# Machine Learning for Survival Analysis: A New Approach

## D. Dooling P. Green A. Kim D. Scroggin L. Stevens J. Webster

Innovative Oncology Business Solutions, 4901 Lang Ave NE, Albuquerque, NM 87109, USA

E-mail: ddooling@innovativeobs.com, pgreen@innovativeobs.com, akim@innovativeobs.com, dscroggin@innovativeobs.com, lstevens@innovativeobs.com

ABSTRACT: We have applied a little-known data transformation on subsets of the Surveillance, Epidemiology, and End Results (SEER) publically available data of the National Cancer Institute (NCI) to make it suitable input to standard machine learning classifiers. This transformation properly treats the right-censored data in the SEER data and the resulting Random Forest and Multi-Layer Perceptron models predict full survival curves. Treating the 6, 12, and 60 months points of the resulting survival curves as 3 binary classifiers, the 18 resulting classifiers have AUC values ranging from .765 to .885. Further evidence that the models have generalized well from the training data is provided by the extremely high levels of agreement between the random forest and neural network models predictions on the 6, 12, and 60 month binary classifiers.

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## 1 Introduction and Background

Extracting actionable information from data is changing the fabric of modern business. A class of techniques that transforms data into actionable information goes by the name of Machine Learning [13]. Machine Learning has recently become a popular method to answer

questions and solve problems that are too complex to solve via traditional methods. The Surveillance, Epidemiolgy, and End Results (SEER) Program of the National Cancer Institute (NCI) has been collecting data because intuitively researchers feel confident that this data is capturing information that has buried within it useful information in the form of relationships between the types of data collected (demographic as well as staging information) and the survival outcomes. Though this relationship evades capture by traditional methods, it is possible to surface it with the two machine learning techniques known as **Random Forests** and **Neural Networks**. These two methods produce very similar results when applied to the SEER dataset, and are based on two almost diametrically opposed learning philosophies, which lends confidence in the validity of the results.

The Surveillance, Epidemiolgy, and End Results (SEER) Program of the National Cancer Institute (NCI) is the most recognized authoritative source of information on cancer incidence and survival in the United States. SEER currently collects and publishes cancer incidence and survival data from population-based cancer registries covering approximately 28 percent of the US population.

Quoting directly from the SEER website [11]:

The SEER program registries routinely collect data on patient demographics, primary tumor site, tumor morphology and stage at diagnosis, first course of treatment, and follow-up for vital status. This program is the only comprehensive source of population-based information in the United States that includes stage of cancer at the time of diagnosis and patient survival data. The mortality data reported by SEER are provided by the National Center for Health Statistics. The population data used in calculating cancer rates is obtained periodically from the Census Bureau. Updated annually and provided as a public service in print and electronic formats, SEER data are used by thousands of researchers, clinicians, public health officials, legislators, policymakers, community groups, and the public.

One characterstic of the SEER data and that is shared by many datasets in the medical field goes by the name of "censored data." The SEER data contains the number of months each patient survived, as well as an indicator variable showing whether or not the patient is still alive at the end of the data collection period. Methods to deal effectively with this kind of "right-censored data" include Kaplan-Meier curves and Cox's Proportional Hazard models [4]. The Kaplan-Meier techniques only give estimates for cohorts of patients and are not applicable for predicting the surival curve for a single patient, and the Cox Proportional Hazard models require a fairly restrictive set of assumptions to be satisifed in order to yield reliable results. In addition, the Cox Proportional Hazard models are not able to capture the nonlinear relationships between the given data fields that go into making predictions; they can only capture the first-order linear relationships.

Previous work applying machine learning methods to subsets of the SEER data include creative attempts to deal with the problems presented by "right-censored data." The authors of [14] use semi-supervised learning techniques to predict 5 year survival, essentially

imputing values for SEER records where the survival months infomation is censored at a value less than 5 years. The authors of [15] investigate the effects of comordbidities; i.e., patients with two different cancer diagnosises, but their treatment of the censored data underestimates the survival probabilities. All records representing patients who survived at least 60 months as well as all those who died earlier than 60 months were considered, but patients alive prior to 60 months but censored out of the study before 60 months were not included. This treatment biases the data and the predictions, leading to overly pessimistic survival probabilities predicted by the trained models.

To overcome these limitations of the traditional methods, IOBS has applied a little-known technique to transform the SEER data to make it amenable to more powerful machine learning methods. The essential idea is to recast the problem to an appropriate discrete classification problem instead of a regression problem (predicting survival months). Treating months after diagnosis as just another discrete feature, the SEER data (or any other right-censored data) can be transformed simply so as to make predictions for the hazard function, probability of dying in the next month, given that the patient has not yet died. The full survival function can then be derived from the hazard function. Details of this transformation can be found in this blog post [1].

## 2 Methodology

#### 2.1 Data acquisition

We used the publically available 1973-2012 SEER incidence data files corresponding to colon, breast and lung cancer contained in the following list. SEER requires that researchers submit a request for the data, which includes an agreement form. Detailed documentation explaining the contents of both the incidence data files used in this study as well as a data dictionary for the 1973-2012 SEER incidence data files are available without the need to register or submit a data request [12].

- incidence\yr1973 2012.seer9\COLRECT.txt
- incidence\yr1973 2012.seer9\BREAST.txt
- incidence $\yr1973\_2012.seer9\RESPIR.txt$
- incidence\yr1992 2012.sj la rg ak\COLRECT.txt
- incidence $\yr1992\_2012.sj_la\_rg\_ak\BREAST.txt$
- incidence\yr2000 2012.ca ky lo nj ga\COLRECT.txt
- incidence\yr2000\_2012.ca\_ky\_lo\_nj\_ga\RESPIR.txt
- incidence\yr2005.lo 2nd half\COLRECT.txt
- incidence\yr2005.lo 2nd half\BREAST.txt
- incidence\yr2005.lo 2nd half\RESPIR.txt

#### 2.2 Data preparation and preprocessing

A great deal of data munging is necessary before using these SEER incidence files as input into machine learning algorithms. A preprocessing step common to each of three cancer

Code	Description
1	Male
2	Female

**Table 1**. Encoding of gender in the SEER incidence files. These types of categorical variables need to be transformed via one-hot-encoding.

types studied involves the STATE-COUNTY RECODE. The STATE-COUNTY RECODE field is a state-county combination where the first two characters represent the state FIPS code and the last three digits represent the FIPS county code. This particular field illustrates an important feature of machine learning, that between categorical features and numeric features. All input into a machine learning algorithm must be numeric, but real numbers carry with them the usually extremely useful property known as the well-ordering property of the real numbers. But if one is tasked with encoding a categorical feature into suitable numeric format for machine learning, it is necessary to do so in a way that removes the well-ordering property [2].

