Prediction

David Puelz
The University of Austin

Regression

Regression analysis is the most widely used statistical tool for understanding relationships among variables

It provides a conceptually simple method for investigating functional relationships between one or more factors and an outcome of interest

The relationship is expressed in the form of an equation or a model connecting the response or dependent variable and one or more explanatory or predictor variables

Why?

Straight-up prediction:

- How much will I sell my house for?

Explanation and understanding:

- What is the impact of adding a bedroom to my sale price?

Example 1: Predicting house prices

Problem:

Predict market price based on observed characteristics

Solution:

- Look at property sales data where we know the price and some observed characteristics.
- Build a decision rule that predicts price as a function of the observed characteristics.

Q: What characteristics do we use?

Q: What characteristics do we use?

We have to define the variables of interest and develop a specific quantitative measure of these variables ...

Many factors or variables affect the price of a house:

- size
- number of baths
- garage, air conditioning, etc
- neighborhood

To keep things super simple, let's focus only on size. The value

that we seek to predict is called the dependent (or output) variable, and we denote this:

-Y =price of house (e.g. thousands of dollars)

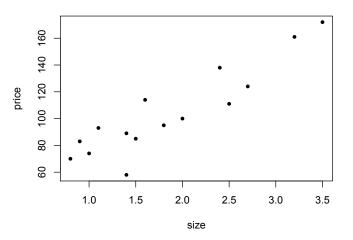
The variable that we use to guide prediction is the explanatory (or input) variable, and this is labeled

-X =size of house (e.g. thousands of square feet)

What does this data look like?

| Size | Price |
|------|-------|
| 0.80 | 70 |
| 0.90 | 83 |
| 1.00 | 74 |
| 1.10 | 93 |
| 1.40 | 89 |
| 1.40 | 58 |
| 1.50 | 85 |
| 1.60 | 114 |
| 1.80 | 95 |
| 2.00 | 100 |
| 2.40 | 138 |
| 2.50 | 111 |
| 2.70 | 124 |
| 3.20 | 161 |
| 3.50 | 172 |

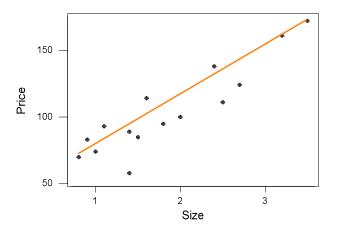
It is much more useful to look at a scatterplot



In other words, view the data as points in the $X \times Y$ plane.

There seems to be a linear relationship between price and size:

As size goes up, price goes up.

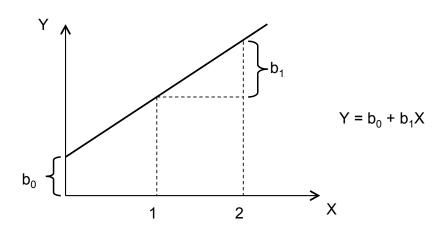


Recall that the equation of a line is:

$$Y = b_0 + b_1 X$$

Where b_0 is the intercept and b_1 is the slope.

- \rightarrow The intercept value is in units of Y (\$1,000)
- \rightarrow The slope is in units of *Y* per units of *X* (\$1,000/1,000 sq ft)



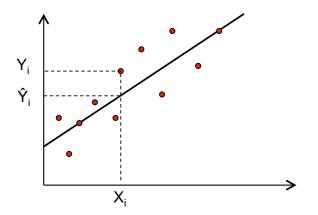
Q: How to find the "best line"?

We desire a strategy for estimating the slope and intercept parameters in the model $\hat{Y} = b_0 + b_1 X$

A reasonable way to fit a line is to minimize the amount by which the fitted value differs from the actual value.

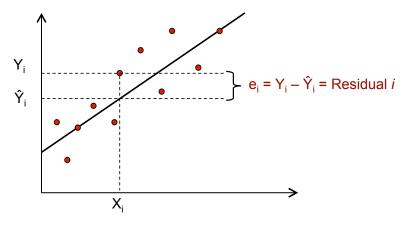
This amount is called the residual.

What is the "fitted value"?



The dots are the observed values and the line represents our fitted values given by $\hat{Y}_i = b_0 + b_1 X_1$.

What is the "residual" for the *i*th observation?



We can write $Y_i = \hat{Y}_i + (Y_i - \hat{Y}_i) = \hat{Y}_i + e_i$.

Least squares

Ideally, we want to minimize the size of all residuals:

- If they were all zero we would have a perfect line.
- Trade-off between moving closer to some points and at the same time moving away from other points.

