

STAT230 HW 7

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1 Lab 11

```
# Compute log likelihood
log_likelihood_aux = function(theta,X){
  n = length(X)
  n*log(theta) - 2*sum(log(theta+X))
}
log_likelihood = Vectorize(log_likelihood_aux)

# Change variable theta<-exp(phi)
log_likelihood_phi_aux = function(phi,X){
  n = length(X)
  n*phi - 2*sum(log(exp(phi)+X))
}

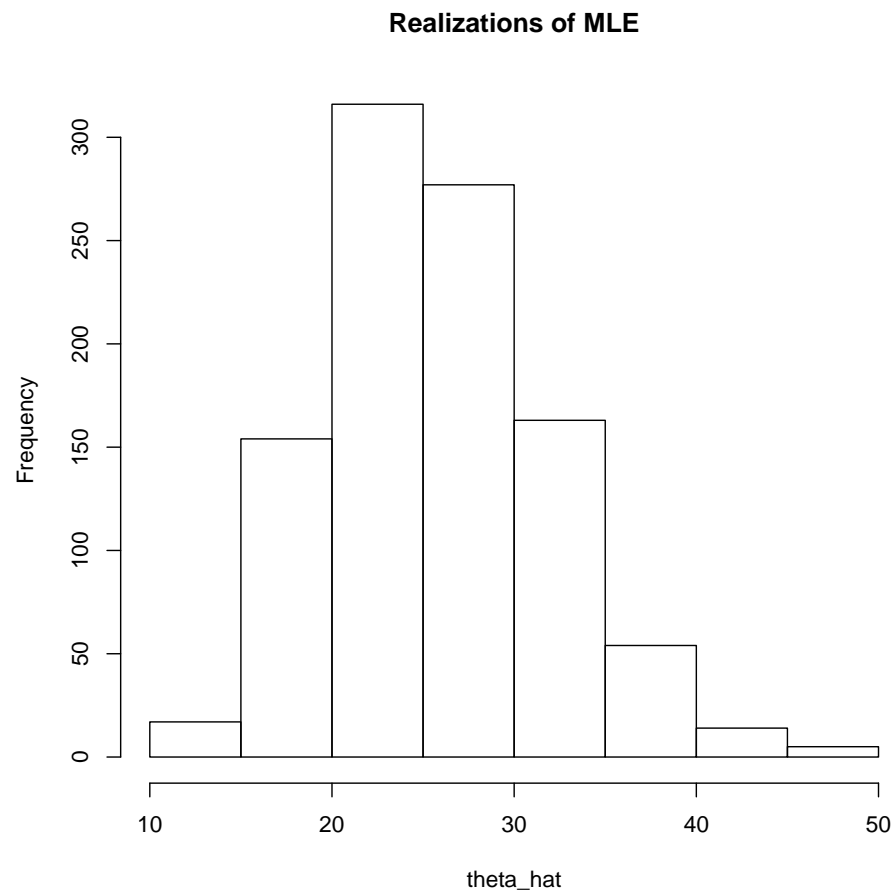
MLE = function (X){
  f = function(phi){-log_likelihood_phi_aux(phi,X)}
  opt = optim(0,f,
             method="Brent",
             lower = 0,
             upper = 10)
  theta_hat = exp(opt$par)
  theta_hat
}
```

1.1 Generate uniform RV

```
set.seed(1)
theta_hat_values = c()
for (i in 1:1000){
  # Generate data
  U = runif(50)
  theta = 25
  X = theta*sapply(U,function(x) x/(1-x))
  # Compute MLE
  theta_hat = MLE(X)
  theta_hat_values = c(theta_hat_values,theta_hat)
}
```

1.2 Plot histogram

```
hist(theta_hat_values,
      xlab = "theta_hat",
      main = "Realizations of MLE")
```



1.3 Mean/SD

```
mu = mean(theta_hat_values)
mu

## [1] 25.72292

sigma = sd(theta_hat_values)
sigma

## [1] 6.04452
```

```

fisher_info = function(theta){
  1/(3*theta^2)
}
# Comparison
asympt_var = 1/sqrt(50*fisher_info(25))
sigma

asympt_var-sigma

```

The asymptotic sd is equal to $1/\sqrt{50I_{\theta}(25)}$. Therefore they should be equal for an infinite number of simulations, c.f. formula Example 4, Chapter 7.

