

Intro to Random processes - Markov Chain

Applications of Markov chain

- Memory management in computer science
- Text generation: Markov chains can be used to generate sentences in a given language → Natural Language Processing
- Artificial intelligence, learning theory and machine learning.

Random process

- A random process is a mathematical model of a probabilistic experiment that evolves in time and generates a sequence of numerical values.
- Each numerical value in the sequence is modeled by a random variable
- A collection of random variables



Example

- the sequence of daily prices of a stock;
- the sequence of scores in a football game;
- the sequence of failure times of a machine;
- the sequence of hourly traffic loads at a node of a communication network;
- the sequence of radar measurements of the position of an airplane

Discrete random process

$$X_0 \longrightarrow X_1 \longrightarrow \cdots \longrightarrow X_{n-1} \longrightarrow X_n \longrightarrow \cdots$$

- Sequence of random variable X_0, X_1, \dots, X_n
- X_n : state of random process at time n
- All possible values of state: **State space**
- State space is countable \rightarrow Discrete random process



Markov property

$$X_0 \rightarrow X_1 \rightarrow \dots \rightarrow X_{n-1} = i \rightarrow X_n = j \rightarrow \dots$$

Memoryless property *Given current state, the past does not matter*

$$\begin{aligned} &P(X_n = j | X_0 = i_0, X_1 = i_1, \dots, X_{n-1} = i) \\ &= \underbrace{P(X_n = j | X_{n-1} = i)}_{\text{1-step transition probability}} \end{aligned}$$

independent of n

Markov chain

- A Markov chain is a random process with Markov property
- Model specification
 - identify all possible states
 - identify the possible transition
 - identify the transition probability



Markov chain

- A Markov chain is a random process with Markov property
- Model specification
 - identify all possible states
 - identify the possible transition
 - identify the transition probability



Transition matrix

(1-step) -Transition probability

$$P_{ij} = P(X_{n+1} = j | X_n = i)$$

independent of n

$$\begin{array}{c} \text{To} \\ \begin{array}{ccc} 1 & 2 & \dots \end{array} \\ \begin{array}{c} \text{From} \\ 1 \\ \vdots \\ i \\ \vdots \end{array} \left[\begin{array}{ccc} P_{11} & P_{12} & \dots \\ \vdots & \vdots & \vdots \\ P_{i1} & P_{i2} & \dots \\ \vdots & \vdots & \vdots \end{array} \right] \end{array}$$

Index in **row**: current state (**from**)

Index in **column**: next/future state (**to**)



Transition matrix

(1-step) -Transition
probability

$$P_{ij} = P(X_{n+1} = j | X_n = i)$$

independent of n

		To		
		1	2	...
From	1	P_{11}	P_{12}	\dots
	\vdots	\vdots	\vdots	\vdots
	i	P_{i1}	P_{i2}	\dots
	\vdots	\vdots	\vdots	\vdots

Index in **row**: current state (**from**)

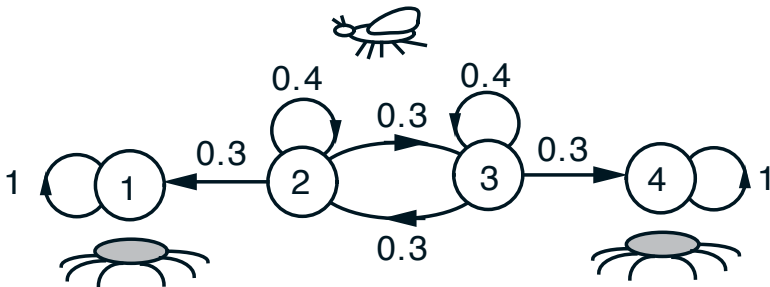
Index in **column**: next/future state
(**to**)



Example

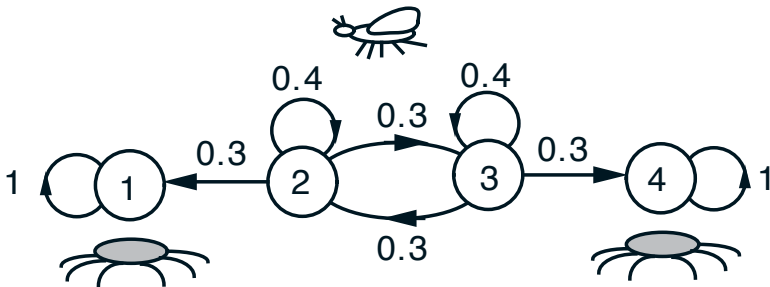
A fly moves along a straight line in unit increments. At each time period, it moves one unit to the left with probability 0.3, one unit to the right with probability 0.3, and stays in place with probability 0.4, independently of the past history of movements. A spider is lurking at positions 1 and 4: if the fly lands there, it is captured by the spider, and the process terminates. Construct the Markov chain model, assuming that the fly starts at position 3





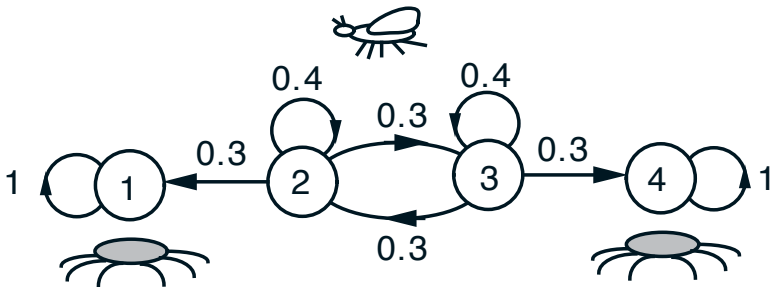
Sample episodes starting from 2:

- $2 \xrightarrow{.3} 1 \xrightarrow{1} 1 \xrightarrow{1} 1$
- $2 \xrightarrow{.3} 3 \xrightarrow{.3} 4 \xrightarrow{1} 4$
- $2 \xrightarrow{.3} 3 \xrightarrow{.4} 3 \xrightarrow{.3} 2 \xrightarrow{.4} 2 \xrightarrow{.3} 3 \xrightarrow{.3} 4$



