

Research presentation for PhD admits

Edgar Dobriban

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Overview

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Uncertainty quantification: calibration

Fairness: Bayes-optimal classifiers

High-dim. statistics & deterministic equivalents

Research Interests

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 - ▶ Usually work closely on first project, then as hands-on/off as you would like.










Research Interests: see my website for details

- the efficient statistical analysis of “big data” using advanced tools, such as those from random matrix theory
 - dimension reduction, PCA: [1], [2], [3], [4], [5], [6]
 - multiple testing: [1], [2], [3]
 - high-dimensional regression: [1], [2]
 - invariance-based randomization tests
- the theoretical foundations of modern machine learning, including deep learning
 - data augmentation: [1], [2]
 - weight normalization
 - (stochastic) gradient descent and flow: [1], [2]
 - overparametrization
 - sketching and random projections, [1], [2], [3], [4], [5]
 - distributed learning: [1], [2], [3]
 - adversarial robustness: [1], [2]
 - retraining of ML models
 - uncertainty quantification: [1], [2]
 - fairness: [1]
 - reinforcement learning inspired by child-like learning
- in addition, we occasionally work on important applications and methods, such as
 - genomics
 - group testing for COVID-19

Uncertainty quantification for ML - My course at Penn

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 dobriban / Topics-In-Modern-Statistical-Learning Public

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STAT 991: Topics In Modern Statistical Learning (UPenn, 2022 Spring)

This class surveys advanced topics in statistical learning based on student presentations.

The core topic of the course is [uncertainty quantification for machine learning methods](#). While modern machine learning methods can have a high prediction accuracy in a variety of problems, it is still challenging to properly quantify their uncertainty. There has been a recent surge of work developing methods for this problem. It is one of the fastest developing areas in contemporary statistics. This course will survey a variety of different problems and approaches, such as calibration, prediction intervals (and sets), conformal inference, OOD detection, etc. We will discuss both empirically successful/popular methods as well as theoretically justified ones. See below for a sample of papers.

About

Materials for STAT 9911303: Topics In Modern Statistical Learning (UPenn, 2022 Spring) - focus on uncertainty quantification

machine-learning deep-learning prediction uncertainty-quantification conformal-prediction tolerance-intervals

Topics in Deep Learning - My course at Penn

STAT 991: Topics in deep learning (UPenn)

STAT 991: Topics in Deep Learning is a seminar class at UPenn started in 2018. It surveys advanced topics in deep learning based on student presentations.

Fall 2019

- [Syllabus](#).
- [Lecture notes](#). (~170 pages, file size ~30 MB, mostly covering notes from previous semesters.)

Lectures

Lectures 1 and 2: Introduction and uncertainty quantification ([jackknife+](#), and [Pearce et al, 2018](#)), presented by Edgar Dobriban.

Lecture 3: [NTK](#) by Jiayao Zhang. [Blog post](#) on the off-convex blog.

Lecture 4: [Adversarial robustness](#) by Yinjun Wu.

Lecture 5: [ELMo and BERT](#) by Dan Deutsch.

Lecture 6: [TCAV](#) by Ben Auerbach (adapted from Been Kim's slides).

Lecture 7: [Spherical CNN](#) by Arjun Guru and Claudia Zhu.

Lecture 8: [DNNs and approximation](#) by Yebiao Jin.

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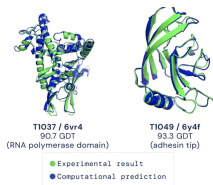
High-dim. statistics & deterministic equivalents

Context

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- ▶ Success stories: AlphaFold, cancer tissue image classification, computer vision, NLP ...

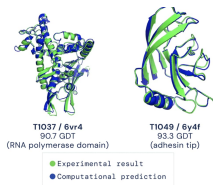


Title: United Methodists Agree to Historic Split
Subtitle: Those who oppose gay marriage will form their own denomination
Article: After two days of intense debate, the United Methodist Church has agreed to a historic split - one that is expected to end in the creation of a new denomination, one that will be "theologically and socially conservative," according to The Washington Post. The majority of delegates attending the church's annual General Conference in May voted to strengthen a ban on the ordination of LGBTQ clergy and to write new rules that will "discipline" clergy who officiate at same-sex weddings. But those who opposed these measures have a new plan: They say they will form a separate denomination by 2020, calling their church the Christian Methodist denomination.

Figure 3.14: The GPT-3 generated news article that humans had the greatest difficulty distinguishing from a human written article (accuracy: 12%).

