## STAT636 - Homework 3

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- 1. Consider the Auto data. These represent a two-factor experiment on cars. The two factors are (i) the number of cylinders (4 and 6 were considered) and (ii) origin (three origins were considered). We have 4 cars under each of the  $2 \times 3 = 6$  factor combinations. For each car, we have measurements of three weight variables:  $X_1$  = displacement,  $X_2$  = horsepower, and  $X_3$  = acceleration. So, in terms of a two-way MANOVA model, q = 2, b = 3, and n = 4.
  - > auto<- read.csv("Auto\_hw.csv")</pre>
  - > auto\$cylinders <- as.factor(auto\$cylinders)
  - > auto\$origin <- as.factor(auto\$origin)
  - (a) Test for a location effect, a variety effect, and a location-variety interaction at  $\alpha = 0.05$ . Do this using the manova function in R. Overall, what do you conclude about these data? As all the p-values are lower than the given  $\alpha$ , I can conclude that there is enough evidence against the  $H_0$ , suggesting that there is a real difference in the means associated to the analyzed factors

> manovaResult

## Call:

manova(cbind(displacement, horsepower, acceleration) ~ origin +
 cylinders + origin \* cylinders, auto)

## Terms:

	origin	cylinders	origin:cylinders	Residuals
resp 1	9843.25	34808.17	5766.58	4692.00
resp 2	787.00	4082.04	1057.33	6857.25
resp 3	23.24	5.04	5.74	105.91
Deg. of Freed	om 2	1	2	18

Residual standard errors: 16.14517 19.51815 2.425673 Estimated effects may be unbalanced

> summary(manovaResult, test = "Wilks")

```
Df
                      Wilks approx F num Df den Df
                                                        Pr(>F)
origin
                  2 0.20304
                                6.503
                                            6
                                                  32 0.0001468 ***
cylinders
                  1 0.11520
                               40.961
                                            3
                                                  16 9.813e-08 ***
origin:cylinders 2 0.25986
                                5.129
                                            6
                                                  32 0.0008635 ***
Residuals
                 18
```

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

(b) Construct the two-way MANOVA table by computing SSP<sub>FAC1</sub>, SSP<sub>FAC2</sub>, SSP<sub>INT</sub>, SSP<sub>RES</sub>, and SSP<sub>COR</sub>. Provide R code that matches the Wilks' statistics computed by manova. Note that

