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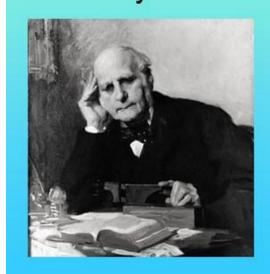
**Simple Linear Regression (SLR)** 

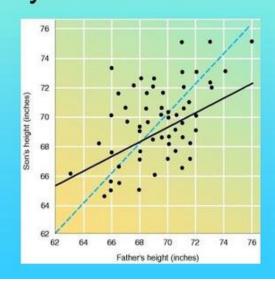
## **Simple Linear Regression**

## The purpose of regression?

- Describe functional relationships between variables
- Control
- Prediction of outcomes

Developed by Sir Francis Galton (1822-1911) in his article "Regression towards mediocrity in hereditary structure"





## **Simple Linear Regression**

#### The basic concepts of regression

- Describe statistical relationships between variables
- The statistical relation has two essential ingredients
  - A tendency of the response variable Y to vary with the predictor variable X
    - There is a probability distribution of Y for each level of X.
  - A scattering of points around the curve of statistical relationship.
    - The means of these probability distributions vary in some systematic fashion with X

# Example (diamonds.csv)

### Variables:

Response Variable: Price in Singapore dollars (Y)

Explanatory Variable: Weight of diamond in carats (X)

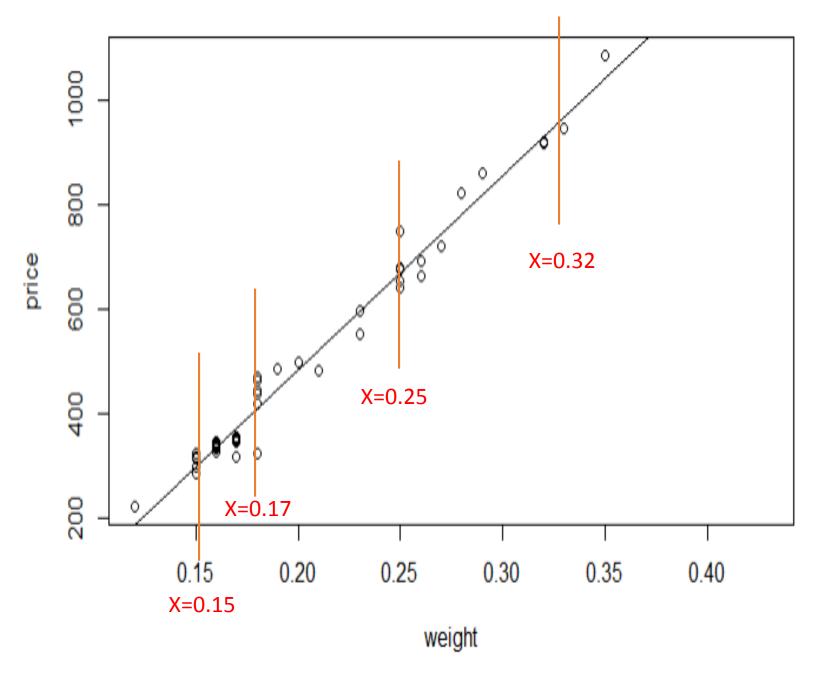
## Goal:

Predict the price of a sale for a 0.43 carat diamond ring

What are the two ingredients in understudying statistical relationship between price and weight?

## Scatter plot

#### Mean price = intercept+ slope (weight)



- The means of the price distributions increase linearly with the weight
- For any given weight, the distribution of price varies, and we can see later that the distribution is Normal (the bell-shape distribution).

## **Notation for Simple Linear Regression (SLR)**

- Observe a pair of variables (explanatory and response) on each of i = 1, 2, ..., n samples
- Each pair often called a case or a data point  $(X_i, Y_i)$
- $Y_i$  is the value of the response for the i-th case
- $X_i$  is the value of the explanatory variable for the i-th case

## **Simple Linear Regression Model**

$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$$
 for  $i = 1, 2, ..., n$ 