1.4 Bonus

6.

The asymptotic sd should be more accurate to estimate SE because it is a limit of the sd of a n-sized sample.

7.

The asymptotic sd doubles and the fisher info is divided by 4, see formula. The standard deviation of the observed info will also double and will still tend to the asymptotic sd as n grows to infinity.

2 Lab 12

```

data = read.table("pac01.dat")
head(data)

##      V1 V2 V3 V4 V5 V6 V7 V8 V9 V10 V11 V12 V13 V14
##      V15 V16
## 1  23   1  4  8  4  2  3 37  1  2  2  40  0  1
##    22   9

```

```

## 2 45 1 1 8 1 1 3 39 6 0 1 -1 1 0
    0 0
## 3 39 2 3 8 1 1 3 40 1 3 2 46 0 0
    18 2
## 4 16 1 3 8 7 0 3 35 4 8 11 -1 6 1
    24 15
## 5 53 2 1 8 5 0 1 41 1 1 2 54 0 0
    17 2
## 6 42 1 1 8 7 0 1 42 1 1 2 50 0 0
    10 1
##      V17      V18 V19 V20 V21 V22 V23 V24      V25
    V26 V27
## 1 15002 15002 1 0 445 1 1 0 216162
    330502 91
## 2 2200 28300 1 0 1451 1 1 2 259495
    324334 91
## 3 26100 28300 1 0 1451 1 2 1 205681
    324334 91
## 4 0 28300 1 0 1451 1 3 0 218787
    356936 91
## 5 23000 23000 1 0 4356 1 1 0 200660
    347911 91
## 6 44297 44297 1 0 4357 1 1 0 206279
    351372 91

names(data)=c("AGE",
              "SEX",
              "RACE",
              "ETHNICITY",
              "MARITAL",
              "NUMKIDS",
              "FAMPERS",
              "EDLEVEL",
              "LABSTAT",
              "CLASSWORK",
              "FULLPART",
              "HOURS",
              "WHYNOTWORK",
              "INSCHOOL",

```

```

"INDUSTRY",
"OCCUPATION",
"PINCOME",
"INCFAM",
"CITIZEN",
"IMMIGYR",
"HHSEQNUM",
"FSEQNUM",
"PERSCODE",
"SPOUCODE",
"FINALWGT",
"MARCHWGT",
"STATE")

```

```
head(data)
```

```

##    AGE SEX RACE ETHNICITY MARITAL NUMKIDS FAMPERS
      EDLEVEL
## 1  23  1   4         8         4         2         3
      37
## 2  45  1   1         8         1         1         3
      39
## 3  39  2   3         8         1         1         3
      40
## 4  16  1   3         8         7         0         3
      35
## 5  53  2   1         8         5         0         1
      41
## 6  42  1   1         8         7         0         1
      42
##    LABSTAT CLASSWORK FULLPART HOURS WHYNOTWORK
      INSCHOOL
## 1         1         2         2    40         0
      1
## 2         6         0         1    -1         1
      0
## 3         1         3         2    46         0
      0
## 4         4         8        11    -1         6

