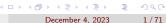
## Notes (To Be Deleted)

- Be able to speak to glasso over neighborhood selection, how sparsity guarantees MLE existence
- Finish missing data results

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#### **Graphical Models**

With a focus towards interimly missing data

Dominic DiSanto

Department of Biostatistics, Harvard University

December 4, 2023

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#### Outline

- Introduction to Graphical Models
- Estimation for Complete Data
  - Neighborhood Selection
  - Graphical Lasso
  - Further Notes
- Stimation with Missingness
  - MissGLasso
  - Additional Missing Data Methods

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#### **Disclaimers**

- Historical coverage is to the best of my ability and time constraint, please correct me with additional information
- Interrupt with any questions, clarification, confusion, etc.
- This is far from a comprehensive treatment, but I attempt to be holistic in my coverage

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Dominic DiSanto Graphical Models

## Outline (Redux)

- 1 Introduction to Graphical Models
- Estimation for Complete Data
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### Graph Theory Origins [8, 11]



Figure: Euler's Bridges Conceptualization (Recreation)

1 Image taken from Wikipedia (https://en.wikipedia.org/wiki/Seven\_Bridges\_of\_Konigsberg) > 4 = > = - < 0 < -

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### Early Applications of Graphs in Mathematics

- $\bullet$  Graph theory attributed to begin with Euler and the "Seven Bridges of Königsberg" ( $\sim\!1736)$
- Random graph theory began developing in  $\sim$ 1940's (Moreno and Jennings) but most notably with the Erdös-Rényi random graph (1958)
- $\bullet$  Ising model ( $\sim$ 1920's) proposed graphical model of interactions of phase transitions (atomic spin)
- Statistical "beginnings"

graphical neural networks

- Arthur Dempster (founding Harvard Stats professor) introduced covariate selection by precision matrix estimation in 1972 [5]
- ullet Judea Pearl  $\sim$ 1980's for causal interretation of Bayesian networks
- Modern interest in related regularized M-estimation problems and

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### Early Applications of Graphs in Mathematics

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- $\bullet$  Ising model ( $\sim\!1920\mbox{'s})$  proposed a graphical model of interactions of phase transitions (atomic spin)
- $\bullet$  Statistical "beginnings"  $^2$  as a subset of methods for contingency tables and log-linear models ( $\sim\!1970\mbox{'s})$
- Arthur Dempster (founding Harvard Stats professor) introduced covariate selection by precision matrix estimation in 1972 [5]
- ullet Judea Pearl  $\sim$ 1980's for causal interretation of Bayesian networks
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<sup>2</sup>Early use in physics were probabilistic, but this may be seen as an early "purse statistics" application > < > >

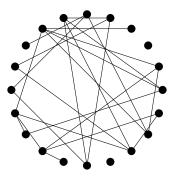
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#### Graphical Model Motivation

Suppose you have 20 random variables\*, how do you model their interrelationship? \*Consider any of the following:

- General -omic data
- Spatial data
- Computational neuroscience data
- Clinical language (see: EHR LLM<sup>a</sup>)
- Time-series data



<sup>&</sup>lt;sup>a</sup>Electronic Healthcare Record Large Language Model

#### Graphs

- Graphs are a natural way to represent interrelationships among our data!
- Present nice properties for estimation of joint distributions
  - Can avail existing graphical algorithms
  - Ability to characterize conditional (in)dependencies
- Probabilistic graphical modelling provide a general formalism of many existing methods in statistics (e.g. Bayesian hierarchical modelling, Hidden Markov Models, Kalman filter)
- Wainwright, Jordan "Graphical Models, Exponential Families, and Variational Inference" (2007) is an excellent reference for further applications (and theory) behind graphical models [12]<sup>3</sup>

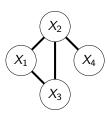
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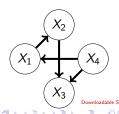
#### Graphs

