

Nearest Neighbor Gaussian Processes for Large Spatial Data

Andrew Finley¹ & Jeffrey Doser²

May 15, 2023

¹Department of Forestry, Michigan State University.

²Department of Integrative Biology, Michigan State University.

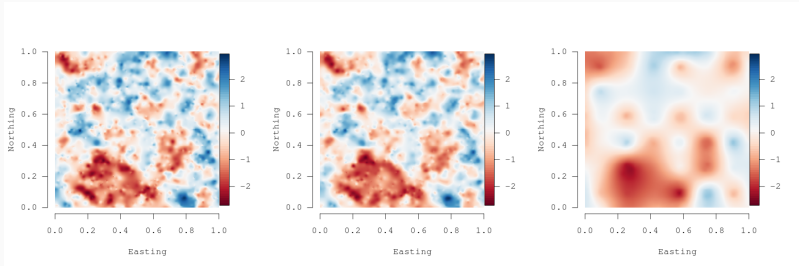
Low rank Gaussian Predictive Process

Pros

- Proper Gaussian process
- Allows for coherent spatial interpolation at arbitrary resolution
- Can be used as prior for spatial random effects in any hierarchical setup for spatial data
- Computationally tractable

Low rank Gaussian Predictive Process

Cons



True w

Full GP

PP 64 knots

Figure: Comparing full GP vs low-rank GP with 2500 locations

- Low rank models like the Predictive Process (PP) often tends to oversmooth
- Increasing the number of knots can fix this but will lead to heavy computation

Sparse matrices

- **Idea:** Use a **sparse** matrix instead of a low rank matrix to approximate the dense full GP covariance matrix
- **Goals:**
 - Scalability: Both in terms of **storage** and computing **inverse** and **determinants**
 - Closely approximate full GP inference
 - Proper Gaussian process model like the Predictive Process

Cholesky factors

- Write a joint density $p(w) = p(w_1, w_2, \dots, w_n)$ as:

$$p(w_1)p(w_2 | w_1)p(w_3 | w_1, w_2) \cdots p(w_n | w_1, w_2, \dots, w_{n-1})$$

- For Gaussian distribution $w \sim N(0, C)$ this \Rightarrow

$$w_1 = 0 + \eta_1;$$

$$w_2 = a_{21}w_1 + \eta_2;$$

$$\dots \quad \dots \quad \dots$$

$$w_n = a_{n1}w_1 + a_{n2}w_2 + \cdots + a_{n,n-1}w_{n-1} + \eta_n;$$

Cholesky factors

- Write a joint density $p(w) = p(w_1, w_2, \dots, w_n)$ as:

$$p(w_1)p(w_2 | w_1)p(w_3 | w_1, w_2) \cdots p(w_n | w_1, w_2, \dots, w_{n-1})$$

- For Gaussian distribution $w \sim N(0, C)$ this \Rightarrow

$$\begin{bmatrix} w_1 \\ w_2 \\ w_3 \\ \vdots \\ w_n \end{bmatrix} = \begin{bmatrix} 0 & 0 & 0 & \dots & 0 & 0 \\ a_{21} & 0 & 0 & \dots & 0 & 0 \\ a_{31} & a_{32} & 0 & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ a_{n1} & a_{n2} & a_{n3} & \dots & a_{n,n-1} & 0 \end{bmatrix} \begin{bmatrix} w_1 \\ w_2 \\ w_3 \\ \vdots \\ w_n \end{bmatrix} + \begin{bmatrix} \eta_1 \\ \eta_2 \\ \eta_3 \\ \vdots \\ \eta_n \end{bmatrix}$$
$$\Rightarrow w = Aw + \eta; \quad \eta \sim N(0, D), \text{ where } D = \text{diag}(d_1, d_2, \dots, d_n).$$