```

| | | | | | | |
|------|----------|------------|----------|----------|----------|--|
| | 1 | | | | | |
| ## 5 | 1 | 1 | 2 | 54 | 0 | |
| | 0 | | | | | |
| ## 6 | 1 | 1 | 2 | 50 | 0 | |
| | 0 | | | | | |
| ## | INDUSTRY | OCCUPATION | PINCOME | INCFAM | CITIZEN | |
| | IMMIGYR | | | | | |
| ## 1 | 22 | 9 | 15002 | 15002 | 1 | |
| | 0 | | | | | |
| ## 2 | 0 | 0 | 2200 | 28300 | 1 | |
| | 0 | | | | | |
| ## 3 | 18 | 2 | 26100 | 28300 | 1 | |
| | 0 | | | | | |
| ## 4 | 24 | 15 | 0 | 28300 | 1 | |
| | 0 | | | | | |
| ## 5 | 17 | 2 | 23000 | 23000 | 1 | |
| | 0 | | | | | |
| ## 6 | 10 | 1 | 44297 | 44297 | 1 | |
| | 0 | | | | | |
| ## | HHSEQNUM | FSEQNUM | PERSCODE | SPOUCODE | FINALWGT | |
| | MARCHWGT | | | | | |
| ## 1 | 445 | 1 | 1 | 0 | 216162 | |
| | 330502 | | | | | |
| ## 2 | 1451 | 1 | 1 | 2 | 259495 | |
| | 324334 | | | | | |
| ## 3 | 1451 | 1 | 2 | 1 | 205681 | |
| | 324334 | | | | | |
| ## 4 | 1451 | 1 | 3 | 0 | 218787 | |
| | 356936 | | | | | |
| ## 5 | 4356 | 1 | 1 | 0 | 200660 | |
| | 347911 | | | | | |
| ## 6 | 4357 | 1 | 1 | 0 | 206279 | |
| | 351372 | | | | | |
| ## | STATE | | | | | |
| ## 1 | 91 | | | | | |
| ## 2 | 91 | | | | | |
| ## 3 | 91 | | | | | |
| ## 4 | 91 | | | | | |
| ## 5 | 91 | | | | | |

```
## 6      91

subset = data[,c("LABSTAT", "AGE", "SEX", "RACE", "EDLEVEL")]
head(subset)

##      LABSTAT AGE SEX RACE EDLEVEL
## 1         1  23  1    4       37
## 2         6  45  1    1       39
## 3         1  39  2    3       40
## 4         4  16  1    3       35
## 5         1  53  2    1       41
## 6         1  42  1    1       42
```

2.1

I chose to only define binary random variables in order to avoid putting more weight on some factors. The baseline individual in the model, choose a person who is male, non- white, age 1619, and did not graduate from high school. It corresponds to zero values variables in the model.

```
# Split age
split_age1 = rep(0,length(data$AGE))
split_age2 = rep(0,length(data$AGE))
split_age3 = rep(0,length(data$AGE))

split_age1[data$AGE>=20 & data$AGE<=39] = 1#"2039"
split_age2[data$AGE>=40 & data$AGE<=64] = 1#"40-64"
split_age3[data$AGE>=65] = 1#"65+"

# Split Race
split_race = rep(0,length(data$RACE))
split_race[data$RACE!=1] = 1#"white"

# Split Education level
split_edlevel1 = rep(0,length(data$EDLEVEL))
split_edlevel2 = rep(0,length(data$EDLEVEL))
split_edlevel1[data$EDLEVEL==39] = 1#"HS "
```



```

split_edlevel2[data$EDLEVEL>=40] = 1#"HS+"

# Split SEX
split_sex = rep(NA,length(data$SEX))
split_sex[data$SEX==1]=0#"M"
split_sex[data$SEX==2]=1#"F"

features = data.frame(LABSTAT = as.numeric(data$LABSTAT==1),
                      SEX = as.numeric(split_sex),
                      AGE1 = as.numeric(split_age1),
                      AGE2 = as.numeric(split_age2),
                      AGE3 = as.numeric(split_age3),
                      RACE = as.numeric(split_race),
                      EDLEVEL1 = as.numeric(split_edlevel1),
                      EDLEVEL2 = as.numeric(split_edlevel2))

head(features)

##      LABSTAT SEX AGE1 AGE2 AGE3 RACE EDLEVEL1
##      EDLEVEL2
## 1          1  0    1    0    0    1          0
##          0
## 2          0  0    0    1    0    0          1
##          0
## 3          1  1    1    0    0    1          0
##          1
## 4          0  0    0    0    0    1          0
##          0
## 5          1  1    0    1    0    0          0
##          1
## 6          1  0    0    1    0    0          0
##          1

any(is.na(features))

## [1] FALSE

## # 1.
design = features[,-1]
design$Intercept = as.numeric(1)
head(design)