- Consider random vector  $X \sim \mathcal{N}(\boldsymbol{\mu}, \Sigma)$  and precision matrix  $\Theta \equiv \Sigma^{-1}$ 
  - Interested in estimating  $\Sigma$  to characterize joint distribution  $f_X$
- Can construct a resulting graph  $\mathcal{G} = (V, E)$ ,  $V = X, E \subseteq V \times V$ 
  - Let ne(x) represent the neighborhood of x, or  $ne(x) = \{b \in V \mid (x, b) \in E\}$
- Can construct adjacency matrix  $A \in \mathbb{R}^{|V| \times |V|}$  describing edge set E
  - $A_{ij} = \mathbb{I}\{(i,j) \in E\}$
  - Let  $D_{max}$  represent the maximum degree

Undirected Graph



Directed Graph



#### Notation/Nomenclature

Omitting some philosophical discrepancies

Directed (Acyclic) Graph 
 ⇔ Bayesian network

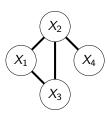
Undirected graph 
 ⇔ Markov network / Markov random field

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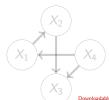
#### Graphs

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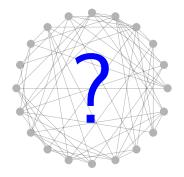
#### Undirected Graph



Directed Graph



### How do we estimate graph structure?



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#### Gaussian Graphical Models

Recall the form and properties of a multivariate Gaussian random vector:

$$f(x; \mu, \Theta) = \frac{|\Theta|^{1/2}}{(2\pi)^{d/2}} \exp\left(-\frac{1}{2}(x - \mu)^T \Theta(x - \mu)\right)$$
$$\mathbb{E}[X_i \mid X_{(-i)}] = \mu_i + (X_{(-i)} - \mu_{(-i)})^T \Theta_{j \neq i} \sigma_{i, j \neq i}$$
$$\operatorname{Var}[X_i \mid X_{(-i)}] = \Sigma_{ii} - \sigma_{i, i \neq i}^T \Theta_{i \neq i} \sigma_{i \neq i, i}$$



#### Gaussian Graphical Models

Gaussianity gives us the nice property that  $\Theta_{ii} = 0 \Leftrightarrow X_i \perp X_i | X_{-\{i,j\}}$ 

$$\mathbb{E}[X_i \mid X_{(-i)}] = \mu_i + (X_{(-i)} - \mu_{(-i)})^T \Theta_{j \neq i} \sigma_{i, j \neq i}$$

$$Var[X_i \mid X_{(-i)}] = \Sigma_{ii} - \sigma_{i,j\neq i}^T \Theta_{j\neq i} \sigma_{j\neq i,i}$$



## Outline (Redux)

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#### Preface

- Consider the setting of graph estimation for  $X = (X_1, ..., X_d)$ 
  - Generally your data may or may not contain a "response variable" of interest
- $\bullet$  Identifying conditional relationships  $\Leftrightarrow$  Estimating/Identifying 0's in  $\Sigma$
- Enforse sparsity for  $\hat{\Theta}$  to account for possible rank-degeneracy of S if  $d\gg N$

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### Neighborhood Selection

- Note that  $E\setminus\{\operatorname{ne}(X_i)\}$  includes all nodes independent of  $X_i$  conditional upon  $\operatorname{ne}(X_i)$
- Proposed by Meinshausen & Bühlmann (2006) [10], for  $X \in \mathbb{R}^d$  concern yourself only with  $(\Theta)_{ij} = 0$ , or  $(\Theta)_{ij} \neq 0$
- Assume sparsity of  $\Theta$  and fit d, element-wise lasso models
  - Regress  $X_i \overset{\mathsf{Lasso}}{\sim} X_1 + ... X_{i-1} + X_{i+1} + ... + X_d$  for all  $i \in [d]$
  - Take  $\hat{\beta}_{(-i)} \in \mathbb{R}^{d-1}$  from each model
  - Conclude<sup>4</sup>  $(\Theta)_{ij} = 0 \Leftrightarrow \hat{\beta}_{(-i)j} = 0 \land \hat{\beta}_{(-j)i} = 0$
- $\bullet$  Admits asymptotic consistency for "zero-selection" of  $\Theta$