Least squares

Ideally, we want to minimize the size of all residuals:

- If they were all zero we would have a perfect line.
- Trade-off between moving closer to some points and at the same time moving away from other points.

The line fitting process:

Minimize the "total" of residuals to get best fit.

Least squares

Ideally, we want to minimize the size of all residuals:

- If they were all zero we would have a perfect line.
- Trade-off between moving closer to some points and at the same time moving away from other points.

The line fitting process:

– Minimize the "total" of residuals to get best fit. Least Squares chooses b_0 and b_1 to minimize $\sum_{i=1}^{N} e_i^2$

$$\sum_{i=1}^{N} e_i^2 = e_1^2 + e_2^2 + \dots + e_N^2 = (Y_1 - \hat{Y}_1)^2 + (Y_2 - \hat{Y}_2)^2 + \dots + (Y_N - \hat{Y}_N)^2$$

Example 2: Offensive performance in baseball

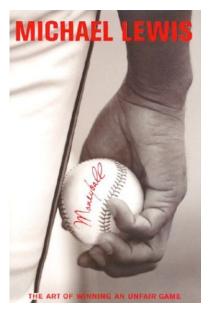
Problems:

- Evaluate/compare traditional measures of offensive performance
- Help evaluate the worth of a player

Solutions:

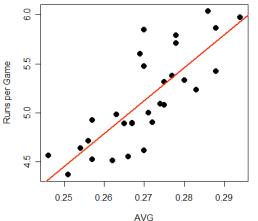
 Compare prediction rules that forecast runs as a function of either AVG (batting average), SLG (slugging percentage – total bases divided by at bats) or OBP (on base percentage)

Example 2: Offensive performance in baseball



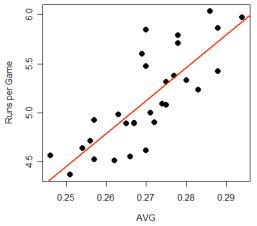
Baseball data – using AVG

Each observation corresponds to a team in MLB. Each quantity is the average over a season.



Y = runs per game; X = AVG (average)LS fit: Runs/Game = -3.93 + 33.57 AVG

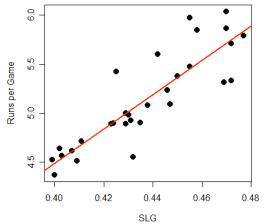
Baseball data – using AVG



Y = runs per game; X = AVG (average)

LS fit: Runs/Game = -3.93 + 33.57 AVG

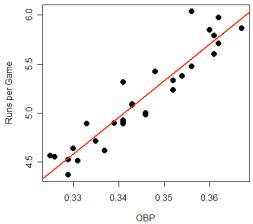
Baseball data – using SLG



Y = runs per game; X = SLG (slugging percentage)

LS fit: Runs/Game = -2.52 + 17.54 SLG

Baseball data – using OBP



Y = runs per game; X = OBP (on base percentage) LS fit: Runs/Game = -7.78 + 37.46 OBP

Baseball data

- What is the best prediction rule?
- Let's compare the predictive ability of each model using the average squared error

$$\sqrt{\frac{1}{N}\sum_{i=1}^{N}e_{i}^{2}} = \left(\frac{\sum_{i=1}^{N}\left(\widehat{\mathsf{Runs}_{i}} - \mathsf{Runs}_{i}\right)^{2}}{N}\right)^{\frac{1}{2}}$$

Place your money on OBP!

| Root Mean Squared Error | |
|-------------------------|------|
| AVG | 0.29 |
| SLG | 0.23 |
| OBP | 0.16 |

More on least squares

Remember how we get the slope (b_1) and intercept (b_0) . We minimize the sum of squared prediction errors.

The formulas for b_0 and b_1 that minimize the least squares criterion are:

$$b_1 = r_{xy} \times \frac{s_y}{s_x} \qquad b_0 = \bar{Y} - b_1 \bar{X}$$

where,

- \bar{X} and \bar{Y} are the sample mean of X and Y
- $corr(x, y) = r_{xy}$ is the sample correlation
- $-s_x$ and s_y are the sample standard deviation of X and Y

What are these numbers in the formula?