Cholesky factors

- Write a joint density $p(w) = p(w_1, w_2, \dots, w_n)$ as:

$$p(w_1)p(w_2 | w_1)p(w_3 | w_1, w_2) \cdots p(w_n | w_1, w_2, \dots, w_{n-1})$$

- For Gaussian distribution $w \sim N(0, C)$ this \Rightarrow

$$\begin{bmatrix} w_1 \\ w_2 \\ w_3 \\ \vdots \\ w_n \end{bmatrix} = \begin{bmatrix} 0 & 0 & 0 & \dots & 0 & 0 \\ a_{21} & 0 & 0 & \dots & 0 & 0 \\ a_{31} & a_{32} & 0 & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ a_{n1} & a_{n2} & a_{n3} & \dots & a_{n,n-1} & 0 \end{bmatrix} \begin{bmatrix} w_1 \\ w_2 \\ w_3 \\ \vdots \\ w_n \end{bmatrix} + \begin{bmatrix} \eta_1 \\ \eta_2 \\ \eta_3 \\ \vdots \\ \eta_n \end{bmatrix}$$

$$\Rightarrow w = Aw + \eta; \quad \eta \sim N(0, D), \text{ where } D = \text{diag}(d_1, d_2, \dots, d_n).$$

- Cholesky factorization:** $C^{-1} = (I - A)'D^{-1}(I - A)$

Cholesky factors

- $w_{<i} = (w_1, w_2, \dots, w_{i-1})'$
- $c_i = \text{Cov}(w_i, w_{<i}), C_i = \text{Var}(w_{<i})$
- i^{th} row of A and $d_i = \text{Var}(\eta_i)$ are obtained from $p(w_i | w_{<i})$ as follows:
 - Solve for a_{ij} 's from $\sum_{j=1}^{i-1} a_{ij} w_j = E(w_i | w_{<i}) = c_i' C_i^{-1} w_{<i}$

Cholesky factors

- $w_{<i} = (w_1, w_2, \dots, w_{i-1})'$
- $c_i = \text{Cov}(w_i, w_{<i}), C_i = \text{Var}(w_{<i})$
- i^{th} row of A and $d_i = \text{Var}(\eta_i)$ are obtained from $p(w_i | w_{<i})$ as follows:
 - Solve for a_{ij} 's from $\sum_{j=1}^{i-1} a_{ij} w_j = E(w_i | w_{<i}) = c_i' C_i^{-1} w_{<i}$
 - $d_i = \text{Var}(w_i | w_{<i}) = \sigma^2 - c_i' C_i^{-1} c_i$

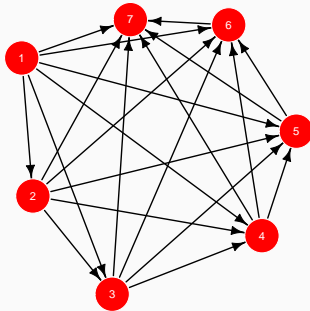
Cholesky factors

- $w_{<i} = (w_1, w_2, \dots, w_{i-1})'$
- $c_i = \text{Cov}(w_i, w_{<i}), C_i = \text{Var}(w_{<i})$
- i^{th} row of A and $d_i = \text{Var}(\eta_i)$ are obtained from $p(w_i | w_{<i})$ as follows:
 - Solve for a_{ij} 's from $\sum_{j=1}^{i-1} a_{ij} w_j = E(w_i | w_{<i}) = c_i' C_i^{-1} w_{<i}$
 - $d_i = \text{Var}(w_i | w_{<i}) = \sigma^2 - c_i' C_i^{-1} c_i$

Cholesky factors

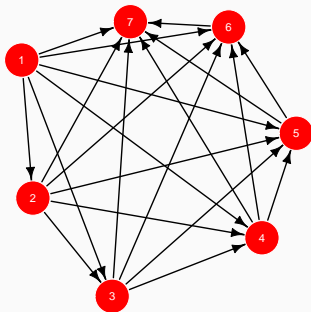
- $w_{<i} = (w_1, w_2, \dots, w_{i-1})'$
- $c_i = \text{Cov}(w_i, w_{<i}), C_i = \text{Var}(w_{<i})$
- i^{th} row of A and $d_i = \text{Var}(\eta_i)$ are obtained from $p(w_i | w_{<i})$ as follows:
 - Solve for a_{ij} 's from $\sum_{j=1}^{i-1} a_{ij} w_j = E(w_i | w_{<i}) = c_i' C_i^{-1} w_{<i}$
 - $d_i = \text{Var}(w_i | w_{<i}) = \sigma^2 - c_i' C_i^{-1} c_i$
- For large i , inverting C_i becomes **slow**
- The Cholesky factor approach for the full GP covariance matrix C **does not** offer any computational benefits

Cholesky Factors and Directed Acyclic Graphs (DAGs)