4Authors both AND or OR rule for final step with similar performance

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#### Neighborhood Selection

#### Potential Drawbacks:

- Fitting d regression models is almost assuredly redundant
- Although consistent, does not exactly compute but approximates the joint likelihood over X, and thus does not necessarily produce MLE [4]
- ullet  $\hat{\Theta}$  is *not* guaranteed to be positive semi-definite
- Requires sparsity assumptions for theoretical guarantees:
  - $\exists \kappa$ ,  $\max_{a \in V} |\operatorname{ne}(a)| = O(n^{\kappa})$
  - For any conected nodes a,b (i.e.  $\forall (a,b) \in E$ ),  $||\theta^{a,\text{ne}(b)\setminus\{a\}}||_1 \leq \vartheta < \infty$

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#### Graphical Lasso

Natural extension, why not just maximize the log-likelihood?

$$\hat{\Theta}_{\textit{MLE}} = \textit{argmax}_{\Theta} \left\{ \log \det \Theta - \text{trace}(S\Theta) \right\}$$

For N < d, we have the empirical covariance matrix  $S = n^{-1} \sum X_i X_i^T$  is rank-degenerate, and the MLE does not exist!

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#### Graphical Lasso

So we assume sparsity and apply the  $\ell_1$  penalty

$$\hat{\Theta}_{\lambda, \mathit{MLE}} = \mathit{argmax}_{\Theta} \left\{ \log \det \Theta - \mathrm{trace}(S\Theta) - \lambda \sum_{i \neq j} |\Theta_{ij}| \right\}$$

What does this give us?

- True graph recovery guaranteed for  $N = \Omega(D_{max}^3 \log p)$
- Convex program, quickly optimizable

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### Graphical Lasso - Algorithm<sup>5</sup>

6: **for** j = 1, 2, ..., d final update: **do** 

1: Initialize  $\mathbf{W} \leftarrow S + \lambda \mathbf{I}$ 

2: **for** j=1,2,...,d,1,2... until convergence **do**:
3: Partition  $\mathbf{W} = \begin{bmatrix} \mathbf{W}_{kk} & \mathbf{w}_{kj} \\ \mathbf{w}_{jk} & w_{jj} \end{bmatrix}$ 4: Solve estimating equations  $\mathbf{W}_{kk}\beta - \mathbf{s}_{jk} + \lambda \mathrm{sign}\beta = 0$ 5: Update  $\mathbf{w}_{ki} = \mathbf{W}_{kk}\hat{\beta}$ 

7: Solve for  $\hat{\theta}_{kj} = -\hat{\beta} \cdot \hat{\theta}_{jj}$  where  $1/\hat{\theta}_{jj} = 2_{ji} - \mathbf{w}_{\iota i}^T \hat{\beta}_{\iota j}$ 

5Pseudocode adopted from Elements of Statistical Learning, Chapter 17 [6] and Statistical Learning with Sparsity Chapter 94cs [7]

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#### Graphical Lasso - Algorithm

- The estimating equations themselves are solved using cyclical coordinate-descent algorithm
- Additional structural checks on S, W at initialization and interim steps have since been implemented
- Replacing  $\mathbf{W}_{kk}$  with  $\mathbf{S}_k k$  iterates once and returns the neighborhood selection algorithm!