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Figure 3.14: The GPT-3 generated news article that humans had the greatest difficulty distinguishing from a human written article (accuracy: 12%).

- ▶ Meanwhile, growing concerns: safety, ethics, energy- and sample-efficiency, **uncertainty**

Calibration

- **Calibration**: construct probability predictions that reflect true probabilities. For binary classification, for all appropriate z ,

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- Modern finding: powerful ML methods (e.g., deep CNNs) are *over-confident* and *mis-calibrated*

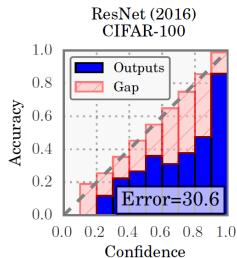
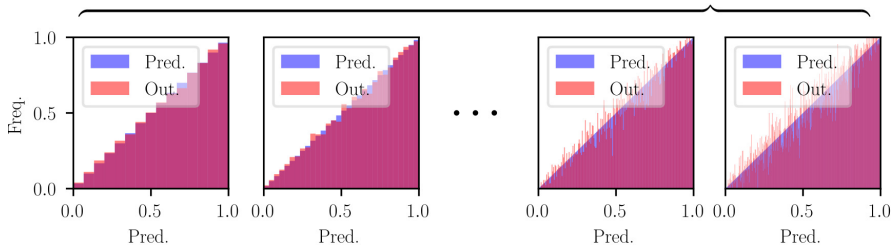
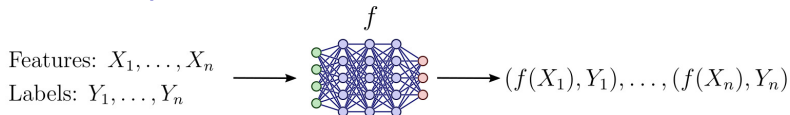


Figure: Guo et al, 2017

T-Cal: An optimal test of calibration



$$T_{m,n}^d = \sum_{\substack{1 \leq i \leq m \\ |\mathcal{I}_{m,i}| \geq 1}} \frac{1}{n|\mathcal{I}_{m,i}|} \sum_{j_1 \neq j_2 \in \mathcal{I}_{m,i}} [Y_{j_1} - f(X_{j_1})][Y_{j_2} - f(X_{j_2})]$$

-7.65×10^{-4} ✓ 1.60×10^{-5} ✓ ... 4.51×10^{-4} ✗ 2.67×10^{-4} ✓

✗ (f is mis-calibrated)

T-Cal

- ▶ Theoretical result: minimax optimal under Hölder smoothness
- ▶ Empirical results: large power in simulations; can use it to detect mis-calibration of state-of-the-art deep networks

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Uncertainty quantification: calibration

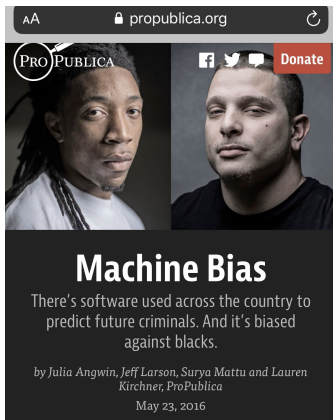
Fairness: Bayes-optimal classifiers

High-dim. statistics & deterministic equivalents

Motivation

- ▶ Machine learning algorithms are becoming integrated into more and more high-stakes decision-making processes.
- ▶ Algorithm-based decision-making systems could retain or even amplify historical unfairness in data.

COMPAS Algorithm



Amazon Recruitment System

RETAIL OCTOBER 10, 2018 / 7:04 PM / UPDATED 3 YEARS AGO

Amazon scraps secret AI recruiting tool that showed bias against women

By Jeffrey Dastin

8 MIN READ



SAN FRANCISCO (Reuters) - Amazon.com Inc's [AMZN.O](#) machine-learning specialists uncovered a big problem: their new recruiting engine did not like women.

Group Fairness

- ▶ Consider a classification problem with two types of feature: the usual feature $X \in \mathcal{X}$, and the protected (or, sensitive) feature $A \in \mathcal{A} = \{0, 1\}$.
- ▶ Binary labels in $\mathcal{Y} = \{0, 1\}$, prediction \hat{Y} .

Fair Bayes-optimal Classifier under Demographic Parity

Several group fairness measures have been proposed. Measure "unfairness" by *Difference in demographic parity*:

$$DDP = P(\hat{Y} = 1|A = 1) - P(\hat{Y} = 1|A = 0).$$

For input x , let $f(x) := P(\hat{Y} = 1|X = x)$.