As a simple example of how to correctly treat categorical variables in a mchine learning context, consider the SEER variable SEX. This variable is encoded with a numeric 1 for males and a numeric 2 for females as shown in Table 1. Values such as "Male" and "Female" encoded as numbers are dangerous because if not handled properly, they can generate bogus results [7]. The proper way to transform the SEER SEX variable is to create two additional variables: sex\_Male and sex\_Female, and then to eliminate the variable SEX. For example,

$$\frac{\text{Sex}}{1} \longrightarrow \frac{\text{sex\_Male sex\_Female}}{1} \qquad (2.1)$$

and

$$\frac{\texttt{Sex}}{2} \longrightarrow \frac{\texttt{sex\_Male} \quad \texttt{sex\_Female}}{0}$$
 (2.2)

The procedure outlined in Equations (2.1, 2.2) needs to be applied to all of the nominal categorical variables in the SEER data that we wish to include in our predictive models. In particular, in order to include the geophgraphical information contained in the SEER categorical variable STATE-COUNTY RECODE, it becomes necessary to create a new feature variable for each of the distinct (state,county) pairs in the data. In the United States, there are approximately 3,000 counties. Clearly, transforming the STATE-COUNTY RECODE data representation into distinct (state\_county) columns will explode the data to become wider than is optimal for machine learning. Adding extra columns to your dataset, making it wider, requires more data rows (making it taller) in order for machine learning algorithms to effectively learn [2]. Because one-hot coding STATE-COUNTY RECODE would cause such

STATE-COUNTY RECODE	address	elevation	lat	lng
35001	Bernalillo+county+NM	5207.579772	35.017785	-106.629130
35003	${\tt Catron+county+NM}$	8089.242628	34.151517	-108.427605
35005	${\tt Chaves+county+NM}$	3559.931671	33.475739	-104.472330
35006	${\it Cibola+county+NM}$	6443.415570	35.094756	-107.858387
35007	$_{\rm Colfax+county+NM}$	6147.749089	36.579976	-104.472330

**Table 2**. Example of the transformation of STATE-COUNTY RECODE to elevation, lat, and lng.

drastic shape changes in our data, we wish to avoid doing so. Fortunately, this variable, though given as a categorical variable, is actually an ordinal variable. There is an ordering among the (state\_county) columns, name longitude, latitude, and elevation. We can transform the data in STATE-COUNTY RECODE into three new numerical columns: lat, lng, and elevation.

For example, Table (2) shows how five entries of STATE-COUNTY RECODE corresponding to counties within New Mexico would can be represented by the elevation, lat, and lng features.

It is a simple exercise to construct the full lookup table from the SEER STATE-COUNTY RECODE variable to the corresponding three values elevation, lat, and lng. Using the publically available dafafile from the United States Census Bureau [3] to construct query strings like the values of the address field in Table (2), it is possible to then programmatically query the Google Maps Geocoding API for the latitude and longitude [6], and the Google Maps Elevation API for the corresponding elevation [5]. An added benefit of this shift from the single categorical variable STATE-COUNTY RECODE to the three continuous numerical variables lat, lng, and elevation is that input into the web applications described later are not restricted to the states and counties covereed in the SEER registries. The full lookup table analogous to Table 2 is available from a GitHub repository containing supplemental information for this study [8].

#### 2.3 Colon Cancer Data

In this section we describe the data processing steps that were specific to the colon cancer model development. The four COLRECT.txt files were imported into a pandas DataFrame object. This data was then filtered according to the conditions in Table 3.

The following categorical features were one-hot encoded as described in section 2.2:

- SEX,
- MARITAL STATUS AT DX,
- RACE/ETHNICITY,
- SPANISH/HISPANIC ORIGIN,
- GRADE ,

Column	Filter
SEQUENCE NUMBER-CENTRAL	eq "Unspecified"
AGE AT DIAGNOSIS	eq "Unknown age"
BIRTHDATE-YEAR	eq "Unknown year of birth"
YEAR OF DIAGNOSIS	$\geq 2004$
SURVIVAL MONTHS FLAG	= "1"
CS TUMOR SIZE EXT/EVAL	<i>≠</i> ""
CS TUMOR SIZE	$\neq 999$
SEER RECORD NUMBER	=1
PRIMARY SITE	= "LARGE INTESTINE, (EXCL. APPENDIX)"
SEQUENCE NUMBER-CENTRAL	=0

**Table 3**. Filters applied to the Colon Cancer data.

- PRIMARY SITE,
- LATERALITY,
- SEER HISTORIC STAGE A,
- HISTOLOGY RECODE-BROAD GROUPINGS,
- MONTH OF DIAGNOSIS,
- VITAL STATUS RECODE.

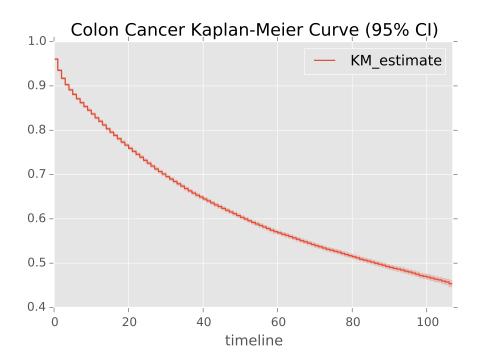
The STATE-COUNTY RECODE variable was dropped and replaced with the elevation, lat, and lng variables as illustrated in Table 2.

With just the above data preparation, it is possible to construct traditional Kaplan-Meier estimates of the survival curves for the colon cancer population represented by this subset of the data. After the above one-hot encoding procedure, the new variable vital\_status\_recode\_Dead indicates that the patient is deceased if this variable = 1, or else that the patient's record is right-censored if this variable = 0. SURVIVAL MONTHS and vital\_status\_recode\_Dead are all that is needed to construct the Kaplan-Meier estimate shown in Figure (1).

#### 2.4 Lung Cancer Data

In this section we describe the data processing steps that were specific to the lung cancer model development. The four RESPIR.txt files were imported into a pandas DataFrame object. This data was then filtered according to the conditions in Table 4. The same list of categorical features as in the colon cancer case were then one-hot encoded.

With just the above data preparation, it is possible to construct traditional Kaplan-Meier estimates of the survival curves for the colon cancer population represented by this subset of the data. After the above one-hot encoding procedure, the new variable vital\_status\_recode\_Dead indicates that the patient is deceased if this variable = 1, or else that the patient's record is right-censored if this variable = 0. SURVIVAL MONTHS and



**Figure 1**. Traditional Kaplan-Meier estimate of the survival curve for all colon cancer patients. Fitted with 113072 observations, 71804 censored.

