

STA 35C: Statistical Data Science III

Lecture 5: Multiple Linear Regression & Polynomial Regression

Dogyoon Song

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Agenda

Last time: Simple linear regression

- Model: $Y = \beta_0 + \beta_1 X + \epsilon$
- Least squares: estimate β_0, β_1 by minimizing $\text{RSS} = \sum_{i=1}^n (y_i - \hat{y}_i)^2$
- Inference on β_0, β_1 : confidence intervals & hypothesis tests using $\text{SE}(\hat{\beta}_i)$
- Model fit: $R^2 = 1 - \frac{\text{RSS}}{\text{TSS}}$ where $\text{TSS} = \sum_{i=1}^n (y_i - \bar{y})^2$

Today: Extending simple linear regression

- What if we have more than one predictor: X_1, X_2, \dots ?
→ Multiple linear regression
- What if X_1 and X_2 interact, or if Y depends on X^2 instead of X ?
→ Polynomial regression

Outline

- Multiple linear regression
- Key statistical questions in multiple linear regression
- Accommodating non-linear relationships

Motivation for multiple linear regression

Recall the **Advertising** dataset and the three separate simple linear regression lines:

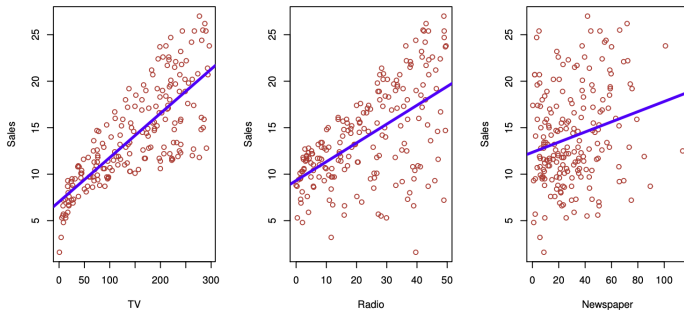


Figure: The **Advertising** data set: **Sales** of a product in 200 different markets against advertising budgets for three media: **TV**, **Radio**, and **Newspaper** [JWHT21, Figure 2.1].

Problem: Each simple linear regression line ignores the other two predictors

Question: Can we extend our analysis to accommodate *all* predictors simultaneously?

Multiple linear regression: Setup

We predict Y using multiple variables X_1, X_2, \dots, X_p , assuming:

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_p X_p + \epsilon,$$

- *Model parameters:* $\beta_0, \beta_1, \dots, \beta_p$ are fixed, unknown constants
- ϵ is an error term, independent of X_1, \dots, X_p

The coefficient β_j is interpreted as the *average effect* on Y of a unit increase in X_j , *holding all other predictors fixed*

Once we *estimate* $\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_p$ from training data, we can predict

$$\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_p x_p$$

Visualizing multiple linear regression

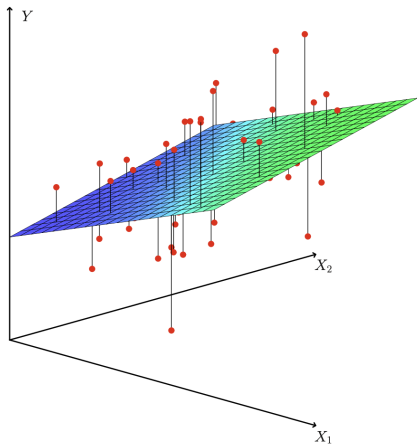


Figure: An illustration of multiple linear regression [JWHT21, Figure 3.4].

Coefficient estimation via least squares

$\beta_0, \beta_1, \dots, \beta_p$ are unknown and must be estimated from data $(x_1, y_1), (x_2, y_2), \dots, (x_n, y_n)$, where $x_i = (x_{i1}, x_{i2}, \dots, x_{ip})$

Again, we use the **least squares** criterion:

- The *least squares* approach chooses $\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_p$ to minimize the RSS:

$$\text{RSS} = \sum_{i=1}^n (\hat{y}_i - y_i)^2 = \sum_{i=1}^n (\hat{\beta}_0 + \hat{\beta}_1 x_{i1} + \hat{\beta}_2 x_{i2} + \dots + \hat{\beta}_p x_{ip} - y_i)^2$$