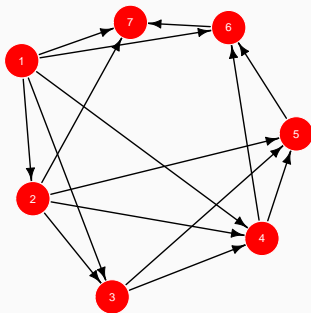
- Number of non-zero entries (**sparsity**) of A equals number of arrows in the graph
- In particular: Sparsity of the i^{th} row of A is same as the number of arrows towards i in the DAG

Introducing sparsity via graphical models



$$p(y_1)p(y_2 | y_1)p(y_3 | y_1, y_2)p(y_4 | y_1, y_2, y_3) \\ \times p(y_5 | y_1, y_2, y_3, y_4)p(y_6 | y_1, y_2, \dots, y_5)p(y_7 | y_1, y_2, \dots, y_6) .$$

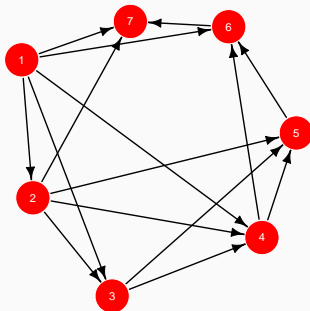
Introducing sparsity via graphical models



$$p(y_1)p(y_2 | y_1)p(y_3 | y_1, y_2)p(y_4 | y_1, y_2, y_3)$$

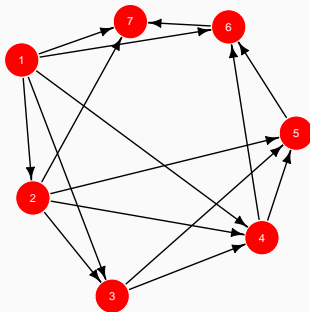
$$p(y_5 | \cancel{y_1}, y_2, y_3, y_4)p(y_6 | y_1, \cancel{y_2}, \cancel{y_3}, y_4, y_5)p(y_7 | y_1, y_2, \cancel{y_3}, \cancel{y_4}, \cancel{y_5}, y_6)$$

Introducing sparsity via graphical models



- Create a **sparse** DAG by keeping **at most m** arrows pointing to each node
- Set $a_{ij} = 0$ for all i, j which has no arrow between them
- Fixing $a_{ij} = 0$ introduces **conditional independence** and w_j drops out from the conditional set in $p(w_i \mid \{w_k : k < i\})$

Introducing sparsity via graphical models



- $N(i)$ denote *neighbor set* of i , i.e., the set of nodes from which there are arrows to i
- $a_{ij} = 0$ for $j \notin N(i)$ and nonzero a_{ij} 's obtained by solving:

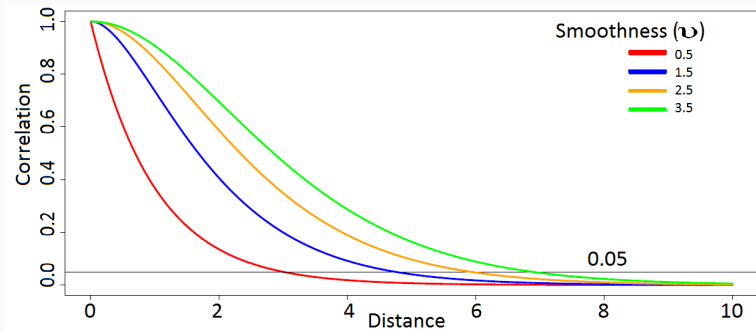
$$E[w_i \mid w_{N(i)}] = \sum_{j \in N(i)} a_{ij} w_j$$

- The above linear system is only $m \times m$

Choosing neighbor sets

Matern Covariance Function:

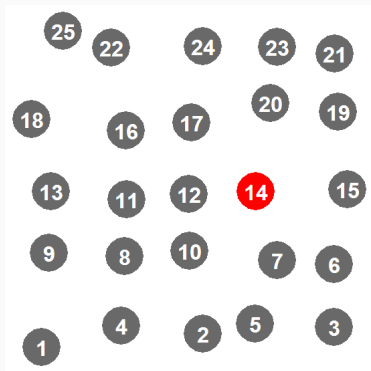
$$C(s_i, s_j) = \frac{1}{2^{\nu-1}\Gamma(\nu)} (\|s_i - s_j\|\phi)^\nu \mathcal{K}_\nu(\|s_i - s_j\|\phi); \phi > 0, \nu > 0,$$



