \in \{0, 1\}, \quad Z \in \{0, 1\}, \quad X \in \mathcal{X} = \{x_1, \dots, x_M\}$$

generated from a distribution satisfying

- (i) Potential outcomes: $Y = DY_1 + (1 - D)Y_0$,
- (ii) Potential treatment statuses: $D = ZD_1 + (1 - Z)D_0$, with $D_z \in \{0, 1\}$,
- (iii) Instrument exogeneity: $(Y_1, Y_0, D_1, D_0) \perp Z \mid X$,
- (iv) Monotonicity: $D_1 \geq D_0$ almost surely,
- (v) Existence of compliers: $P(D_1 > D_0, X = x) > 0$ for each x , and
- (vi) $P(X = x, Z = z) > 0$ for each (x, z)

► Terminology: always-taker, complier, defier, never-taker.

	$D_0 = 1$	$D_0 = 0$
$D_1 = 1$	Always-takers	Compliers
$D_1 = 0$	Defiers	Never-takers

► Monotonicity rules out defiers. Focus on distribution of compliers.

Appendix: identification of $P(Y_1 - Y_0 \leq \delta)$

- ▶ $OT_c(P_1, P_0)$ is well behaved (attained, strong duality holds, etc) when $c(y_1, y_0)$ is bounded and lower semicontinuous
- ▶ If $c(y_1, y_0) = \mathbb{1}\{y_1 - y_0 \leq \delta\}$, let

$$c_L(y_1, y_0) = \mathbb{1}\{y_1 - y_0 < \delta\},$$

$$\theta_x^L = OT_{c_L}(P_{1|x}, P_{0|x}),$$

$$c_H(y_1, y_0) = \mathbb{1}\{y_1 - y_0 > \delta\}$$

$$\theta_x^H = 1 - OT_{c_H}(P_{1|x}, P_{0|x})$$

- ▶ The form of the bounds remains the same:

$$\theta^L = E[\theta_X^L],$$

$$\gamma^L = \min_{t \in [\theta^L, \theta^H]} g(t, \eta),$$

$$\theta^H = E[\theta_X^H]$$

$$\gamma^H = \max_{t \in [\theta^L, \theta^H]} g(t, \eta)$$

- ▶ Identified sets are still sharp when CDFs are continuous:

$$F_{d|x}(y) = P(Y_d \leq y \mid X = x)$$

Appendix: aside, CDF results are conservative when continuity fails

$$OT_c(P_1, P_0) = \inf_{\pi \in \Pi(P_1, P_0)} E_{\pi}[c(Y_1, Y_0)]$$

- Bounds on $\theta = P(Y_1 - Y_0 \leq \delta)$ are found with

$$\begin{aligned} c_L(y_1, y_0) &= \mathbb{1}\{y_1 - y_0 < \delta\}, & c_H(y_1, y_0) &= \mathbb{1}\{y_1 - y_0 > \delta\}, \\ \theta^L &= OT_{c_L}(P_1, P_0), & \theta^H &= 1 - OT_{c_H}(P_1, P_0) \end{aligned}$$

Using OT results, show that if marginal CDFs F_d are continuous then $\Theta_{ID} = [\theta^L, \theta^H]$.

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- As a byproduct, recover the famed **Makarov bounds** studied by Fan and Park (2010)

$$\theta^L = \sup_y \{F_1(y) - F_0(y - \delta)\}, \quad \theta^H = 1 + \inf_y \{F_1(y) - F_0(y - \delta)\}$$

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- ▶ **Furthermore**, $\mathbb{1}\{y_1 - y_0 < \delta\} \leq \mathbb{1}\{y_1 - y_0 \leq \delta\}$ implies **the bounds are conservative**: $\Theta_{ID} \subseteq [\theta^L, \theta^H]$ **whether or not F_d are continuous**.

Appendix: full assumption 3

- Parameter of interest:

$$\gamma = g(\theta, \eta) \in \mathbb{R}$$

where $\theta = E[c(Y_1, Y_0)] \in \mathbb{R}$ and $\eta = (E[\eta_1(Y_1)], E[\eta_0(Y_0)]) \in \mathbb{R}^{K_1+K_0}$.