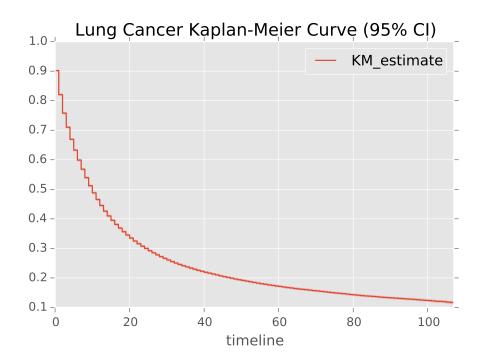
Column	Filter
SEQUENCE NUMBER-CENTRAL	$ \neq$ "Unspecified"
AGE AT DIAGNOSIS	eq "Unknown age"
BIRTHDATE-YEAR	eq "Unknown year of birth"
YEAR OF DIAGNOSIS	$\geq 2004$
SURVIVAL MONTHS FLAG	= "1"
CS TUMOR SIZE EXT/EVAL	≠ ""
CS TUMOR SIZE	$\neq 999$
SEER RECORD NUMBER	=1
PRIMARY SITE	= "LUNG & BRONCHUS"
SEQUENCE NUMBER-CENTRAL	=0

Table 4. Filters applied to the Lung Cancer data.

vital\_status\_recode\_Dead are all that is needed to construct the Kaplan-Meier estimate shown in Figure (2).

## 2.5 Breast Cancer Data

In this section we describe the data processing steps that were specific to the lung cancer model development. The four BREAST.txt files were imported into a pandas DataFrame



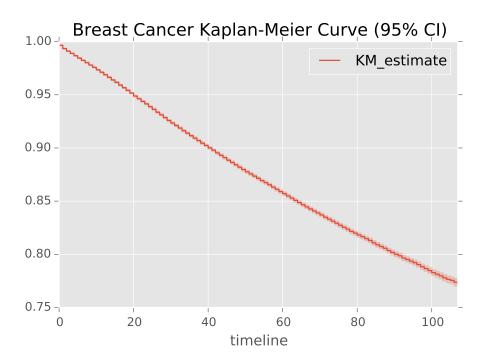
**Figure 2**. Traditional Kaplan-Meier estimate of the survival curve for all lung cancer patients. Fitted with 177089 observatins, 47409 censored.

Column	Filter
SEQUENCE NUMBER-CENTRAL	eq "Unspecified"
AGE AT DIAGNOSIS	eq "Unknown age"
BIRTHDATE-YEAR	eq "Unknown year of birth"
YEAR OF DIAGNOSIS	$\geq 2004$
SURVIVAL MONTHS FLAG	= "1"
CS TUMOR SIZE EXT/EVAL	≠ " "
CS TUMOR SIZE	$\neq 999$
SEER RECORD NUMBER	= 1
SEQUENCE NUMBER-CENTRAL	=0

Table 5. Filters applied to the Breast Cancer data.

object. This data was then filtered according to the conditions in Table 5. The same list of categorical features as in the colon cancer case were then one-hot encoded.

With just the above data preparation, it is possible to construct traditional Kaplan-Meier estimates of the survival curves for the colon cancer population represented by this subset of the data. After the above one-hot encoding procedure, the new variable vital\_status\_recode\_Dead indicates that the patient is deceased if this variable = 1, or



**Figure 3**. Traditional Kaplan-Meier estimate of the survival curve for all breast cancer patients. Fitted with 329949 observatins, 292279 censored.

else that the patient's record is right-censored if this variable = 0. SURVIVAL MONTHS and vital\_status\_recode\_Dead are all that is needed to construct the Kaplan-Meier estimate shown in Figure (3).

Before applying machine learning models trained with these data sets, we review in section (3) the sailent features of survival analysis and censored data. We then describe in detail a method that takes full advantage of all the data, including the right-censored data, and which involves a simple and intuitive transformation, culminating in the full set of features and target variables listed in sections (A.1, A.2, A.3).

## 3 Machine Learning Survival Analysis with Censored Data

The above Kaplan-Meier estimates of the survival curves for colon (Figure (1), lung (Figure (2), and breast cancer (Figure (3) are constructed from the full population of cancer patients in the respective datasets. An unsatisfactory consequence is that these estimates are highly course-grained, and not very meaningful to an indivual. Patients with very disparate characteristics are given the same prognosis by these Kaplan-Meier survival curve estimates. Therefore it is desirable to find robust predictors for survival curves of individual where the input is an individual record where the predictors are trained on larger populations.

#### 3.1 Survival Analysis

To understand survival analysis, you first have to understand survival data - that survival times are *intervals* between certain kinds of events, that these intervals are often affected by a peculiar kind of "partial missingness" called *censoring*, and that censored data must be analyzed in a special way to avoid biased estimates and incorrect conclusions.

In the case of the SEER data, the starting point of the time interval is the diagnosis date. Even though survival times are continuous or nearly continuous numerial quantities, they're never almost never normally distributed. If non-normality were the only problem with survival data, you would be able to summarize survival times as medians and centiles instead of means and standard deviations, and you could compare survival between groups with nonparametric Mann-Whitney and Kruksal-Wallis testse instead of t tests and ANOVAs. But time-to-event data is susceptible to a special situation called *censoring*, which the usual parametric and non-parametric methods cannot handle. Therefore special methods have been developed to analyze censored data properly.

With survival data, including the SEER data considered in this study, you may not know the exact time of death for some subjects. Some of the SEER subjects are still alive at the time of the latest SEER data release. When the VITAL STATUS RECODE variable indicates that the subject is still alive, the SURVIVAL MONTHS variable is only a lower bound on the true number of survival months; this is called the *date of last contact* mode of censoring. You know that each subject either died on a certain date or was definitely alive up to some last-seen date (and you don't know how far beyond that date he or she may ultimately have lived). The latter situation is called a *censored* observation.

Statisticians have developed some traditional techniques to utilize the partial information contained in censored observations: the life-table method and the Kaplan-Meier method. To understand these methods, you need to understand two fundamental concepts: - hazard and survival:

- The hazard rate is the probability of dying in the next small interval of time, assuming that the subject is alive right now.
- The survival rate is the probability of living for a certain amount of time after some starting point.

The first task when analyzing survival data is usually to describe how the hazard and survival rates vary with time. Here are two ways *not* to handle censored survival data:

- You shouldn't excude subjects with a censored surivival time from any survival analysis.
- You shouldn't *impute* (replace) the censored (last-seen) date with some reasonable substitute value. One commonly used imputation scheme is to replace a missing value with the last observed value for that subject (called *last observation carried forward*, or LOCF imputation).