- The solutions have more complicated forms in this multiple-variable case¹:

$$\bullet \hat{\beta} = \begin{bmatrix} \hat{\beta}_0 \\ \vdots \\ \hat{\beta}_p \end{bmatrix} = (\mathbf{X}^\top \mathbf{X})^{-1} \mathbf{X}^\top \mathbf{y} \text{ where } \mathbf{X} = \begin{bmatrix} 1 & x_{11} & \cdots & x_{1p} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_{n1} & \cdots & x_{np} \end{bmatrix} \text{ and } \mathbf{y} = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}$$

¹They can be derived by setting the partial derivatives of RSS to zero

Pop-up quiz: Coefficients in multiple linear regression

Scenario: We fit a multiple linear regression model on the **Advertising** dataset:

$$\text{Sales} = \beta_0 + \beta_1 \text{TV} + \beta_2 \text{Radio} + \beta_3 \text{Newspaper} + \varepsilon.$$

Suppose that we obtain:

$$\hat{\beta}_1 = 0.04, \quad \hat{\beta}_2 = 0.18, \quad \hat{\beta}_3 = -0.02.$$

Question: Which statement *best* describes the meaning of $\hat{\beta}_1 = 0.04$ in this model?

Multiple-choice answers:

- A) TV advertising alone explains 4% of the variation in **Sales**.
- B) For every additional dollar spent on TV, **Sales** increases by 0.04 units, assuming **Radio** and **Newspaper** are both zero.
- C) For every additional dollar spent on TV advertising, **Sales** increases by 0.04 units on average, controlling for **Radio** and **Newspaper**.
- D) If TV advertising goes up by \$100, **Sales** is guaranteed to go up by 4 units, regardless of **Radio** or **Newspaper**.

Some key questions with multiple predictors

When we perform multiple linear regression, we often want to answer questions like:

- Are predictors X_1, \dots, X_p related to Y (i.e., do they help predict Y)?
- Which subset of X_1, \dots, X_p is most important?
- How well does the model fit the data?
- Given new predictor values, what response value should we predict and how accurate is that prediction?

Let's address these questions one by one

Hypothesis testing for relationship between Y and each X_j

Recall from simple linear regression that we conduct a hypothesis test using a t -statistic:

- $H_0 : \beta_1 = 0$ (no relationship) vs. $H_1 : \beta_1 \neq 0$ (some relationship)
- We reject H_0 or not, based on the value of

$$t = \frac{\hat{\beta}_1 - 0}{\text{SE}(\hat{\beta}_1)}$$

In multiple linear regression, we can do the same test to see if Y is related to each X_j

However:

- Would we get the same conclusions from simple vs. multiple regressions?
- What if we want to test whether Y is related to *any* of the X_j 's?

Advertising example: Simple vs. multiple regressions

Q: Is **newspaper** useful in predicting **sales**?

	Coefficient	Std. error	t-statistic	p-value
Intercept	7.0325	0.4578	15.36	< 0.0001
TV	0.0475	0.0027	17.67	< 0.0001

	Coefficient	Std. error	t-statistic	p-value
Intercept	9.312	0.563	16.54	< 0.0001
radio	0.203	0.020	9.92	< 0.0001

	Coefficient	Std. error	t-statistic	p-value
Intercept	12.351	0.621	19.88	< 0.0001
newspaper	0.055	0.017	3.30	0.00115

Figure: Separate simple regressions suggest **TV**, **radio**, and **newspaper** are all significant [JWHT21, Tables 3.1 & 3.3].

	Coefficient	Std. error	t-statistic	p-value
Intercept	2.939	0.3119	9.42	< 0.0001
TV	0.046	0.0014	32.81	< 0.0001
radio	0.189	0.0086	21.89	< 0.0001
newspaper	-0.001	0.0059	-0.18	0.8599

Figure: Multiple regression suggests **newspaper** is not significant [JWHT21, Table 3.4].

	TV	radio	newspaper	sales
TV	1.0000	0.0548	0.0567	0.7822
radio		1.0000	0.3541	0.5762
newspaper			1.0000	0.2283
sales				1.0000

Figure: Correlation matrix for **TV**, **radio**, **newspaper**, and **sales** [JWHT21, Table 3.5].