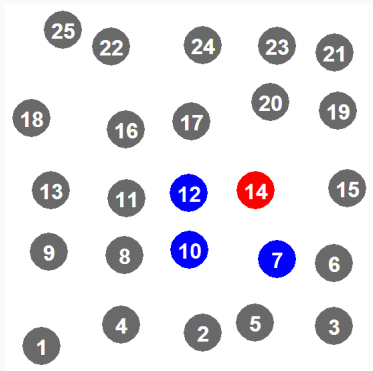
Choosing neighbor sets

- Spatial covariance functions decay with distance
- Vecchia (1988): $N(s_i) = m\text{--nearest neighbors}$ of s_i in s_1, s_2, \dots, s_{i-1}
 - Nearest points have highest correlations
 - Theory: "Screening effect" – Stein, 2002
- We use Vecchia's choice of m -nearest neighbor
- Other choices proposed in Stein et al. (2004); Gramacy and Apley (2015); Guinness (2016) can also be used

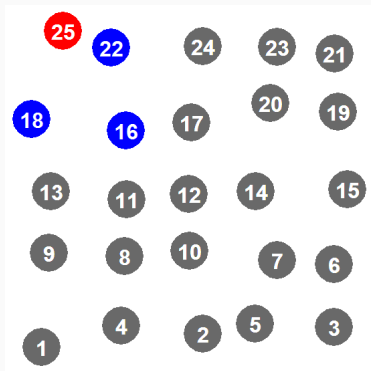
Nearest neighbors



Nearest neighbors

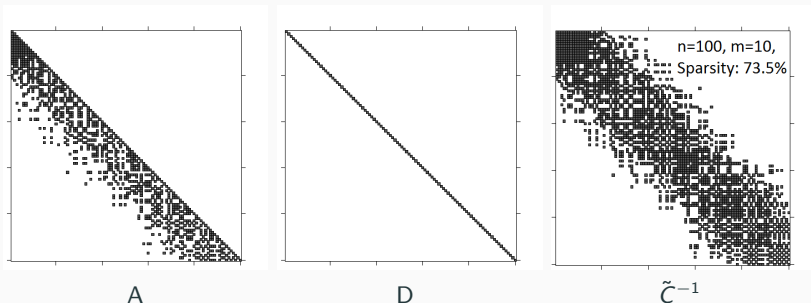


Nearest neighbors



Sparse precision matrices

- The neighbor sets and the covariance function $C(\cdot, \cdot)$ define a sparse Cholesky factor A
- $N(w | 0, C) \approx N(w | 0, \tilde{C})$; $\tilde{C}^{-1} = (I - A)^{\top} D^{-1} (I - A)$

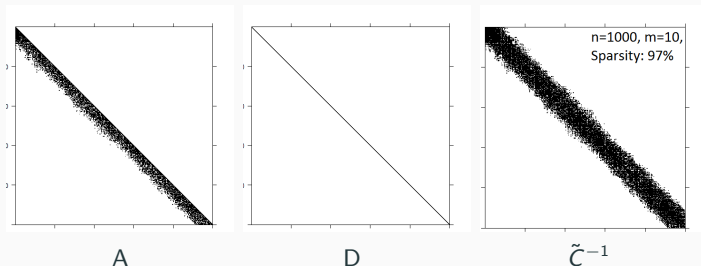


- $\det(\tilde{C}) = \prod_{i=1}^n D_i$,
- \tilde{C}^{-1} is sparse with $O(nm^2)$ entries
- Explore some A and \tilde{C}^{-1} sparsity patterns

https://github.com/finleya/NNGP_LDL

Sparse precision matrices

- The neighbor sets and the covariance function $C(\cdot, \cdot)$ define a sparse Cholesky factor A
- $N(w | 0, C) \approx N(w | 0, \tilde{C}) ; \tilde{C}^{-1} = (I - A)^{\top} D^{-1} (I - A)$



- $\det(\tilde{C}) = \prod_{i=1}^n D_i,$
- \tilde{C}^{-1} is sparse with $O(nm^2)$ entries
- Explore some A and \tilde{C}^{-1} sparsity patterns

https://github.com/finleya/NNGP_LDL

Extension to a Process

- We have defined $w \sim N(0, \tilde{C})$ over the set of data locations $S = \{s_1, s_2, \dots, s_n\}$
- For $s \notin S$, define $N(s)$ as set of m -nearest neighbors of s in S
- Define $w(s) = \sum_{i:s_i \in N(s)} a_i(s)w(s_i) + \eta(s)$ where $\eta(s) \stackrel{ind}{\sim} N(0, d(s))$
 - $a_i(s)$ and $d(s)$ are once again obtained by solving $m \times m$ system
- Well-defined GP over entire domain
 - **Nearest Neighbor GP (NNGP)** – Datta et al., JASA, (2016)