## **Simple Linear Regression Model Parameters**

- $\beta_0$  is the intercept.
- $\beta_1$  is the slope.
- $\varepsilon_i$  are independent, normally distributed random errors with mean 0 and variance  $\sigma^2$ ,

$$\varepsilon_i \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$$

## **Features of Simple Linear Regression Model**

- Individual observations:  $Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$
- Since  $\varepsilon_i$  are random,  $Y_i$  are also random and

$$E(Y_i) = \beta_0 + \beta_1 X_i + E(\varepsilon_i) = \beta_0 + \beta_1 X_i$$

$$Var(Y_i) = 0 + 0 + Var(\varepsilon_i) = \sigma^2.$$

Since  $\varepsilon_i$  is Normally distributed,  $Y_i \sim N(\beta_0 + \beta_1 X_i, \sigma^2)$ 

## **Fitted Regression Equation and Residuals**

The parameters  $\beta_0$ ,  $\beta_1$ , and  $\sigma^2$  are unknown and must be estimated from the data.

The "hat" symbol



is "point estimation" 
$$\Rightarrow \hat{Y} = b_0 + b_1 X$$

•  $b_0$  estimates  $\beta_0$  (intercept)

$$\widehat{\beta}_0 = b_0 \text{ or } \widehat{\beta}_1 = b_1$$

- $b_1$  estimates  $\beta_1$  (slope)
- $\hat{Y}_i = b_0 + b_1 X_i$  gives the estimated mean of Y when the predictor is  $X_i$ .
- The *residual* for the *i*-th case is  $e_i = Y_i \hat{Y}_i = Y_i (b_0 + b_1 X_i)$
- $s^2 = \text{Var}(e_i)$  estimates the error variance  $\sigma^2$  $\widehat{\sigma^2} = s^2$

The residual  $e_i$  (in one sample) is NOT the same as the error  $\varepsilon_i$  (in population)!

$$\hat{\varepsilon} = e$$

## Estimating the parameters with Least Squares (LS) Solution

- We want to find the "best" estimates,  $b_0$  and  $b_1$ .
- Minimize the sum of the squared residuals,  $\sum_{i=1}^{n} e_i^2$ , i.e., find

$$\underset{(b_0,b_1)}{\operatorname{arg\,min}} = \sum_{i=1}^{n} [Y_i - (b_0 + b_1 X_i)]^2$$

- How? Calculus!
  - 1. Take derivatives with respect to  $b_0$  and with respect to  $b_1$ .
  - 2. Set equations equal to zero and solve for both  $b_0$  and b.

## Estimating the parameters with Least Squares (LS) Solution

• The best estimates of  $\beta_1$  and  $\beta_0$  given the data (X, Y) are:

$$b_1 = \frac{\sum (X_i - \overline{X})(Y_i - \overline{Y})}{\sum (X_i - \overline{X})^2} = \frac{SS_{XY}}{SS_X}$$
 SS is "sum of squares"

• 
$$b_0 = \bar{Y} - b_1 \bar{X}$$

- This estimate is the "best" because it
  - 1. is *unbiased* (its expected value is equal to the true value)
  - 2. has minimum variance

## Estimate the parameters with Maximum Likelihood Estimation (MLE)

Our model says that  $Y_i \sim N(\beta_0 + \beta_1 X_i, \sigma^2)$ .

Given  $X_i$ , the probability of data point i is,

$$f_i = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}\left(\frac{Y_i - \beta_0 - \beta_1 X_i}{\sigma}\right)^2}$$

 $\beta_0$  and  $\beta_1$  are unknown, but the *likelihood* of the proposed values  $(\beta_0^*, \beta_1^*)$  given the data is,

$$L(\beta_i^*, \beta_i^*|X, Y) = f_1 \times f_2 \times \ldots \times f_n = \prod_{i=1}^n f_i$$

L is maximized when  $\beta_0^* = b_0$  and  $\beta_1^* = b_1$ . Thus, the LS estimates,  $b_0$  and  $b_1$ , are also

the estimated parameter values that are most (probabilistically) consistent with the data!