```

```
##      SEX AGE1 AGE2 AGE3 RACE EDLEVEL1 EDLEVEL2
      Intercept
## 1    0     1    0    0    1         0         0
      1
## 2    0     0    1    0    0         1         0
      1
## 3    1     1    0    0    1         0         1
      1
## 4    0     0    0    0    1         0         0
      1
## 5    1     0    1    0    0         0         1
      1
## 6    0     0    1    0    0         0         1
      1

names_features = names(design)
design = as.matrix(design)

# Size of design matrix
dim(design)

## [1] 13803      8

Y=features[,1]
```

2.2

I use pracma library in order to find the maximum likelihood estimator.

```
library(pracma)

Eta = function(x){
  sapply(x,function(x) 1/(1+exp(-x)))
}

# Negative log likelihood
loss = function(beta){
```

```

x = design %*% beta
-(sum(Y*log(Eta(x))+(1-Y)*log(1-Eta(x))))
}

beta0 = as.matrix(rep(0,dim(design)[2]))
loss(beta0)

## [1] 9567.511

mini = fminsearch(loss,beta0)
beta_hat = mini$xval
names(beta_hat) <- names_features
beta_hat

##          SEX          AGE1          AGE2          AGE3
##          RACE
## -0.7176359  1.3289470  1.2365739 -1.6667350
## -0.1826282
## EDLEVEL1 EDLEVEL2 Intercept
##  0.7369803  0.9931386 -0.5787865

```

We can also fit the logit with the glm function directly:

```

## Alternative
glm.fit <- glm(LABSTAT ~. , data = features, family = "binomial")
summary(glm.fit)

##
## Call:
## glm(formula = LABSTAT ~ ., family = "binomial",
##      data = features)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.9517  -0.8754   0.5925   0.8144   2.5251
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)

```

```
## (Intercept) -0.57880      0.06953   -8.324   < 2e-16
## ***
## SEX          -0.71753      0.04088  -17.551   < 2e-16
## ***
## AGE1         1.32888      0.07487   17.750   < 2e-16
## ***
## AGE2         1.23652      0.07563   16.350   < 2e-16
## ***
## AGE3        -1.66690      0.10143  -16.434   < 2e-16
## ***
## RACE         -0.18265      0.05034   -3.628  0.000285
## ***
## EDLEVEL1      0.73705      0.05657   13.029   < 2e-16
## ***
## EDLEVEL2      0.99318      0.05067   19.601   < 2e-16
## ***
## ---
## Signif. codes:
## 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken
## to be 1)
##
##      Null deviance: 18319   on 13802   degrees of
##      freedom
## Residual deviance: 14879   on 13795   degrees of
##      freedom
## AIC: 14895
##
## Number of Fisher Scoring iterations: 4
```

2.3

```
# Standard errors
summary(glm.fit)$coefficients[,2]
```

| ## | (Intercept) | SEX | AGE1 | AGE2 |
|----|-------------|------------|------------|------------|
| ## | AGE3 | | | |
| ## | 0.06953246 | 0.04088345 | 0.07486672 | 0.07562729 |
| | 0.10143142 | | | |
| ## | RACE | EDLEVEL1 | EDLEVEL2 | |
| ## | 0.05034075 | 0.05657165 | 0.05066913 | |

2.4

When looking at the sign of the coefficients we can first say that employment is positively correlated with having been to high school or above, since the baseline is no high school and the coefficients for EDLEVEL1 and EDLEVEL2 are positive. Similarly, we can also conclude that being either a woman or non-white has a bad impact on employment. Moreover, it is important to notice that the p-values are all very small, which shows the importance of all the features used to predict the outcome.

2.5

First of all there is no way to correctly quantify the education, so we cannot use a real valued variable and perform regression. As a matter of fact we have categorical variables. We use dummy variables in order to avoid giving more weight to some variables: the way of assigning factors matters in the regression in the sense that the linear model gives more importance to categories with a bigger factor.

2.6

The fact that most of women give birth may impact their employment since the employers know that they might not be able to work for a while. Moreover, the SEX variable might be correlated with the education level for example. It is known that women have less access to education than men for some reasons and it might impact their employment. LABSTAT codes more than 4 are relevant because women are more likely not to work than men do. Therefore they might not be looking for a job, which is generally not the case for men.