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### Simulations (Complete Data)

- glasso package in R can fit Graphical Lasso as well as neighborhood-selection approximation
- huge is a very nice extension of glasso with algorithmic/convergence fixes, computation in C, additional flexibility, graph generating functions
- sklearn has similar sklearn.covariance.graphicallasso command
- skggm extends Gaussian Graphical Model methods

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## Simulations (Complete Data)

- Theory-suggested penalty  $\lambda = 2\sqrt{\frac{\log d}{N}}$ , but implementations often supply a range similar to glmnet default behavior
- Graph Recovery (accuracacy by proportion of correct edge recovery)
- Operator Norm Distance  $||\hat{\mathbf{\Theta}} \mathbf{\Theta}||_2 \lesssim \sqrt{\frac{D_{max}^2 \log d}{N}}$



## Simulations (Set-Up)

- Generated multivariate normal data for  $d = \{64, 128, 256\}$  with  $AR(n,\rho)$  adjacency structures
- ullet Assessed edge-selection performance for N ranging from 10 to 2000
  - TPR = (# of true edges selected) / (# of true edges)
  - TNR = (# of true non-edges not selected) / (# of true non-edges)
  - Operator norm  $||\hat{\mathbf{\Theta}} \mathbf{\Theta}||_2$

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## Simulations (Set-Up)

$$AR(3,\rho) = \begin{bmatrix} 1 & \rho^1 & \rho^2 & \rho^3 & 0 & \dots & \dots & 0 \\ \rho^1 & 1 & \rho^1 & \rho^2 & \rho^3 & 0 & \dots & \dots & 0 \\ \rho^2 & & \rho^1 & 1 & \rho^1 & \rho^2 & \rho^3 & 0 & \dots \\ \vdots & \vdots \\ 0 & \dots & \rho^3 & \rho^2 & \rho^1 & 1 & \rho^1 & \rho^2 & \rho^3 \end{bmatrix}$$

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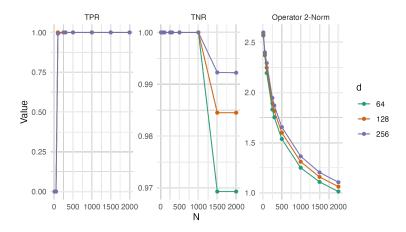


Figure: AR(1),  $\rho = 0.4$  adjacency structure

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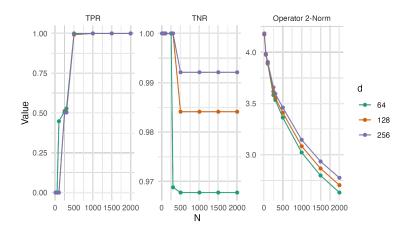


Figure: AR(2),  $\rho = 0.4$  adjacency structure

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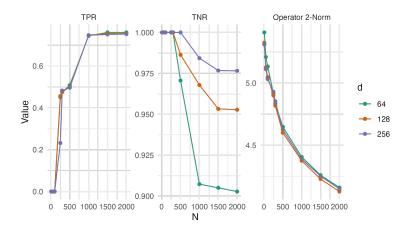


Figure: AR(4),  $\rho = 0.4$  adjacency structure

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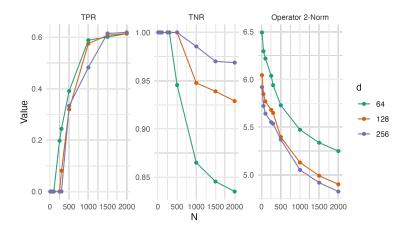


Figure: AR(8),  $\rho = 0.4$  adjacency structure

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# Simulations (Neighborhood Selection - Complete Data)

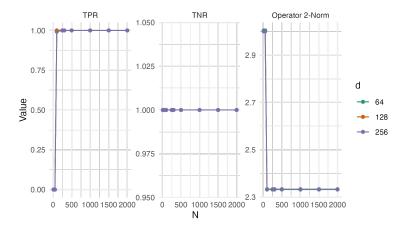


Figure: AR(1),  $\rho = 0.4$  adjacency structure

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# Simulations (Neighborhood Selection - Complete Data)

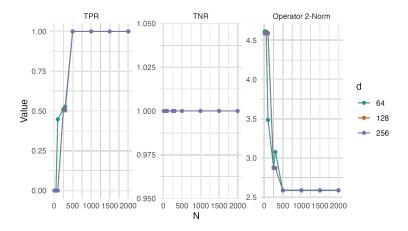


Figure: AR(2),  $\rho = 0.4$  adjacency structure

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# Simulations (Neighborhood Selection - Complete Data)