Sample Mean: measure of centrality

$$\bar{Y} = \frac{1}{n} \sum_{i=1}^{n} Y_i$$

- Sample Variance: measure of spread

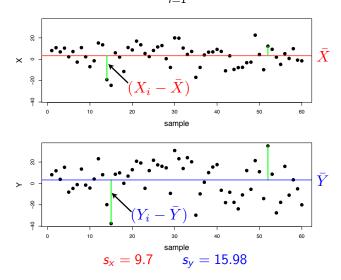
$$s_y^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2$$

Sample Standard Deviation:

$$s_y = \sqrt{s_y^2}$$

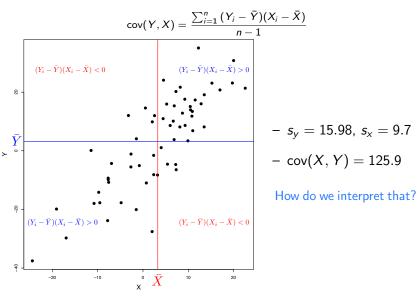
Visual: standard deviation

$$s_y^2 = \frac{1}{n-1} \sum_{i=1}^n (Y_i - \bar{Y})^2$$



Visual: Covariance

Measure the direction and strength of the linear relationship between Y and X



A standardized measure: Correlation

Correlation is the standardized covariance:

$$\operatorname{corr}(X,Y) = \frac{\operatorname{cov}(X,Y)}{\sqrt{s_x^2 s_y^2}} = \frac{\operatorname{cov}(X,Y)}{s_x s_y}$$

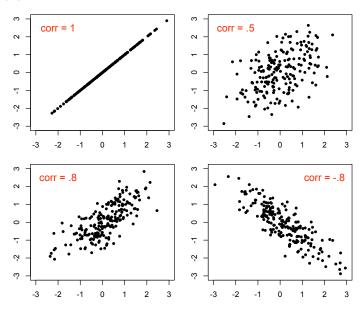
The correlation is scale invariant and the units of measurement don't matter: It is always true that $-1 \le corr(X, Y) \le 1$.

This gives the direction (negative or positive) and strength $(0 \to 1)$ of the linear relationship between X and Y.

Correlation

$$corr(Y,X) = \frac{cov(X,Y)}{\sqrt{s_X^2 s_y^2}} = \frac{cov(X,Y)}{s_X s_y} = \frac{125.9}{15.98 \times 9.7} = 0.812$$

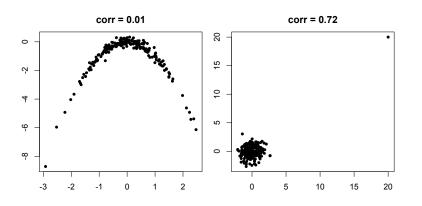
Correlation



Correlation

Only measures linear relationships:

corr(X, Y) = 0 does not mean the variables are not related!



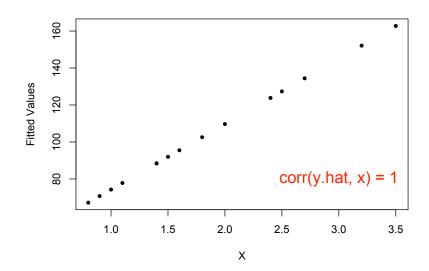
Also be careful with influential observations. Check out cor() in R.

More on least squares

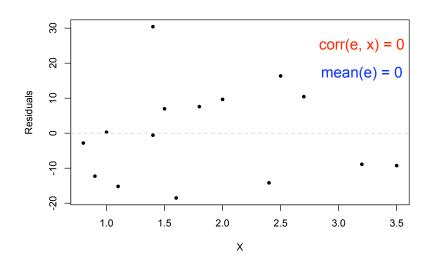
From now on, terms "fitted values" (\hat{Y}_i) and "residuals" (e_i) refer to those obtained from the least squares line.

The fitted values and residuals have some special properties. Let's look at the housing data analysis to figure out what these properties are...

The fitted values and X

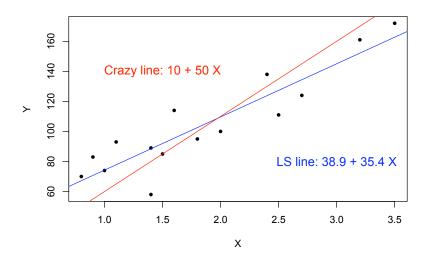


The residuals and X



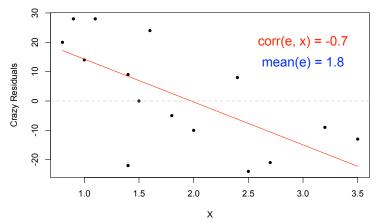
Why?

What is the intuition for the relationship between \hat{Y} and e and X? Lets consider some "crazy" alternative line:



Fitted values and residuals

This is a bad fit! We are underestimating the value of small houses and overestimating the value of big houses.



Clearly, we have left some predictive ability on the table!