These techniques for dealing with missing data don't work for censored data. If you simply exclude all subjects with censored death dates from your analysis, you may be left

	Survival Time (Years)	Censored Status
0	0.75	1
1	6.10	1
2	7.00	0
3	2.40	1
4	0.50	0
5	4.50	1
6	3.50	0
7	5.80	0
8	2.30	1
9	5.20	1

Table 6. Example data to illustate traditional Survival Analysis.

with too few analyzable subjects, which weakens (underpowers) youe study. Worse, it will also bias your results in subtle and unpredictable ways. The problem is that a censored observation time isn't really missing. If you know that a person was last seen alive three years after treatment, you have partial information for that patient. You don't know exactly what the patient's true survival time is, but you do know that it's at least three years.

To estimate survival and hazard rates in a population from a set of observed survival times, some of which are censored, you must combine the information from censored and uncensored observations properly. You have to think of the process in terms of a series of small slices of time, and think of the probability of making it through each time slice, assuming that the subject is alive at the start of that slice. The cumulative survival probability can then be obtained by successively multiplying all these individual time-slice survival probabities together. For example, to survive three years, first the subject has to make it through Year 1, then she has to make it through Year 2, and then she has to make it through Year 3. The probability of making it through all three years is the product of the probabilities of making it through Year 1, Year 2, and Year 3. These calculations can be laid out very systematically in a life table, sometimes called an actuarial life table because of its early use by insurance comparies. The calculations involce only addition, subtraction, multiplication, and division and are simple enough to do by hand.

To create a life table from your survival-time data, first break the entire range of survival times into convenient time slices (months, quarters, or years, depending on the time scale of the event you're studying). You should try to have at least five slices, otherwise, your survival and hazard estimates will be too coarse to show any useful features. Having very fine slices doesn't hurt the calculations, although the table will have more rows and may become unweildly. Next, count how many people died during each slice and how many were *censored* (that is, last seen alive during that slice, either because they became lost to follow-up or were still alive at the end of the study). For the survival times shown in Table (6), a natural choice would be to use seven one-year time slices.

	Died	Censored	Alive at Start	At Risk	Prob of Dying	Prob of Surviving	Cum Survival
0-1 yr	1	1	10	9.5	0.105263	0.894737	0.894737
1-2 yr	0	0	8	8.0	0.000000	1.000000	0.894737
$23~\mathrm{yr}$	2	0	8	8.0	0.250000	0.750000	0.671053
$3-4~\mathrm{yr}$	0	1	6	5.5	0.000000	1.000000	0.671053
$4-5~\mathrm{yr}$	1	0	5	5.0	0.200000	0.800000	0.536842
$5\text{-}6~\mathrm{yr}$	1	1	4	3.5	0.285714	0.714286	0.383459
$6\text{-}7~\mathrm{yr}$	1	1	2	1.5	0.666667	0.333333	0.127820

**Table 7**. Lifetable corresponding to the example data in Table (6).

Next, count how many people died during each slice and how many were *censored* (that is, last seen alive during that slice, either because they became lost to follow-up or were still alive at the end of the study). From Table (6) you see that

- During the first year after surgery, one subjects died (#0) and one subject was censored (#4)
- During the second year, nothing happened (no deaths, no censoring)
- During the third year, two subjects died (#8 and #3)
- During the fourth year, one subject was censored (#6)
- During the fifth year, one subject died (#5)
- During the sixth year, one subject died (#9) one subject was censored (#7)
- During the seventh year, one subjectdied (#1) and one was censored (#2)

To construct a lifetable, one proceeds as follows:

- Enter the total number of subjects alive at the start into column Alive at Start, in the 0-1 yr row
- Enter the counts of people who died within each time slice into column **Died**
- Enter the counts of people who were censored during each time slice into **Censored**
- The column **At Risk** shows the number of subjects known to be alive at the start of each year after surgery. This is equal to the number of subjects alive at the start of the preceding year minus the number of subjects who died (**Died**) or were censored (**Censored**) during the preceding year.
- The Column At Risk shows the number of subjects "at risk for dying" during each year. You may guess that this is the number of people alive at the start of the interval, but there's one minor correction. (COMPLETELY DISAGREE WITH THIS; A WEAKNESS OF THE METHOD) If any people were censured during that year, then they weren't really "available to die" (to use an awful expression) for the entire year. If you don't know exactly when, during that year, they became censored, then it's reasonable to "split the difference" and consider them at risk for aonly half the year. So the number at risk can be estimated as the number alive at the start

of the year, minus one-half of the number who became censored during that year. (ONLY MAKES SENSE FOR LOST TO FOLLOW UP).

- The column **Prob of Dying** shows the probability of dying during each interval, assuming the subject has survived up to the start of that interval. This is simply the number of people who died divided by the number of people at risk during each interval.
- The column **Prob of Surviving** shows the probability of surviving during each interval, assuming the subject has survived up to the start of that interval. Surviving means not dying, so the probability of surviving is simply 1 the probability of dying.
- The column **Cum Survival** shows the cumulative probability of surviving from the time of the operation all the way through the end of this time slice. To survive from the time of the operation through the end of any given year (year N), the subject must survive each of the years from Year 1 through Year N. Because surviving each year is an independent accomplishment, the probability of surviving all N of the years is the product of the individual years' probabilities.

The sample hazard and survival values obtained from a life table are only sample estimates (in this example, at 1-year time slices) of the true population hazard and survival functions. The hazard rate obtained from a life table is equal to the probability of dying during each time slice (column **Prob of Dying**) divided by the width of the slice, so the hazard rate for the first year would be expressed as .105 per year, or 10.5 percent per year. The cumulative survival probability in column **Cum Survival**, is the probability of surviving from the operation date through to the end of the interval. It has no units, and it can be expressed as a fraction or as a percentage.

Using very narrow time slices doesn't hurt life-table calculations. In fact, you can define slices so narrow that each subject's survival time falls within its own private little slice. With N subjects, N rows would have one subject each; all the rest of the rows would be empty. And because empty rows don't affect the life-table calculations, you can delete them entirely, leaving a table with only N rows, one for each subject. (If you happen to have two or more subjects with exactly the same survival or censoring time, it's okay to put each of the subjects in a separate row). The life-table calculations work fine with only one subject per row and produce what's called Kaplan-Meier (K-M) Survival estimates. You can think of the K-M method as a very fine-grained life table or a life table as a grouped K-M calculation.

The Kaplan-Meier survival estimate corresponding to the data given in Table (6) is shown in Table (8).

Shouldn't the column **Alive at Start** only decrease when the previous row has definitely died?? Why is it decreasing for the censored data?? Seems like a poor estimate of the hazard rate and survival curve. A knock against the cases where the data has no lost to follow up censored cases, such as the SEER data. Again, leads to overly pessimistic estimates of survival.

