Least-Squares Fitting

James R. Graham 2014/9/21

A straight line fit

Suppose that we have a set of N observations (x_i, y_i) where we believe that the measured value, y, depends linearly on x, i.e.,

$$y = mx + c$$
.

For example, suppose a body is moving with constant velocity what is the speed (m) and initial (c) position of the object?

Given our data, what is the best estimate of m and c? Assume that the independent variable, x_i , is known exactly, and the dependent variable, y_i , is drawn from a Gaussian probability distribution function with constant standard deviation, $\sigma_i = const$. Under these circumstances the most likely values of m and c are those corresponding to the straight line with the total minimum square deviation, i.e., the quantity

$$\chi^2 = \sum_{i} \left[y_i - (mx_i + c) \right]^2$$

is minimized when m and c have their most likely values. Figure 1 shows a typical deviation.

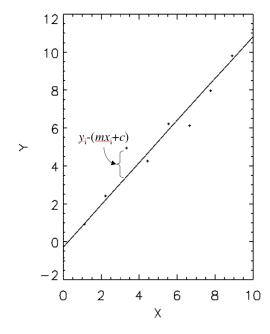


Figure 1: Some data with a least squares fit to a straight line. A typical deviation is illustrated.

The best values of m and c are found by solving the simultaneous equations,

$$\frac{\partial}{\partial m}\chi^2 = 0, \quad \frac{\partial}{\partial c}\chi^2 = 0.$$

Evaluating the derivatives yields

$$\frac{\partial}{\partial m}\chi^2 = \frac{\partial}{\partial m}\sum_{i} \left[y_i - (mx_i + c)\right]^2 = 2m\sum_{i}x_i^2 + 2c\sum_{i}x_i - 2\sum_{i}x_iy_i = 0$$

$$\frac{\partial}{\partial c}\chi^2 = \frac{\partial}{\partial c}\sum_{i} \left[y_i - (mx_i + c)\right]^2 = 2m\sum_{i}x_i + 2cN - 2\sum_{i}y_i = 0.$$

Which can conveniently be expressed in matrix form,

$$\left(\begin{array}{cc} \sum x_i^2 & \sum x_i \\ \sum x_i & N \end{array}\right) \left(\begin{array}{c} m \\ c \end{array}\right) = \left(\begin{array}{c} \sum x_i y_i \\ \sum y_i \end{array}\right)$$

and solved by multiplying both sides by the inverse,

$$\binom{m}{c} = \left(\sum_{i=1}^{n} x_i^2 \sum_{i=1}^{n} x_i \right)^{-1} \left(\sum_{i=1}^{n} x_i y_i \right)$$

The inverse can be computed analytically, or in Python it is trivial to compute the inverse numerically, as follows.

Example Python

```
# Test least squares fitting with simulated data.
import numpy as np
import matplotlib.pyplot as plt
nx = 20  # Number of data points m = 1.0  # Gradient c = 0.0  # Intercept
x = np.arange(nx,dtype=float)
                                      # Independent variable
y = m * x + c
                                      # dependent variable
# Generate Gaussian errors
sigma = 1.0
                                      # Measurement error
np.random.seed(1)
                                      # init random no. generator
errors = sigma*np.random.randn(nx) # Gaussian distributed errors
                                      # Add the noise
ye = y + errors
plt.plot(x,ye,'o',label='data')
plt.xlabel('x')
plt.ylabel('y')
# Construct the matrices
ma = np.array([ [np.sum(x**2), np.sum(x)], [np.sum(x), nx ] ] )
mc = np.array([ [np.sum(x*ye)], [np.sum(ye)]])
```

```
# Compute the gradient and intercept
mai = np.linalg.inv(ma)
print 'Test matrix inversion gives identity',np.dot(mai,ma)
md = np.dot(mai,mc)  # matrix multiply is dot

# Overplot the best fit
mfit = md[0,0]
cfit = md[1,0]
plt.plot(x, mfit*x + cfit)
plt.axis('scaled')
plt.text(5,15,'m = {:.3f}\nc = {:.3f}'.format(mfit,cfit))
plt.savefig('lsq1.png')
```

See Figure 2 for the output of this program.

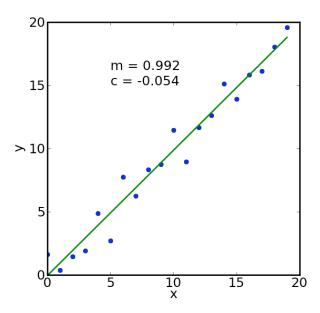


Figure 2—Least squares straight line fit. The true values are m = 1 and c = 0.