your p-values (computed according to the notes) will not match those of manova, because the distributional results we have learned for two-way MANOVA are large-sample approximations. That said, how do your p-values compare to those of manova? All the p-values allow performing the same inference about the effects of the factors analyzed. Those computed manually tend to be smaller than the reported by the manova function

```
> n <- 4
> p <- 3
> g <- 2
> b <- 3
> xBar <- apply(auto[,2:4],2,mean)</pre>
> SSPfac1 <- SSPfac2 <- SSPint <- SSPres <- SSPcor <- 0
> for (l in unique(auto$cylinders)){
    xBar_1 \leftarrow apply(auto[auto[,1] == 1,2:4],2,mean)
    SSPfac1 \leftarrow SSPfac1 + ((b * n * (xBar_1-xBar)^2))
    for(k in unique(auto$origin)){
      xBar_k \leftarrow apply(auto[auto[,5] == k,2:4],2,mean)
      xBar_1k \leftarrow apply(auto[auto[,1] == 1 \& auto[,5] == k,2:4],2,mean)
      SSPint \leftarrow SSPint + (n * ((xBar_lk - xBar_l - xBar_k + xBar)^2))
+
      for(r in seq_len(n)){
        x \leftarrow auto[auto[,1] == 1 \& auto[,5] == k \& seq_len(n) == r,2:4]
        SSPres \leftarrow SSPres + ((x - xBar_1k)^2)
        SSPcor \leftarrow SSPcor + ((x - xBar)^2)
      }
    }
+
+ }
> for(k in unique(auto$origin)){
      xBar_k \leftarrow apply(auto[auto[,5] == k,2:4],2,mean)
      SSPfac2 \leftarrow SSPfac2 + (g * n * ((xBar_k - xBar)^2))
+ }
> manovaTable <- cbind(SSPfac1,SSPfac2, SSPint,</pre>
                         t(SSPres), t(SSPcor))
> colnames(manovaTable) <- c("SSP_fac1", "SSP_fac2", "SSP_int",
                                "SSP_res", "SSP_cor")
> manovaTable
                  SSP_fac1
                              SSP_fac2
                                            SSP_int SSP_res
                                                                SSP_cor
displacement 34808.166667 9843.25000 5766.583333 4692.00 55110.0000
               4082.041667
                            787.00000 1057.333333 6857.25 12783.6250
horsepower
acceleration
                  5.041667
                              23.24333
                                          5.743333 105.91
                                                               139.9383
> xBar <- as.numeric(apply(auto[,2:4],2,mean))</pre>
> SSPfac1 <- SSPfac2 <- SSPint <- SSPres <- SSPcor <- 0
> for (l in unique(auto$cylinders)){
    xBar_1 \leftarrow as.numeric(apply(auto[auto[,1] == 1,2:4],2,mean))
    SSPfac1 \leftarrow SSPfac1 + ((b * n * (xBar_1-xBar) %*% t(xBar_1 - xBar)))
    for(k in unique(auto$origin)){
      xBar_k \leftarrow as.numeric(apply(auto[auto[,5] == k,2:4],2,mean))
      xBar_lk <- as.numeric(apply(auto[auto[,1] == 1 &
                                            auto[,5] == k,2:4],2,mean))
      SSPint <- SSPint + (n * ((xBar_lk - xBar_l - xBar_k + xBar) %*%
                                    t(xBar_1k - xBar_1 - xBar_k + xBar)))
+
      for(r in seq_len(n)){
```

```
x \leftarrow as.numeric(auto[auto[,1] == 1 \& auto[,5] == k \&
                                seq_len(n) == r, 2:4])
        SSPres <- SSPres + ((x - xBar_lk) %*% t(x - xBar_lk))
        SSPcor \leftarrow SSPcor + ((x - xBar) \% * \% t(x - xBar))
    }
+ }
> for(k in unique(auto$origin)){
      xBar_k \leftarrow apply(auto[auto[,5] == k,2:4],2,mean)
      SSPfac2 \leftarrow SSPfac2 + (g * n * ((xBar_k - xBar) %*% t(xBar_k - xBar)))
+ }
> # Cylinders
> Lambda <- det(SSPres) / det(SSPfac1 + SSPres)
> Lambda
[1] 0.1152037
> 1 - pf((((g * b * (n - 1) - p + 1) / 2) / ((abs((g - 1) - p) + 1) / 2)) *
    (1 - Lambda) / Lambda, abs((g - 1) - p) + 1, g * b * (n - 1) - p + 1)
[1] 9.813233e-08
> # Origin
> Lambda <- det(SSPres) / det(SSPfac2 + SSPres)
> Lambda
[1] 0.2030373
> 1 - pf((((g * b * (n - 1) - p + 1) / 2) / ((abs((b - 1) - p) + 1) / 2)) *
    (1 - Lambda) / Lambda, abs((b - 1) - p) + 1, g * b * (n - 1) - p + 1)
[1] 2.888066e-06
> # Interaction
> Lambda <- det(SSPres) / det(SSPint + SSPres)
> Lambda
[1] 0.2598601
> 1 - pf((((g * b * (n - 1) - p + 1) / 2) /
            ((abs((g-1)*(b-1)-p)+1)/2))*
    (1 - Lambda) / Lambda, abs((g - 1) * (b - 1) - p) +
      1, g * b * (n - 1) - p + 1)
[1] 2.079298e-05
```