## Estimation of stochastic variance, $\sigma^2$

We estimate  $\sigma^2$  as the sum of the squared residuals, SSE, divided by the degrees of freedom:

$$s^{2} = \frac{\Sigma(Y_{i} - \hat{Y}_{i})^{2}}{n - 2} = \frac{\Sigma e_{i}^{2}}{n - 2} = \frac{SSE}{n - 2} = \frac{SSE}{DFE} = MSE$$

**SSE** stands for "sum of squares error"

**DFE** stands for "degree of freedom of error"

MSE stands for "mean squared error"

$$E\{MSE\}=\sigma^2$$
  $\longrightarrow$  MSE is an unbiased estimator of  $\sigma^2$ 

$$S=\sqrt{MSE}$$
 This is the residual standard error, which estimates the residual standard deviation ( $\sigma$ )

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MSE measures variability around the fitted regression line,

A\_\_\_\_\_(A. smaller/B larger) MSE is preferred and often used as a criterion for model selection

#### A comment on the notation

We will also estimate variances for other quantities.

These will also be denoted  $S^2$ , but will have a subscript to identify them, e.g.  $S^2_{\{b_i\}}$ .

Without any subscript,  $\mathbf{S}^2$  refers to the the estimated variance of the residuals.

And S refers to the standard error of the residuals.

# Identifying statistics and estimates in the R output

```
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) -259.63 17.32 -14.99 <2e-16 ***
weight
           3721.02
                   81.79 45.50 <2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 31.84 on 46 degrees of freedom
  (1 observation deleted due to missingness)
Multiple R-squared: 0.9783, Adjusted R-squared: 0.9778
F-statistic: 2070 on 1 and 46 DF, p-value: < 2.2e-16
Analysis of Variance Table
Response: price
          Df Sum Sq Mean Sq F value Pr(>F)
           1 2098596 209<u>8596</u> 2070 < 2.2e-16 ***
Residuals 46 46636
                         1014
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

diamond.mod<-lm(price~weight, diamond)
summary(diamond.mod)
anova(diamond.mod)</pre>

$$Y = \beta_0 + \beta_1 X + \epsilon$$

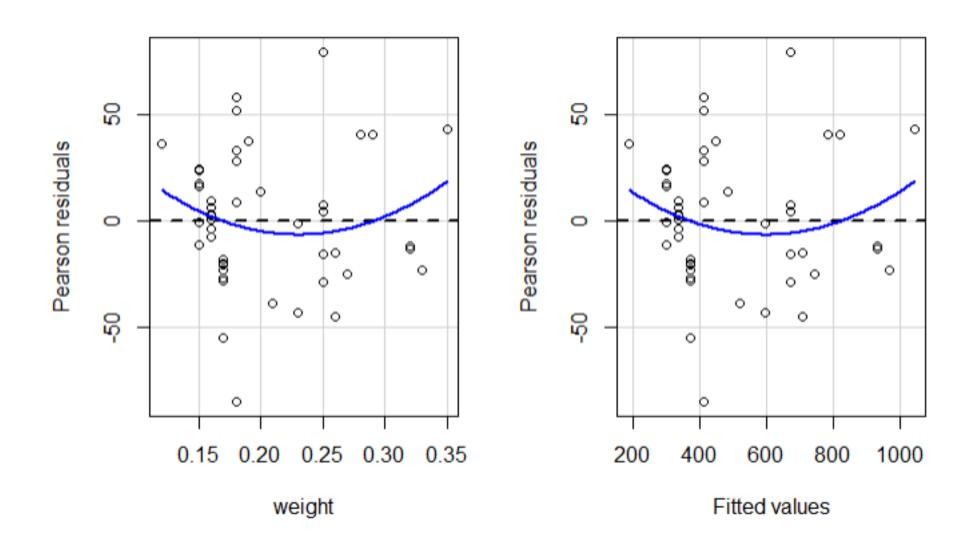
$$MSE =$$

$$s = \sqrt{MSE} =$$

$$DF = n - 2 =$$
after remove 1 observation

# Residual plots

## residualPlots(diamond.mod)



Residuals show a random pattern.

## **Properties of the LS Line**

- The least-squares line always passes through the point (X, Y).
- The residuals always sum to zero:

$$\sum e_i = \sum [Y_i - (b_0 + b_1 X_i)]$$

$$= \sum Y_i - b_0 - b_1 X_i$$

$$= n\overline{Y} - nb_0 - nb_1 \overline{X}$$

$$= n[(\overline{Y} - b_1 \overline{X}) - b_0]$$

$$= 0$$

- $\sum Y_i = \sum \widehat{Y}_i$
- $\sum X_i e_i = 0$
- $\sum \hat{Y}_i e_i = 0$