Survival analysis allows us to model the time until an event happens. Time-to-event measures pose unique problems for the analyst. Suppose that you want to predict the

	Censored Status	Survival Times	Alive at Start	Prob of Dying	Prob of Surv	Cum Survival
4	0	0.50	10	0.000000	1.000000	1.000000
0	1	0.75	9	0.111111	0.888889	0.888889
8	1	2.30	8	0.125000	0.875000	0.777778
3	1	2.40	7	0.142857	0.857143	0.666667
6	0	3.50	6	0.000000	1.000000	0.666667
5	1	4.50	5	0.200000	0.800000	0.533333
9	1	5.20	4	0.250000	0.750000	0.400000
7	0	5.80	3	0.000000	1.000000	0.400000
1	1	6.10	2	0.500000	0.500000	0.200000
2	0	7.00	1	0.000000	1.000000	0.200000

**Table 8.** Kaplan-Meier table corresponding to the example data in Table (6).

survival time for patients receiving an experimental cancer treatment. After three years, some of the patients in the study have died, and you can compute the survival time for each of these patients. However, many of the patients are still living at the end of three years; you do not know their ultimate survival time. Statisticians call this problem *censoring*, a problem that surfaces when you try to model time-to-event reopnse measures using data captured over a limited time period.

The two kinds of censoring are right censoring and left censoring. If you only know that the pertinent event is after some date, as is the case for patients in the preceding example who survive to the end of the study, the data is right-censored. On the other hand, if you only know that the beginning of the pertinent time-to-event took place before a certain date, the data is left-censored. For example, if you know that every patient in the study received the experimental treatment before the study started but do not know the exact date of treatment, the data is left-censored. Data can be both right-censored and left-censored.

Survival analysis is a family of techniques developed to work with censored time-to-event response measures. Note that if censoring is not present, you may be able to model time-to-event using standard modeling techniques. For some studies, however, you would have to wait a very long time before every sampled observation has a terminal event; in the case of the experimental cancer treatment, some patients might live another 20 years. Hence, survival analysis techniques enable the analyst to take full advantage of available data without waiting until every treated patient dies, every sampled part fails, or every tracked account closes.

The fundamental concept in survival analysis is the survival curve, S(t), which is a

$$S(t) = 1 - CDF(t) \tag{3.1}$$

where CDF(t) is the probability of a lifetime less than or equal to t. From the survival curve we can derive the **hazard function**; for pregnancy lengths, hazard function maps

from a time, t, to the fraction of pregnancies that continue until t

$$\lambda(t) = \frac{S(t) - S(t+1)}{S(t)} \tag{3.2}$$

The numerator is the fraction of lifetimes that end at t, which is also PMF(t). If someone gives you the CDF of lifetimes, it is easy to compute the survival and hazard functions. But in many real-world scenarios, we can't measure the distribution of lifetimes directly. We have to infer it. For example, suppose you are following a group of patients to see how long they survive after diagnosis. Not all patients are diagnosed on the same day, so at any point in time, some patients have survived longer than others. If some patients have died, we know their survival

If we wait until all patients are dead, we can compute the survival curve, but it we are evaluating the effectiveness of a new treatment, we can't wait that long! We need a way to estimate survival curves using incomplete information.

The general idea is that we can use the data to estimate the hazard function, then convert the hazard function to a survival curve. Once we have the hazard function, we can estimate the survival curve. The chance of surviving past time t is the chance of surviving all times up through t, which is the cumulative product of the complementary hazard function:

$$[1 - \lambda(0)][1 - \lambda(1)] \cdots [1 - \lambda(t)]$$
 (3.3)

#### 3.2 Transformation of Censored Data for Machine Learning

In this section we describe an inuitive way to transform right-censored data appropriately so that it may be used as input to machine learning algorithms that learn the hazard fuction described in section 3.1. The full details of this transformation, and a large inspiration for this study, can be flound in this blog post [1].

The overall philosophy of the Kaplan-Meier estimate of the survival curve for a population differs fundamentally from the methods described below and used in this study. The Kaplan-Meier estimate of the survival curve is given by

$$\hat{S}(t) = \prod_{t_i < t} \frac{n_i - d_i}{n_i} \tag{3.4}$$

where  $d_i$  are the number of death events at time t and  $n_t$  is the number of subjects at risk of death just prior to time t. Equation 3.4 uses the entire data set to arrive at an estimate of the entire population survival curve. In contrast, the method described below uses the entire data set to learn a model so as to predict hazard and survival curves for all of the individual records in the data set.

The key observation is to note that the hazard function in Equation 3.2 can be readily learned via machine learning methods. It can be rewritten as

$$h(\mathbf{X}, t) = P(Y = t | Y \ge t, \mathbf{X}), \tag{3.5}$$

the probability that, if someone has survived up until month t, they will die in that month. where **X** represents all of the data for that particular record, and in our case Y represents the

	cs_tumor_size	year_of_birth	survival_months	vital_status_recode_Dead
newindex				
205	60	1951	3	1

Table 9. Example of four columns in an uncensored record in the untransformed dataset.

	cs_tumor_size	year_of_birth	survival_months	vital_status_recode_Dead
newindex				
205	40	1950	3	0

Table 10. Example of four columns in a censored record in the untransformed dataset.

	cs_tumor_size	year_of_birth	month	newtarget
newindex				
205	60	1951	0	0
205	60	1951	1	0
205	60	1951	2	0
205	60	1951	3	1

Table 11. Example of four columns in an uncensored record in the transformed dataset.

true, uncensored number of survival months of the patient. What is actually provided in the SEER data is the related variable SURVIVAL MONTHS T (how long each subject was in the study), and whether they exited by dying or being censored (D), VITAL STATUS RECODE. D is a Boolean variable, so D=1 if T=Y, and D=0 if T<Y.

Treating T is just another covariate is the key to the transformation. Each datapoint in the hidden classification problem is the combination of an  $\mathbf{X}_i$  in hte original dataset plut some month t, and the classification problem is "did point  $\mathbf{X}_i$  die in month t." We will call this new variable  $D_{it}$  (newtarget). We can transformation our original data set into a new one, with one row for each month that each  $\mathbf{X}_i$  is in the sample; train a standard classifier on this new dataset with  $D_{it}$  as the target, and derive a survival model from the original dataset. Psuedocode for this transformation is found in section A.4.

Explicit examples will help make this transformation clear. The untransformed datapoint represented Table (9) is transformed to the multiple records shown in Table (11). All uncensored data is transformed in this way. All censored data is similar transformed. The untransformed datapoint represented Table (10) is transformed to the multiple records shown in Table (12).

One obvious side effect of this transformation is that it explodes the data size. For this study, the original, untransformed colon cancer DataFrame has shape (113072, 106),

	cs_tumor_size	year_of_birth	month	newtarget
newindex				
205	40	1950	0	0
205	40	1950	1	0
205	40	1950	2	0
205	40	1950	3	0

Table 12. Example of four columns in a censored record in the transformed dataset.

and the total transformed colon cancer DataFrame has shape (4165251, 106). Similary, the original, untransformed lung cancer DataFrame has shape (177089, 118), and the total transformed colon cancer DataFrame has shape (3079931, 118). The biggest explosion in data size occured with the breast cancer data. The original, untransformed breast cancer DataFrame has shape (329949, 70), and the total transformed breast cancer DataFrame has shape (15085711, 70). Traning machine learning algorithms on such large datasets, even after splitting into training and testing sets described below, require large RAM. All computations for this study were performed on a Dell XPS 8700 Desktop with 32BG of RAM.