Goal: Find **δ -fair Bayes-optimal classifier** with respect to demographic parity; defined as

$$f_{D,\delta}^* \in \operatorname{argmin}_{f: |DDP(f)| \leq \delta} [P(Y \neq \hat{Y})].$$

Main Theorem

Denote

$$p_a := P(A = a)$$

$$\eta_a(x) := P(Y = 1|A = a, X = x)$$

$$S_a(t) := P(\eta_a(X) > t|A = a)$$

Theorem (Fair Bayes-optimal Classifier under Demographic Parity)

Let $D^* = \text{DDP}(f^*)$, where f^* is unconstrained Bayes-optimal classifier. For any $\delta > 0$, all δ -fair Bayes optimal classifiers $f_{D,\delta}^*$ have the following form:

- ▶ When $|D^*| \leq \delta$, $f_{D,\delta}^* = f^*$.
- ▶ When $|D^*| > \delta$, for all $x \in \mathcal{X}$ and $a \in \mathcal{A}$,

$$\begin{aligned} f_{D,\delta}^*(x, a) = & I \left(\eta_a(x) > \frac{1}{2} + \frac{(2a-1)t_{D,\delta}^*}{2p_a} \right) \\ & + at_{D,\delta}^* I \left(\eta_a(x) = \frac{1}{2} + \frac{(2a-1)t_{D,\delta}^*}{2p_a} \right), \end{aligned} \tag{1}$$

Main Theorem

Theorem (continued)

where $t_{D,\delta}^*$ is defined as

$$t_{D,\delta}^* = \sup \left\{ t : S_1 \left(\frac{1}{2} + \frac{t}{2p_1} \right) > S_0 \left(\frac{1}{2} - \frac{t}{2p_0} \right) + \frac{D^*}{|D^*|} \delta \right\}. \quad (2)$$

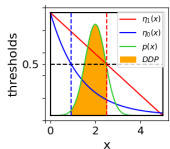
Here, $\tau_{D,\delta}^* \in [0, 1]$ can be an arbitrary constant if

$P_{X|A=1}(\eta_1(X) = \frac{1}{2} + \frac{t}{2p_1}) = 0$, and otherwise

$$\tau_{D,\delta}^* = \frac{S_1 \left(\frac{1}{2} + \frac{t}{2p_1} \right) - S_0 \left(\frac{1}{2} - \frac{t}{2p_0} \right) - \frac{D^*}{|D^*|}}{P_{X|A=1}(\eta_1(X) = \frac{1}{2} + \frac{t}{2p_1})}. \quad (3)$$

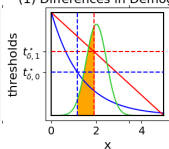
Illustration of Theorem

(a) Unconstrained Bayes-optimal Classifier

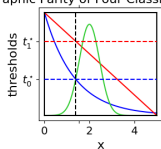


(b) Bayes-optimal Classifier with $DDP \leq \delta$

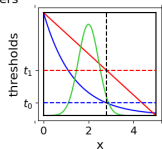
(1) Differences in Demographic Parity of Four Classifiers



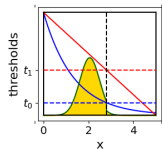
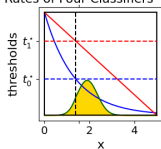
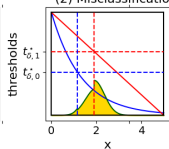
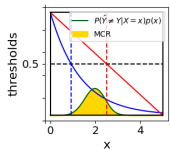
(c) Bayes-optimal Classifier with $DDP=0$



(d) Fair Classifier



(2) Misclassification Rates of Four Classifiers



Proof Sketch

Lemma (Generalized Neyman-Pearson lemma)

Let f_0, f_1, \dots, f_m be $m + 1$ real-valued functions defined on a Euclidean space \mathcal{X} . Assume they are ν -integrable for a σ -finite measure ν . Let ϕ_0 be any function of the form

$$\phi_0(x) = \begin{cases} 1, & f_0(x) > \sum_{i=1}^m c_i f_i(x); \\ \gamma(x) & f_0(x) = \sum_{i=1}^m c_i f_i(x); \\ 0, & f_0(x) < \sum_{i=1}^m c_i f_i(x), \end{cases} \quad (4)$$

where $0 \leq \gamma(x) \leq 1$ for all $x \in \mathcal{X}$.