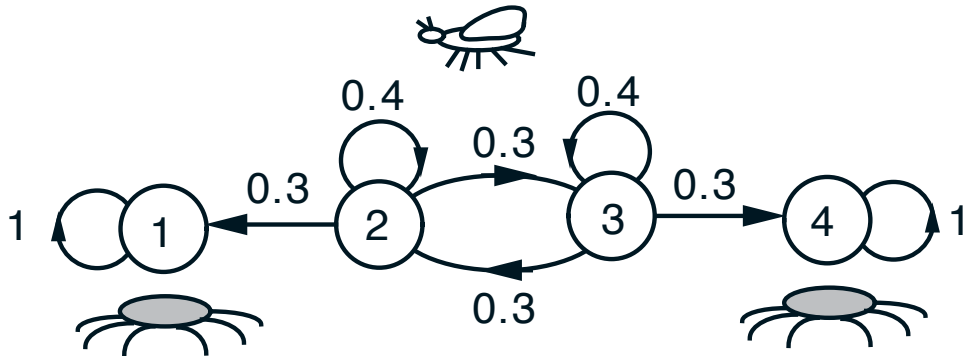
Sample episodes starting from 2:

- $2 \xrightarrow{.3} 1 \xrightarrow{1} 1 \xrightarrow{1} 1$
- $2 \xrightarrow{.3} 3 \xrightarrow{.3} 4 \xrightarrow{1} 4$
- $2 \xrightarrow{.3} 3 \xrightarrow{.4} 3 \xrightarrow{.3} 2 \xrightarrow{.4} 2 \xrightarrow{.3} 3 \xrightarrow{.3} 4$



Sample episodes starting from 2:

- $2 \xrightarrow{.3} 1 \xrightarrow{1} 1 \xrightarrow{1} 1$
- $2 \xrightarrow{.3} 3 \xrightarrow{.3} 4 \xrightarrow{1} 4$
- $2 \xrightarrow{.3} 3 \xrightarrow{.4} 3 \xrightarrow{.3} 2 \xrightarrow{.4} 2 \xrightarrow{.3} 3 \xrightarrow{.3} 4$



1 and 4 are **absorbing states** that once entered, cannot left



Solution

- All possible states: 1, 2, 3, 4

- Transition probability

- $p_{11} = 1, p_{44} = 1$



$$p_{ij} = \begin{cases} 0.3 & \text{if } j = i + 1 \\ 0.4 & \text{if } j = i \\ 0.3 & \text{if } j = i - 1 \end{cases} \quad \text{for } i = 2, 3, \dots, m - 1$$

Solution

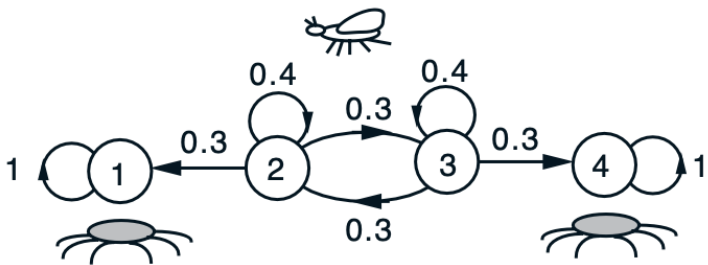
- All possible states: 1, 2, 3, 4
- Transition probability
 - $p_{11} = 1$, $p_{44} = 1$

$$p_{ij} = \begin{cases} 0.3 & \text{if } j = i + 1 \\ 0.4 & \text{if } j = i \\ 0.3 & \text{if } j = i - 1 \end{cases} \quad \text{for } i = 2, 3, \dots, m - 1$$

Solution

- All possible states: 1, 2, 3, 4
- Transition probability
 - $p_{11} = 1$, $p_{44} = 1$
 -

$$p_{ij} = \begin{cases} 0.3 & \text{if } j = i + 1 \\ 0.4 & \text{if } j = i \\ 0.3 & \text{if } j = i - 1 \end{cases} \quad \text{for } i = 2, 3, \dots, m - 1$$



	1	2	3	4
1	1.0	0	0	0
2	0.3	0.4	0.3	0
3	0	0.3	0.4	0.3
4	0	0	0	1.0

p_{ij}

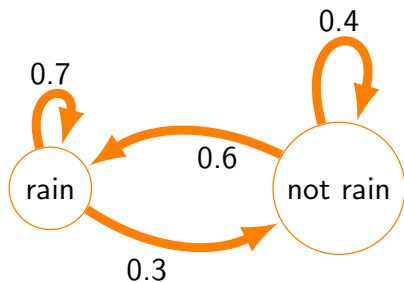


Example: Weather forecast

Suppose that the chance of rain tomorrow depends on previous weather conditions only through whether or not it is raining today and not on past weather conditions. Suppose also that if it rains today, then it will rain tomorrow with probability .7; and if it does not rain today, then it will rain tomorrow with probability .6. Find a Markov chain that modeling the system.



Solution

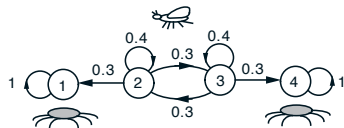


- State 1 (rain)
- State 2 (no rain)

- Transition matrix

		Next	
		Rain	No rain
Current	Rain	.7	.3
	No rain	.6	.4

Probability of a path



$2 \rightarrow 3 \rightarrow 3$

$$P(X_2 = 3, X_1 = 3 | X_0 = 2)$$

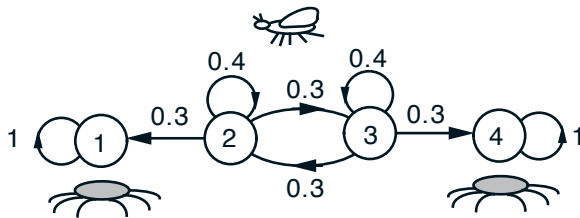
$$= \underbrace{P(X_1 = 3 | X_0 = 2) P(X_2 = 3 | X_1 = 3, X_0 = 2)}_{\text{(multiple law)}}$$

$$= P(X_1 = 3 | X_0 = 2) \underbrace{P(X_2 = 3 | X_1 = 3)}_{\text{(Memoryless property)}}$$

$$= \underbrace{p_{23}}_{\text{from 2 to 3}} \underbrace{p_{33}}_{\text{then from 3 to 3}} = (0.3)(0.4)$$