Error propagation

What are the uncertainties in the slope and the intercept? To begin the process of error propagation we need the inverse

$$\begin{pmatrix}
\sum_{i} x_{i}^{2} & \sum_{i} x_{i} \\
\sum_{i} x_{i} & N
\end{pmatrix}^{-1} = \begin{pmatrix}
N/\left[N\sum_{i} x_{i}^{2} - \left(\sum_{i} x_{i}\right)^{2}\right] & \sum_{i} x_{i}/\left[N\sum_{i} x_{i}^{2} - \left(\sum_{i} x_{i}\right)^{2}\right] \\
\sum_{i} x_{i}/\left[\left(\sum_{i} x_{i}\right)^{2} - N\sum_{i} x_{i}^{2}\right] & \sum_{i} x_{i}/\left[N\sum_{i} x_{i}^{2} - \left(\sum_{i} x_{i}\right)^{2}\right]
\end{pmatrix},$$

so that we can compute analytic expressions for m and c,

$$\begin{pmatrix} m \\ c \end{pmatrix} = \begin{pmatrix} \sum x_i^2 & \sum x_i \\ \sum x_i & N \end{pmatrix}^{-1} \begin{pmatrix} \sum x_i y_i \\ \sum y_i \end{pmatrix} = \begin{pmatrix} \frac{\sum x_i \sum y_i - N \sum x_i y_i}{\left(\sum x_i\right)^2 - N \sum x_i^2} \\ \frac{\sum x_i \sum x_i y_i - \sum x_i^2 \sum y_i}{\left(\sum x_i\right)^2 - N \sum x_i^2} \end{pmatrix}.$$

The results of error propagation show that if $z = z(x_1, x_2, ... x_N)$

$$\sigma_z^2 = \sum_i (\partial z/\partial x_i)^2 \sigma_i^2 ,$$

assuming uncorrelated data (i.e., zero covariance). Thus,

$$\sigma_m^2 = \sum_j (\partial m/\partial y_j)^2 \sigma_j^2$$
 and $\sigma_c^2 = \sum_j (\partial c/\partial y_j)^2 \sigma_j^2$.

The expression for the derivative of the gradient, m, is

$$\frac{\partial m}{\partial y_j} = \frac{\partial}{\partial y_j} \left(\frac{\sum x_i \sum y_i - N \sum x_i y_i}{\left(\sum x_i\right)^2 - N \sum x_i^2} \right) = \frac{\sum x_i - N x_j}{\left(\sum x_i\right)^2 - N \sum x_i^2}$$

because $(\partial y_i/\partial y_j) = \delta_{ij}$. If we assume that the measurement error is the same for each point then

$$\sigma_{m}^{2} = \sigma^{2} \sum_{j} \left(\frac{\sum_{i} x_{i} - Nx_{j}}{\left(\sum_{i} x_{i}\right)^{2} - N\sum_{i} x_{i}^{2}} \right)^{2}$$

$$= \frac{\sigma^{2}}{\left[\left(\sum_{i} x_{i}\right)^{2} - N\sum_{i} x_{i}^{2} \right]^{2}} \sum_{j} \left[\left(\sum_{i} x_{i}\right)^{2} - 2Nx_{j} \sum_{i} x_{i} + N^{2}x_{j}^{2} \right]$$

$$= \frac{\sigma^{2}}{\left[\left(\sum_{i} x_{i}\right)^{2} - N\sum_{i} x_{i}^{2} \right]^{2}} \left[N\left(\sum_{i} x_{i}\right)^{2} - 2N\left(\sum_{i} x_{i}\right)^{2} + N^{2}\sum_{i} x_{i}^{2} \right]$$

$$= \frac{N\sigma^{2}}{N\sum_{i} x_{i}^{2} - \left(\sum_{i} x_{i}\right)^{2}}$$

Similarly,

$$\frac{\partial c}{\partial y_j} = \frac{\partial}{\partial y_j} \left(\frac{\sum x_i \sum x_i y_i - \sum x_i^2 \sum y_i}{\left(\sum x_i\right)^2 - N \sum x_i^2} \right) = \frac{x_j \sum x_i - \sum x_i^2}{\left(\sum x_i\right)^2 - N \sum x_i^2}$$

and

$$\sigma_{c}^{2} = \sigma^{2} \sum_{j} \left(\frac{x_{j} \sum x_{i} - \sum x_{i}^{2}}{\left(\sum x_{i} \right)^{2} - N \sum x_{i}^{2}} \right)^{2}$$

$$= \frac{\sigma^{2}}{\left[\left(\sum x_{i} \right)^{2} - N \sum x_{i}^{2} \right]^{2}} \sum_{j} \left[x_{j}^{2} \left(\sum x_{i} \right)^{2} - 2x_{j} \sum x_{i} \sum x_{i}^{2} + \left(\sum x_{i}^{2} \right)^{2} \right]$$

$$= \frac{\sigma^{2} \sum x_{i}^{2}}{\left[\left(\sum x_{i} \right)^{2} - N \sum x_{i}^{2} \right]^{2}} \left[N \sum x_{i}^{2} - \left(\sum x_{i} \right)^{2} \right]$$

$$= \frac{\sigma^{2}}{\left[\left(\sum x_{i} \right)^{2} - N \sum x_{i}^{2} \right]^{2}} \left[\sum x_{i}^{2} \sum x_{i} - 2 \sum x_{i} \sum x_{i}^{2} + N \sum x_{i}^{2} \right]$$

$$= \frac{\sigma^{2} \sum x_{i}^{2}}{N \sum x_{i}^{2} - \left(\sum x_{i} \right)^{2}}.$$

If we do not know, *a priori*, the standard deviation of the measurements, σ , the best estimate is derived from the deviations from comparing the data to the fit, i.e.,

$$\sigma^{2} = \frac{1}{N-2} \sum_{i} [y_{i} - (mx_{i} + c)]^{2}.$$