Proof Sketch

Lemma (continued)

For given constants $t_1, \dots, t_m \in \mathbb{R}$, let \mathcal{T} be the class of Borel functions $\phi : \mathcal{X} \mapsto \mathbb{R}$ satisfying

$$\int_{\mathcal{X}} \phi f_i d\nu \leq t_i, \quad i = 1, 2, \dots, m. \quad (5)$$

and \mathcal{T}_0 be the set of ϕ s in \mathcal{T} satisfying (5) with all inequalities replaced by equalities. If $\phi_0 \in \mathcal{T}_0$, then $\phi_0 \in \operatorname{argmax}_{\phi \in \mathcal{T}_0} \int_{\mathcal{X}} \phi f_0 d\nu$. Moreover, if $c_i \geq 0$ for all $i = 1, \dots, m$, then $\phi_0 \in \operatorname{argmax}_{\phi \in \mathcal{T}} \int_{\mathcal{X}} \phi f_0 d\nu$.

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Motivation

- ▶ Standard linear model $Y = X\beta + \varepsilon$, where
 1. Y is $n \times 1$ outcome, X is $n \times p$ feature matrix.
 2. β is p -dim parameter
- ▶ Ordinary least squares

$$\hat{\beta} = (X^\top X)^{-1} X^\top Y$$

- ▶ Mean squared error of OLS, assuming $\mathbb{E}\varepsilon = 0$, $\text{cov}(\varepsilon) = \sigma^2 I_n$

$$\mathbb{E}\|\hat{\beta} - \beta\|^2 = \sigma^2 \text{tr}[(X^\top X)^{-1}]$$

- ▶ How large is this? (How hard? How much error?)

Motivation ctd

- ▶ When $X_{ij} \sim \mathcal{N}(0, 1)$ are iid standard normal,

$$\mathbb{E} \operatorname{tr}[(X^\top X)^{-1}] = \frac{p}{n - p - 1}.$$

- ▶ More general data distributions? There are only approximate expressions.

Deterministic equivalents

- ▶ We have sequences of (not necessarily symmetric) $p_n \times p_n$ random matrices A_n and deterministic matrices B_n of growing dimensions
- ▶ **Definition:** B_n is a *deterministic equivalent* for A_n ,

$$A_n \asymp B_n$$

if

$$\lim_{n \rightarrow \infty} |\text{tr}(C_n A_n) - \text{tr}(C_n B_n)| = 0$$

almost surely, for any $p_n \times p_n$ sequence C_n of (not necessarily symmetric) deterministic real matrices with bounded trace norm, i.e.,

$$\limsup_{n \rightarrow \infty} \|C_n\|_{tr} = \limsup_{n \rightarrow \infty} \sum_i \sigma_i(C_n) < \infty.$$

e.g, $C_n = c_n c_n^\top$, $\|c_n\|_2$ bounded

Sample covariance matrices

Example (Mestre et al., 2011)

Let $\hat{\Sigma} = X^\top X/n$, where $X = Z\Sigma^{1/2}$ and Z is an $n \times p$ random matrix with iid entries of zero mean, unit variance and finite $8 + \eta$ moment. Also, $\Sigma^{1/2}$ is any sequence of $p \times p$ positive semi-definite matrices satisfying $\sup \|\Sigma\|_2 < \infty$. As $n, p \rightarrow \infty$ proportionally, for any $\lambda > 0$

$$(\hat{\Sigma} + \lambda I_p)^{-1} \asymp (q_p \Sigma + \lambda I_p)^{-1},$$

where q_p is the solution of a fixed point equation.

- This is the simplest way I know how to think of a broad class of results in random matrix theory.

Distributed linear regression

- ▶ Standard linear model $Y = X\beta + \varepsilon$
- ▶ Data distributed across k machines. The i -th machine has matrix X_i ($n_i \times p$) and outcomes Y_i .

$$X = \begin{bmatrix} X_1 \\ \dots \\ X_k \end{bmatrix}, Y = \begin{bmatrix} Y_1 \\ \dots \\ Y_k \end{bmatrix}$$

- ▶ Global least squares - infeasible
- ▶ *Local* least squares estimator $\hat{\beta}_i = (X_i^\top X_i)^{-1} X_i^\top Y_i$ (assume $n_i > p$)
- ▶ Send to parameter server, average
- ▶ How does this compare to OLS on full data?