2. Conduct a simulation study to investigate the coverage probabilities of different confidence interval types with multivariate data. Let the sample size be n = 30, the number of variables p = 5, the number of simulations B = 10000, and

$$\mu' = (0, 0, 0, 0, 0)$$

and

$$\Sigma = \begin{pmatrix} 1.0 & 0.6 & 0.6 & 0.6 & 0.6 \\ 0.6 & 1.0 & 0.6 & 0.6 & 0.6 \\ 0.6 & 0.6 & 1.0 & 0.6 & 0.6 \\ 0.6 & 0.6 & 0.6 & 1.0 & 0.6 \\ 0.6 & 0.6 & 0.6 & 0.6 & 1.0 \end{pmatrix}$$

For each of B times, simulate a dataset of size n from the  $N_p(\mu, \Sigma)$  distribution, compute 95% confidence intervals of types one-at-a-time,  $T^2$  simultaneous, and Bonferroni simultaneous, and record

whether each interval contains its corresponding population mean component value. Report a  $3 \times 5$  matrix of estimated coverage probabilities. The rows of your matrix should correspond to the 3 different interval types, and the columns should correspond to the p mean components; be sure to clearly indicate which row goes with which interval type. Comment on the performance of the different interval types. The confidence intervals using the quantiles by resampling tend to be more relaxed in detecting differences with respect to the mean, meanwhile, the Bonferroni ones seem to be more accurate. T2 based confidence intervals are between the other two type of intervals in terms of accuracy.

```
> simmulationFunction <- function(n, p, rho, alpha, B = 1000) {
    mu \leftarrow rep(0, p)
    Sigma <- matrix(rho, nrow = p, ncol = p); diag(Sigma) <- 1
+
    F_{crit} \leftarrow (n - 1) * p * qf(1 - alpha, p, n - p) / (n - p)
+
+
    bon_crit \leftarrow qt(1 - alpha / (2 * p), n - 1)
    cov_Q <- cov_T2 <- cov_bon <- matrix(NA, nrow = B, ncol = p)</pre>
+
+
    for(i in seq_len(B)) {
+
      X <- MASS::mvrnorm(n, mu, Sigma)</pre>
      x_bar <- colMeans(X)</pre>
+
      S \leftarrow var(X)
+
+
      for(k in 1:p) {
         ci_Q \leftarrow quantile(X[,k],c((alpha/2),(1-(alpha/2))))
+
         ci_T2 \leftarrow x_bar[k] + c(-1, 1) * sqrt(F_crit * S[k, k] / n)
+
         ci_bon \leftarrow x_bar[k] + c(-1, 1) * bon_crit * sqrt(S[k, k] / n)
+
+
         cov_Q[i, k] \leftarrow (ci_Q[1] \leftarrow mu[k]) & (mu[k] \leftarrow ci_Q[2])
         cov_T2[i, k] \leftarrow (ci_T2[1] \leftarrow mu[k]) & (mu[k] \leftarrow ci_T2[2])
+
         cov_bon[i, k] \leftarrow (ci_bon[1] \leftarrow mu[k]) & (mu[k] \leftarrow ci_bon[2])
+
      }
+
    }
+
+
    out <- matrix(data = NA, nrow = 3, ncol = p)</pre>
    rownames(out) <- c("95%:", "T2:", "Bonferroni:")
+
+
    out[1,] <- apply(cov_Q, 2, mean)
+
    out[2,] <- apply(cov_T2, 2, mean)
    out[3,] <- apply(cov_bon, 2, mean)</pre>
+
    return(out)
+
+ }
> simmulationFunction(n = 30, p = 5, alpha = 0.05, rho = 0.6, B = 10000)
                [,1]
                        [,2]
                                [,3]
                                        [,4]
                                                [,5]
95%:
             1.0000 1.0000 1.0000 1.0000 1.0000
T2:
             0.9997 0.9996 0.9998 0.9999 0.9996
Bonferroni: 0.9909 0.9912 0.9915 0.9909 0.9908
```