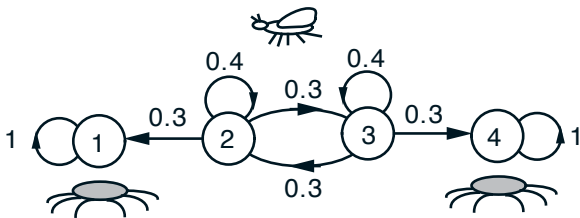
Exercise



Compute

$$P(X_4 = 1, X_3 = 2, X_2 = 2 | X_1 = 3)$$

Exercise



Given that the fly starts at position 2, find all the paths and then compute the probability that

- 1 the fly is at the position 1 after 3 steps.
- 2 the fly visits position 1 for the first time after 3 steps

n-steps transition

Given process initial state i , want to know probability that it will be in state j after n steps

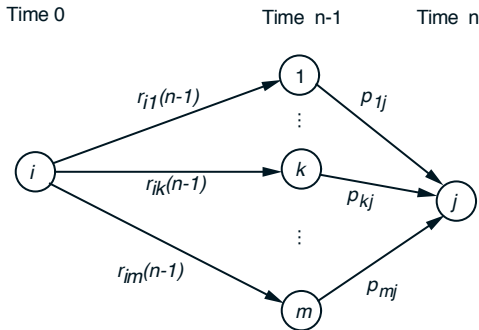
$$r_{ij}^{(n)} = P(X_n = j | X_0 = i)$$

Remark

$$r_{ij}^{(1)} = p_{ij}$$

Chapman-Kolmogorov Equation for n -step transition proba

Key recursion

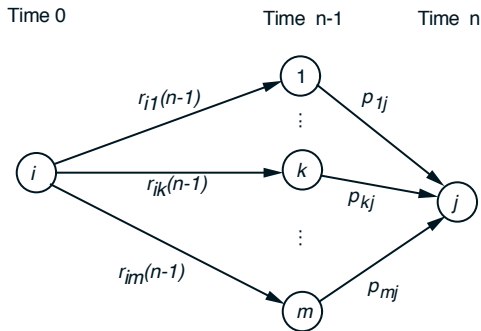


$$r_{ij}^{(n)} = \sum_k r_{ik}^{(n-1)} p_{kj}$$

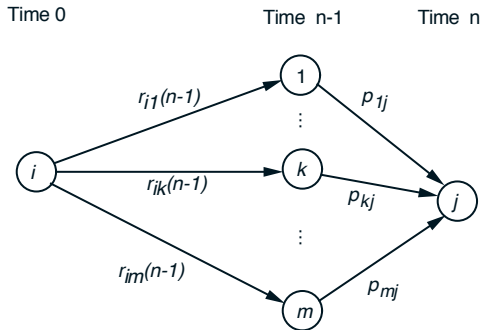
starting with

$$r_{ij}(1) = p_{ij}$$





All the path that start from i and visit j after n steps can be divided into some subsets based on the state that it visits at time $n - 1$



Case 1: Starting from state i , it visits state 1 at time $n - 1$ and in the last transition, it moves from state 1 to state j at time n

Case 2:

state $i \xrightarrow{(n-1)\text{steps}}$ state 2 $\xrightarrow{\text{last step}}$ state j

...

Case k :

state $i \xrightarrow{(n-1)\text{steps}}$ state $k \xrightarrow{\text{last step}}$ state j

...



Thanks to total probability rule

$$r_{ij}^{(n)} = P(X_n = j | X_0 = i) = \sum_{k=1}^m P(X_n = j, X_{n-1} = k | X_0 = i)$$

$$\text{state } i \xrightarrow[r_{ik}^{(n-1)}]{(n-1)\text{ steps}} \text{state } k \xrightarrow[p_{kj}]{\text{last step}} \text{state } j$$

By multiple law

$$\begin{aligned} & P(X_n = j, X_{n-1} = k | X_0 = i) \\ &= P(X_{n-1} = k | X_0 = i) \underbrace{P(X_n = j | X_{n-1} = k, X_0 = i)}_{\text{memoryless property}} \\ &= P(X_{n-1} = k | X_0 = i) P(X_n = j | X_{n-1} = k) \\ &= r_{ik}^{(n-1)} p_{kj} \end{aligned}$$

Hence

$$\begin{aligned} r_{ij}^{(n)} &= \sum_{k=1}^m P(X_n = j, X_{n-1} = k | X_0 = i) \\ &= \sum_{k=1}^m r_{ik}^{(n-1)} p_{kj} \end{aligned}$$



Matrix multiplication

$$\text{row}_i(A) \begin{bmatrix} a_{11} & a_{12} & \cdots & a_{1p} \\ a_{21} & a_{22} & \cdots & a_{2p} \\ \vdots & \vdots & & \vdots \\ a_{i1} & a_{i2} & \cdots & a_{ip} \\ \vdots & \vdots & & \vdots \\ a_{m1} & a_{m2} & \cdots & a_{mp} \end{bmatrix} \quad \text{col}_j(B) \begin{bmatrix} b_{11} & b_{12} & \cdots & b_{1j} & \cdots & b_{1n} \\ b_{21} & b_{22} & \cdots & b_{2j} & \cdots & b_{2n} \\ \vdots & \vdots & & \vdots & & \vdots \\ b_{p1} & b_{p2} & \cdots & b_{pj} & \cdots & b_{pn} \end{bmatrix}$$

$$= \begin{bmatrix} c_{11} & c_{12} & \cdots & c_{1n} \\ c_{21} & c_{22} & \cdots & c_{2n} \\ \vdots & \vdots & & \vdots \\ c_{m1} & c_{m2} & \cdots & c_{mn} \end{bmatrix}$$

$$(\text{row}_i(A))^T \cdot \text{col}_j(B) = \sum_{k=1}^p a_{ik} b_{kj} = c_{ij}$$



