Statistical analysis of EEG signals

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1. Introduction

In this paper, the analysis and statistical modelling of two electroencephalogram (EEG) signals is reported and its results interpreted and discussed. We are working with two EEG signals, X and Y, that were measured brain activity of two regions. The goal of this analysis is to analyze the relationship between these signals and identify good model structure that can predict one signal from the other.

For the modelling purposes we refer to the signal X as an input (independent variable) and signal Y as output signal (dependent variable). In other words, signal Y is predicted by signal X.

First an exploratory data analysis was performed in
two approaches were used to identify the best model structure; forward subset selection algorithm and AIC
model selection. After the best model structure is identified, we validate the model
. Then we also quantify the uncertainty around both the model parameters and its
predictions. Finally, rejection approximate Bayesian computation (ABC) is applied to estimate the posterior
distribution of the model parameters.

The whole analysis was performed using R 3.6.1 Action of the Toes (ref).

1.1 Data

We are working with two separate EEG signals. The data sample consists of 250 data points. We assume that both signals are independent. Furthermore, we assume that the signals contain additive gaussian noise with zero mean and unknow variance.

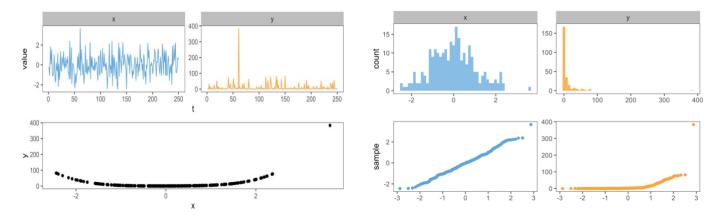


Figure 2 Top: Time series plots of signal X and Y.

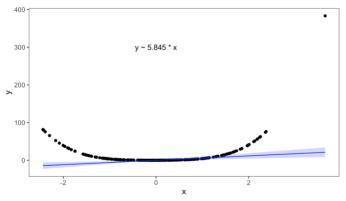
Figure 2 Top: Histograms of signal X and Y. Bottom: QQ-plots of signal X and Y

2. Exploratory data analysis

The aim of this initial exploratory analysis is to gain	
We begin with time series plots of the signals, i.e. plot each signal on y-axis and tir	me (order in which data
were recorded) on x-axis (figure 1, top).	
	-
Next, we plot the signals against each other in order to observe the relationship b	etween them (figure 1,
bottom). We can assume that the relationship between X and Y is	
. The best model structure is likely	
Next, we inspect the distribution of the signals. The histograms of both signals are sl	howed in figure 2 (top).
We also create QQ-plots to see (figure 2, b)	oottom). The QQ-plot is
	efore, if the data follows
normal distribution then the QQ plot will show	
Based on the histogram we might say that signal X is	
. To test whether signal is normally distributed we used Shapiro-W	Vilk test
	Results of this test
suggest that signal X is	(W = 0.993, p-value >
0.05).	
From the histogram of signal Y, we can clearly see that it is	Based on the shape it
might be following	
The QQ-plot confirms that signal Y is	
2.1 Linear model	

From the scatterplot we drew conclusion that the relationship between X and Y is probably However, to explore all options we fitted a simple linear model (1).

$$y \sim \beta_1 x$$



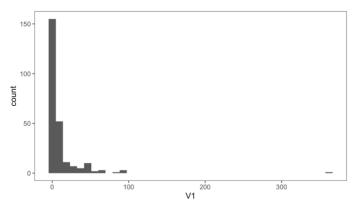


Figure 4 Linear model $y \sim x$, estimated parameters and predictions (blue) with 95% confidence interval (blue ribbon) plotted together with true y values (black points)

Figure 4 Distribution of residuals of the linear model $y \sim x$

To assess the performance of the fitted model we use mean squared error (MSE) which is calculated using formula given by (2). In the formula n is the number of data points, Y is the true Y value and the \hat{Y} is the model prediction. We also use adjusted R₂ defined in formula (3).

Adjusted R₂ is

$$MSE = \frac{1}{n} \sum_{i=1}^{n} (Y_i - \hat{Y}_i)^2$$

$$R_{adjusted}^2 = 1 - \frac{SS_{residual}/df_e}{SS_{total}/df_t}$$

(3)

(2)

Two measurements of the model performance are used in order to

We use least squares to estimate the parameter of the model.