#### 4 Prediction Models

With the datasets transformed as described in section (3.2), we are now able to split them into training and testing sets in the usual manner. The classifier models described in this section are learning the hazard function: given all of the data given in sections (A.2, A.1, A.3), which includes the field months (the months after diagnosis), the models predict the target variable newtarget, which represents the probability of dying in that month, given that the patient represented by the record has survived up to that month. This prediction task should not be confused with the regression problem of trying to predict precisely in what month a patient will die. The hazard functions thus learned and predicted are intermediary products; what we are really pursuing are the survival functions for each patient that are derived from the learned and predicted hazard functions. From the resulting hazard functions for each unique patient, we can construct the resulting survival functions as presented in section (A.4) and explicitly given in python code in the notebooks at the github repository containing supplemental material for this study [8].

#### 4.1 Decision Trees and Random Forests

#### 4.2 MLP Neural Networks

#### 5 Performance Metrics

#### 5.1 Model Agreement

TO DO: Check if comorbidites are contributing to the outliers in the agreement boxplots that follow. Could mesh with the prevous work [15].

Model	6 Months AUC	12 Months AUC	60 Months AUC
Breast RF	.846	.885	.844
Breast $NN$	.855	.867	.836
Colon RF	.804	.806	.828
Colon NN	.797	.804	.841
Lung RF	.772	.796	.874
Lung NN	.765	.796	.875

**Table 13**. AUC values for the Random Forest and Neural Networks model binary classifiers derived from the full survival curve predictions; see text for details.

Cancer Type	% agreement 6 months	% agreement 12 months	% agreement 60 months
Colon	.981	.971	.915
$\operatorname{Breast}$	.994	.984	.938
Lung	.861	.883	.900

**Table 14**. Percentage agreement for the Random Forest and Neural Network classifiers for 6, 12, and 60 month survival predictions on the test data for each cancer type.

## 6 Web Applications

CF this guy http://kmplot.com/analysis/index.php?p=service&cancer=lung

#### 7 Further Directions

Discussion of causality. A certain Marital status is not a "cause" of a better prognosis; c.f. Simpson's Paradox. Implementation of Judea Pearl's Causality Calculus.

## A Selected Features

In this Appendix we explicitly list the features chosen for each of the Colon, Breast and Lung cancer predictive models. For each cancer type, the features chosen for the random forest and neural network models were the same, so as to be best able to compare the two models. IPython notebooks explicitly providing all code, as well as html versions of the notebooks, are available from a GitHub repository providing supplemental material for thus study [8].

## A.1 Colon Cancer Feature Selection

The feature set used as input into both the Random Forest and Neural Network models, after the transformation described in Section is given below and also available in full detail in the file NewPatientColonML.html.

- cs tumor size
- elevation
- grade\_cell type not determined
- grade moderately differentiated
- grade poorly differentiated
- grade\_undifferentiated; anaplastic
- grade well differentiated
- histology recode broad groupings acinar cell neoplasms
- histology recode broad groupings adenomas and adenocarcinomas
- histology recode broad groupings blood vessel tumors
- $\bullet \ \ histology\_recode\_broad\_groupings\_complex\ epithelial\ neoplasms$
- histology recode broad groupings complex mixed and stromal neoplasms
- histology\_recode\_broad\_groupings\_cystic, mucinous and serous neoplasms
- histology\_recode\_broad\_groupings\_ductal and lobular neoplasms
- histology\_recode\_broad\_groupings\_epithelial neoplasms, NOS
- histology\_recode\_broad\_groupings\_fibromatuos neoplasms
- histology recode broad groupings germ cell neoplasms
- histology recode broad groupings lipomatous neplasms
- histology recode broad groupings miscellaneous bone tumors
- histology recode broad groupings myomatous neoplasms
- histology recode broad groupings neuroepitheliomatous neoplasms
- histology recode broad groupings nevi and melanomas
- histology recode broad groupings paragangliomas and glumus tumors
- histology recode broad groupings soft tissue tumors and sarcomas, NOS
- histology recode broad groupings squamous cell neoplasms
- histology recode broad groupings synovial-like neoplasms
- histology recode broad groupings transistional cell papillomas and carcinomas
- histology\_recode\_broad\_groupings\_unspecified neoplasms
- lai
- laterality Left: origin of primary
- laterality\_Not a paired site
- laterality Only one side involved, right or left origin unspecified
- laterality Paired site, but no information concerning laterality; midline tumor
- laterality\_Right: origin of primary
- lng
- marital\_status\_at\_dx\_Divorced
- marital status at dx Married (including common law)
- marital status at dx Separated
- marital status at dx Single (never married)
- marital status at dx Unknown
- marital\_status\_at\_dx Unmarried or domestic partner
- marital status at dx Widowed
- month of diagnosis Apr

- month of diagnosis Aug
- month of diagnosis Dec
- month of diagnosis Feb
- month of diagnosis Jan
- month of diagnosis Jul
- $\bullet$  month\_of\_diagnosis\_Jun
- month\_of\_diagnosis\_Mar
- month\_of\_diagnosis\_May
- month of diagnosis Nov
- month\_of\_diagnosis\_Oct
- $\bullet \hspace{0.1cm} \text{month\_of\_diagnosis\_Sep}$
- number\_of\_primaries
- race\_ethnicity\_Amerian Indian, Aleutian, Alaskan Native or Eskimo
- race\_ethnicity\_Asian Indian
- race\_ethnicity\_Asian Indian or Pakistani
- race ethnicity Black
- race ethnicity Chinese
- race ethnicity Fiji Islander
- race ethnicity Filipino
- race ethnicity Guamanian
- race ethnicity Hawaiian
- race ethnicity Hmong
- race ethnicity Japanese
- race ethnicity Kampuchean
- race\_ethnicity\_Korean
- race ethnicity Laotian
- race ethnicity Melanesian
- race ethnicity Micronesian
- race ethnicity New Guinean
- $\bullet \ \ {\rm race\_ethnicity\_Other}$
- race\_ethnicity\_Other Asian
- race ethnicity Pacific Islander
- race ethnicity Pakistani
- race\_ethnicity\_Polynesian
- race ethnicity Samoan
- race\_ethnicity\_Thai
- race ethnicity Tongan
- race ethnicity Unknown
- race ethnicity Vietnamese
- race ethnicity White
- seer historic stage a Distant
- seer historic stage a In situ
- seer historic stage\_a\_Localized