Fitted values and residuals

As long as the correlation between e and X is non-zero, we could always adjust our prediction rule to do better.

We need to exploit all of the predictive power in the X values and put this into \hat{Y} , leaving no "Xness" in the residuals.

In summary: $Y = \hat{Y} + e$ where:

- \hat{Y} is "made from X"; $\operatorname{corr}(X, \hat{Y}) = 1$.
- e is unrelated to X; $\operatorname{corr}(X, e) = 0$.

Multiple linear regression

Multiple linear regression

Many problems involve more than one independent variable or factor which affects the dependent or response variable.

- More than size to predict house price!
- Demand for a product given prices of competing brands, advertising, household attributes, etc.

In SLR, the conditional mean of Y depends on X. The Multiple Linear Regression (MLR) model extends this idea to include more than one independent variable.

The MLR model

Same as always, but with more covariates.

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_p X_p + \epsilon$$

Recall the key structural assumption of our linear regression model:

 \rightarrow The conditional mean of Y is linear in the X_j variables.

The MLR model

Our interpretation of regression coefficients can be extended from the simple single covariate regression case:

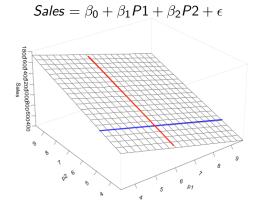
$$\beta_j = \frac{\partial E[Y|X_1, \dots, X_p]}{\partial X_j}$$

Holding all other variables constant, β_j is the average change in Y per unit change in X_j .

The MLR model

If p = 2, we can plot the regression surface in 3D.

Consider sales of a product as predicted by price of this product (P1) and the price of a competing product (P2).



Least squares again!

$$Y = \beta_0 + \beta_1 X_1 \dots + \beta_p X_p + \varepsilon$$

How do we estimate the MLR model parameters?

The principle of least squares is exactly the same as before:

- Define the fitted values
- Find the best fitting plane by minimizing the sum of squared residuals

Least squares again!

The data...

```
p1
                p2
                        Sales
5.1356702 5.2041860
                     144.48788
3.4954600 8.0597324
                     637.24524
7.2753406 11.6759787
                     620.78693
4.6628156 8.3644209
                     549.00714
3.5845370 2.1502922
                      20,42542
5.1679168 10.1530371 713.00665
3.3840914 4.9465690
                     346.70679
4.2930636 7.7605691
                     595.77625
4.3690944 7.4288974 457.64694
7.2266002 10.7113247
                     591,45483
```

The model: Sales_i = $\beta_0 + \beta_1 P1_i + \beta_2 P2_i + \epsilon_i$

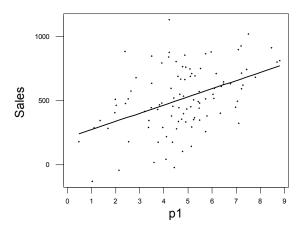
But what about the right-hand-side?

It is also important to understand and interpret the coefficients, i.e., what is is happening on the "right-hand-side" of our model ...

- Sales: units sold in excess of a baseline
- P1: our price in \$ (in excess of a baseline price)
- P2: competitors price (again, over a baseline)

But what about the right-hand-side?

If we regress Sales on our own price, we obtain a somewhat surprising conclusion... the higher the price the more we sell!



 \rightarrow It looks like we should just raise our prices, right?

The regression equation for Sales on own price (P1) is:

$$Sales = 211 + 63.7P1$$

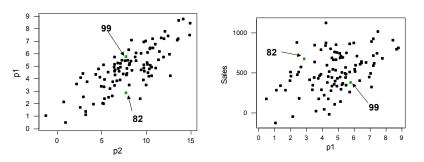
If now we add the competitors price to the regression we get

$$Sales = 116 - 97.7P1 + 109P2$$

Does this look better? How did it happen? Remember: -97.7 is the affect on sales of a change in P1 with P2 held fixed!

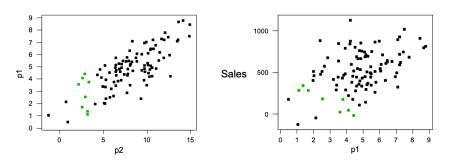
How can we see what is going on? Let's compare Sales in two different observations: weeks 82 and 99.

We see that an increase in P1, holding P2 constant, corresponds to a drop in Sales!



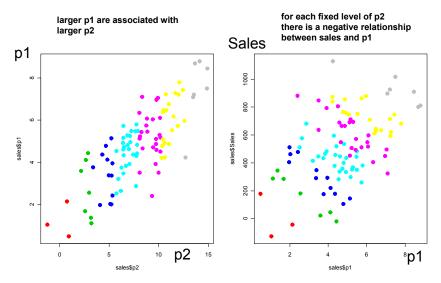
Note the strong relationship (dependence) between P1 and P2!