Spatial linear model

$$y(\mathbf{s}) = x(\mathbf{s})'\beta + w(\mathbf{s}) + \epsilon(\mathbf{s})$$

- $w(s)$ modeled as *NNGP* derived from a $GP(0, C(\cdot, \cdot, | \sigma^2, \phi))$
- $\epsilon(s) \stackrel{\text{iid}}{\sim} N(0, \tau^2)$ contributes to the nugget
- Priors for the parameters β , σ^2 , τ^2 and ϕ
- **Only** difference from a full GP model is the NNGP prior $w(s)$

Hierarchical spatial regression with NNGP

Full Bayesian Model

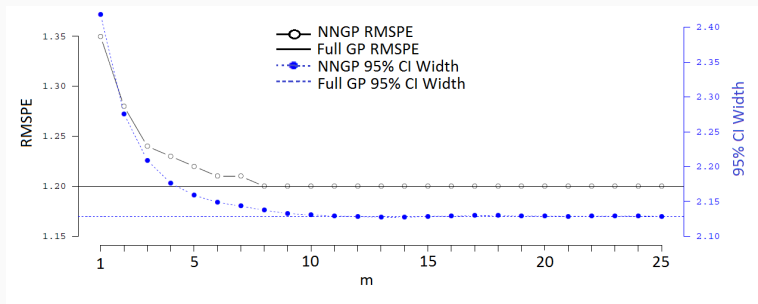
$$N(y | X\beta + w, \tau^2 I) \times N(w | 0, \tilde{C}(\sigma^2, \phi)) \times N(\beta | \mu_\beta, V_\beta) \\ \times IG(\tau^2 | a_\tau, b_\tau) \times IG(\sigma^2 | a_\sigma, b_\sigma) \times Unif(\phi | a_\phi, b_\phi)$$

Gibbs sampler:

- Conjugate full conditionals for β , τ^2 , σ^2 and $w(s_i)$'s
- Metropolis step for updating ϕ
- **Posterior predictive distribution** at any location using composition sampling:

$$\int N(y(s) | x(s)' \beta + w(s), \tau^2 I) \times N(w(s) | a(s)' w_R, d(s)) \times \\ p(w, \beta, \tau^2, \sigma^2, \phi | y) d(w, \beta, \tau^2, \sigma^2, \phi)$$

Choosing m



- Run NNGP in parallel for few values of m
- Choose m based on model evaluation metrics
- Our results suggested that typically $m \approx 20$ yielded excellent approximations to the full GP

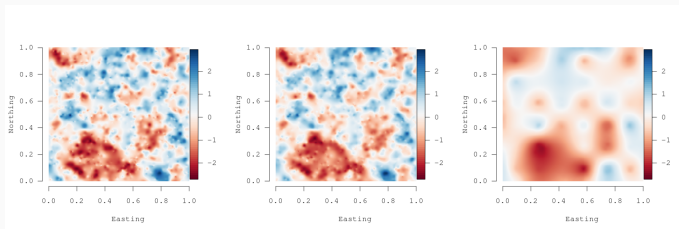
Storage and computation

- Storage:
 - **Never** needs to store $n \times n$ distance matrix
 - Stores smaller $m \times m$ matrices
 - Total storage requirements $O(nm^2)$
- Computation:
 - Only involves inverting small $m \times m$ matrices
 - Total flop count per iteration of Gibbs sampler is $O(nm^3)$
- Since $m \ll n$, NNGP offers great **scalability** for large datasets

Simulation experiments

- 2500 locations on a unit square
- $y(s) = \beta_0 + \beta_1 x(s) + w(s) + \epsilon(s)$
- Single covariate $x(s)$ generated as iid $N(0, 1)$
- Spatial effects generated from $GP(0, \sigma^2 R(\cdot, \cdot | \phi))$
- $R(\cdot, \cdot | \phi)$ is exponential correlation function with decay ϕ
- Candidate models: Full GP, Gaussian Predictive Process (GPP) with 64 knots and NNGP