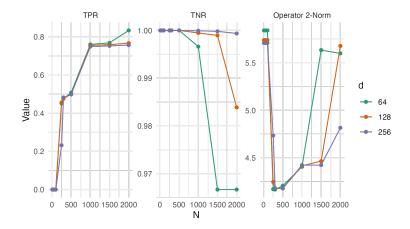


Figure: AR(4),  $\rho = 0.4$  adjacency structure

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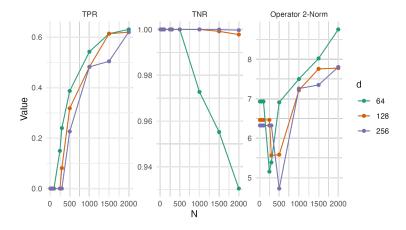


Figure: AR(8),  $\rho = 0.4$  adjacency structure

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### Further Notes

- Neighborhood selection (slightly) seems to outperform GLasso in edge-selection, more notably in our less-sparse settings<sup>6</sup>
- GLasso "better approximate" Θ with asymptotic guarantees not provided by neighborhood selection
- Error scaling more vulnerable to  $D_{max}$  than d-dimensionalty of random vector
- Omitted results for random Erdös Rényi graphs yield similar results, conclusions

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## Outline (Redux)

- Introduction to Graphical Models
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- 3 Estimation with Missingness
  - MissGLasso
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### Motivation

- Methods above largely assume complete data
- Networks change, measurement availability (and quality) varies
- Measurement is also often differential between nodes
- Complete case analysis can drastically reduce sample size if requiring complete data on all nodes

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### MissGLasso

- Suppose for  $X^{(i)} \sim N(\mu, \Sigma)$ , we can partition  $X = (X_o^{(i)}, X_m^{(i)})$
- Consider all observed data  $\mathbf{X}_o = (X_o^{(1)}, X_o^{(2)}, ..., X_o^{(N)})$  and maximize the resulting likelihood (with penalty):

$$\begin{split} \hat{\Theta} &= argmax_{\Theta \succ 0} \ell_{\lambda}(\Sigma; \mathbf{X}_{o}) = argmin_{\Theta \succ 0} - \ell_{\lambda}(\Sigma; \mathbf{X}_{o}) \\ -\ell_{\lambda}(\Sigma; \mathbf{X}_{o}) &= -\sum_{i=1}^{N} \frac{1}{2} \log \det \Theta_{o} + \frac{1}{2} (X_{o}^{(i)} - \mu_{o})^{T} \Theta_{o} (X_{o}^{(i)} - \mu_{o}) + \lambda ||\Theta||_{1} \end{split}$$

• Despite similar structure to our typical Gaussian log-likelihood, because of missingness, **this is not necessarily convex**!

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• We can consider our unobserved  $X_m$  latent, and implement EM!

$$-\ell_{\lambda}(\Sigma; \mathbf{X}) = -\frac{n}{2} \log \det \Theta + \frac{n}{2} \mu^T \Theta \mu - \mu^T \Theta \underbrace{\mathbf{X}^T \mathbf{1}}_{\mathbf{T}_1} + \frac{1}{2} \mathrm{tr}(\Theta \underbrace{\mathbf{X}^T \mathbf{X}}_{\mathbf{T}_2}) + \lambda ||\Theta||_1$$

- $\mathbb{E}(\ell_{\lambda}|X_o,\mu,\Theta)$  is only random in  $\mathsf{T_1},\mathsf{T_2}$
- The E-step reduces to  $\mathbb{E}(\mathsf{T}_1|X_o,\mu,\Theta)$  and  $\mathbb{E}(\mathsf{T}_2|X_o,\mu,\Theta)$ 
  - More simply,  $\mathbb{E}(x_k^{(i)}|X_o^{(i)}, \mu, \Theta)$  and  $\mathbb{E}(x_j^{(i)}x_k^{(i)}|X_o^{(i)}, \mu, \Theta)$ , j, k = 1, ..., d; i = 1, ..., N