- seer historic stage a Regional
- seer\_historic\_stage\_a\_Unstaged
- $\bullet \text{ sex} \_\text{Female}$
- spanish hispanic origin Cuban
- spanish hispanic origin Dominican Republic
- spanish\_hispanic\_origin\_Mexican
- spanish hispanic origin Non-Spanish/Non-hispanic
- spanish\_hispanic\_origin\_Other specified Spanish/Hispanic origin (excludes Dominican Repuclic)
- spanish hispanic origin Puerto Rican
- spanish hispanic origin South or Central American (except Brazil)
- spanish hispanic origin Spanish surname only
- spanish\_hispanic\_origin\_Spanish, NOS; Hispanic, NOS; Latino, NOS
- spanish\_hispanic\_origin\_Uknown whether Spanish/Hispanic or not
- year\_of\_birth
- year of diagnosis
- month
- newtarget

#### A.2 Lung Cancer Feature Selection

The feature set used as input into both the Random Forest and Neural Network models, after the transformation described in Section is given below and also available in full detail in the file NewPatientLungML.html.

- $\bullet$  cs\_tumor\_size
- elevation
- grade cell type not determined
- grade moderately differentiated
- grade poorly differentiated
- grade undifferentiated; anaplastic
- grade well differentiated
- histology recode broad groupings acinar cell neoplasms
- histology recode broad groupings adenomas and adenocarcinomas
- histology recode broad groupings blood vessel tumors
- histology recode broad groupings complex epithelial neoplasms
- histology\_recode\_broad\_groupings\_complex mixed and stromal neoplasms
- histology recode broad groupings cystic, mucinous and serous neoplasms
- histology recode broad groupings ductal and lobular neoplasms
- histology recode broad groupings epithelial neoplasms, NOS
- histology recode broad groupings fibroepithelial neoplasms
- histology recode broad groupings fibromatuos neoplasms
- histology recode broad groupings germ cell neoplasms
- histology recode broad groupings gliomas

- histology\_recode\_broad\_groupings\_granular cell tumors & alveolar soft part sarcomas
- $\bullet \ histology\_recode\_broad\_groupings\_lipomatous\ neplasms$
- histology recode broad groupings miscellaneous bone tumors
- histology recode broad groupings miscellaneous tumors
- histology recode broad groupings mucoepidermoid neoplasms
- histology recode broad groupings myomatous neoplasms
- histology recode broad groupings myxomatous neoplasms
- histology recode broad groupings nerve sheath tumors
- $\bullet \ \ histology\_recode\_broad\_groupings\_neuroepitheliomatous\ neoplasms$
- histology recode broad groupings nevi and melanomas
- histology recode broad groupings osseous and chondromatous neoplasms
- $\bullet$  histology\_recode\_broad\_groupings\_paragangliom as and glumus tumors
- histology\_recode\_broad\_groupings\_soft tissue tumors and sarcomas, NOS
- histology\_recode\_broad\_groupings\_squamous cell neoplasms
- histology recode broad groupings synovial-like neoplasms
- histology recode broad groupings thymic epithelial neoplasms
- histology\_recode\_broad\_groupings\_transistional cell papillomas and carcinomas
- histology recode broad groupings trophoblastic neoplasms
- histology\_recode\_broad\_groupings\_unspecified neoplasms
- lat
- laterality Bilateral involvement, lateral origin unknown; stated to be single primary
- laterality Left: origin of primary
- laterality Not a paired site
- laterality\_Only one side involved, right or left origin unspecified
- laterality Paired site, but no information concerning laterality; midline tumor
- laterality Right: origin of primary
- lng
- marital status at dx Divorced
- marital status at dx Married (including common law)
- marital status at dx Separated
- marital status at dx Single (never married)
- marital status at dx Unknown
- marital\_status\_at\_dx\_Unmarried or domestic partner
- marital status at dx Widowed
- month\_of\_diagnosis\_Apr
- month of diagnosis Aug
- month\_of\_diagnosis\_Dec
- $\bullet \hspace{0.1cm} \text{month\_of\_diagnosis\_Feb}$
- month of diagnosis Jan
- month of diagnosis Jul
- month of diagnosis Jun
- month of diagnosis Mar

- month of diagnosis May
- $\bullet$  month\_of\_diagnosis\_Nov
- month of diagnosis Oct
- month of diagnosis Sep
- number of primaries
- race\_ethnicity\_Amerian Indian, Aleutian, Alaskan Native or Eskimo
- race ethnicity Asian Indian
- race ethnicity Asian Indian or Pakistani
- race ethnicity Black
- race ethnicity Chamorran
- $\bullet \ \ {\rm race\_ethnicity\_Chinese} \\$
- race\_ethnicity\_Fiji Islander
- race ethnicity Filipino
- $\bullet$  race\_ethnicity\_Guamanian
- race\_ethnicity\_Hawaiian
- race\_ethnicity\_Hmong
- race ethnicity Japanese
- race ethnicity Kampuchean
- race ethnicity Korean
- race ethnicity Laotian
- race ethnicity Melanesian
- race ethnicity Micronesian
- race ethnicity New Guinean
- race ethnicity Other
- race\_ethnicity\_Other Asian
- race ethnicity Pacific Islander
- race ethnicity Pakistani
- race ethnicity Polynesian
- race ethnicity Samoan
- race ethnicity Thai
- $\bullet$  race\_ethnicity\_Tongan
- race ethnicity Unknown
- race ethnicity Vietnamese
- $\bullet$  race\_ethnicity\_White
- seer historic stage a Distant
- seer\_historic\_stage\_a\_In situ
- seer historic stage a Localized
- seer historic stage a Regional
- seer historic stage a Unstaged
- $\bullet$  sex Female
- spanish hispanic origin Cuban
- spanish hispanic origin Dominican Republic
- spanish hispanic origin Mexican

- spanish hispanic origin Non-Spanish/Non-hispanic
- spanish\_hispanic\_origin\_Other specified Spanish/Hispanic origin (excludes Dominican Repuclic)
- spanish hispanic origin Puerto Rican
- spanish hispanic origin South or Central American (except Brazil)
- spanish hispanic origin Spanish surname only
- spanish hispanic origin Spanish, NOS; Hispanic, NOS; Latino, NOS
- spanish hispanic origin Uknown whether Spanish/Hispanic or not
- year of birth
- year of diagnosis
- month
- newtarget

#### A.3 Breast Cancer Feature Selection

The feature set used as input into both the Random Forest and Neural Network models, after the transformation described in Section is given below and also available in full detail in the file NewPatientBreastMl.html.