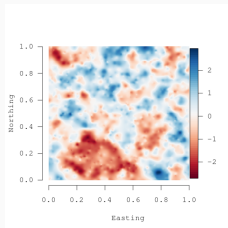
Fitted Surfaces



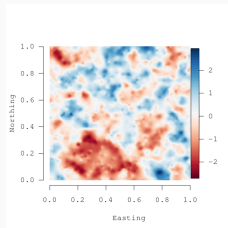
True w

Full GP

GPP 64 knots



NNGP, $m = 10$



NNGP, $m = 20$

Figure: Univariate synthetic data analysis

Parameter estimates

		NNGP		Predictive Process	Full
	True	$m = 10$	$m = 20$	64 knots	Gaussian Process
β_0	1	1.00 (0.62, 1.31)	1.03 (0.65, 1.34)	1.30 (0.54, 2.03)	1.03 (0.69, 1.34)
β_1	5	5.01 (4.99, 5.03)	5.01 (4.99, 5.03)	5.03 (4.99, 5.06)	5.01 (4.99, 5.03)
σ^2	1	0.96 (0.78, 1.23)	0.94 (0.77, 1.20)	1.29 (0.96, 2.00)	0.94 (0.76, 1.23)
τ^2	0.1	0.10 (0.08, 0.13)	0.10 (0.08, 0.13)	0.08 (0.04, 0.13)	0.10 (0.08, 0.12)
ϕ	12	12.93 (9.70, 16.77)	13.36 (9.99, 17.15)	5.61 (3.48, 8.09)	13.52 (9.92, 17.50)

Model evaluation

	NNGP		Predictive Process	Full
	$m = 10$	$m = 20$	64 knots	Gaussian Process
DIC score	2390	2377	13678	2364
RMSPE	1.2	1.2	1.68	1.2
Run time (Minutes)	14.40	46.47	43.36	560.31

- NNGP performs at par with Full GP
- GPP oversmooths and performs much worse both in terms of parameter estimation and model comparison
- NNGP yields huge computational gains

Spatially varying coefficients

- Often the relationship between the (univariate) spatial response and covariates vary across the space
- The regression coefficients can then be modeled as spatial processes
- Spatially varying coefficient (SVC) model:
$$y(s) = x(s)' \beta(s) + \epsilon(s)$$
- Even though the response can be univariate, $\beta(s)$ is modeled as a p -variate GP

- $\text{Cov}(w(s_i), w(s_j)) = C(s_i, s_j | \theta)$ – a $q \times q$ cross-covariance matrix
- Choices for the function $C(\cdot, \cdot | \theta)$
 - Multivariate Matérn
 - Linear model of co-regionalization
- For data observed at n locations, all choices lead to a dense $nq \times nq$ matrix $C = \text{Cov}(w(s_1), w(s_2), \dots, w(s_n))$
- Not scalable when nq is large

Multivariate NNGPs

- Cholesky factor approach similar to the univariate case

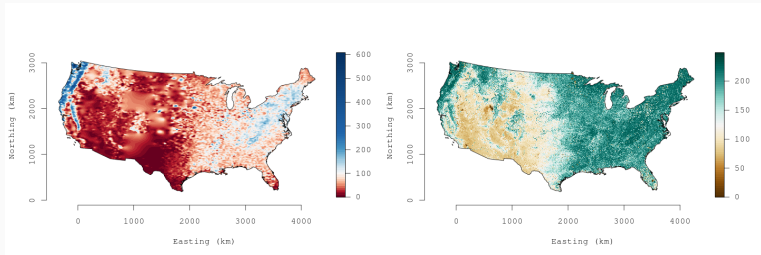
$$\begin{bmatrix} w(s_1) \\ w(s_2) \\ w(s_3) \\ \vdots \\ w(s_n) \end{bmatrix} = \begin{bmatrix} 0 & 0 & 0 & \dots & 0 & 0 \\ A_{21} & 0 & 0 & \dots & 0 & 0 \\ A_{31} & A_{32} & 0 & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots \\ A_{n1} & A_{n2} & A_{n3} & \dots & A_{n,n-1} & 0 \end{bmatrix} \begin{bmatrix} w(s_1) \\ w(s_2) \\ w(s_3) \\ \vdots \\ w(s_n) \end{bmatrix} + \begin{bmatrix} \eta(s_1) \\ \eta(s_2) \\ \eta(s_3) \\ \vdots \\ \eta(s_n) \end{bmatrix}$$
$$\implies w = Aw + \eta; \quad \eta \sim N(0, D), \quad D = \text{diag}(D_1, D_2, \dots, D_n).$$

