

3.66 Let  $X \sim \mathcal{N}(8, 3)$ ,  $Z \sim \mathcal{N}(1, 0)$ . Therefore

$$\begin{aligned}
 \mathbb{P}(X > \alpha) &= 0.15 \\
 1 - \mathbb{P}(X \leq \alpha) &= 0.15 \\
 \mathbb{P}(X \leq \alpha) &= 0.85 \\
 \mathbb{P}(\sqrt{3}Z + 8 \leq \alpha) &= 0.85 \\
 \mathbb{P}(Z \leq \frac{\alpha - 8}{\sqrt{3}}) &= 0.85 \\
 \Phi(\frac{\alpha - 8}{\sqrt{3}}) &= 0.85 \\
 \frac{\alpha - 8}{\sqrt{3}} &= \Phi^{-1}(0.85) \\
 \frac{\alpha - 8}{\sqrt{3}} &\approx 1.04 \\
 \alpha &\approx 9.8
 \end{aligned}$$

3.67 Let  $X \sim \mathcal{N}(\mu, \sigma^2)$ ,  $Z \sim \mathcal{N}(1, 0)$

(a) Since  $x^3$  is odd and  $\phi(x)$  is even then  $x^3\phi(x)$  is odd. Thus the integral across all of  $\mathbb{R}$  is 0.

(b)

$$\begin{aligned}
 \mathbb{E}[X^3] &= \mathbb{E}[(\sigma Z + \mu)^3] \\
 &= \sigma^3 \mathbb{E}[Z^3] + 3\mu\sigma^2 \mathbb{E}[Z^2] + 3\mu^2\sigma \mathbb{E}[Z] + \mu^3 \mathbb{E}[1] \\
 &= \sigma^3 * 0 + 3\mu\sigma^2(1 + 0^2) + 3\mu^2\sigma * 0 + \mu^3 * 1 \\
 &= \mu\sigma^2 + \mu^3
 \end{aligned}$$

3.68 (a)

$$\begin{aligned}
 \int_{-\infty}^{\infty} x^4 e^{-x^2/2} dx &= -x^3 e^{-x^2/2} \Big|_{-\infty}^{\infty} + 3 \int_{-\infty}^{\infty} x^2 e^{-x^2/2} dx \\
 &= 3(-x e^{-x^2/2}) \Big|_{-\infty}^{\infty} + \int_{-\infty}^{\infty} e^{-x^2/2} dx \\
 &= 3\sqrt{2\pi}
 \end{aligned}$$

Thus  $\mathbb{E}[Z^4] = 3$ .

(b) Same  $X$  and  $Z$  as stated in the previous problem

$$\begin{aligned}
 \mathbb{E}[X^4] &= \mathbb{E}[(\sigma Z + \mu)^4] \\
 &= \sigma^4 \mathbb{E}[Z^4] + 4\sigma^3\mu \mathbb{E}[Z^3] + 6\sigma^2\mu^2 \mathbb{E}[Z^2] + \sigma\mu^3 \mathbb{E}[Z] + \mu^4 \\
 &= 3\sigma^4 + 0 + 6\sigma^2\mu^2 + 0 + \mu^4 \\
 &= 3\sigma^4 + 6\sigma^2\mu^2 + \mu^4
 \end{aligned}$$

4.4 Let  $X_{90}$  be the total steps taken after 90 rolls. Let  $S_{90}$  be the number of rolls yielding 1 step forward. Then our desired probability can be written as  $\mathbb{P}(X_{90} \geq 160) = \mathbb{P}(S_{90} \leq 20)$ . Since each roll has probability  $1/3$  of hitting a number which advances us one tile, then clearly  $S_{90} \sim \text{Bin}(90, 1/3)$ . Therefore if we normalize we find that our probability becomes  $\mathbb{P}(\frac{S_{90}-30}{\sqrt{20}} \leq -\sqrt{5})$ . Thus by the central limit theorem this can be approximated by  $\Phi(-\sqrt{5}) = 1 - \Phi(\sqrt{5})$ . Thus the probability is approximately 0.013.

4.18 Since we can hit a point uniformly on the dart board, then the probability of hitting the center is simply dividing the areas yielding  $p = \frac{\pi 1^2}{\pi 5^2} = \frac{1}{25}$ . Therefore the variable  $H_{2000}$  denoting the number of bullseyes is a binomial variable  $H_{2000} \sim \text{Bin}(2000, \frac{1}{25})$  with  $\mathbb{E}H_{2000} = 80, \text{Var}(H_{2000}) = 76.8$ .