Assumption 3 (Function of moments)

- (i) $E[\|\eta_d(Y)\|^2] < \infty$ for $d = 1, 0$,
- (ii) $g(\cdot, \eta)$ is continuous, and
- (iii) the functions

$$g^L(t^L, t^H, e) = \min_{t \in [t^L, t^H]} g(t, e), \quad g^H(t^L, t^H, e) = \max_{t \in [t^L, t^H]} g(t, e)$$

are continuously differentiable at $(t^L, t^H, e) = (\theta^L, \theta^H, \eta)$.

Remark: A3 (ii), (iii) implied by g continuously differentiable and $g(\cdot, \eta)$ monotonic

Appendix: quantiles

- Suppose the parameter of interest is q_τ solving

$$P(Y_1 - Y_0 \leq q_\tau) = \tau$$

- View CDF bounds as a function: $\theta(\delta) = P(Y_1 - Y_0 \leq \delta)$

$$c_{L,\delta}(y_1, y_0) = \mathbb{1}\{y_1 - y_0 < \delta\},$$

$$c_{H,\delta}(y_1, y_0) = \mathbb{1}\{y_1 - y_0 > \delta\},$$

$$\theta_x^L(\delta) = OT_{c_L}(P_{1|x}, P_{0|x}),$$

$$\theta_x^H(\delta) = 1 - OT_{c_H}(P_{1|x}, P_{0|x})$$

$$\theta^L(\delta) = E[\theta_X^L(\delta)]$$

$$\theta^H(\delta) = E[\theta_X^H(\delta)]$$

and let $Q_{I,\tau}$ be the sharp identified set for q_τ .

Lemma (Identification of q_τ). Suppose assumptions 1 and 2(ii) hold. Then $q \in Q_{I,\tau}$ if and only if $\theta^L(q) \leq \tau \leq \theta^H(q)$.

Examples

Appendix: aside, IV

- Identification extends easily to IV.

Ident. Thm.

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$$Y_d \mid D_1 > D_0, X = x$$

- ▶ Same identification applies to parameters conditional on compliance. E.g.,

$$P(Y_1 > Y_0 \mid D_1 > D_0)$$

Appendix: definition of \mathcal{T}

- Proof defines a set of universally bounded functions

$$\mathcal{F} \subseteq \{f : \mathcal{Y} \times \{0, 1\} \times \mathcal{X} \rightarrow \mathbb{R}\}$$

Weak convergence theorem

Appendix: definition of T

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- View \mathbb{P}_n, P as bounded functions on \mathcal{F} :

$$\ell^\infty(\mathcal{F}) = \left\{ g : \mathcal{F} \rightarrow \mathbb{R} ; \|g\|_\infty = \sup_{f \in \mathcal{F}} |g(f)| < \infty \right\}$$

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- The map $T : \ell^\infty(\mathcal{F}) \rightarrow \mathbb{R}^2$ is described by $P \mapsto (P_{1|x}, P_{0|x}, \eta)$ and

$$\begin{aligned} \theta_x^L &= OT_c(P_{1|x}, P_{0|x}), & \theta_x^H &= -OT_{-c}(P_{1|x}, P_{0|x}) \\ \theta^L &= E[\theta_X^L], & \theta^H &= E[\theta_X^H] \\ \gamma^L &= \min_{t \in [\theta^L, \theta^H]} g(t, \eta), & \gamma^H &= \max_{t \in [\theta^L, \theta^H]} g(t, \eta) \end{aligned}$$

Weak convergence theorem

Appendix: proof sketch (1/3)

1. Will view P, \mathbb{P} as maps in $\ell^\infty(\mathcal{F})$ for Donsker set \mathcal{F} (defined later), and $T : \ell^\infty(\mathcal{F}) \rightarrow \mathbb{R}^2$.

Weak convergence theorem

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2. To show $T(\cdot)$ is (Hadamard) directionally differentiable, suffices to show OT_c is directionally differentiable.
3. By strong duality,

$$OT_c(P_{1|x}, P_{0|x}) = \sup_{(\varphi, \psi) \in \Phi_c} E_{P_{1|x}}[\varphi(Y_1)] + E_{P_{0|x}}[\psi(Y_0)]$$
$$\Phi_c = \{(\varphi, \psi) : \varphi(y_1) + \psi(y_0) \leq c(y_1, y_0)\}$$

Weak convergence theorem

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$$\sup_{(\varphi, \psi) \in \Phi_c} E_{P_1}[\varphi(Y_1)] + E_{P_0}[\psi(Y_0)] = \sup_{(\varphi, \psi) \in \Phi_c \cap (\mathcal{F}_c \times \mathcal{F}_c^c)} E_{P_1}[\varphi(Y_1)] + E_{P_0}[\psi(Y_0)] \quad (1)$$

- (i) if $c(y_1, y_0)$ is L -Lip. and \mathcal{Y} is compact, \mathcal{F}_c and \mathcal{F}_c^c are L -Lip. and universally bounded.
- (ii) if $c(y_1, y_0) = \mathbb{1}\{y_1 - y_0 \leq \delta\}$, \mathcal{F}_c is the set of intervals, \mathcal{F}_c^c the complements of intervals.