- cs tumor size
- $\bullet$  elevation
- grade moderately differentiated
- grade poorly differentiated
- grade ndifferentiated; anaplastic
- grade well differentiated
- histology\_recode\_broad\_groupings\_adenomas and adenocarcinomas
- histology recode broad groupings adnexal and skin appendage neoplasms
- histology recode broad groupings basal cell neoplasms
- histology recode broad groupings complex epithelial neoplasms
- histology recode broad groupings cystic, mucinous and serous neoplasms
- histology recode broad groupings ductal and lobular neoplasms
- histology recode broad groupings epithelial neoplasms, NOS
- $\bullet \ histology\_recode\_broad\_groupings\_nerve \ sheath \ tumors$
- histology recode broad groupings unspecified neoplasms
- lat
- laterality Bilateral involvement, lateral origin unknown; stated to be single primary
- laterality\_Paired site, but no information concerning laterality; midline tumor
- laterality Right: origin of primary
- lng
- marital stats at dx Divorced
- marital\_stats\_at\_dx\_Married (inclding common law)
- marital stats at dx Separated
- marital stats at dx Single (never married)
- marital stats at dx Unknown

- marital stats at dx Unmarried or domestic partner
- marital stats at dx Widowed
- month\_of\_diagnosis\_Apr
- month of diagnosis Aug
- month of diagnosis Dec
- month\_of\_diagnosis\_Feb
- month of diagnosis Jan
- month of diagnosis Jul
- month of diagnosis Jun
- month of diagnosis Mar
- month of\_diagnosis\_May
- month of diagnosis Nov
- month of diagnosis Oct
- month of diagnosis Sep
- race\_ethnicity\_Amerian Indian, Aletian, Alaskan Native or Eskimo
- race\_ethnicity\_Asian Indian
- race ethnicity Black
- race ethnicity Chinese
- race ethnicity Japanese
- race ethnicity Melanesian
- race ethnicity Other
- race ethnicity Other Asian
- race ethnicity Pacific Islander
- race ethnicity Thai
- race\_ethnicity\_Unknown
- race ethnicity Vietnamese
- race ethnicity White
- seer historic stage a Distant
- seer historic stage a In sit
- seer historic stage a Localized
- seer historic stage a Unstaged
- sex Female
- spanish hispanic origin Cuban
- $\bullet \ \, spanish\_hispanic\_origin\_Mexican \\$
- spanish hispanic origin Non-Spanish/Non-hispanic
- spanish\_hispanic\_origin\_Other specified Spanish/Hispanic origin (excldes Dominican Republic)
- spanish\_hispanic\_origin\_Spanish surname only
- spanish hispanic origin Spanish, NOS; Hispanic, NOS; Latino, NOS
- year of birth
- year of diagnosis
- month
- newtarget

#### A.4 Pseudocode for the Data Transformation

```
def train(X, T, D)
    // X, T, D are the original dataset
    X ' = []
    D' = \lceil \rceil
    // the transformation
    for each index i in X:
        for t=1 to T[i]:
            new_D = (0 if t < T[i], else D[i])
            append new_D to D'
            new_X = (X[i], t)
            append new_X to X'
    return a decision tree trained on (X', D')
def pmf(h, X)
    // X is a single datapoint
    // returns an array A where A[i] = P(Y = i | X)
    A = []
    p_so_far = 1 // this is p(T >= t | X)
    for t = 1 to (the last month where h has any data):
        // h knows p(T = t | T >= t, X), we call this p_cur
        p_cur = h's prediction for (X, t)
        append (p_so_far * p_cur) to A
        p_so_far *= (1 - p_cur)
```

## B Model Architecture and Python Code

## C GitHub Repositories

Please always give a title also for appendices.

## Acknowledgments

This is the most common positions for acknowledgments. A macro is available to maintain the same layout and spelling of the heading.

**Note added.** This is also a good position for notes added after the paper has been written.

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Figure 4. The top levels of a decision tree trained on the Lung Cancer training data.

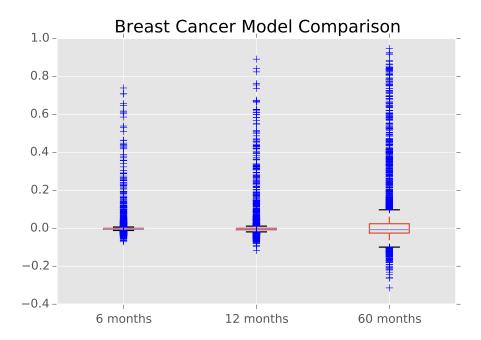
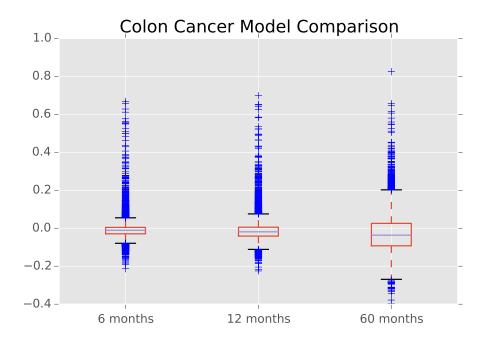


Figure 5. Box plots showing the distributions of the signed difference between the MLP model's prediction for the probability of surviving 6 months and the Random Forest model's prediction of the same quantity for breast cancer. The plot shows the same quantity for the 12 and 60 months classifiers. It is apparent from the figures that the outliers are due to the neural network models predicting higher survival probablitlies than the random forest for some few cases. These differences were evaluated for the 3300 test patients in the breast cancer data.



**Figure 6**. Box plots showing the distributions of the signed difference between the MLP model's prediction for the probability of surviving 6 months and the Random Forest model's prediction of the same quantity for colon cancer. The plot shows the same quantity for the 12 and 60 months classifiers. It is apparent from the figures that the outliers are due to the neural network models predicting higher survival probablitlies than the random forest for some few cases. These differences were evaluated for the 5654 test patients in the colon cancer data.

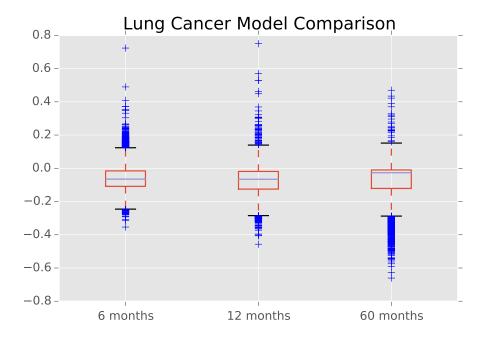


Figure 7. Box plots showing the distributions of the signed difference between the MLP model's prediction for the probability of surviving 6 months and the Random Forest model's prediction of the same quantity for lung cancer. The plot shows the same quantity for the 12 and 60 months classifiers. These differences were evaluated for the 5654 test patients in the colon cancer data. The Interquartile Ranges for lung cancer are visibly larger than those for breast cancer and colon cancer shown in fig 5 and fig 6.