Therefore our desired probability is

$$\mathbb{P}(H_{2000} \geq 100) = 1 - \mathbb{P}(H_{2000} < 100) = 1 - \mathbb{P}\left(\frac{H_{2000} - 80}{\sqrt{76.8}} < \frac{20}{\sqrt{76.8}}\right) \approx 1 - \Phi(2.282) \approx 0.0113$$

AE1  $\mathbb{E}[e^{cZ}] = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{cx} e^{-\frac{x^2}{2}} dx = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{cx - \frac{x^2}{2}} dx$ . Thus by homework 1 problem 1a we have that the integral evaluates to  $e^{-\frac{c^2}{2}}$

AE2 The only  $c$  which works is  $c < \frac{1}{2}$ , because if  $c \geq \frac{1}{2}$  then  $c - \frac{1}{2} \geq 0$ , and since for every  $|x| > 1$   $e^{(c-\frac{1}{2})|x|} < e^{(c-\frac{1}{2})x^2}$ , then for a sufficiently large  $a$ ,  $\int_{-a}^a e^{(c-\frac{1}{2})|x|} dx < \int_{-a}^a e^{(c-\frac{1}{2})x^2} dx$ . Since  $\int_{-\infty}^{\infty} e^{(c-\frac{1}{2})|x|} dx = \infty$ , then  $\int_{-\infty}^{\infty} e^{(c-\frac{1}{2})x^2} = \infty$ . (Note if  $c = \frac{1}{2}$  then we have the integral  $\int_{-\infty}^{\infty} 1 dx$ , which clearly diverges.)

AE3 (a) Let  $X = \sigma Z + \mu$  where  $Z \sim \mathcal{N}(0, 1)$

$$\begin{aligned} \mathbb{P}(Y \geq K) &= \mathbb{P}(e^X \geq K) \\ &= \mathbb{P}(X \geq \log(K)) \\ &= \mathbb{P}\left(Z \geq \frac{\log(K) - \mu}{\sigma}\right) \\ &= 1 - \mathbb{P}\left(Z \leq \frac{\log(K) - \mu}{\sigma}\right) \\ &= 1 - \Phi\left(\frac{\log(K) - \mu}{\sigma}\right) \\ &= \Phi\left(\frac{\mu - \log(K)}{\sigma}\right) \end{aligned}$$

(b)

$$\begin{aligned}
\mathbb{E}[\max(Y - K, 0)] &= \mathbb{E}[(Y - K)\mathbb{I}_{X \geq \log(K)}] \\
&= \mathbb{E}[Y\mathbb{I}_{X \geq \log(K)}] - \mathbb{E}[K\mathbb{I}_{X \geq \log(K)}] \\
&= \mathbb{E}[Y\mathbb{I}_{X \geq \log(K)}] - K\mathbb{E}[\mathbb{I}_{X \geq \log(K)}] \\
&= \mathbb{E}[Y\mathbb{I}_{X \geq \log(K)}] - K\Phi\left(\frac{\mu - \log(K)}{\sigma}\right)
\end{aligned}$$

by the hint

$$\begin{aligned}
&= \frac{1}{\sqrt{2\pi}\sigma^2} \int_{\log(K)}^{\infty} e^r e^{\frac{(r-\mu)^2}{2\sigma^2}} dr - K\Phi\left(\frac{\mu - \log(K)}{\sigma}\right) \\
&= \frac{e^\mu}{\sqrt{2\pi}} \int_{\frac{\log(K)-\mu}{\sigma}}^{\infty} e^{\sigma l - \frac{l^2}{2}} dl - K\Phi\left(\frac{\mu - \log(K)}{\sigma}\right)
\end{aligned}$$

$$\begin{aligned}
\text{Let } l &= \frac{r - \mu}{\sigma}, dl = \frac{dr}{\sigma} \\
&= \frac{e^\mu}{\sqrt{2\pi}} \left( \sqrt{2\pi} e^{\frac{\sigma^2}{2}} - \int_{-\infty}^{\frac{\log(K)-\mu}{\sigma}} e^{\sigma l - \frac{l^2}{2}} dl \right) - K\Phi\left(\frac{\mu - \log(K)}{\sigma}\right)
\end{aligned}$$

By 1a on hw 1

$$= \frac{e^\mu}{\sqrt{2\pi}} \left( \sqrt{2\pi} e^{\frac{\sigma^2}{2}} - \int_{-\infty}^{\frac{\log(K)-\mu}{\sigma}} e^{-\frac{(l-\sigma)^2}{2} + \frac{\sigma^2}{2}} dl \right) - K\Phi\left(\frac{\mu - \log(K)}{\sigma}\right)$$

completing the square

$$= \frac{e^{\mu + \frac{\sigma^2}{2}}}{\sqrt{2\pi}} \left( \sqrt{2\pi} - \int_{-\infty}^{\frac{\log(K)-\mu}{\sigma} - \sigma} e^{-\frac{u^2}{2}} du \right) - K\Phi\left(\frac{\mu - \log(K)}{\sigma}\right)$$

Let  $u = l - \sigma, du = dl$ 

$$= e^{\mu + \frac{\sigma^2}{2}} \left( 1 - \Phi\left(\frac{\log(K) - \mu}{\sigma} - \sigma\right) \right) - K\Phi\left(\frac{\mu - \log(K)}{\sigma}\right)$$

Definition of  $\Phi$ 

$$= e^{\mu + \frac{\sigma^2}{2}} \left( \Phi\left(\frac{\mu - \log(K)}{\sigma} + \sigma\right) \right) - K\Phi\left(\frac{\mu - \log(K)}{\sigma}\right)$$