In multiple regression, β_j measures the effect of X_j on Y , *holding all other predictors fixed*

Advertising example: Single vs. any predictor

Q: Is “any” of TV, radio, newspaper useful in predicting sales?

We now test a different, *joint* hypothesis:

$$H_0 : \beta_1 = \beta_2 = \cdots = \beta_p = 0 \quad \text{vs.} \quad H_1 : \text{at least one } \beta_j \neq 0$$

This can be tested using the *F-statistic*:

$$F = \frac{(\text{TSS} - \text{RSS})/p}{\text{RSS}/(n - p - 1)}$$

Rationale: If H_0 is true,

- $\mathbb{E}[\text{RSS}/(n - p - 1)] = \mathbb{E}[(\text{TSS} - \text{RSS})/p] = \sigma^2$
- F follows an F-distribution with $(p, n - p - 1)$ degrees of freedom

\Rightarrow If H_0 is true, F will likely take a typical value; if F is very large, then we reject H_0

Selecting “important” predictors²

Suppose we are confident that *at least some* predictors are related to Y

Variable selection: “Which subset of predictors is most useful or important?”

- Naive approach: Try all $2^p - 1$ possible combinations of predictors
 - Evaluate each model by some criterion
 - **Challenge:** Intractable for large p (exponential number of subsets)
- Practical approaches:
 - *Greedy methods:* Forward, backward, or stepwise (mixed) selection
 - *Regularization methods:* Modify the least squares criterion, e.g., LASSO

We will discuss these methods in more detail in future lectures

²We will revisit this question later

Evaluating the model fit

The quality of a multiple linear regression fit can be measured by the RSE or the R^2

- **Residual standard error (RSE):** “average deviation of Y from the regression line”

$$\text{RSE} = \sqrt{\frac{\text{RSS}}{n - p - 1}} \quad \text{where} \quad \text{RSS} = \sum_{i=1}^n (y_i - \hat{y}_i)^2$$

- **The R^2 :** “the proportion of variance in Y explained by X ”

$$R^2 = \frac{\text{TSS} - \text{RSS}}{\text{TSS}} = 1 - \frac{\text{RSS}}{\text{TSS}}, \quad \text{where} \quad \text{TSS} = \sum_{i=1}^n (y_i - \bar{y})^2$$

- R^2 always increases when more predictors are added to the model
- “**Adjusted**” R^2 compensates for adding predictors:

$$R_{\text{adj}}^2 = 1 - \frac{\text{RSS}/(n - p - 1)}{\text{TSS}/(n - 1)}$$

Pop-up quiz: R^2 vs. adjusted R^2

Scenario: We fit a model on $n = 100$ data points using a single predictor X_1 :

$$R^2 = 0.80, \quad R^2_{\text{adj}} = 0.79.$$

After adding a second predictor X_2 (suspected to be mostly noise), we get:

$$R^2 = 0.82, \quad R^2_{\text{adj}} = 0.78.$$

Question: Why did R^2 go up while R^2_{adj} went down?

Multiple-choice answers:

- a) There must be a calculation error; if R^2 increases, R^2_{adj} must also increase.
- b) X_2 adds a tiny improvement to the fit by chance, raising R^2 , but not enough to offset the penalty for extra parameters, so R^2_{adj} drops.
- c) Adjusted R^2 always decreases whenever you add predictors, no matter how useful they are.
- d) R^2 does not measure model fit at all, whereas R^2_{adj} is the only valid measure of fit.

Confidence intervals and prediction intervals

With $\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_p$, we predict $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_p x_p$

How certain are we about this prediction?

- $\hat{y} = \hat{f}_{\hat{\beta}}(x)$ only *estimates* $f_{\beta}(x) = \beta_0 + \sum_{j=1}^p \beta_j x_j$.
- $y = f(x) + \epsilon$ has an error term, so additional variability.