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Moreover, by conditional properties of the Gaussian, for the  $\gamma$ th iteration, we have simple forms for these expectations!

$$\mathsf{Let}\ c = \mu_m^{(\gamma)} - (\Theta_{mm}^{(-1)})^{(\gamma)} \Theta_{mo}^{(\gamma)}(X_o - \mu_o)$$

$$\mathbf{T}_1^{(\gamma+1)} = \mathbb{E}(x_j | X_o, \mu, \Theta) = \begin{cases} x_j, & \text{if } x_j \text{ observed} \\ c_j & \text{otherwise} \end{cases}$$

$$\mathbf{T}_2^{(\gamma+1)} = \mathbb{E}(x_j x_k | X_o, \mu, \Theta) = \begin{cases} x_j x_k, & \text{if } x_j, x_k \text{ observed} \\ x_j c_k, & \text{if only } x_j \text{ observed} \\ \left(\Theta_{mm}^{(\gamma)}\right)_{jk}^{-1} + c_j c_k, & \text{if } x_j, x_k \text{ missing} \end{cases}$$

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M-step also follows from first-order optimality conditions on the log-likelihood with these nice expectations:

$$\begin{split} \mu^{(\gamma+1)} &= \frac{1}{n} \mathbf{T}_1^{(\gamma+1)} \\ \Theta^{(\gamma+1)} &= \textit{argmin}_{\Theta \succ 0} \Big\{ - \log \det \Theta + \text{tr} \left( \Theta \left[ \frac{1}{n} \mathbf{T}_2^{(\gamma+1)} - \mu^{(\gamma+1)} (\mu^{(\gamma+1)})^T \right] \right) \\ &+ \frac{2\lambda}{n} ||\Theta||_1 \Big\} \end{split}$$

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- Estimate the expectations of sufficient statistics over complete data
  - Conditional Gaussian properties give simple, closed forms
- Maximize likelihood (GLasso) and update parameter estimates by optimality conditions

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### MissGlasso - Notes

- Fairly intuitive implementation which allows us to use all data, including rarely observed nodes
- Outperforms complete case analysis and mean-imputation in graph selection
- Assumes MAR and empirically underperforms in simulations with MNAR data

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# What is the state of methods for partially-observed, MNAR graphs?

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### Censored Data

- Methods for censored data included in cglasso using recent papers from the package authors [1, 2]
- Include an extension of the MissGLasso EM algorithm for truncated Gaussian distribution
- Clearly extends beyond the MAR framework but requires strict, known missing data mechanism (e.g. detection limits)

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#### Erose Data

- *Erose* data is a term coined by Zheng, Allen (2023) for data with irregular availability [13]
  - Leads to "drastically different" sample size for a small subset of nodes
  - Erose data almost certainly violate MAR/MCAR assumptions of existing methods
  - Motivated by neuroscience but with applications in genetic expression data.
- Authors propose the Graph Inference when Joint Observations are Erose (GI-JOE) method for edge-wise inference in erosely observed edges
  - Adaptation of debiased neighborhood selection algorithm
  - Does not require imputation of missing values

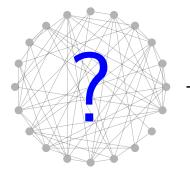
## **Concluding Notes**

- Graphs are a powerful representation of your multivariate data (intuitively and algorithmically)
- These extensions tend to distill to regularized M-estimation problems, an area with great theoretical contributions and guarantees
- Important to consider sparsity/degree scaling of your graph
  - ullet Early methods quickly recover graph structure, even when  $N\ll d$
- huge, cglasso packages are poweful implementations in R





## Graphs, how and why (revisited)?



→ Regularized M-estimation(+)!