Example

$$\underbrace{\begin{bmatrix} 1 & 2 \\ 3 & 4 \end{bmatrix}}_A \underbrace{\begin{bmatrix} 0 & 1 \\ -1 & 5 \end{bmatrix}}_B = \begin{bmatrix} \underbrace{1 \times 0 + 2 \times (-1)}_{\text{row}_1(A)^T \bullet \text{col}_1(B)} & \underbrace{1 \times 1 + 2 \times 5}_{\text{row}_1(A)^T \bullet \text{col}_2(B)} \\ \underbrace{3 \times 0 + 4 \times (-1)}_{\text{row}_2(A)^T \bullet \text{col}_1(B)} & \underbrace{3 \times 1 + 4 \times 5}_{\text{row}_2(A)^T \bullet \text{col}_2(B)} \end{bmatrix}$$

Shortly

$$\begin{bmatrix} 1 & 2 \\ 3 & 4 \end{bmatrix} \begin{bmatrix} 0 & 1 \\ -1 & 5 \end{bmatrix} = \begin{bmatrix} -2 & 11 \\ -4 & 23 \end{bmatrix}$$



Matrix representation

Let

$$P^{(n)} = \begin{bmatrix} r_{11}^{(n)} & \dots & r_{1m}^{(n)} \\ \vdots & \vdots & \vdots \\ r_{m1}^{(n)} & \dots & r_{mm}^{(n)} \end{bmatrix}$$

with $P^{(1)} = P$ then from key recursive, we can verify that

$$P^{(n)} = P^{(n-1)}P$$

In consequence for $n = 2, 3, \dots$, we have

$$P^{(2)} = P^{(1)}P = P.P = P^2, \quad P^{(3)} = P^{(2)}P = P^2.P = P^3$$



Matrix representation

Let

$$P^{(n)} = \begin{bmatrix} r_{11}^{(n)} & \cdots & r_{1m}^{(n)} \\ \vdots & \vdots & \vdots \\ r_{m1}^{(n)} & \cdots & r_{mm}^{(n)} \end{bmatrix}$$

with $P^{(1)} = P$ then from key recursive, we can verify that

$$P^{(n)} = P^{(n-1)}P$$

In consequence for $n = 2, 3, \dots$, we have

$$P^{(2)} = P^{(1)}P = P.P = P^2, \quad P^{(3)} = P^{(2)}P = P^2P = P^3$$

n-step transition matrix

$$P^{(n)} = P^n = \underbrace{P.P \dots P}_{n \text{ times}}$$

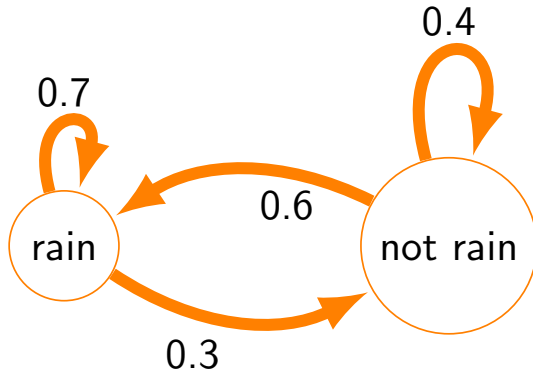
- Element in the row i and column j of the matrix

$$P_{ij}^{(n)} = P(X_n = j | X_0 = i) = r_{ij}^{(n)}$$

- Row i of $P^{(n)}$ provides the conditional distribution of X_n given $X_0 = i$



Example - Weather forecast



If it rains today, calculate the probability that it will rain 4 days from now.



- Transition matrix of the Markov chain

$$\begin{array}{c}
 \text{Next day} \\
 \begin{array}{cc}
 \text{Rain} & \text{No rain} \\
 \text{Current} \begin{array}{c} \text{Rain} \\ \text{No rain} \end{array} & \begin{bmatrix} 0.7 & 0.3 \\ 0.6 & 0.4 \end{bmatrix} = P
 \end{array}
 \end{array}$$

- Want to find $r_{11}^{(4)}$
- Need to calculate 4-step transition probability

$$P^{(4)} = P^4 = P \cdot P \cdot P \cdot P = (P \cdot P) \cdot (P \cdot P)$$

4-step transition matrix is given by

$$\begin{array}{c} \text{After 4 days} \\ \text{Current} \end{array} \begin{array}{cc} & \begin{array}{cc} \text{Rain} & \text{No rain} \end{array} \\ \begin{array}{c} \text{Rain} \\ \text{No rain} \end{array} & \left[\begin{array}{cc} 0.6667 & 0.3333 \\ 0.6666 & 0.3334 \end{array} \right] = P^{(4)} \end{array}$$

$$\text{So } r_{11}^{(4)} = P_{11}^{(4)} = 0.6667$$



Calculating $P^{(4)} = P^4$

First we compute

$$P^2 = P \cdot P = \begin{bmatrix} 0.7 & 0.3 \\ 0.6 & 0.4 \end{bmatrix} \begin{bmatrix} 0.7 & 0.3 \\ 0.6 & 0.4 \end{bmatrix} = \begin{bmatrix} 0.67 & 0.33 \\ 0.66 & 0.34 \end{bmatrix}$$

Then

$$P^{(4)} = (P^2) \cdot (P^2) = \begin{bmatrix} 0.67 & 0.33 \\ 0.66 & 0.34 \end{bmatrix} \begin{bmatrix} 0.67 & 0.33 \\ 0.66 & 0.34 \end{bmatrix}$$

$$\text{or } P^{(4)} = \begin{bmatrix} 0.6667 & 0.3333 \\ 0.6666 & 0.3334 \end{bmatrix}$$



Example

Suppose that X is a Markov chain with two state 0, 1. Its transition matrix is given by

$$\begin{array}{c} \text{After 1 step} \\ \begin{array}{cc} & \begin{array}{cc} 0 & 1 \end{array} \\ \begin{array}{c} \text{Current} \\ 0 \\ 1 \end{array} & \begin{bmatrix} 0.6 & 0.4 \\ 0.2 & 0.8 \end{bmatrix} \end{array}$$