Confidence interval for $f(x)$:

- Reflects uncertainty in prediction due to the coefficient estimates
- A 95% CI should contain $f(x)$ with probability 0.95

Prediction interval for y (given x):

- Includes both uncertainty in $\hat{y} = \hat{f}(x)$ and the noise ϵ
- A 95% PI should contain the actual $y = f(x) + \epsilon$ with probability 0.95

Exact formulas are beyond our scope, but in **R**:

```
predict(model, newdata = x0, interval = "confidence", level = 0.95)
```


What if there is a non-linear relationship?

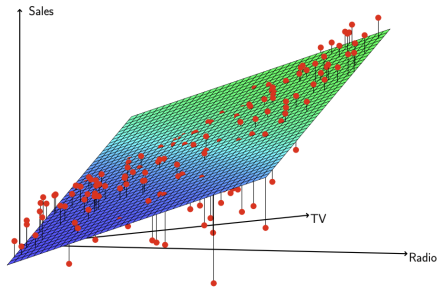


Figure: Pronounced synergy between **TV** and **Radio**; positive residuals cluster along the 45-degree line [JWHT21, Figure 3.5].

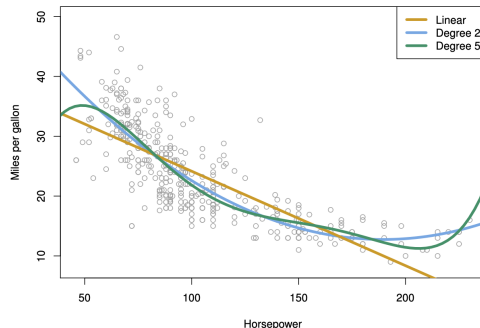


Figure: A non-linear relationship between **mpg** and **horsepower** is noticeable [JWHT21, Figure 3.8].

→ We can add **interaction** terms (**TV** \times **Radio**) or **non-linear** terms (**horsepower**²) to capture these effects

Polynomial regression

Polynomial regression extends the linear model by including powers of predictors³:

$$Y = \beta_0 + \beta_1 X + \beta_2 X^2 + \dots + \beta_d X^d + \epsilon$$

- Treated as multiple linear regression on transformed predictors (X, X^2, \dots, X^d)
- Although non-linear in X , the model is still linear in the coefficients β_j

Example: Interaction effect (synergy between **TV** and **Radio**)

$$\begin{aligned}\text{Sales} &= \beta_0 + \beta_1 \text{TV} + \beta_2 \text{Radio} + \beta_3 \underbrace{\text{TV} \times \text{Radio}}_{\text{interaction term}} + \epsilon \\ &= \beta_0 + (\beta_1 + \beta_3 \text{Radio}) \text{TV} + \beta_2 \text{Radio} + \epsilon\end{aligned}$$

Example: Quadratic model

$$\text{mpg} = \beta_0 + \beta_1 \text{horsepower} + \beta_2 \text{horsepower}^2 + \epsilon$$

³More generally, $Y = \sum_{\alpha: |\alpha| \leq d} \beta_{\alpha} \mathbf{X}^{\alpha} + \epsilon$ where $\alpha = (\alpha_1, \dots, \alpha_p)$ and $\mathbf{X}^{\alpha} = X_1^{\alpha_1} X_2^{\alpha_2} \dots X_p^{\alpha_p}$

Wrap-up

Multiple linear regression assumes a model:

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \cdots + \beta_p X_p + \epsilon$$

with the parameters typically estimated by least squares

We can address the following questions:

- Is X_j related to Y ? \Rightarrow Test $H_0 : \beta_j = 0$
- Is any of X_1, \dots, X_p related to Y ? \Rightarrow Test $H_0 : \beta_1 = \cdots = \beta_p = 0$
- How does Y change per unit change in X_j (others fixed)? $\Rightarrow \hat{\beta}_j$
- How well does the model fit data? $\Rightarrow \text{RSE}, R^2, R^2_{\text{adj}}$
- How certain is our prediction of y ? $\Rightarrow \text{CI \& PI}$
- What if there is a non-linear relationship? \Rightarrow Add non-linear terms

Next lecture: Dummy variables, pitfalls in linear regression, etc.

References



Gareth James, Daniela Witten, Trevor Hastie, and Robert Tibshirani.

An Introduction to Statistical Learning: with Applications in R, volume 112 of *Springer Texts in Statistics*.

Springer, New York, NY, 2nd edition, 2021.