The results of this model are showed in figure 3. The MSE of the fitted model is 886.242. We also generated predictions and 95% confidence intervals. Next residual analysis was performed. In ideal case we would expect the residuals to be ... However, residuals of this model are

(figure 4).

We can conclude that linear model structure is rather fit as

Furthermore, from the fitted line (figure 3) we can see that the model captures of the variance.

3. Model selection

To identify the best model structure, two approaches were used. This was done to in order

First approach is the forward subset selection and second model selection using Akaike information criterion (AIC).

Table 1: Results of forward subset selection

Model structure	Testing MSE	
β_1 x4	5.232	
$\beta_1 x_2 + \beta_2 x_4$	0.308	
$\beta_1 \mathbf{x} + \beta_2 \mathbf{x} 2 + \beta_3 \mathbf{x} 4$	0.008	

Table 2: Parameter covariance matrix

	$oldsymbol{eta_1}$	$oldsymbol{eta}_2$	β_3
$\boldsymbol{\beta_1}$	5.1e-05	1.0e-05	-2e-06
$oldsymbol{eta}_2$	1.0e-05	4.5e-05	-5e-06
β_3	-2.0e-06	-5.0e-06	1e-06

3.1 Forward subset selection

Forward subset selection is iterative approach for finding the best model structure. First, the data is split to training and testing sets. We used 80% of data as training.

The training set is used to estimate the model parameters using least squares. Predictions are then generated only for the testing set and performance metric is calculated using these predictions. Performance metric is MSE (formula 2).

Then list of all possible terms is created. In our case this is X raised to power of up to 5 and intercept term. Then all single term models are estimated and their MSE computed. The model that yields the lowest MSE is selected, the parameter stored in the final model and removed from the list of possible terms. Next all two term models are fitted using the parameter of the final model and the remaining terms in the term list. Again, model with lowest MSE is selected and used as the final model. This process is repeated until satisfactory MSE is achieved or all terms are in the final model.

In our analysis, we limited the number of model terms to 3. The results of each iteration of the subset selection is showed in table 1.

3.2 AIC selection

AIC is intended to give an estimate of The formula for calculating AIC is given in (4). It is based on the model likelihood and penalty for number of parameters.

$$AIC = 2k - 2 \ln (L_{model})$$

(4)

This approach does not require splitting the data, instead the models are fitted using all available data. Then we simply fit all possible combinations of terms and compute AIC for each of these models. The model that provides the lowest AIC is selected.

The best model structure according to AIC is:

$$y = \beta_1 x + \beta_2 x^2 + \beta_3 x^4 + \varepsilon$$
 (AIC = -432.326)

3.3 Evaluation of model selection

Both approaches converged on the same model structure.

4. Model evaluation

The selected model which yields the lowest MSE was then evaluated. The estimated model using the training set is:

$$y = 0.499x + 2.004x^2 + 2x^4 + \varepsilon$$

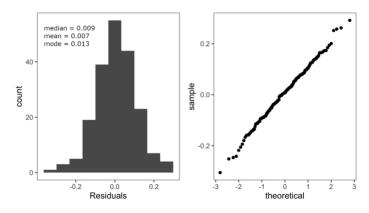


Figure 5 Residuals of the selected model with lowest testing MSE. Right: Histogram of residuals and central descriptive statistics. Left: QQ-plot of residuals

First, we inspect the residuals, especially their distribution. Histogram and QQ-plot are used to investigate the residual distribution (figure 4). The residuals follow approximately

Next, the covariance matrix of the model parameters was calculated (table 2). The covariance of all parameters

The parameter estimate uncertainty probability density function was calculated for each combination of two parameters and plotted in 3D with the distribution contours superimposed, resulting in 3 plots (figure 6).

4.1 Prediction on training data

Predictions of the data used in fitting the model were computed. Furthermore, 95% confidence intervals for each data point were calculated. The predictions were plotted with confidence intervals and the true values in figure 7 (left). For better understanding a zoomed-in view of the same data is showed figure 7 (right).

As expected the predictions are very close to the true values and the confidence intervals are very narrow. This is caused by

A better approach would be

4.2 Model validation