 $\rightarrow$  Consistency

 $\rightarrow$  Fast computation even for  $p \gg N$ 

→ MLE via GLasso

→ EM Extensions for MAR



### References I

- Some diagrams generated in conjunction with ChatGPT 3.5
- [1]Luigi Augugliaro, Antonino Abbruzzo, and Veronica Vinciotti. "L1 -Penalized censored Gaussian graphical model". en. In: *Biostatistics* (Sept. 2018).
- [2] Luigi Augugliaro, Gianluca Sottile, and Veronica Vinciotti. "The conditional censored graphical lasso estimator". en. In: Statistics and Computing 30.5 (Sept. 2020), pp. 1273-1289.
- Luigi Augugliaro et al. cglasso: Conditional Graphical LASSO for [3] Gaussian Graphical Models with Censored and Missing Values. Jan. 2023.



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- A. P. Dempster. "Covariance Selection". In: Biometrics 28.1 (1972). [5] Publisher: [Wiley, International Biometric Society], pp. 157–175.
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- [7] Trevor Hastie, Robert Tibshirani, and Martin Wainwright. Statistical Learning with Sparsity: the Lasso and Generalizations. Dec. 2016.

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- [10] Nicolai Meinshausen and Peter Bühlmann. "High-dimensional graphs and variable selection with the Lasso". In: The Annals of Statistics 34.3 (June 2006). Publisher: Institute of Mathematical Statistics, pp. 1436–1462.
- [11] Rob Shields. "Cultural Topology: The Seven Bridges of Königsburg, 1736". en. In: Theory, Culture & Society 29.4-5 (July 2012). Publisher: SAGE Publications Ltd, pp. 43–57.

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- [12] Martin J. Wainwright and Michael I. Jordan. "Graphical Models, Exponential Families, and Variational Inference". en. In: Foundations and Trends® in Machine Learning 1.1–2 (2007), pp. 1–305.
- [13] Lili Zheng. Gl-JOE: Graph Inference when Joint Observations are Erose. Mar. 2023.

## **Appendix Slides**





## Erdös Rényi Graph Results (Complete Data)



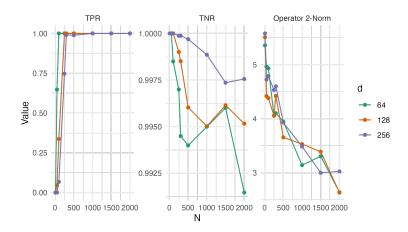


Figure: ER(p=0.01),  $\rho = 0.4$  adjacency structure

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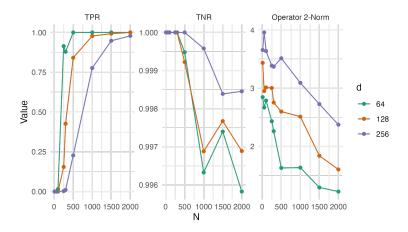


Figure: ER(p=0.05),  $\rho = 0.4$  adjacency structure

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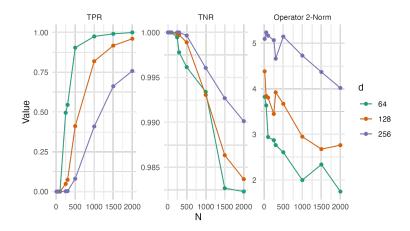


Figure: ER(p=0.1),  $\rho = 0.4$  adjacency structure

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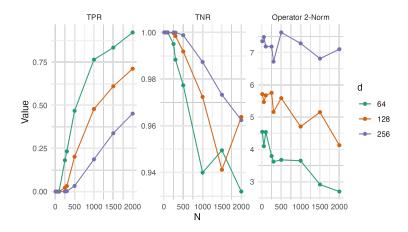


Figure: ER(p=0.2),  $\rho = 0.4$  adjacency structure

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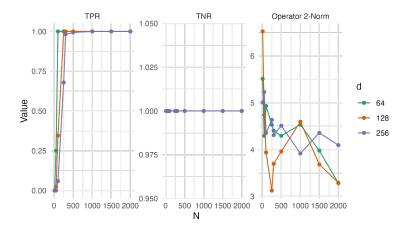