Let's look at a subset of points where P1 varies and P2 is held approximately constant...



For a fixed level of *P*2, variation in *P*1 is negatively correlated with Sales!

Below, different colors indicate different ranges for P2...



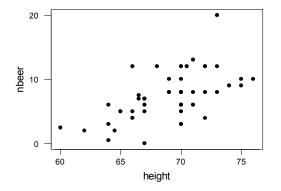
Summary:

- ightarrow A larger P1 is associated with larger P2 and the overall effect leads to bigger sales
- \rightarrow With P2 held fixed, a larger P1 leads to lower sales
- \rightarrow MLR does the trick and unveils the ${\bf correct}$ economic relationship between Sales and prices!

Example: Beers, height, weight, and getting drunk

Beer data (from PRL class last year)

- nbeer number of beers before getting drunk
- height and weight



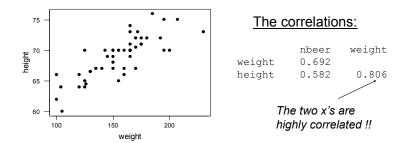
Is number of beers related to height?

R output: Yes!

```
data = read.csv('nbeer.csv')
fit = lm(nbeer~height,data)
summary(fit)
##
## Call:
## lm(formula = nbeer ~ height, data = data)
##
## Residuals:
##
     Min 1Q Median
                          30
                                Max
## -6.164 -2.005 -0.093 1.738 9.978
##
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) -36.9200 8.9560 -4.122 0.000148 ***
## height 0.6430 0.1296 4.960 9.23e-06 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.109 on 48 degrees of freedom
## Multiple R-squared: 0.3389, Adjusted R-squared: 0.3251
## F-statistic: 24.6 on 1 and 48 DF, p-value: 9.23e-06
```

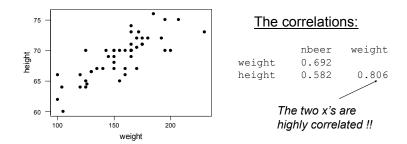
R output: What about now?

```
data = read.csv('nbeer.csv')
fit = lm(nbeer~height+weight,data)
summary(fit)
##
## Call:
## lm(formula = nbeer ~ height + weight, data = data)
##
## Residuals:
      Min
            10 Median
##
                              30
                                    Max
## -8.5080 -2.0269 0.0652 1.5576 5.9087
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) -11.18709 10.76821 -1.039 0.304167
## height 0.07751 0.19598 0.396 0.694254
## weight 0.08530 0.02381 3.582 0.000806 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 2.784 on 47 degrees of freedom
## Multiple R-squared: 0.4807, Adjusted R-squared: 0.4586
## F-statistic: 21.75 on 2 and 47 DF, p-value: 2.056e-07
```



If we regress "beers" only on height we see an effect. Taller heights go with more beers.

However, when height goes up weight tends to go up as well... in the first regression, height was a proxy for the real cause of drinking ability. Bigger people can drink more and weight is a more accurate measure of "bigness."



In the multiple regression, when we consider only the variation in height that is not associated with variation in weight, we see no relationship between height and beers.

R output: Why is this a better model than height + weight?

```
data = read.csv('nbeer.csv')
fit = lm(nbeer~weight,data)
summary(fit)
##
## Call:
## lm(formula = nbeer ~ weight, data = data)
##
## Residuals:
##
      Min 10 Median 30
                                    Max
## -8.7709 -2.0304 -0.0742 1.6580 5.6556
##
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) -7.02070 2.21329 -3.172 0.00264 **
## weight 0.09289 0.01399 6.642 2.6e-08 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 2.76 on 48 degrees of freedom
## Multiple R-squared: 0.4789, Adjusted R-squared: 0.4681
## F-statistic: 44.12 on 1 and 48 DF, p-value: 2.602e-08
```

Summary slide

In general, when we see a relationship between y and x (or x's), that relationship may be driven by variables "lurking" in the background which are related to your current x's.

This makes it hard to reliably find "causal" relationships. Any correlation (association) you find could be caused by other variables in the background... correlation is NOT causation

Any time a report says two variables are related and there's a suggestion of a "causal" relationship, ask yourself whether or not other variables might be the real reason for the effect.

Multiple regression allows us to control for all important variables by including them into the regression. "Once we control for weight, height and beers are NOT related"!

Correlation is NOT causation