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6. Finally, $\Phi_c \cap (\mathcal{F}_c \times \mathcal{F}_c^c)$ is compact and $E_{P_{1|x}}[\varphi(Y_1)] + E_{P_{0|x}}[\psi(Y_0)]$ is continuous

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$\implies OT_c$, and therefore $T(\cdot)$, are Hadamard directionally differentiable.

Appendix: proof sketch (3/3)

7. Define \mathcal{F} to be union of \mathcal{F}_c and \mathcal{F}_c^c (and nuisance moments, all \times indicators).

Weak convergence theorem

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8. \mathcal{F} is Donsker $\implies \sqrt{n}(\mathbb{P}_n - P) \xrightarrow{L} \mathbb{G}$ in $\ell^\infty(\mathcal{F})$.
9. Functional delta method implies the result,

$$\sqrt{n}((\hat{\gamma}^L, \hat{\gamma}^H) - (\gamma^L, \gamma^H)) = \sqrt{n}(T(\mathbb{P}_n) - T(P)) \xrightarrow{L} T'_P(\mathbb{G}).$$

Weak convergence theorem

Appendix: c -concavity

$$OT_c(P_1, P_0) = \sup_{(\varphi, \psi) \in \Phi_c} \underbrace{E_{P_1}[\varphi(Y_1)] + E_{P_0}[\psi(Y_0)]}_{J(\varphi, \psi)},$$

► Define the c -transforms:

$$\varphi^c(y_0) = \inf_{y_1} \{c(y_1, y_0) - \varphi(y_1)\}, \quad \psi^c(y_1) = \inf_{y_0} \{c(y_1, y_0) - \psi(y_0)\}$$

call φ^c (and ψ^c) **c -concave** functions.

Appendix: c -concavity

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⇒ The dual problem can be restricted to c -concave functions.

- c -concave functions often **inherit properties of c** :

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- These properties are used to define \mathcal{F}_c and \mathcal{F}_c^c

Proof sketch

Weak convergence theorem

Appendix: formal assumption 4

- ▶ Let P be the distribution of an observation: $(Y, D, Z, X) \sim P$.
- ▶ Let $\mathcal{Y}_{d,x}$ be the support of $Y \mid D = d, X = x$, and $\mathbb{1}_{\mathcal{Y}_{d,x}}(y) = \mathbb{1}\{y \in \mathcal{Y}_{d,x}\}$
- ▶ Define c_L, c_H :
 - (i) If assumption 2 (i) holds, let $c_L = c(y_1, y_0)$ and $c_H(y_1, y_0) = -c(y_1, y_0)$.
 - (ii) If assumption 2 (ii) holds, let $c_L(y_1, y_0) = \mathbb{1}\{y_1 - y_0 < \delta\}$ and $c_H(y_1, y_0) = \mathbb{1}\{y_1 - y_0 > \delta\}$.

Assumption 4 (Unique solutions) For each $x \in \mathcal{X}$, each $c \in \{c_L, c_H\}$, and any

$$(\varphi_1, \psi_1), (\varphi_2, \psi_2) \in \arg \max_{(\varphi, \psi) \in \Phi_c \cap (\mathcal{F}_c \times \mathcal{F}_c^c)} E_{P_{1|x}}[\varphi(Y_1)] + E_{P_{0|x}}[\psi(Y_0)],$$

there exists $s \in \mathbb{R}$ such that

$$\mathbb{1}_{\mathcal{Y}_{1,x}} \times \varphi_1 = \mathbb{1}_{\mathcal{Y}_{1,x}} \times (\varphi_2 + s), \quad P - a.s., \quad \mathbb{1}_{\mathcal{Y}_{0,x}} \times \psi_1 = \mathbb{1}_{\mathcal{Y}_{0,x}} \times (\psi_2 - s), \quad P - a.s.$$

Assumption 4

Why c_L, c_H ?