Compute $P(X_5 = 1, X_4 = 0 | X_1 = 0)$

Solution

We have

$$\begin{aligned} &P(X_5 = 1, X_4 = 0 | X_1 = 0) \\ &= P(X_4 = 0 | X_1 = 0) P(X_5 = 1 | X_0 = 1, X_4 = 0) \quad \text{multiple rule} \\ &= P(X_4 = 0 | X_1 = 0) P(X_5 = 1 | X_4 = 0) \quad \text{memoryless property} \\ &= r_{00}^{(3)} p_{01} \end{aligned}$$

$$X_1 = 0 \xrightarrow[3 \text{ steps}]{r_{00}^{(3)}} X_4 = 0 \xrightarrow[1 \text{ step}]{p_{01}} X_5 = 1$$



Solution

We have

$$\begin{aligned} &P(X_5 = 1, X_4 = 0 | X_1 = 0) \\ &= P(X_4 = 0 | X_1 = 0) P(X_5 = 1 | X_0 = 1, X_4 = 0) \quad \text{multiple rule} \\ &= P(X_4 = 0 | X_1 = 0) P(X_5 = 1 | X_4 = 0) \quad \text{memoryless property} \\ &= r_{00}^{(3)} p_{01} \end{aligned}$$

$$X_1 = 0 \xrightarrow[3 \text{ steps}]{r_{00}^{(3)}} X_4 = 0 \xrightarrow[1 \text{ step}]{p_{01}} X_5 = 1$$



From matrix P , we have $p_{01} = 0.4$. In order to compute $r_{00}^{(3)}$, we first compute 3-step transition matrix $P^{(3)} = P \cdot P \cdot P = (P^2) \cdot P$

After 3 steps

$$\begin{array}{c} \text{Current} \\ 0 \\ 1 \end{array} \begin{array}{cc} 0 & 1 \\ \left[\begin{array}{cc} 0.376 & 0.624 \\ 0.312 & 0.688 \end{array} \right] \end{array} = P^{(3)}$$

So $r_{00}^{(3)} = 0.376$.



$$X_1 = 0 \xrightarrow[3 \text{ steps}]{0.376} X_4 = 0 \xrightarrow[1 \text{ step}]{0.04} X_5 = 1$$

Therefore

$$P(X_5 = 1, X_4 = 0 | X_1 = 0) = (0.4) \cdot (0.376)$$

Unconditional distribution

- Distribution of random initial state X_0

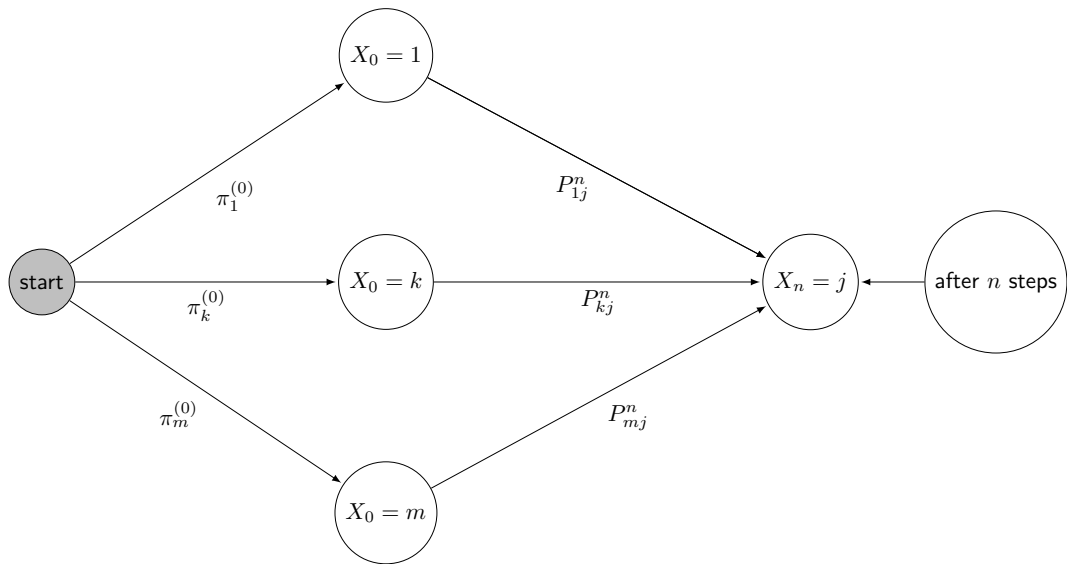
$$\pi^{(0)}(i) = P(X_0 = i)$$

- Distribution of X_n

$$\pi^{(n)}(i) = P(X_n = i)$$

Information about state X_n of Markov chain after n steps when you don't know the starting point of the process at initial time 0





Unconditional distribution of X_n

$$\pi^{(n)} = \pi^{(0)} P^{(n)}$$

where

$$\pi^{(0)} = \left(\pi_1^{(0)} \quad \pi_2^{(0)} \quad \dots \quad \pi_m^{(0)} \right)$$

and

$$\pi^{(n)} = \left(\pi_1^{(n)} \quad \pi_2^{(n)} \quad \dots \quad \pi_m^{(n)} \right)$$

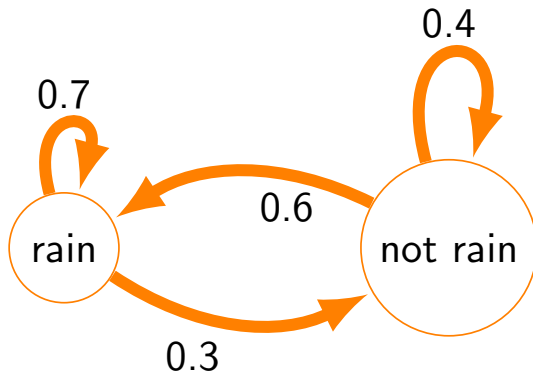


Thanks to Total rule probability

$$\begin{aligned}P(X_n = j) &= \sum_i P(X_n = j | X_0 = i) P(X_0 = i) \\&= \sum_i P_{ij}^n P(X_0 = i) \\&= \sum_i P(X_0 = i) P_{ij}^n \\&= \sum_i \pi_i^{(0)} P_{ij}^n\end{aligned}$$



Example - Weather forecast



Suppose probability rain today is .4, what is the probability that it will rain 4 days from now



