

Modeling non-Gaussian spatial data

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Non-Gaussian spatial data

- Often data sets preclude Gaussian modeling: $y(\mathbf{s})$ may not even be continuous
- Examples:
 - Binary: presence or absence of a species at location \mathbf{s} .
 - Count: abundance of a species at location \mathbf{s} .
 - Categorical: counts of trees by size class at location \mathbf{s} .
- Replace Gaussian likelihood by exponential family member (Diggle, Tawn, and Moyeed (1998)).

Hierarchical Bayesian approach

- **First stage:** $y(\mathbf{s}_i)$ are conditionally independent given β and $w(\mathbf{s}_i)$. Here we use a canonical link function, say $g(E[y(\mathbf{s}_i)]) = \eta(\mathbf{s}_i) = \mathbf{x}(\mathbf{s}_i)^\top \beta + w(\mathbf{s}_i)$.

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- **Third stage:** Priors and hyperpriors.

MCMC sampling for spatial GLMMs

- Additional GLMM flexibility comes at a computational cost: lose conjugacy of β, \mathbf{w}
- Requires more Metropolis steps. Particularly costly for \mathbf{w}
- Practical consequence: slower, less efficient algorithms
- Prediction and interpolation proceed as with the Gaussian case

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- Define $\kappa(\mathbf{s}_i) = y(\mathbf{s}_i) - N(\mathbf{s}_i)/2$
- Resulting Gibbs sampler is remarkably similar to that of a Gaussian model with response $y(\mathbf{s}_i)^* = \kappa(\mathbf{s}_i)/\omega(\mathbf{s}_i)$ and heteroskedastic variances $\tau^2(\mathbf{s}_i) = 1/\omega(\mathbf{s}_i)$.

Pólya-Gamma data augmentation

- Suppose $y(\mathbf{s}_i) \sim \text{Bernoulli}(\psi(\mathbf{s}_i))$.

$$\begin{aligned}\psi(\mathbf{s}_i)^{y(\mathbf{s}_i)}(1 - \psi(\mathbf{s}_i))^{1-y(\mathbf{s}_i)} &= \frac{\exp(\mathbf{x}(\mathbf{s}_i)^\top \boldsymbol{\beta} + w(\mathbf{s}_i))^{y(\mathbf{s}_i)}}{1 + \exp(\mathbf{x}(\mathbf{s}_i)^\top \boldsymbol{\beta} + w(\mathbf{s}_i))} \\ &= \exp(\kappa(\mathbf{s}_i)(\mathbf{x}(\mathbf{s}_i)^\top \boldsymbol{\beta} + w(\mathbf{s}_i))) \times \\ &\quad \int \exp\left(-\frac{\omega(\mathbf{s}_i)}{2}(\mathbf{x}(\mathbf{s}_i)^\top \boldsymbol{\beta} + w(\mathbf{s}_i))\right)^2 \times \\ &\quad p(\omega(\mathbf{s}_i) \mid 1, 0) d\omega(\mathbf{s}_i),\end{aligned}$$

- $p(\omega(\mathbf{s}_i) \mid 1, 0)$ is the Pólya-Gamma PDF with parameters 1 and 0
- With Gaussian priors on $\boldsymbol{\beta}$ and IG prior on σ^2 , full conditionals for $\boldsymbol{\beta}$, σ^2 , and \mathbf{w} are available in closed form. ϕ updated with MH.
- **See Polson, Scott, Windle (2013) JASA**

Example: Binary spatial regression

- Objective: predict the distribution of Loggerhead Shrike across the US

$$y(\mathbf{s}_i) \sim \text{Bernoulli}(\psi(\mathbf{s}_i))$$

$$\text{logit}(\psi(\mathbf{s}_i)) = \mathbf{x}(\mathbf{s}_i)^\top \boldsymbol{\beta} + w(\mathbf{s}_i)$$

$$\mathbf{w} \sim N(\mathbf{0}, \sigma^2 \mathbf{R}(\phi))$$

$$\boldsymbol{\beta} \sim N(\boldsymbol{\mu}_\beta, \boldsymbol{\Sigma}_\beta)$$

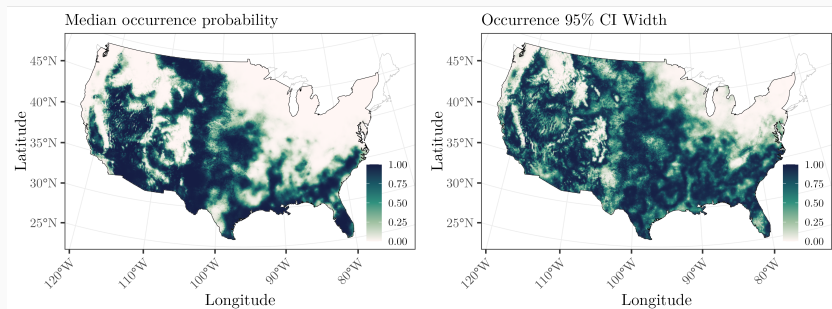
$$\sigma^2 \sim IG(a_\sigma, b_\sigma)$$

$$\phi \sim \text{Uniform}(l, u)$$

$$\omega(\mathbf{s}_i) \sim \text{PG}(1, 0)$$

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Posterior predictive inference proceeds as with the Gaussian case



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- Be careful with non-identity link functions when thinking about priors.
- Pólya-Gamma data augmentation works really for binomial data. Computational cost increases as Binomial weights increases.
- Pólya-Gamma data augmentation also applicable for Negative Binomial count data, but slow for large counts and can be unstable.

- spBayes
 - Univariate and multivariate, full GPs or predictive processes
 - Gaussian, Binomial (no Pólya-Gamma data augmentation), Poisson
- spNNGP
 - Univariate, NNGPs
 - Gaussian, Binomial
- spOccupancy
 - Univariate and multivariate, focus on modeling wildlife distributions, full GPs or NNGPs
 - Bernoulli
- spAbundance
(<https://github.com/doserjef/spAbundance>)
 - Univariate and multivariate, focus on modeling wildlife/plant abundance, NNGPs
 - Gaussian, Poisson, Negative Binomial