- Only differences:** $w(s_i)$ and $\eta(s_i)$'s are $q \times 1$ vectors and A_{ij} and D_i 's are $q \times q$ matrix

Multivariate NNGPs

- Choose neighbor sets $N(i)$ for each location s_i
- Set $A_{ij} = 0$ if $j \notin N(i)$
- Solve for non-zero A_{ij} 's from the $m_q \times m_q$ linear system:
$$\sum_{j \in N(i)} A_{ij} w(s_j) = E(w(s_i) | \{w(s_j) | j \in N(i)\})$$
- **Multivariate NNGP:** $w \sim N(0, \tilde{C})$ where
$$\tilde{C}^{-1} = (I - A)' D^{-1} (I - A)$$
- \tilde{C}^{-1} is sparse with $O(nm^2)$ non-zero $q \times q$ blocks
- $\det(\tilde{C}) = \prod_{i=1}^n \det(D_i)$
- Storage and computation needs remains **linear** in n

U.S. Forest biomass data



Observed biomass

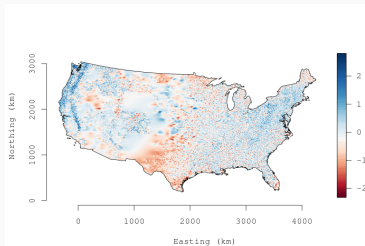
NDVI

- Forest biomass data from measurements at 114,371 plots
- NDVI (greenness) is used to predict forest biomass

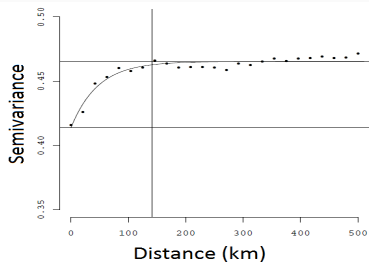
U.S. Forest biomass data

Non Spatial Model

$$\text{Biomass} = \beta_0 + \beta_1 \text{NDVI} + \text{error}, \quad \hat{\beta}_0 = 1.043, \quad \hat{\beta}_1 = 0.0093$$



Residuals



Variogram of residuals

Strong spatial pattern among residuals

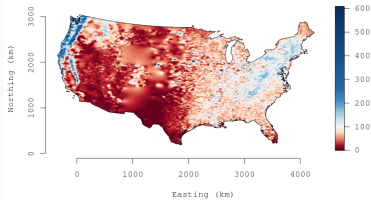
Forest biomass dataset

- $n \approx 10^5$ (Forest Biomass) \Rightarrow full GP requires storage $\approx 40Gb$ and time ≈ 140 hrs per iteration.
- We use a spatially varying coefficients NNGP model

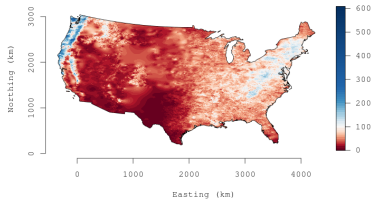
Model

- $Biomass(s) = \beta_0(s) + \beta_1(s)NDVI(s) + \epsilon(s)$
- $w(s) = (\beta_0(s), \beta_1(s))^T \sim \text{Bivariate NNGP}(0, \tilde{C}(\cdot, \cdot | \theta)),$
 $m = 5$
- Time ≈ 6 seconds per iteration
- Full inferential output: 41 hours (25000 MCMC iterations)

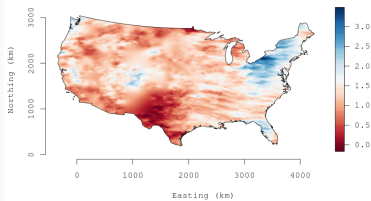
Forest biomass data



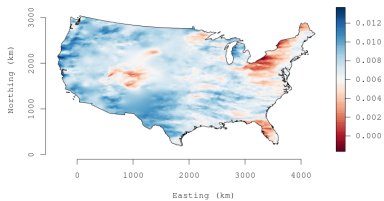
Observed biomass



Fitted biomass



$\beta_0(s)$



$\beta_{NDVI}(s)$

Reducing parameter dimensionality

- The Gibbs sampler algorithm for the NNGP updates $w(s_1), w(s_2), \dots, w(s_n)$ sequentially
- Dimension of the MCMC for this sequential algorithm is $O(n)$
- If the number of data locations n is very large, this high-dimensional MCMC can converge slowly
- Although each iteration for the NNGP model will be very fast, many more MCMC iterations may be required