Figure: ER(p=0.01),  $\rho = 0.4$  adjacency structure

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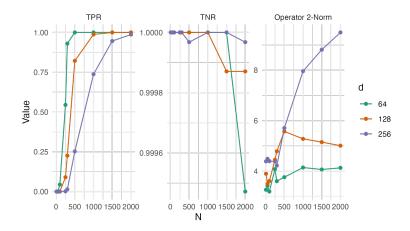


Figure: ER(p=0.05),  $\rho = 0.4$  adjacency structure

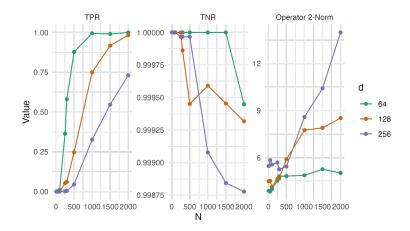


Figure: ER(p=0.1),  $\rho = 0.4$  adjacency structure

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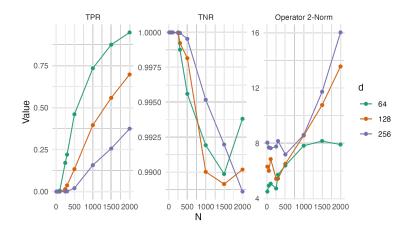


Figure: ER(0.2),  $\rho = 0.4$  adjacency structure

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### Misc. Notes



### Conditional Graphical Lasso

- Can consider random vectors Y, X
  - Y is a random vector of interest (i.e. our graph structure)
  - X is a common set of covariates included in all conditioning
- See cglasso package [3] for implementation
- Package includes conditional MissGLasso implementation
- Methods are extended for censoring



## Forgoing Sparsity Assumptions

- In the above methods, we have almost uniformly assumed some sparsity and applied a penalty  $(\ell_1)$ 
  - How often is this a viable assumption?
  - ② What do we do (or what happens) if we don't meet this sparsity requirement, more severely than our  $AR(\cdot)$  extension sims?
- Mazumder (2012) [9] offers an updated algorithm and insight into performance for p close to but larger than N
- Interplay between d, N, and graph-connectedeness affect computation time and convergence

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### Time-Series Data

- Consider that our repeated observations are time-indexed:
  - $\{X_j(t), t \in \mathcal{T}, j = 1, ..., N\}, X_j \in \mathbb{R}^d$
- Graphical perspective of vector auto-regressive models
  - $X_d(t) = \varepsilon_d(t) + \sum_{j \neq d} \sum_{t \in \mathcal{T}} \alpha_t X_j(t)$
- Can infer "Granger causal" relationships
  - Causal relationships for some time-series using prior data from a different time series

See Michael Eichler's "Granger-causality graphs for multivariate time series" (2007) and Dahlhaus's and Eichler's (2003) "Causality and graphical models in time series" for further discussion

### Inference with Debiased Lasso

- The typical lasso estimator  $\hat{\beta}_{\lambda} = \operatorname{argmin} \beta ||Y X\beta||_2^2 + \lambda ||\beta||_1$  is biased for true  $beta^*$
- ullet Can construct debiased estimator  $\hat{eta}_{\lambda}^{oldsymbol{d}}$  with asymptotic normality
- What inference does this permit in graphical models that use  $\ell_1$  penalization?
- Debiased neighborhood selection is partial motivation for Zheng, Allen's GI-JOE paper [13]

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### "Nothing new under the sun"

My (likely useless and certaintly non-falsifiable) conspiracy theory: Did Euler *really* originate graph theory? For how intuitive graphs seem to understanding interrelationships, this much have existed in some primitive form? Or for how financially relevant this seems, I'm sure some BCE gambler had an idea of "interconnectedness"

For our historical blinders, see Babylonian and Chinese origins of the Pythagorean Theorem

Thoughts, possible leads? Let me know!