- State: 1 = rain, 2 = not rain
- Initial probability for weather today

$$\pi^{(0)} = [0.4 \quad 0.6]$$

- Transition matrix

$$P = \begin{bmatrix} 0.7 & 0.3 \\ 0.6 & 0.4 \end{bmatrix}$$



- Distribution for weather 4 days from now

$$\begin{aligned}\pi^{(4)} &= \pi^{(0)} P^{(4)} = [0.4 \quad 0.6] \begin{bmatrix} 0.6667 & 0.3333 \\ 0.6666 & 0.3334 \end{bmatrix} \\ &= [.6666 \quad .3334]\end{aligned}$$

- Probability that it will rain 4 days from now

$$P(X_4 = 1) = \pi^{(4)}(1) = 0.6666$$

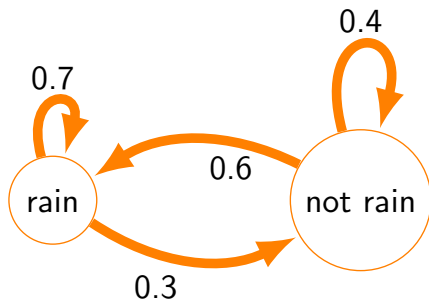


Long term behavior of Markov chain

- Does $r_{ij}(n)$ converge to something?
- Does the limit depend on initial state?

Applications: Google Page's rank problem ...

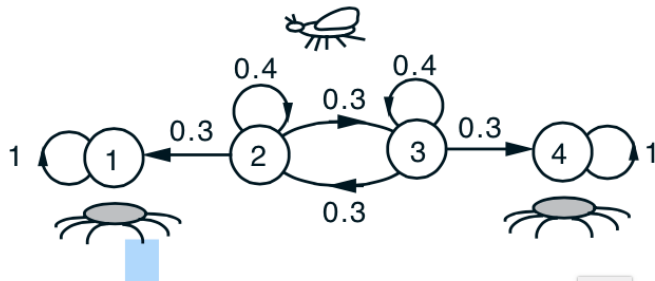




$$r_{ij}(1) = P = \begin{bmatrix} .7 & .3 \\ .6 & .4 \end{bmatrix}, \quad r_{ij}^{(\infty)} = \begin{bmatrix} .6667 & .3333 \\ .6667 & .3333 \end{bmatrix}$$

In long term, it will rain with probability .67 whatever the weather today is





CLOSE

	1	2	3	4
1	1.0	0	0	0
2	0.3	0.4	0.3	0
3	0	0.3	0.4	0.3
4	0	0	0	1.0

$r_{ij}(1)$

...

	1	2	3	4
1	1.0	0	0	0
2	2/3	0	0	1/3
3	1/3	0	0	2/3
4	0	0	0	1.0

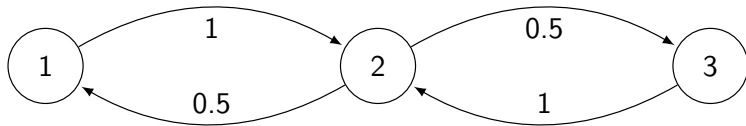
$r_{ij}(\infty)$

After a lot of transition, the fly is at position 4 with probability

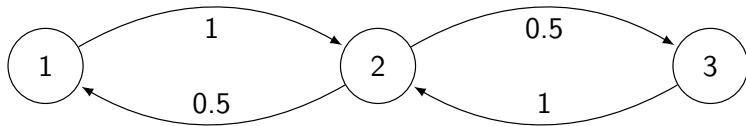
- $1/3$ if it starts at position 2
- $2/3$ if it starts at state 3
- 0 if it starts at other state

Probability that the fly is at position j after long time depends on initial state





- n odd then $r_{22}^{(n)} = 0$
- n even then $r_{22}^{(n)} = 1$
- $r_{ij}^{(n)}$ diverges



- n odd then $r_{22}^{(n)} = 0$
- n even then $r_{22}^{(n)} = 1$
- $r_{ij}^{(n)}$ diverges

Does $r_{ij}^{(n)}$ converge to π_j which is independent of the initial state i ?

- 1 Under which condition?
- 2 How to find π_j if it exists?

Answer for question 1

If the Markov chain has the following properties

- recurrent states are all in a single class
- single recurrent class is not periodic

then the limit of $r_{ij}^{(n)}$ exists and independent of initial state



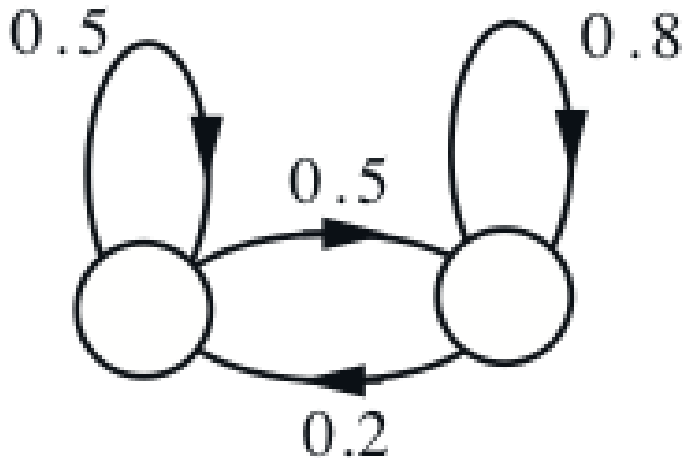
Classification of states

Accessible and communicate

- State j is accessible from state i if $r_{ij}^{(n)} > 0$ for some $n \geq 0$
- Two states that are accessible from each other are said to *communicate*
- If i communicates with j and j communicates with k then i communicates with k .
- Markov chain is *irreducible* if all states communicate with each other.