A general framework

- ▶ Important to study not only estimation, but also prediction/test error, residual error, confidence intervals etc
- ▶ Predict the linear functional

$$L_A = A\beta + Z$$

- ▶ Using the plug-in estimator

$$\hat{L}_A(\hat{\beta}_0) = A\hat{\beta}_0$$

- ▶ A - fixed $d \times p$ matrix; mean and covariance of Z has the structure:
 $Z \sim (0, h\sigma^2 I_d)$, $h \geq 0$
- ▶ The noise can be correlated with ε : $\text{Cov}[\varepsilon, Z] = N$ (e.g., to study residuals)
- ▶ Relative efficiency:

$$E(A; X_1, \dots, X_k) := \frac{\mathbb{E}\|L_A - \hat{L}_A(\hat{\beta})\|^2}{\mathbb{E}\|L_A - \hat{L}_A(\hat{\beta}_{dist})\|^2}.$$

Examples: Predict $L_A = A\beta + Z$ by $\hat{L}_A(\hat{\beta}_0) = A\hat{\beta}_0$

Statistical learning problem	L_A	\hat{L}_A	A	h	N
Parameter estimation	β	$\hat{\beta}$	I_p	0	0
Regression function estimation	$X\beta$	$X\hat{\beta}$	X	0	0
Confidence interval for marginal effect	β_j	$\hat{\beta}_j$	E_j^\top	0	0
Test error	$\mathbf{x}_t^\top \beta + \varepsilon_t$	$\mathbf{x}_t^\top \hat{\beta}$	\mathbf{x}_t^\top	1	0
Training error/Residual	$X\beta + \varepsilon$	$X\hat{\beta}$	X	1	$\sigma^2 I_n$

Finite sample results

- ▶ When $h = 0$ (no noise), the MSE of estimating $L_A = A\beta$ by OLS $\hat{L}_A = A\hat{\beta} = A(X^\top X)^{-1}X^\top Y$ is

$$M(\hat{\beta}) = \sigma^2 \cdot \text{tr} \left[(X^\top X)^{-1} A^\top A \right].$$

- ▶ For the distributed estimator $\hat{\beta}_{dist}(w) = \sum_i w_i \hat{\beta}_i$, $\sum_i w_i = 1$

$$M(\hat{\beta}_{dist}) = \sigma^2 \cdot \sum_{i=1}^k w_i^2 \cdot \text{tr} \left[(X_i^\top X_i)^{-1} A^\top A \right].$$

- ▶ So optimal efficiency is

$$E(A; X_1, \dots, X_k) = \text{tr} \left[(X^\top X)^{-1} A^\top A \right] \cdot \sum_{i=1}^k \frac{1}{\text{tr} \left[(X_i^\top X_i)^{-1} A^\top A \right]}.$$

$$\text{CDE: } \text{tr}[(X_i^\top X_i)^{-1} A^\top A] \asymp \frac{p}{n_i - p} \cdot \text{tr}[\Sigma^{-1} A^\top A] / p.$$

Plot efficiencies

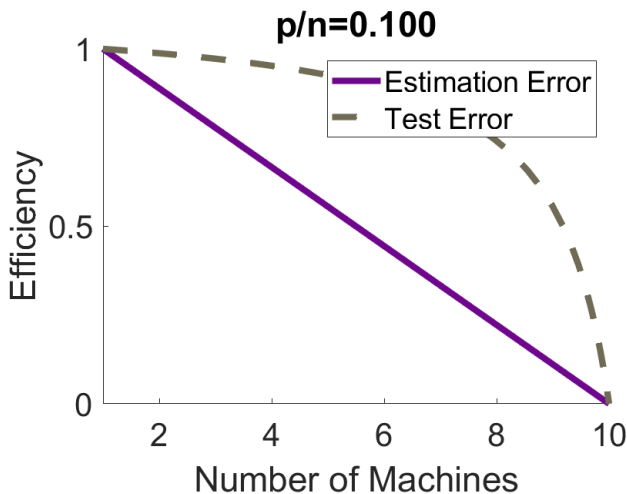


Figure: The loss of efficiency is much worse for estimation ($\frac{\mathbb{E}\|\hat{\beta}-\beta\|^2}{\mathbb{E}\|\hat{\beta}_{dist}-\beta\|^2}$) than for test error ($\frac{\mathbb{E}(x_t^\top \hat{\beta}-y_t)^2}{\mathbb{E}(x_t^\top \hat{\beta}_{dist}-y_t)^2}$).