- Same model:

$$y(s) = x(s)' \beta + w(s) + \epsilon(s)$$

$$w(s) \sim NNGP(0, C(\cdot, \cdot | \theta))$$

$$\epsilon(s) \stackrel{\text{iid}}{\sim} N(0, \tau^2)$$

- Vector form $y \sim N(X\beta + w, \tau^2 I)$; $w \sim N(0, \tilde{C}(\theta))$
- Collapsed model:** Marginalizing out w , we have $y \sim N(X\beta, \tau^2 I + \tilde{C}(\theta))$

Model

$$y \sim N(X\beta, \tau^2 I + \tilde{C}(\theta))$$

- Only involves few parameters β , τ^2 and $\theta = (\sigma^2, \phi)'$
- Drastically **reduces** the MCMC dimensionality
- Gibbs sampler updates are based on sparse linear systems using \tilde{C}^{-1}
- **Improved** MCMC convergence
- Can **recover** posterior distribution of $w \mid y$
- Complexity of the algorithm depends on the design of the data locations and is **not guaranteed to be $O(n)$**

Response NNGP

- $w(s) \sim GP(0, C(\cdot, \cdot | \theta)) \Rightarrow y(s) \sim GP(x(s)' \beta, \Sigma(\cdot, \cdot | \tau^2, \theta))$
- $\Sigma(s_i, s_j) = C(s_i, s_j | \theta) + \tau^2 \delta(s_i = s_j)$ (δ is Kronecker delta)
- We can directly derive the NNGP covariance function corresponding to $\Sigma(\cdot, \cdot)$
- $\tilde{\Sigma}$ is the NNGP covariance matrix for the n locations
- **Response model:** $y \sim N(X\beta, \tilde{\Sigma})$
- Storage and computations are guaranteed to be $O(n)$
- Low dimensional MCMC \Rightarrow Improved convergence
- **Cannot** coherently recover $w | y$

Comparison of NNGP models

	Latent	Collapsed	Response
$O(n)$ time	Yes	No	Yes
Recovery of $w \mid y$	Yes	Yes	No
Parameter dimensionality	High	Low	Low

Comparison of NNGP models

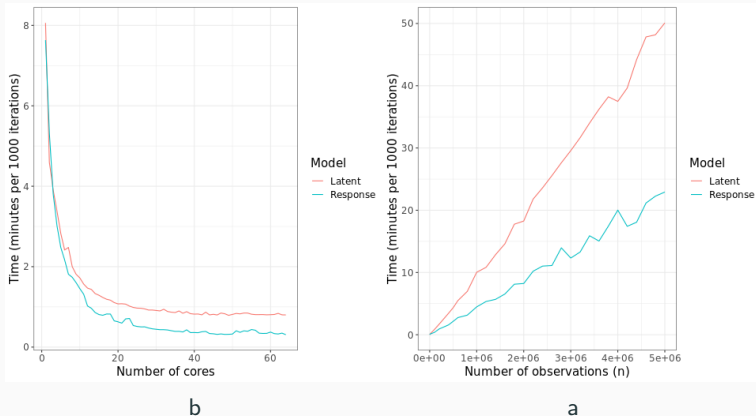


Figure: (a) Runtime for 1000 MCMC iterations for $n = 100000$ and different number of cores. (b) Runtime for 1000 MCMC iterations using 40 cores and n from 1000 to 5 million. Model type (latent and response) refers to different NNGP parameterizations.

Summary of Nearest Neighbor Gaussian Processes

- **Sparsity** inducing Gaussian process
- Constructed from sparse Cholesky factors based on m nearest neighbors
- **Scalability**: Storage, inverse and determinant of NNGP covariance matrix are all $O(n)$
- **Proper Gaussian process**, allows for inference using hierarchical spatial models and predictions at **arbitrary spatial resolution**
- Closely approximates full GP inference, does not oversmooth like low rank models
- Extension to **multivariate NNGP**
- Collapsed and response NNGP models with improved MCMC convergence
- **spNNGP package in R** for analyzing large spatial data using NNGP models