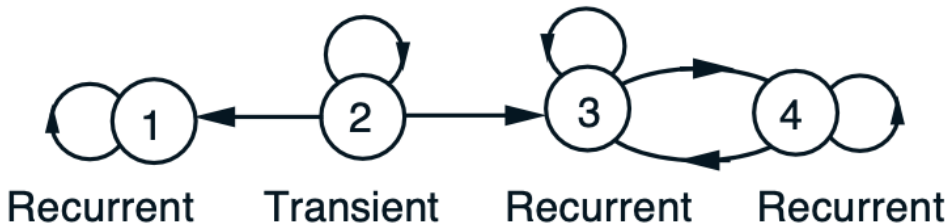


Example



Recurrent and Transient State

- State i is **recurrent** if: starting from i , and from wherever you can go, there is a way of returning to i
- If not recurrent, called **transient**

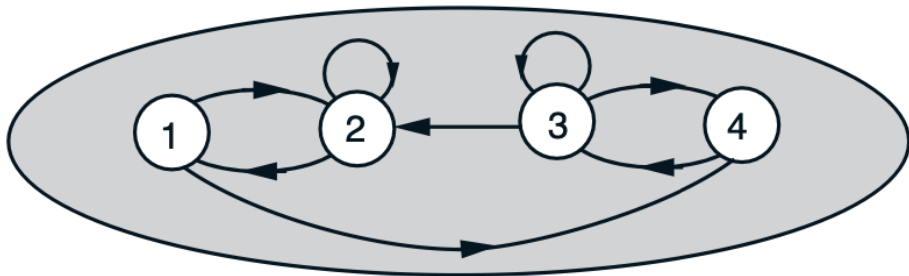


- If a recurrent state is visited once, it will be visited infinitely numbers of time
- a transient state will only be visited a finite number of times.

Reccurent Class

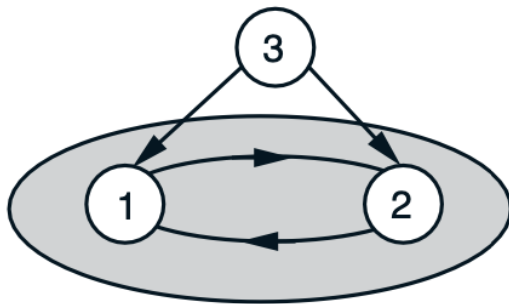
collection of recurrent states that “communicate” to each other and to no other state

Example



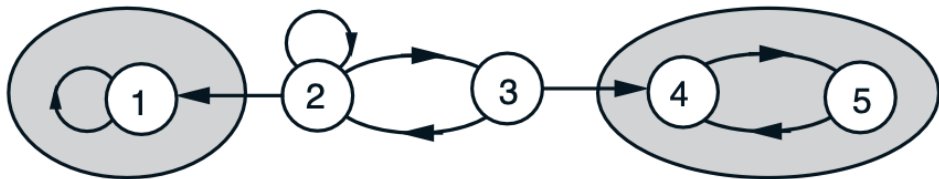
Single class of recurrent states

Example



Single class of recurrent states (1 and 2)
and one transient state (3)

Example

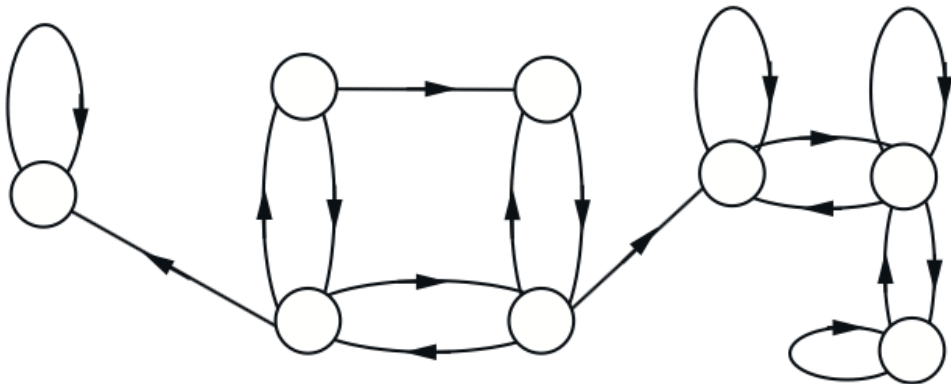


Two classes of recurrent states
(class of state 1 and class of states 4 and 5)
and two transient states (2 and 3)



Practice

Determine classes of recurrent states of the Markov chain



Markov chain decomposition

- Transient states
- Recurrent classes

- once the state enters (or starts in) a class of recurrent states, it stays within that class; since all states in the class are accessible from each other, all states in the class will be visited an infinite number of times;
- if the initial state is transient, then the state trajectory contains an initial portion consisting of transient states and a final portion consisting of recurrent states from the same class



Analyze long - term behavior

- The Markov chain stays forever at a recurrent class that it visits first
- Need to analyze chains that consist of a single recurrent class

Analyze long - term behavior

- The Markov chain stays forever at a recurrent class that it visits first
- Need to analyze chains that consist of a single recurrent class

Periodicity

Consider a recurrent class \mathcal{R}

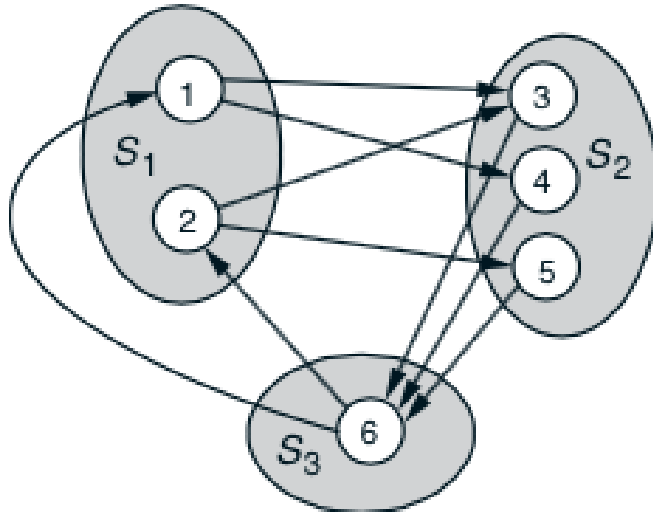
- 1 \mathcal{R} is said to be **periodic** if its states can be grouped in $d > 1$ disjoint subsets S_1, \dots, S_d so that all transitions from one subset lead to the next subset

$$\text{If } i \in S_k \text{ and } p_{ij} > 0 \text{ then } \begin{cases} j \in S_{k+1} & \text{if } k \leq d-1 \\ j \in S_1 & \text{if } k = d \end{cases}$$

- 2 \mathcal{R} is aperiodic if not periodic, i.e there exist a state s and a number n such that $r_{is}(n) > 0$ for all $i \in \mathcal{R}$

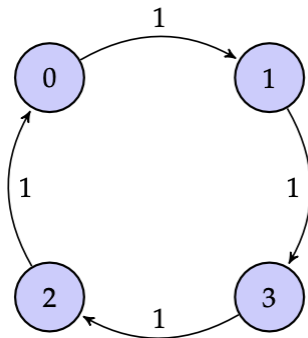


Structure of a periodic recurrent class



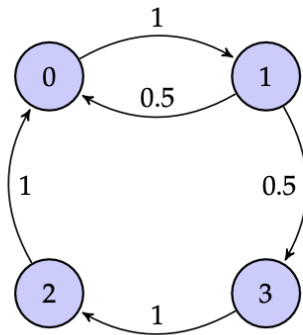
Example

All states have period 4



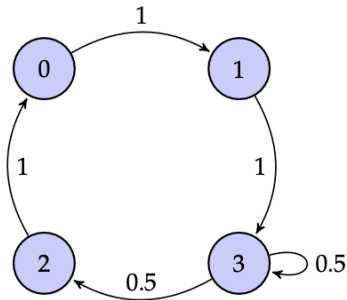
Example

All states have period 2



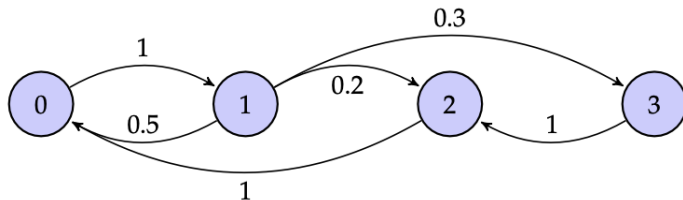
Example

All states have period 1



Example

All states have period 1



- a periodic recurrent class, a positive time n , and a state j in the class, there must exist some state i such that $r_{ij}^{(n)} = 0$ because the subset to which j belongs can be reached at time n from the states in only one of the subsets.
- thus a way to verify aperiodicity of a given recurrent class \mathcal{R} , is to check whether there is a special time $n \geq 1$ and a special state $s \in \mathcal{R}$ that can be reached at time n from all initial states in \mathcal{R} , i.e., $r_{is}^{(n)} > 0$ for all $i \in \mathcal{R}$

- a periodic recurrent class, a positive time n , and a state j in the class, there must exist some state i such that $r_{ij}^{(n)} = 0$ because the subset to which j belongs can be reached at time n from the states in only one of the subsets.
- thus a way to verify aperiodicity of a given recurrent class \mathcal{R} , is to check whether there is a special time $n \geq 1$ and a special state $s \in \mathcal{R}$ that can be reached at time n from all initial states in \mathcal{R} , i.e., $r_{is}^{(n)} > 0$ for all $i \in \mathcal{R}$

Theorem

Let $\{X_n\}$ be a Markov chain with a single recurrent class and aperiodic. The steady-state probability π_j associated with the state j satisfies the following properties

1

$$\lim_{n \rightarrow \infty} P_{ij}^{(n)} = \pi_j$$

2 π_j are the unique nonnegative solution of the **balance equation**

$$\pi_j = \sum_{i=1}^{\infty} \pi_i p_{ij}, \quad \sum_{j=1}^{\infty} \pi_j = 1$$

$\{\pi_j\}$ is called the **stationary distribution**



Answer for question 2

- Start from key recursion $r_{ij}^{(n)} = \sum_k r_{ik}^{(n-1)} p_{kj}$
- let $n \rightarrow \infty$

$$\pi_j = \sum_k \pi_k p_{kj} \text{ for all } j$$

- Addition equation $\sum_j \pi_j = 1$
- (π_j) is called the **stationary distribution** of the Markov chain



Interpretation

After some steps, the distribution of X_n is approximately $\{\pi_j\}$ and will not change much

$$P(X_n = j) \approx \pi_j \text{ for } n \text{ large enough}$$

π_j : steady - state probability

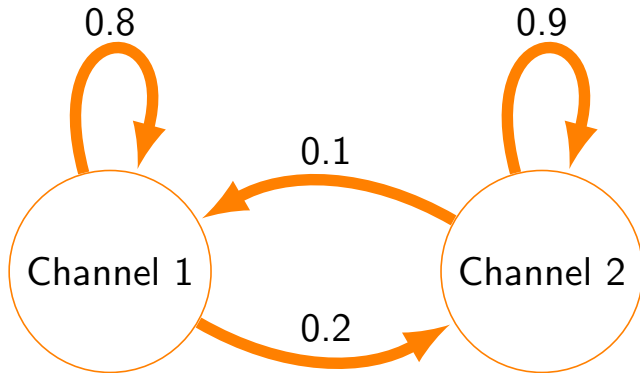


Find stationary distribution

Solve

$$\begin{cases} \pi P = \pi \\ \sum \pi_i = 1 \end{cases}$$

Example



What will be the market share after a long time?



Solution

- Transition matrix $P = \begin{bmatrix} .8 & .2 \\ .1 & .9 \end{bmatrix}$
- Stationary distribution $\pi = [\pi_1 \ \pi_2]$ satisfies

$$\begin{cases} \pi P = \pi \\ \pi_1 + \pi_2 = 1 \end{cases} \quad \text{or} \quad \begin{cases} .8\pi_1 + .1\pi_2 = \pi_1 \\ .2\pi_1 + .9\pi_2 = \pi_2 \\ \pi_1 + \pi_2 = 1 \end{cases}$$

- Result $\pi_1 = 1/3, \pi_2 = 2/3$



After a long time, the market is stable. Each year, there is about

- 33% of customers watch channel 1
- 67% of customers watch channel 2

Find stationary distribution of the Markov chain with transition probability

$$P = \begin{bmatrix} 0.8 & 0.2 \\ 0.6 & 0.4 \end{bmatrix}$$